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# **Information Flow in Interactive Systems**

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**Abstract.** We consider the problem of defining the information leakage in interactive systems where secrets and observables can alternate during the computation. We show that the information-theoretic approach which interprets such systems as (simple) noisy channels is not valid anymore. However, the principle can be retrieved if we consider more complicated types of channels, that in Information Theory are known as channels with memory and feedback. We show that there is a complete correspondence between interactive systems and such kind of channels. Furthermore, we show that the capacity of the channels associated to such systems is a continuous function of the Kantorovich metric.

### **1** Introduction

Information leakage refers to the problem that the observable parts of the behavior of a system may reveal information that we would like to keep secret. In recent years, there has been a growing interest in the quantitative aspects of this problem, partly because it is convenient to represent the partial knowledge of the secrets as a probability distribution, and partly because the mechanisms to protect the information may use randomization to obfuscate the relation between the secrets and the observables.

Among the quantitative approaches, some of the most popular ones are based on Information Theory. The idea is that the system can be interpreted as an information-theoretic *channel*, where the secrets are the input and the observables are the output. The leakage is represented by the *mutual information*, and the worst-case leakage by the *capacity* of the channel. See for example [3, 10, 2, 14]. This interpretation, however, works well only when (the actions that represent) the secrets occur at the beginning of the computation. In this case the channel matrix is constituted by the conditional probabilities  $p(b \mid a)$ , given by the probability that the computation fragment *starting from the node immediately after the choice of the secret a* gives the observable *b*.

By *interactive systems* we mean systems in which secrets and observables can alternate during the computation, and influence each other. Many real-world protocols and programs are interactive. Examples of such protocols include the *auction protocols* like [19,16,15], to cite a few. Some of these have become very popular thanks to their implementation and integration in electronic commerce applications. Online auctions are an effective approach to buying and selling activities, employed in the emerging Internet-based electronic commerce platforms [7,8,12]. As for interactive programs, examples include web servers, GUI applications, and command-line programs.

In this paper we investigate the applicability of the information-theoretic approach to interactive systems. One approach, proposed in [6], consisted in defining the conditional probabilities  $p(b \mid a)$  as the ratio between the probability of the traces with

(secret, observable)-projection (a, b), and the probability of the trace with secret projection a. This looks natural, as it follows the definition of conditional probability in terms of joint and marginal probability. However, it does not help to define an information-theoretic channel. In fact, by definition a channel should be invariant with respect to the input distribution, and such construction is not, as shown by the following example.

*Example 1*. Figure 1 represents a web-based interaction between one seller and two possible buyers, *rich* and *poor*. The seller offers two different products, *cheap* and *expensive*, with given probabilities. Once the product is offered, each buyer may try to buy the product, with a certain probability. For simplicity we assume that the buyers offers are exclusive (for instance they could represent the final offers in a bidding process). We assume that the offers are observables, in the sense that they are made public in the website, while the buyers should be secret to an external observer. The symbols  $r, s, t, \overline{r}, \overline{s}, \overline{t}$  represent the probabilities, with the convention that  $\overline{r} = 1 - r$ .

Following [6] we can compute the conditional probabilities as  $p(b|a) = \frac{p(a,b)}{p(a)}$ , thus obtaining the matrix on Table 1. However, the matrix is not invariant with respect to the input distribution. For instance, if we fix  $r = \overline{r} = 0.5$  and consider two different input distributions, obtained by varying the values of (s, t), we get two different matrices of conditional probabilities, which are represented in Table 2. Hence when the secrets occur after the observables we cannot consider the conditional probabilities as representing a (classical)



Fig.1. Inter. System

channel, and we cannot apply the standard information-theoretic concepts. In particular, we cannot adopt the (classical) capacity to represent the worst-case leakage, since the capacity is defined using a fixed channel matrix over all possible input distributions.

The first contribution of this paper is to consider an extension of the theory of channels which makes the informationtheoretic approach applicable also the case of interactive systems. It turns out that a richer notion of channels, known in Information Theory as *channels with memory and feedback*, serves our purposes. The dependence of inputs on previous outputs corresponds to feedback, and the dependence of outputs on previous inputs and outputs corresponds to memory.

	cheap	expensive
poor	$\frac{rs}{rs+\overline{r}t}$	$\frac{\overline{rt}}{rs+\overline{rt}}$
rich	$\frac{r\overline{s}}{r\overline{s}+\overline{r}\overline{t}}$	$\frac{\overline{rt}}{r\overline{s}+\overline{rt}}$

**Table 1.** Cond. probabilities of Example 1

	cheap	expensive	Marginal $p_A(\cdot)$			cheap	expensive	Marginal $p_A(\cdot)$
poor	$\frac{2}{5}$	$\frac{3}{5}$	$\frac{1}{4}$		poor	$\frac{1}{4}$	$\frac{3}{4}$	$\frac{1}{2}$
rich	$\frac{8}{15}$	$\frac{7}{15}$	$\frac{3}{4}$		rich	$\frac{9}{16}$	$\frac{7}{16}$	$\frac{1}{2}$
(a) $r = \frac{1}{2}, s = \frac{2}{5}, t = \frac{3}{5}$						(b) <i>r</i>	$=\frac{1}{2}, s=\frac{1}{10}$	$t_{1}, t = \frac{3}{10}$

Table 2. Two different channel matrices depending on the input distribution

A second contribution of our work is the proof that the channel capacity is a continuous function of the Kantorovich metric on interactive systems. This was pointed out also in [6], however their construction does not work in our case due to the fact that (as far as we understand) it assumes that the probability of a secret action, in any point of the computation, is not 0. This assumption is not guaranteed in our case and therefore we had to proceed differently. The fact that our proof does not need this assumption shows that the intuition of [6] concerning the continuity of capacity is valid in general.

# 2 Preliminaries

In this section we briefly review some basic notions that we will need along the paper.

#### 2.1 Concepts from Information Theory

For more detailed information on this part we refer to [4]. Let A, B denote two random variables with corresponding probability distributions  $p_A(\cdot), p_B(\cdot)$ , respectively. We shall omit the subscripts when they are clear from the context. Let  $\mathcal{A} = \{a_1, \ldots, a_n\}, \mathcal{B} = \{b_1, \ldots, b_m\}$  denote, respectively, the sets of possible values for A and for B.

The entropy of A is defined as  $H(A) = -\sum_{A} p(a_i) \log p(a_i)$  and it measures the uncertainty of A. It takes its minimum value H(A) = 0 when  $p_A(\cdot)$  is all concentrated on one value (i.e. when  $p_A(\cdot)$  is a delta of Dirac). The maximum value  $H(A) = \log |\mathcal{A}|$  is obtained when  $p_A(\cdot)$  is the uniform distribution. Usually the base of the logarithm is set to be 2 and, correspondingly, the entropy is measured in *bits*.

The conditional entropy of A given B is  $H(A|B) = -\sum_{\mathcal{B}} p(b_i) \sum_{\mathcal{A}} p(a_j|b_i)$ log  $p(a_j|b_i)$ , and it measures the uncertainty of A when B is known. We can prove that  $0 \le H(A|B) \le H(A)$ . The minimum value, 0, is obtained when A is completely determined by B. The maximum value H(A) is obtained when B reveals no information about A.

The *mutual information* between A and B is defined as I(A; B) = H(A) - H(A|B), and it measures the amount of information that one variable carries about the other. In other words, it measures the amount of uncertainty about A that we lose by observing B. It can be shown that I(A; B) = I(B; A) and  $0 \le I(A; B) \le H(A)$ .

The entropy and mutual information respect the *chain laws*. Namely, given a sequence of random variables  $A_1, A_2, \ldots, A_k$  and B, we have:

$$H(A_1, A_2, \dots, A_k) = \sum_{i=1}^k H(A_i | A_1, \dots, A_{i-1})$$
(1)

$$I(A_1, A_2, \dots, A_k; B) = \sum_{i=1}^k I(A_i; B | A_1, \dots, A_{i-1})$$
(2)

A (discrete memoryless) information-theoretic channel is a tuple  $(\mathcal{A}, \mathcal{B}, p(\cdot|\cdot))$ , where  $\mathcal{A}, \mathcal{B}$  are the sets of input and output symbols, respectively, and  $p(b_j|a_i)$  is the probability of observing the output symbol  $b_j$  when the input symbol is  $a_i$ . An input distribution  $p(a_i)$  over  $\mathcal{A}$  determines, together with the channel, the joint distribution

 $p(a_i, b_j) = p(a_i|b_j) \cdot p(a_i)$  and consequently the mutual information I(A; B). The maximum I(A; B) over all possible input distributions is known as the channel's *capacity*. The famous result by Shannon states that the capacity of a channel coincides with the maximum rate by which information can be transmitted using the channel.

In this paper we consider input and output sequences instead of just symbols.

**Convention 1.** Let  $\mathcal{A} = \{a_1, \ldots, a_n\}$  be a finite set of *n* different symbols (alphabet). When we have a sequence of symbols (ordered in time), we use a Greek letter  $\alpha_t$  to denote the symbol at time *t*. The notation  $\alpha^t$  stands for the sequence  $\alpha_1 \alpha_2 \ldots \alpha_t$ . For instance, in the sequence  $a_3 a_7 a_5$ , we have  $\alpha_2 = a_7$  and  $\alpha^2 = a_3 a_7$ .

**Convention 2.** Let X be a random variable.  $X^t$  denotes the sequence of t consecutive occurrences  $X_1, \ldots, X_t$  of the random variable X.

When the channel is used repeatedly, the discrete memoryless channel described above represents the case in which the behavior of the channel at the present time does not depend upon the past history of inputs and outputs. If this assumption does not hold, then we have a channel *with memory*. Furthermore, if the outputs from the channel can be fed back to the encoder, thus influencing the generation of the next input symbol, then the channel is said to be *with feedback*; otherwise it is *without feedback*.

Equation 3 makes explicit the probabilistic behavior of channels regarding those classifications. Suppose a general channel from  $\mathcal{A}$  to  $\mathcal{B}$  with the associated random variables A for input and B for output. Using the notation introduced in Convention 1, the channel behavior after T uses can be fully described by the joint probability  $p(\alpha^T, \beta^T)$ .

Using probability laws we derive:

$$p(\alpha^T, \beta^T) = \prod_{t=1}^T p(\alpha_t | \alpha^{t-1}, \beta^{t-1}) p(\beta_t | \alpha^t, \beta^{t-1}) \qquad \text{(by the expansion law)} \quad (3)$$

The first term  $p(\alpha_t | \alpha^{t-1}, \beta^{t-1})$  indicates that the probability of  $\alpha_t$  depends not only on  $\alpha^{t-1}$ , but also on  $\beta^{t-1}$  (*feedback*). The second term  $p(\beta_t | \alpha^t, \beta^{t-1})$  indicates that the probability of each  $\beta_t$  depends on previous history of inputs  $\alpha^{t-1}$  and outputs  $\beta^{t-1}$  (*memory*).

If the channel is without feedback, then we have that  $p(\alpha_t | \alpha^{t-1}, \beta^{t-1}) = p(\alpha_t | \alpha^{t-1})$ , and if the channel is without memory, then we have also  $p(\beta_t | \alpha^t, \beta^{t-1}) = p(\beta_t | \alpha_t)$ . From these we derive  $p(\beta^T | \alpha^T) = \prod_{t=1}^T p(\beta_t | \alpha_t)$ , which is the classic equation for discrete memoryless channels without feedback.

Let  $(\mathcal{V}, \mathcal{K})$  be a Borel space and let  $(\mathcal{X}, \mathcal{B}_{\mathcal{X}})$  and  $(\mathcal{Y}, \mathcal{B}_{\mathcal{Y}})$  be Polish spaces equipped with their Borel  $\sigma$ -algebras. Let  $\rho(dx|r)$  be a family of measures on  $\mathcal{X}$  given  $\mathcal{V}$ . Then  $\rho(dx|r)$  is a *stochastic kernel* if and only if and only if  $\rho(\cdot|r)$  is a random variable from  $\mathcal{V}$  into the power set  $\mathcal{P}(\mathcal{X})$ .

#### 2.2 Probabilistic automata

A function  $\mu: S \to [0, 1]$  is a *discrete probability distribution* on a countable set S if  $\sum_{s \in S} \mu(s) = 1$ . The set of all discrete probability distributions on S is  $\mathcal{D}(S)$ .

A probabilistic automaton [13] is a quadruple  $M = (S, \mathcal{L}, \hat{s}, \vartheta)$  where S is a countable set of states,  $\mathcal{L}$  a finite set of labels or actions,  $\hat{s}$  the initial state, and  $\vartheta$  a transition function  $\vartheta : S \to \wp_f(\mathcal{D}(\mathcal{L} \times S))$ . Here  $\wp_f(X)$  is the set of all finite subsets of X. If  $\vartheta(s) = \emptyset$  then s is a terminal state. We write  $s \to \mu$  for  $\mu \in \vartheta(s), s \in S$ . Moreover, we write  $s \to r$  for  $s, r \in S$  whenever  $s \to \mu$  and  $\mu(\ell, r) > 0$ . A fully probabilistic automaton is a probabilistic automaton satisfying  $|\vartheta(s)| \leq 1$  for all states. When  $\vartheta(s) \neq \emptyset$  we overload the notation and denote  $\vartheta(s)$  the distribution outgoing from s.

A path in a probabilistic automaton is a sequence  $\sigma = s_0 \stackrel{\ell_1}{\to} s_1 \stackrel{\ell_2}{\to} \cdots$  where  $s_i \in S, \ell_i \in \mathcal{L}$  and  $s_i \stackrel{\ell_{i+1}}{\to} s_{i+1}$ . A path can be *finite* in which case it ends with a state. A path is *complete* if it is either infinite or finite ending in a terminal state. Given a finite path  $\sigma$ ,  $last(\sigma)$  denotes its last state. Let  $Paths_s(M)$  denote the set of all paths,  $Paths^*_s(M)$  the set of all finite paths, and  $CPaths_s(M)$  the set of all complete paths of an automaton M, starting from the state s. We will omit s if  $s = \hat{s}$ . Paths are ordered by the prefix relation, which we denote by  $\leq$ . The *trace* of a path is the sequence of actions in  $\mathcal{L}^* \cup \mathcal{L}^\infty$  obtained by removing the states, hence for the above  $\sigma$  we have  $trace(\sigma) = l_1 l_2 \dots$  If  $\mathcal{L}' \subseteq \mathcal{L}$ , then  $trace_{\mathcal{L}'}(\sigma)$  is the projection of  $trace(\sigma)$  on the elements of  $\mathcal{L}'$ .

Let  $M = (S, \mathcal{L}, \hat{s}, \vartheta)$  be a (fully) probabilistic automaton,  $s \in S$  a state, and let  $\sigma \in \operatorname{Paths}_{s}^{\star}(M)$  be a finite path starting in s. The *cone* generated by  $\sigma$  is the set of complete paths  $\langle \sigma \rangle = \{\sigma' \in \operatorname{CPaths}_{s}(M) \mid \sigma \leq \sigma'\}$ . Given a fully probabilistic automaton  $M = (S, \mathcal{L}, \hat{s}, \vartheta)$  and a state s, we can calculate the *probability value*, denoted by  $\mathbf{P}_{s}(\sigma)$ , of any finite path  $\sigma$  starting in s as follows:  $\mathbf{P}_{s}(s) = 1$  and  $\mathbf{P}_{s}(\sigma \xrightarrow{\ell} s') = \mathbf{P}_{s}(\sigma) \mu(\ell, s')$ , where  $\operatorname{last}(\sigma) \to \mu$ .

Let  $\Omega_s \triangleq \text{CPaths}_s(M)$  be the sample space, and let  $\mathcal{F}_s$  be the smallest  $\sigma$ -algebra generated by the cones. Then **P** induces a unique *probability measure* on  $\mathcal{F}_s$  (which we will also denote by  $\mathbf{P}_s$ ) such that  $\mathbf{P}_s(\langle \sigma \rangle) = \mathbf{P}_s(\sigma)$  for every finite path  $\sigma$  starting in s. For  $s = \hat{s}$  we write **P** instead of  $\mathbf{P}_{\hat{s}}$ .

Given a probability space  $(\Omega, \mathcal{F}, P)$  and two events  $A, B \in \mathcal{F}$  with P(B) > 0, the *conditional probability* of A given  $B, P(A \mid B)$ , is defined as  $P(A \cap B)/P(B)$ .

## **3** Discrete channels with memory and feedback

We adopt the model proposed in [17] for discrete channels with memory and feedback. Such model, represented in Figure 2, can be decomposed in sequential components as follows. At time t the channel's behavior is represented by the conditional probabilities  $p(\beta_t | \alpha^t, \beta^{t-1})$ . The output is fed back to the encoder with delay one. At time t - 1 the encoder takes the message and the past output symbols  $\beta^{t-1} = \beta_1, \ldots, \beta_{t-1}$ , and produces a channel input symbol  $\alpha_t$ . At final time T the decoder takes all the channel outputs  $\beta^T$  and produces the decoded message  $\hat{W}$ . The whole process respects the following order:

Message W,  $\alpha_1, \beta_1, \alpha_2, \beta_2, \ldots, \alpha_T, \beta_T$ , Decoded Message  $\hat{W}$ 

Let us describe such channel in more detail. Let  $\mathcal{A}$  and  $\mathcal{B}$  be two finite sets. Let  $\{A_t\}_{t=1}^T$  (channel's input) and  $\{B_t\}_{t=1}^T$  (channel's output) be families of random variables in  $\mathcal{A}$ 



Fig. 2. Model for discrete channel with memory and feedback

and  $\mathcal{B}$  respectively. Moreover, let  $\mathcal{A}^T$  and  $\mathcal{B}^T$  represent their T-fold product spaces. A *channel* is a family of stochastic kernels  $\{p(\beta_t | \alpha^t, \beta^{t-1})\}_{t=1}^T$ . Let  $\mathcal{F}_t$  be the set of all measurable maps  $\varphi_t : \mathcal{B}^{t-1} \to \mathcal{A}$  endowed with a probability

Let  $\mathcal{F}_t$  be the set of all measurable maps  $\varphi_t : \mathcal{B}^{t-1} \to \mathcal{A}$  endowed with a probability distribution, and let  $F_t$  be the corresponding random variable. Let  $\mathcal{F}^T$ ,  $F^T$  denote the Cartesian product on the domain and the random variable, respectively. A *channel code function* is an element  $\varphi^T = (\varphi_1, \ldots, \varphi_T) \in \mathcal{F}^T$ .

Note that, by probability laws,  $p(\varphi^T) = \prod_{t=1}^T p(\varphi_t | \varphi^{t-1})$ . Hence the distribution on  $\mathcal{F}^T$  is uniquely determined by a sequence  $\{p(\varphi_t | \varphi^{t-1})\}_{t=1}^T$ . We will use the notation  $\varphi^t(\beta^{t-1})$  to represent the  $\mathcal{A}$ -valued t-tuple  $(\varphi_1, \varphi_2(\beta^1), \dots, \varphi_t(\beta^{t-1}))$ .

In Information Theory this kind of channels are used to encode and transmit messages. If  $\mathcal{W}$  is a message set of cardinality M with typical element w, endowed with a probability distribution, a *channel code* is a set of M channel code functions  $\varphi^T[w]$ , interpreted as follows: for message w, if at time t the channel feedback is  $\beta^{t-1}$ , then the channel encoder outputs  $\varphi_t[w](\beta^{t-1})$ . A *channel decoder* is a map from  $\mathcal{B}^T$  to  $\mathcal{W}$ which attempts to reconstruct the input message after observing all the output history  $\beta^T$  from the channel.

#### 3.1 Directed information and capacity of channels with feedback

In classical Information Theory, the channel capacity, which is related to the channel's transmission rate by Shannon's fundamental result, can be obtained as the supremum of the mutual information over all possible input's distributions. In presence of feedback, however, this correspondence does not hold anymore. More specifically, mutual information does not represent any longer the information flow from  $\alpha^T$  to  $\beta^T$ . Intuitively, this is due to the fact that mutual information expresses correlation, and therefore it is increased by feedback. But the feedback, i.e the way the output influences the next input, is part of the a priori knowledge, and therefore should not be counted when we measure the output's contribution to the reduction of the uncertainty about the input. If we want to maintain the correspondence with the transmission rate and with information flow, we need to replace mutual information with *directed information* [11].

**Definition 1.** In a channel with feedback, the directed information from input  $A^T$  to output  $B^T$  is defined as  $I(A^T \to B^T) = \sum_{t=1}^T I(\alpha^t; \beta_t | \beta^{t-1})$ . In the other di-

rection, the directed information from  $B^T$  to  $A^T$  is defined as:  $I(B^T \to A^T) = \sum_{t=1}^{T} I(\alpha_t; \beta^{t-1} | \alpha^{t-1}).$ 

Note that the directed information defined above are not symmetric: the flow from  $A^T$  to  $B^T$  takes into account the correlation between  $\alpha^t$  and  $\beta_t$ , while the flow from  $B^T$  to  $A^T$  is based on the correlation between  $\beta^{t-1}$  and  $\alpha_t$ . Intuitively, this is because  $\alpha^t$  influences  $\beta_t$ , but, in the other direction, it is  $\beta^{t-1}$  that influences  $\alpha_t$ . It can be proved [17] that  $I(A^T; B^T) = I(A^T \to B^T) + I(B^T \to A^T)$ . If a channel

It can be proved [17] that  $I(A^T; B^T) = I(A^T \to B^T) + I(B^T \to A^T)$ . If a channel does not have feedback, then  $I(B^T \to A^T) = 0$  and  $I(A^T; B^T) = I(A^T \to B^T)$ .

The following example should help understanding why in a channel with feedback it is the directed information, and not the mutual information, that represents the information transmitted by the channel.

*Example 2.* Consider the discrete memoryless channel with input  $\mathcal{A} = \{a_1, a_2\}$  and output  $\mathcal{B} = \{b_1, b_2\}$  whose matrix is represented in Table 3. Suppose that the channel is used with feedback, in such a way that, for all *t*'s,

$\int a_1$ , if $\beta_t = b_1$		$b_1$	$b_2$
$\alpha_{t+1} = \begin{cases} a_2, \text{ if } \beta_t = b_2 \end{cases}$	$a_1$	1/2	1/2
	$a_2$	1/2	1/2

It is easy to show that if  $t \ge 2$  then  $I(A^t; B^t) \ne 0$ . However, there is no leakage from from  $A^t$  to  $B^t$ , since the rows of the matrix **Ta** are all equal. We have indeed that  $I(A^t \rightarrow B^t) = 0$ , and the mutual **ma** information  $I(A^t; B^t)$  is only due to the feedback information flow **ple**  $I(B^t \rightarrow A^t)$ .

**Table 3.** Channel matrix for Example 2

The concept of capacity is generalized for channels with feedback as follows. Let  $\mathcal{D}_T = \{\{p(\alpha_t | \alpha^{t-1}, \beta^{t-1})\}_{t=1}^T\}$  be the set of all input distributions. For finite T, the capacity of a channel  $\{p(\beta_t | \alpha^t, \beta^{t-1})\}$  is:

$$C_T = \sup_{\mathcal{D}_T} \frac{1}{T} I(A^T \to B^T) \tag{4}$$

#### **4** Interactive systems as channels with memory and feedback

General Interactive Information Hiding Systems (general IIHSs, [1]), are a variant of probabilistic automata in which we indicate explicitly that some action are secret or observable. The attribute "interactive" means that secret and observable actions can interleave and influence each other.

**Definition 2.** A general IIHS is a quadruple  $\mathfrak{I} = (M, \mathcal{A}, \mathcal{B}, \mathcal{L}_{\tau})$ , where M is a probabilistic automaton  $(\mathcal{S}, \mathcal{L}, \hat{s}, \vartheta), \mathcal{L} = \mathcal{A} \cup \mathcal{B} \cup \mathcal{L}_{\tau}$  where  $\mathcal{A}, \mathcal{B}, and \mathcal{L}_{\tau}$  are pairwise disjoint sets of secret, observable, and internal actions respectively, and  $\vartheta(s) \subseteq \mathcal{D}(\mathcal{B} \cup \mathcal{L}_{\tau} \times \mathcal{S})$  implies  $|\vartheta(s)| \leq 1$ , for all s. The condition on  $\vartheta$  ensures that all observable transitions are fully probabilistic.

Assumption In this paper we assume that general IIHSs are *normalized*, i.e. once unfolded, all the transitions between two consecutive levels have either secret labels only,

or observable labels only. Moreover, the occurrences of secret and observable labels alternate between levels. We will call *secret states* the states from which only secrets-labeled transitions are possible, and *observable states* the others. Under this assumption we have that the traces of a computation determine the final state, as expressed by the next proposition. In the following  $trace_A$  and  $trace_B$  indicate the projection of the traces on secret and observable actions, respectively. Given a general IIHS, it is always possible to find an equivalent one that satisfies this assumptions. The interested reader can find in Appendix the formal definition of the transformation.

**Proposition 1.** Let  $\mathfrak{I} = (M, \mathcal{A}, \mathcal{B}, \mathcal{L}_{\tau})$  be a general IIHS. Consider two paths  $\sigma$  and  $\sigma'$ . Then,  $trace_{\mathcal{A}}(\sigma) = trace_{\mathcal{A}}(\sigma')$  and  $trace_{\mathcal{B}}(\sigma) = trace_{\mathcal{B}}(\sigma')$  implies  $\sigma = \sigma'$ .

In the following, we will consider two particular cases: the *fully probabilistic* IIHSs, where there is no nondeterminism, and the *secret -nondeterministic* IIHSs, where each secret choice is fully nondeterministic. The latter will be called simply IIHSs.

**Definition 3.** Let  $\mathfrak{I} = ((\mathcal{S}, \mathcal{L}, \hat{s}, \vartheta), \mathcal{A}, \mathcal{B}, \mathcal{L}_{\tau})$  be a general IIHS. Then  $\mathfrak{I}$  is:

- fully probabilistic if  $\vartheta(s) \subseteq \mathcal{D}(\mathcal{A} \times \mathcal{S})$  implies  $|\vartheta(s)| \leq 1$  for each  $s \in \mathcal{S}$ .
- secret-nondeterministic if  $\vartheta(s) \subseteq \mathcal{D}(\mathcal{A} \times \mathcal{S})$  implies that for each  $s \in \mathcal{S}$  there exist  $s_i$ ' such that  $\vartheta(s) = \{\delta(a_i, s_i)\}_{i=1}^n$ .

We show now how to construct a channel with memory and feedback from IIHSs and fully probabilistic IIHSs. We will see that an IIHS corresponds precisely to a channel as determined by its stochastic kernel, while a fully probabilistic IIHS determines, additionally, the input distribution. In the following, we consider an IIHS  $\mathcal{I} = ((\mathcal{S}, \mathcal{L}, \hat{s}, \vartheta), \mathcal{A}, \mathcal{B}, \mathcal{L}_{\tau})$  is in *normalized form*. Given a path  $\sigma$  of length 2t - 1, we denote  $trace_{\mathcal{A}}(\sigma)$  by  $\alpha^{t}$ , and  $trace_{\mathcal{B}}(\sigma)$  by  $\beta^{t-1}$ .

**Definition 4.** The channel's stochastic kernel corresponding to  $\mathfrak{I}$  is defined as  $p(\beta_t | \alpha^t, \beta^{t-1}) = \vartheta(q)(\beta_t, q')$ , where q is the state reached from the root via the path  $\sigma$  whose input-trace is  $\alpha^t$  and output trace  $\beta^{t-1}$ .

Note that q and q' in previous definitions are well defined: by Proposition 1, q is unique, and since the choice of  $\beta_t$  is fully probabilistic, q' is also unique.

If  $\mathcal{I}$  is fully probabilistic, then it determines also the input distribution and the dependency of  $\alpha_t$  upon  $\beta^{t-1}$  (feedback) and  $\alpha^{t-1}$ .

**Definition 5.** If  $\mathfrak{I}$  is fully probabilistic, the associated channel has a conditional input distribution defined as  $p(\alpha_t | \alpha^{t-1}, \beta^{t-1}) = \vartheta(q)(\alpha_t, q')$ , where q is the state reached from the root via the path  $\sigma$  whose input-trace is  $\alpha^{t-1}$  and output trace is  $\beta^{t-1}$ .

In Proposition 1, q is unique, and since  $\mathcal{I}$  is fully probabilistic, also q' is unique.

#### 4.1 Lifting the channel inputs to reaction functions

Definitions 4 and 5 completely define the joint probabilities  $p(\alpha^t, \beta^t)$  for a fully probabilistic IIHS. However, we still need to show in what sense these define a channel in the information-theoretic sense.

The  $p(\beta_t | \alpha^t, \beta^{t-1})$  determined by the IIHS correspond naturally to a channel's stochastic kernel. The problem resides in the conditional probability of  $p(\alpha_t | \alpha^{t-1}, \beta^{t-1})$ . In an information-theoretic channel, the value of  $\alpha_t$  is determined in the encoder by a deterministic function  $\varphi_t(\beta^{t-1})$ . However, inside the encoder there is no possibility for a probabilistic description of  $\alpha_t$ . Furthermore, in our setting the concept of encoder makes no sense as there is no information to encode. A solution to this problem is to externalize the probabilistic behavior of  $\alpha_t$ : the code functions become simple *reaction functions*  $\varphi_t$  that depend only on  $\beta^{t-1}$  (the message w does not play a role any more), and these reaction functions are endowed with a probability distribution that generates the probabilistic behavior of the values of  $\alpha_t$ .

**Definition 6.** A reactor is a distribution on reaction functions, i.e., a stochastic kernel  $\{p(\varphi_t | \varphi^{t-1})\}_{t=1}^T$ . A reactor R is consistent with a fully probabilistic IIHS  $\mathcal{I}$  if it induces the compatible distribution  $Q(\varphi^T, \alpha^T, \beta^T)$  such that  $Q(\alpha_t | \alpha^{t-1}, \beta^{t-1}) =$  $p(\alpha_t | \alpha^{t-1}, \beta^{t-1})$ , where the latter is the probability distribution induced by  $\mathfrak{I}$ .

The main result of this section states that for any fully probabilistic IIHS there is a reactor that generates the probabilistic behavior of the IIHS.

**Lemma 1.** Let  $\mathcal{X}, \mathcal{Y}$  be finite sets, and let  $\tilde{x} \in \mathcal{X}, \tilde{y} \in \mathcal{Y}$ . Let  $p : \mathcal{X} \times \mathcal{Y} \to [0, 1]$  be a function such that, for every  $x \in \mathcal{X}$ , we have:  $\sum_{y \in \mathcal{Y}} p(x, y) = 1$ . Then:

$$\sum_{\substack{f \in \mathcal{X} \to \mathcal{Y} \\ f(\tilde{x}) = \tilde{y}}} \prod_{x \in \mathcal{X}} p(x, f(x)) = p(\tilde{x}, \tilde{y})$$

**Theorem 3.** Given a fully probabilistic IIHS  $\mathfrak{I}$ , we can construct a channel with memory and feedback, and probability distribution  $Q(\varphi^T, \alpha^T, \beta^T)$ , which corresponds to  $\mathfrak{I}$  in the sense that, for every  $t, \alpha^t$  and  $\beta^t$  we have:

$$Q(\alpha^t, \beta^t) \stackrel{\text{def}}{=} \sum_{\varphi^T} Q(\varphi^T, \alpha^t, \beta^t) = p(\alpha^t, \beta^t)$$

where  $p(\alpha^t, \beta^t)$  is the joint probability of input and output traces induced by  $\mathfrak{I}$ .

The proof of the above theorem (see appendix) shows also how to construct the stochastic kernel  $p(\varphi_t | \varphi^{t-1})_{t=1}^T$  that leads to the compatible distribution  $Q(\varphi^T, \alpha^T, \beta^T)$ .

**Corollary 1.** Let a  $\mathbb{J}$  be a fully probabilistic IIHS. Let  $\{p(\beta_t | \alpha^t, \beta^{t-1})\}_{t=1}^T$  be a sequence of stochastic kernels and  $\{p(\alpha_t | \alpha^{t-1}, \beta^{t-1})\}_{t=1}^T$  a sequence of input distributions defined by  $\mathbb{J}$  according to Definitions 4 and 5. Then the reactor  $R = \{p(\varphi_t | \varphi^{t-1})\}_{t=1}^T$  compatible with respect to the  $\mathbb{J}$  is given by:

$$p(\varphi_1) = p(\alpha_1 | \alpha^0, \beta^0) = p(\alpha_1)$$
(5)

$$p(\varphi_t | \varphi^{t-1}) = \prod_{\beta^{t-1}} p(\varphi_t(\beta^{t-1}) | \varphi^{t-1}(\beta^{t-2}), \beta^{t-1}), \quad 2 \le t \le T$$
(6)

Figure 3 depicts the model for IIHS. Note that, in relation to Figure 2, there are some simplifications: (1) no message w is needed; (2) the decoder is not used. At the beginning, a reaction function sequence  $\varphi^T$  is chosen and then the channel is used T times. At each usage t, the encoder decides the next input symbol  $\alpha_t$  based on the code function  $\varphi_t$  and the output fed back  $\beta^{t-1}$ . Then the channel produces an output  $\beta_t$  based on the stochastic kernel  $\{p(\beta_t | \alpha^t, \beta^{t-1})\}$ . The output is then fed back to the encoder with a delay one.



Fig.3. Channel with memory and feedback model for IIHS

We conclude this section by remarking an intriguing coincidence: The notion of reaction function sequence  $\varphi^T$ , on the IIHSs, corresponds to the notion of deterministic scheduler. In fact, each reaction function  $\varphi_t$  selects the next step,  $\alpha_t$ , on the basis of the  $\beta^{t-1}$  and  $\alpha^{t-1}$  (generated by  $\varphi^{t-1}$ ), and  $\beta^{t-1}$ ,  $\alpha^{t-1}$  represent the path until that state.

# 5 Leakage in Interactive Systems

In this section we propose a notion of information flow based on our model. We follow the idea of defining leakage and maximum leakage using the concepts of mutual information and capacity (see for instance [2]), making the necessary adaptations.

Since the directed information  $I(A^T \to B^T)$  is a measure of how much information flows from  $A^T$  to  $B^T$  in a channel with feedback (cfr. Section 3.1), it is natural to consider it as a measure of leakage of information by the protocol.

**Definition 7.** The information leakage of an IIHS is defined as:  $I(A^T \rightarrow B^T) = \sum_{t=1}^{T} H(A_t|A^{t-1}, B^{t-1}) - H(A^T|B^T).$ 

Note that  $\sum_{t=1}^{T} H(A_t | A^{t-1}, B^{t-1})$  can be seen as the entropy  $H_R$  of reactor R.

Compare this definition with the classical Information-theoretic approach to information leakage: when there is no feedback, the leakage is defined as:

$$I(A^{T}; B^{T}) = H(A^{T}) - H(A^{T}|B^{T})$$
(7)

This definition is based on the principle that the leakage is equal to the difference between the *a priori uncertainty*  $H(A^T)$  (the lack of knowledge about the secret before observing the outcome of the protocol) and the *a posteriori uncertainty*  $H(A^T|B^T)$ (the residual lack of knowledge about the secret after observing the outcome). Our definition maintains the same principle, with the proviso that the a priori uncertainty is now represented by  $H_R$ .

#### 5.1 Maximum leakage as capacity

In the case of secret-nondeterministic IIHS, we have a stochastic kernel but no distribution on the code functions. In this case it seems natural to define as leakage the worst leakage taken over all possible distributions on code functions. This is exactly the concept of capacity.

**Definition 8.** The maximum leakage of an IIHS is defined as the capacity  $C_T$  of the associated channel with memory and feedback.

## 6 Modeling IIHSs as channels: An example

In this section we show the application of our approach to the *Cocaine Auction Protocol* [15]. Let us imagine a situation where several mob individuals are gathered around a table. An auction is about to be held in which one of them offers his next shipment of cocaine to the highest bidder. The seller describes the merchandise and proposes a starting price. The others then bid increasing amounts until there are no bids for 30 consecutive seconds. At that point the seller declares the auction closed and arranges a secret appointment with the winner to deliver the goods.

The basic protocol is fairly simple and is organized as a succession of rounds of bidding. Round *i* starts with the seller announcing the bid price  $b_i$  for that round. Buyers have *t* seconds to make an offer (i.e. to say yes, meaning "I'm willing to buy at the current bid price  $b_i$ "). As soon as one buyer anonymously says yes, he becomes the winner  $w_i$  of that round and a new round begins. If nobody says anything for *t* seconds, round *i* is concluded by timeout and the auction is won by the winner  $w_{i-1}$  of the previous round, if one exists. If the timeout occurs during round 0, this means that nobody made any offers at the initial price  $b_0$ , so there is no sale.

We are going to instantiate this example in a table where there are two potential buyers, *Candlemaker* and *Scarface*. We assume that the initial bid is always 1 dollar, so the first bid does not need to be announced by the seller. In each turn the seller can choose how much he wants to increase the actual bid. This is done by adding an increment to the last bid. There are two options of increments, namely *inc*+ and *inc*++ (the latter is the biggest one). In that way,  $b_{i+1}$  is either  $b_i + inc$ + or  $b_i + inc$ ++. We can describe this protocol as a *normalized* IIHS  $\mathcal{I} = (M, \mathcal{A}, \mathcal{B}, \mathcal{L}_{\tau})$ , where  $\mathcal{A} =$ {Candlemaker, Scarface,  $a_*$ } is the set of secret actions,  $\mathcal{B} = \{inc+, inc++, b_*\}$  is the set of observable actions,  $\mathcal{L}_{\tau} = \emptyset$  is the set of hidden actions, and the probabilistic automaton M is represented in Figure 4. For clarity reasons, we omit transitions with probability 0 in the automaton. Note that the special secret action  $a_*$  represents the situation where neither *Candlemaker* nor *Scarface* bid. The special observable action  $b_*$  is only possible after no one has bidded, and signalizes the end of the auction and, therefore, no bid is allowed anymore.

Table 4 shows all the stochastic kernels for this example. In Appendix C, we show also how to compute the reactors and the leakage.



Fig.4. Cocaine Auction example

				$\alpha_1, \beta_1, \alpha_2 \to \beta_2$	Cneap	Expensive	0*
				Candlemaker,inc+,Candlemaker	$q_{22}$	$q_{23}$	0
				Candlemaker,inc+,Scarface	$q_{24}$	$q_{25}$	0
				Candlemaker, inc+, $a_*$	0	0	1
				Candlemaker,inc++,Candlemaker	$q_{27}$	$q_{28}$	0
$\alpha_1 \rightarrow \beta_1$	inc+	inc++	$b_{*}$	Candlemaker,inc++,Scarface	$q_{29}$	$q_{30}$	0
Candlomakor	<i>a</i> .		• *	Candlemaker, inc++, $a_*$	0	0	1
Saarfaaa	$q_4$	$q_5$	0	Scarface,inc+,Candlemaker	$q_{32}$	$q_{33}$	0
scurjuce	$q_6$	$q_7$	1	Scarface,inc+,Scarface	$q_{34}$	$q_{35}$	0
$a_*$	0	0	1	Scarface, inc+, $a_*$	0	0	1
(a) $t = 1$	$1 m(\beta   \alpha)$	$1 a^{0}$		Scarface,inc++,Candlemaker	$q_{37}$	$q_{38}$	0
(a) $\iota = 1$	$p(p_1 \alpha)$	$, \rho$ )		Scarface,inc++,Scarface	$q_{39}$	$q_{40}$	0
				Scarface,inc++,a <sub>*</sub>	0	0	1
				$a_*, b_*, a_*$	0	0	1
				All other lines	0	0	1

(b) 
$$t = 2, p(\beta_2 | \alpha^2, \beta^1)$$

**Table 4.** Stochastic kernels for the Cocaine Auction example.

# 7 Topological properties of IIHSs and their Capacity

In this section we show how to extend to IIHSs the notion of pseudometric defined in [6] for Concurrent Labelled Markov Chains, and we prove that the capacity of the corresponding channels is a continuous function on this pseudometric. The metric construction is sound for general IIHSs, but the result on capacity is only valid for secretnondeterministic IIHSs.

Given a set of states S, a pseudometric (or distance) is a function d that yields a non-negative real number for each pair of states and satisfies the following: d(s,s) = 0; d(s,t) = d(t,s), and  $d(s,t) \le d(s,u) + d(u,t)$ . We say that a pseudometric d is c-bounded if  $\forall s, t : d(s,t) \le c$ , where c is a positive real number.

Note that, in contrast to metrics, in pseudometrics two elements can have distance 0 without being identical. The reason for considering pseudometrics instead than metrics is because the purpose is to extend the notion of (probabilistic) bisimulation: having distance 0 will correspond to being bisimilar.

We now define a complete lattice on pseudometrics, in order to define the distance between IIHSs as the greatest fixpoint of a particular transformation, in line with the coinductive theory of bisimilarity. Since larger bisimulations identify more, the natural extension of the ordering to metrics must shorten the distances as we go up in the lattice:

**Definition 9.**  $\mathcal{M}$  is the class of 1-bounded pseudometrics on states with the ordering

$$d \leq d'$$
 if  $\forall s, s' \in S : d(s, s') \geq d'(s, s')$ 

It is easy to see that  $(\mathcal{M}, \preceq)$  is a complete lattice. In order to define pseudometrics on IIHSs, we now need to lift the pseudometrics on states to pseudometrics on distributions in  $\mathcal{D}(\mathcal{L} \times S)$ . Following standard lines [18,6,5], we apply the construction based on the Kantorovich metric [9].

**Definition 10.** For  $d \in \mathcal{M}$ , and  $\mu, \mu' \in \mathcal{D}(\mathcal{L} \times S)$ , we define  $d(\mu, \mu')$  (overloading the notation d) as  $d(\mu, \mu') = \max \sum_{(\ell_i, s_i) \in \mathcal{L} \times S} (\mu(\ell_i, s_i) - \mu'(\ell_i, s_i)) x_i$  where the maximization is on all possible values of the  $x_i$ 's, subject to the constraints  $0 \le x_i \le 1$  and  $x_i - x_j \le \hat{d}((\ell_i, s_i), (\ell_j, s_j))$ , where

$$\hat{d}((\ell_i, s_i), (\ell_j, s_j)) = \begin{cases} 1 & \text{if } \ell_i \neq \ell_j \\ d(s_i, s_j) & \text{otherwise} \end{cases}$$

It can be shown that with this definition m is a pseudometric on  $\mathcal{D}(\mathcal{L} \times S)$ .

**Definition 11.**  $d \in \mathcal{M}$  is a bisimulation metric *if*, for all  $\epsilon \in [0, 1)$ ,  $d(s, s') \leq \epsilon$  implies that if  $s \to \mu$ , then there exists some  $\mu'$  such that  $s' \to \mu'$  and  $d(\mu, \mu') \leq \epsilon$ .

Note that it is not necessary to require the converse of the condition in Definition 11 to get a complete analogy with bisimulation: the converse is indeed implied by the symmetry of d as a pseudometric. Note also that we prohibit  $\epsilon$  to be 1 because throughout this paper 1 represents the maximum distance, which includes the case where one state may perform a transition and the other may not.

The greatest bisimulation metric is  $d_{max} = \bigsqcup \{ d \in \mathcal{M} \mid d \text{ is a bisimulation metric} \}$ . We now characterize  $d_{max}$  as a fixed point of a monotonic function  $\Phi$  on  $\mathcal{M}$ . Eventually we are interested in the distance between IIHSs, and for the sake of simplicity, from now on we consider only the distance between states belonging to different IIHSs. The extension to the general case is trivial. For clarity purposes, we assume that different IIHSs have disjoint sets of states.

**Definition 12.** Given two IIHSs with transition relations  $\theta$  and  $\theta'$  respectively, and a preudometric d on states, define  $\Phi : \mathcal{M} \to \mathcal{M}$  as:

$$\Phi(d)(s,s') = \begin{cases} \max_i d(s_i,s'_i) \text{ if } \quad \vartheta(s) = \{\delta_{(a_1,s_1)}, \dots, \delta_{(a_m,s_m)}\} \\ \text{ and } \vartheta'(s') = \{\delta_{(a_1,s'_1)}, \dots, \delta_{(a_m,s'_m)}\} \\ d(\mu,\mu') \quad \text{ if } \vartheta(s) = \{\mu\} \text{ and } \vartheta'(s') = \{\mu'\} \\ 0 \quad \text{ if } \vartheta(s) = \vartheta'(s') = \emptyset \\ 1 \quad \text{ otherwise} \end{cases}$$

It is easy to see that the definition of  $\Phi$  is a particular case of the function F defined in [6,5]. Hence it can be proved, by adapting the proofs of the analogous results in [6, 5], that F(d) is a pseudometric, and that the following property holds.

**Lemma 2.** For  $\epsilon \in [0, 1)$ ,  $\Phi(d)(s, s') \leq \epsilon$  holds if and only if whenever  $s \to \mu$ , there exists some  $\mu'$  such that  $s' \to \mu'$  and  $d(\mu, \mu') \leq \epsilon$ .

**Corollary 2.** *d* is a bisimulation metric iff  $d \leq \Phi(d)$ .

As a consequence of Corollary 2, we have that  $d_{max} = \bigsqcup \{ d \in \mathcal{M} \mid d \preceq \Phi(d) \}$ , and still as a particular case of F in [6,5], we have that  $\Phi$  is monotonic on  $\mathcal{M}$ .

We can now apply Tarski's fixed point theorem, which ensures that  $d_{max}$  is the greatest fixed point of  $\Phi$ . Furthermore, by Corollary 2 we know that  $d_{max}$  is indeed a bisimulation metric, and that it is the greatest bisimulation metric. In addition, the finite branchingness of IIHSs ensures that the closure ordinal of  $\Phi$  is  $\omega$  (cf. Lemma 3.10 in the full version of [6], available on the authors' web pages). Therefore one can proceed in a standard way to show that  $d_{max} = \prod {\Phi^i(\top) \mid i \in \mathbb{N}}$ , where  $\top$  is the greatest pseudometric (i.e.  $\top(s, s') = 0$  for every s, s'), and  $\Phi^0(\top) = \top$ .

Given two IIHSs  $\mathfrak{I}$  and  $\mathfrak{I}'$ , with initial states s and s' respectively, we define the distance between  $\mathfrak{I}$  and  $\mathfrak{I}'$  as  $d(\mathfrak{I}, \mathfrak{I}') = d_{max}(s, s')$ . The following properties are auxiliary to the theorem which states the continuity of the capacity.

**Lemma 3.** Consider two IIHSs  $\mathfrak{I}$  and  $\mathfrak{I}'$  with transition functions  $\vartheta$  and  $\vartheta'$  respectively. Given  $t \geq 2$  and two sequences  $\alpha^t$  and  $\beta^t$ , assume that both  $\mathfrak{I}(\alpha^{t-1}, \beta^{t-1})$  and  $\mathfrak{I}'(\alpha^{t-1}, \beta^{t-1})$  are defined, that  $d_{max}(\mathfrak{I}(\alpha^{t-1}, \beta^{t-1}), \mathfrak{I}'(\alpha^{t-1}, \beta^{t-1})) < p(\beta_t \mid \alpha^t, \beta^{t-1})$ , and  $\vartheta(\mathfrak{I}(\alpha^t, \beta^{t-1})) \neq \emptyset$ . Then:

1.  $\vartheta'(\mathfrak{I}'(\alpha^t, \beta^{t-1})) \neq \emptyset$  holds as well,

2.  $\mathfrak{I}(\alpha^t, \beta^t)$  and  $\mathfrak{I}'(\alpha^t, \beta^t)$  are both defined,  $p(\beta_t \mid \alpha^t, \beta^{t-1}) > 0$ , and

$$d_{max}(\mathfrak{I}(\alpha^{t},\beta^{t}),\mathfrak{I}'(\alpha^{t},\beta^{t})) \leq \frac{d_{max}(\mathfrak{I}(\alpha^{t-1},\beta^{t-1}),\mathfrak{I}'(\alpha^{t-1},\beta^{t-1}))}{p(\beta_{t} \mid \alpha^{t},\beta^{t-1}).}$$

**Lemma 4.** Consider two IIHSs  $\mathfrak{I}$  and  $\mathfrak{I}'$ , and let  $p(\cdot | \cdot, \cdot)$  and  $p'(\cdot | \cdot, \cdot)$  be their distributions on the output nodes. Given T > 0, and two sequences  $\alpha^T$  and  $\beta^T$ , assume that  $p(\beta_t | \alpha^t, \beta^{t-1}) > 0$  for every t < T. Let  $m = \min_{1 \le t < T} p(\beta_t | \alpha^t, \beta^{t-1})$  and let  $\epsilon \in (0, m^{T-1})$ . Assume  $d(\mathfrak{I}, \mathfrak{I}') < \epsilon$ . Then, for every  $t \le T$ , we have

$$p(\beta_t \mid \alpha^t, \beta^{t-1}) - p'(\beta_t \mid \alpha^t, \beta^{t-1}) < \frac{\epsilon}{m^{T-1}}.$$

The main contribution of this section, stated in next theorem, is the continuity of the capacity w.r.t. the metric on IIHSs. For this theorem, we assume that the IIHSs are normalized. Furthermore, it is crucial that they are secret-nondeterministic (while the definition of the metric holds in general).

**Theorem 4.** Consider two normalized IIHSs  $\mathfrak{I}$  and  $\mathfrak{I}'$ , and fix a T > 0. For every  $\epsilon > 0$  there exists  $\nu > 0$  such that if  $d(\mathfrak{I},\mathfrak{I}') < \nu$  then  $|C_T(\mathfrak{I}) - C_T(\mathfrak{I}')| < \epsilon$ .

We conclude this section with an example showing that the continuity result for the capacity does not hold if the construction of the channel is done starting from a system in which the secrets are endowed with a probability distribution. This is also the reason why we could not simply adopt the proof technique of the continuity result in [6] and we had to come up with a different reasoning.

*Example 3.* Consider the two following programs, where  $a_1, a_2$  are secrets,  $b_1, b_2$  are observable,  $\parallel$  is the parallel operator, and  $+_p$  is a binary probabilistic choice that assigns probability p to the left branch, and probability 1 - p to the right one.

s)  $(send(a_1) +_p send(a_2)) \parallel receive(x).output(b_2)$ t)  $(send(a_1) +_q send(a_2)) \parallel receive(x).if \ x = a_1 \ then \ output(b_1) \ else \ output(b_2).$ 

Table 5 shows the fully probabilistic IIHSs corresponding to these programs, and their associated channels, which in this case (since the secret actions are all at the toplevel) are classic channels, i.e. memoryless and without feedback. As usual for classic channels, they do not depend on p and q. It is easy to see that the capacity of the first channel is 0 and the capacity of the second one is 1. Hence their difference is 1, independently from p and q.

Let now p = 0 and  $q = \epsilon$ . It is easy to see that the distance between s and t is  $\epsilon$ . Therefore (when the automata have probabilities on the secrets), the capacity is not a continuous function of the distance.



Table 5. The IIHSs of Example 3 and their corresponding channels

#### 8 **Conclusion and future work**

In this paper we have investigated the problem of information leakage in interactive systems, and we have proved that these systems can be modeled as channels with memory and feedback.

IIHSs <b>as automata</b>	IIHSs <b>as channels</b>	Notion of leakage
Normalized IIHSs with nondeterministic	Sequence of stochastic kernels	Leakage as capacity
inputs and probabilistic outputs	$\{p(\beta_t   \alpha^t, \beta_{t-1})\}_{t=1}^T$	
Normalized IIHSs with a deterministic	Sequence of stochastic kernels	
scheduler solving the nondeterminism	$\{p(\beta_t   \alpha^t, \beta_{t-1})\}_{t=1}^T + $	
	reaction function seq. $\varphi^T$	
Fully probabilistic normalized IIHSs	Sequence of stochastic kernels	Leakage as directed
	$\{p(\beta_t   \alpha^t, \beta^{t-1})\}_{t=1}^T +$	information $I(A^T \to B^T)$
	reactor $\{p(\varphi_t   \varphi^{t-1})\}_{t=1}^T$	
	Tabla 6	

Table	e 6.
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Classical channels	Channels with memory and feedback
The protocol is modeled in independent uses of	The protocol is modeled in several
the channel, often a unique use.	consecutive uses of the channel.
The channel is from $\mathcal{A}^T \to \mathcal{B}^T$ , i.e., its input	The channel is from $\mathcal{F} \to \mathcal{B}$ , i.e. its
is a single string $\alpha^T = \alpha_1 \dots \alpha_T$ of secret	input is a reaction function $\varphi_t$ and its
symbols and its output is a single string $\beta^T =$	output is an observable $\beta_t$ .
$\beta_1 \dots \beta_T$ . of observable symbols.	
The channel is memoryless and in general	The channel has memory. Despite the fact that the
implicitly it is assumed the absence of	channel from $\mathcal{F}  ightarrow \mathcal{B}$ does not have
feedback.	feedback, the internal stochastic kernels
	do.
The capacity is calculated using information	The capacity is calculated using mutual
$I(\alpha^T; \beta^T).$	directed information $I(\alpha^T \to \beta^T)$ .

Table	7	
	-	-

The situation is summarized in Table 6. The comparison with the classical situation of non-interactive systems is represented in Table 7. Furthermore, we have proved that the channel capacity is a continuous function of the kantorovich metric.

For future work we would like to provide algorithms to compute the leakage and maximum leakage of interactive systems. These problems result very challenging given the exponential growth of reaction functions (needed to compute the leakage) and the quantification over infinitely many reactors (given by the definition of maximum leakage in terms of capacity). One possible solution is to study the relation between deterministic schedulers and sequence of reaction functions. In particular, we believe that for each sequence of reaction functions and distribution over it there exists a probabilistic scheduler for the automata representation of the secret-nondeterministic IIHS. In this way, the problem of computing the leakage and maximum leakage would reduce to a standard probabilistic model checking problem (where the challenge is to compute probabilities ranging over infinitely many schedulers).

In addition, we plan to investigate measures of leakage for interactive systems other than mutual information and capacity.

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# Appendix

### A: Normalization of IIHS trees

In this section we will consider the problem of *normalizing* an IIHS in such a way it is compatible with the assumptions made along the paper. The process of normalization described bellow is general enough to be applied to any IIHS without loss of generality or expression power.

Consider a general IIHS  $\mathfrak{I} = (M, \mathcal{A}, \mathcal{B}, \mathcal{L}_{\tau})$ , with  $M = (Q, \mathcal{L}, \hat{s}, \vartheta)$ , where  $\mathcal{L} = \mathcal{A} \cup \mathcal{B} \cup \mathcal{L}_{\tau}$  representing a protocol. Let us consider that we are interested only in a finite execution of the protocol, so the automaton tree is already unfolded up to a certain level in such a way that the longest input trace is  $\alpha^{T'}$  and the longest output trace is  $\beta^{T''}$ .

Let us now present some conventions that we will need for the normalization process. We will introduce a new input symbol  $a_* \notin \mathcal{A}$  to represent the absence of an input symbol and, in the same way, a new symbol  $b_* \notin \mathcal{B}$  to represent the absence of an output symbol. We will extend the input and output alphabets in such a way that  $\mathcal{A}' = \mathcal{A} \cup \{a_*\}$  and  $\mathcal{B}' = \mathcal{B} \cup \{b_*\}$ .

Let us define  $T = \max(T', T'')$ , i.e., the maximum length of any input or output trace in the unfolded tree of the automaton. The function Labels $(\mathfrak{I}, t)$  : IIHS  $\times$  $\{1, \ldots, T\} \rightarrow \wp(\mathcal{L})$  from an IIHS  $\mathfrak{I}$  and a given level  $1 \leq t \leq T$  of its unfolded tree to the set  $\mathcal{L}$  of input symbols, output symbols and unobservable symbols of  $\mathfrak{I}$ . Informally, Labels $(\mathfrak{I}, t)$  is the set of all labels that can be executed from the  $t^{th}$  level of the automaton of  $\mathfrak{I}$ .

It is possible to construct an equivalent IIHS  $\mathcal{I}' = \mathcal{I} = (M', \mathcal{A}', \mathcal{B}', \mathcal{L}_{\tau})$ , where  $M' = (Q', \mathcal{L}', \hat{s}', \vartheta')$  such that  $\mathcal{L}' = \mathcal{A}' \cup \mathcal{B}' \cup \mathcal{L}_{\tau}$  and its unfolded tree up to depth 2T respects, for every  $1 \leq t \leq T$ :

- 1. Labels $(\mathfrak{I}',t) \cap \mathcal{A}' = \emptyset$  or Labels $(\mathfrak{I}',t) \cap \mathcal{B}' = \emptyset$ ;
- 2.  $\mathcal{A}' \subseteq \texttt{Labels}(\mathcal{I}', t)$  or  $\mathcal{B}' \subseteq \texttt{Labels}(\mathcal{I}', t)$ ;
- 3.  $\mathcal{A}' \subseteq \texttt{Labels}(\mathcal{I}', t)$  iff  $\mathcal{B}' \subseteq \texttt{Labels}(\mathcal{I}', t+1)$ , where we consider the arithmetic on t modulo 2T;
- 4.  $\mathcal{A}' \subseteq \texttt{Labels}(\mathcal{I}', 1);$
- 5.  $trace_{\mathcal{A}'}(\sigma) = trace_{\mathcal{B}'}(\sigma) = T$ , for all path  $\sigma$  in the unfolded tree

Condition 1 states that each level can admit input actions or output actions, but not both. Condition 2 states that all input actions must be listed in an input level, and the same for output levels (as we will see soon, even if we need to associate probability zero to an action). Condition 3 states that input and output levels must necessarily alternate. Condition 4 assures that we always start with an input level. Cond 5 assures that all the leaves of the unfolded tree are in the same level, i.e., the tree is *balanced*.

The proof is straightforward, but we shall give an intuition of it. First, the new symbols  $a_*$  and  $b_*$  are place holders for the absence of an input and output symbol, respectively. Now, if in a given level t we want to have only input symbols, we can postpone output symbols by adding  $a_*$  to the level and "moving" all the output symbols to the subtree of  $a_*$ . Figure 5 exemplifies the local transformations we desire in a tree.



Fig. 5. Local transformation on an IIHS tree

Note that in 5(b) the introduction of new nodes changed the probabilities. In general, if we are in an input level, we need to introduce  $a_*$  to postpone the output symbols, and the probabilities change as follows:

1. 
$$p'_{a_i} = p_{a_i}$$
  
2.  $p_{a_*} = \sum_{i=0}^{m} p_{b_i}$   
3.  $p'_{b_i} = \frac{p_{b_i}}{p_{b_*}} = \frac{p_{b_i}}{\sum_{i=0}^{m} p_{b_i}}$ 

If a node does not have descendants, we complete the tree by adding all the possible actions in A with probability 0, and the action  $a_*$  with probability 1.

If we are in an output level, the same rules apply, guarding the proper symmetry between input and outputs. Figure 6 shows an example of a full transformation on a tree (for the sake of readability, we omit the levels where only  $a_* = 1$  or  $b_* = 1$ ).



(a) Tree before transformation

(b) Tree after transformation

Fig. 6. Transformation on an IIHS tree

# **B:** Proofs of the main results

In this section we prove the results stated in the paper. We start by showing a transformation from channels with feedback to channels without feedback, that will be useful for the proofs. As showed in [17], the original channel with feedback from input symbols  $\mathcal{A}^T$  to output symbols  $\mathcal{B}^T$  can be lifted to an equivalent channel without feedback from code functions  $\mathcal{F}^T$  to output symbols  $\mathcal{B}^T$ . This transformation allows us to calculate the channel capacity. Let  $\{p(\varphi_t|\varphi^{t-1})\}_{t=1}^T$  be a sequence of code function stochastic kernels and let  $\{p(\beta_t|\alpha^t,\beta^{t-1})\}_{t=1}^T$  be a channel with memory and feedback. The channel from  $F^T$  to  $B^T$  is constructed using a joint measure  $Q(\varphi^T, \alpha^T, \beta^T)$  that respects the following constraints:

**Definition 13.** A measure  $Q(\varphi^T, \alpha^T, \beta^T)$  is said to be consistent with respect to the code function stochastic kernels  $\{p(\varphi_t|\varphi^{t-1})\}_{t=1}^T$  and the channel  $\{p(\beta_t|\alpha^t, \beta^{t-1})\}_{t=1}^T$  if, for each t:

- 1. There is no feedback to the code functions:  $Q(\varphi_t | \varphi^{t-1}, \alpha^{t-1}, \beta^{t-1}) = p(\varphi_t | \varphi^{t-1}).$
- 2. The input is a function of the past outputs:  $Q(\alpha_t | \varphi^t, \alpha^{t-1}, \beta^{t-1}) = \delta_{\{\varphi_t(\beta^{t-1})\}}(\alpha_t)$ where  $\delta$  is the Dirac measure.
- 3. The properties of the underlying channel are preserved:

$$Q(\beta_t | F^t = \varphi^t, A^t = \alpha^t, B^{t-1} = \beta^{t-1}) = p(\beta_t | \alpha^t, \beta^{t-1})$$

The following result states that there is only one consistent measure  $Q(\varphi^T, \alpha^T, \beta^T)$ :

**Theorem 5** ([17]). Given  $\{p(\varphi_t | \varphi^{t-1})\}_{t=1}^T$  and a channel  $\{p(\beta_t | \alpha^t, \beta^{t-1})\}_{t=1}^T$ , there exists only one consistent measure  $Q(\varphi^T, \alpha^T, \beta^T)$ . Furthermore the channel from  $\mathcal{F}^T$  to  $\mathcal{B}^T$  is given by:

$$Q(\beta_t | \varphi^t, \beta^{t-1}) = p(\beta_t | \varphi^t(\beta^{t-1}), \beta^{t-1})$$
(8)

**Lemma 5** (Lemma 1 in the paper). Let  $\mathcal{X}, \mathcal{Y}$  be finite sets, and let  $\tilde{x} \in \mathcal{X}, \tilde{y} \in \mathcal{Y}$ . Let  $p : \mathcal{X} \times \mathcal{Y} \to [0, 1]$  be a function such that, for every  $x \in \mathcal{X}$ , we have:  $\sum_{y \in \mathcal{Y}} p(x, y) = 1$ . Then:

$$\sum_{\substack{f \in \mathcal{X} \to \mathcal{Y} \\ f(\tilde{x}) = \tilde{y}}} \prod_{x \in \mathcal{X}} p(x, f(x)) = p(\tilde{x}, \tilde{y})$$

*Proof.* By induction on the number of elements of  $\mathcal{X}$ .

**Base case:**  $X = {\tilde{x}}$ . In this case:

$$\sum_{\substack{f \in \mathcal{X} \to \mathcal{Y} \\ f(\tilde{x}) = \tilde{y}}} \prod_{x \in \mathcal{X}} p(x, f(x)) = p(\tilde{x}, f(\tilde{x})) = p(\tilde{x}, \tilde{y})$$

**Inductive case:** Let  $\mathcal{X} = \mathcal{X}' \cup \{\hat{x}\}$ , with  $\hat{x} \neq \tilde{x}$ , and  $\tilde{x} \in \mathcal{X}$ . Then:

$$\sum_{\substack{f \in \mathcal{X}' \cup \{\mathring{x}\} \to \mathcal{Y} \\ f(\widehat{x}) = \widetilde{y}}} \prod_{x \in \mathcal{X}' \cup \{\mathring{x}\}} p(x, f(x))$$

= (by distributivity)

$$\left(\sum_{\substack{f\in\mathcal{X}'\rightarrow\mathcal{Y}\\f(\tilde{x})=\tilde{y}}}\prod_{x\in\mathcal{X}'}p(x,f(x))\right)\cdot\sum_{g\in\{\hat{x}\}\rightarrow\mathcal{Y}}p(\mathring{x},g(\mathring{x}))$$

(by the assumption)

$$\sum_{\substack{f \in \mathcal{X}' \to \mathcal{Y} \\ f(\tilde{x}) = \tilde{y}}} \prod_{x \in \mathcal{X}'} p(x, f(x))$$

= (by the induction hypothesis)

 $p(\tilde{x}, \tilde{y})$ 

=

**Theorem 6** (Theorem 3 in the paper). Given a fully probabilistic IIHS  $\mathcal{I}$ , we can construct a channel with memory and feedback, and probability distribution  $Q(\varphi^T, \alpha^T, \beta^T)$ , which corresponds to  $\mathcal{I}$  in the sense that, for every  $t, \alpha^t$  and  $\beta^t$  we have:

$$Q(\boldsymbol{\alpha}^t,\boldsymbol{\beta}^t) \stackrel{\text{def}}{=} \sum_{\boldsymbol{\varphi}^T} Q(\boldsymbol{\varphi}^T,\boldsymbol{\alpha}^t,\boldsymbol{\beta}^t) = p(\boldsymbol{\alpha}^t,\boldsymbol{\beta}^t)$$

where  $p(\alpha^t, \beta^t)$  is the joint probability of input and output traces induced by  $\mathfrak{I}$ .

*Proof.* First of all we note that, by probability laws,  $Q(\alpha^t, \beta^t) = \sum_{\varphi^t} Q(\varphi^t, \alpha^t, \beta^t)$ . So we need to show that  $\sum_{\varphi^t} Q(\varphi^t, \alpha^t, \beta^t) = p(\alpha^t, \beta^t)$  by induction on t. Base case: t=1. Let us define  $Q(\varphi_1|\epsilon)=p(\varphi_1(\epsilon))$  and  $Q(\beta_1|\alpha^1,\epsilon)=p(\beta_1|\alpha_1).$  Then:

$$\begin{split} \sum_{\varphi^1} Q(\varphi^1, \alpha^1, \beta^1) &= \sum_{\varphi_1} Q(\varphi_1, \alpha_1, \beta_1) \\ &= \sum_{\varphi_1} Q(\varphi_1 | \epsilon, \epsilon, \epsilon) Q(\alpha_1 | \varphi_1, \epsilon, \epsilon) Q(\beta_1 | \varphi_1, \alpha_1, \epsilon) \text{ (by the chain rule)} \\ &= \sum_{\varphi_1} Q(\varphi_1 | \epsilon) \delta_{\{\varphi_1(\epsilon)\}}(\alpha_1) Q(\beta_1 | \alpha^1, \epsilon) \qquad \text{(by Definition 13)} \\ &= \sum_{\varphi_1} p(\varphi_1(\epsilon)) \delta_{\{\varphi_1(\epsilon)\}}(\alpha_1) p(\beta_1 | \alpha_1) \\ &= p(\alpha_1) p(\beta_1 | \alpha_1) \qquad \text{(by definition of } \delta) \\ &= p(\alpha_1, \beta_1) \\ &= p(\alpha^1, \beta^1) \end{split}$$

Inductive case: Let us define  $Q(\beta_t|\alpha^t,\beta^{t-1})=p(\beta_t|\alpha^t,\beta^{t-1}),$  and

$$Q(\varphi_t|\varphi^{t-1}) = \prod_{\beta^{t-1}} p(\varphi_t(\beta^{t-1})|\varphi^{t-1}(\beta^{t-2}), \beta^{t-1})$$

Note that, if we consider  $\mathcal{X} = \{\beta^{t-1} \mid \beta_i \in \mathcal{B}, 1 \leq i \leq t-1\}, \mathcal{Y} = \mathcal{A}$ , and  $p(\beta^{t-1}, \alpha_t) = p(\alpha_t | \varphi^{t-1}(\beta^{t-2}), \beta^{t-1})$ , then  $\mathcal{X}, \mathcal{Y}$  and p satisfy the hypothesis of Lemma 1.

Then:  

$$\begin{split} &\sum_{q^{i}} Q(\varphi^{t}, \alpha^{t}, \beta^{t}) \\ &= \text{ (by the chain Rule)} \\ &\sum_{q^{i}} Q(\varphi^{t-1}, \alpha^{t-1}, \beta^{t-1})Q(\varphi_{t}|\varphi^{t-1}, \alpha^{t-1}, \beta^{t-1})Q(\alpha_{t}|\varphi^{t}, \alpha^{t-1}, \beta^{t-1})Q(\beta_{t}|\varphi^{t}, \alpha^{t}, \beta^{t-1}) \\ &= \text{ (by Definition 13)} \\ &\sum_{q^{i}} Q(\varphi^{t-1}, \alpha^{t-1}, \beta^{t-1})Q(\varphi_{t}|\varphi^{t-1}), \delta_{\{\varphi_{t}(\beta^{t-1})\}}(\alpha_{t})Q(\beta_{t}|\alpha^{t}, \beta^{t-1}) \\ &= \text{ (by construction of } Q) \\ &\sum_{q^{i}} Q(\varphi^{t-1}, \alpha^{t-1}, \beta^{t-1}) \left( \prod_{\beta'^{t-1}} p(\varphi_{t}(\beta'^{t-1})|\varphi^{t-1}(\beta'^{t-2}), \beta'^{t-1}) \right) \delta_{\{\varphi_{t}(\beta^{t-1})\}}(\alpha_{t})p(\beta_{t}|\alpha^{t}, \beta^{t-1}) \\ &= \text{ (by definition of } \delta) \\ &\sum_{q^{i}(\beta^{t-1})=\alpha_{t}} Q(\varphi^{t-1}, \alpha^{t-1}, \beta^{t-1}) \left( \prod_{\beta'^{t-1}} p(\varphi_{t}(\beta'^{t-1})|\varphi^{t-1}(\beta'^{t-2}), \beta'^{t-1}) \right) p(\beta_{t}|\alpha^{t}, \beta^{t-1}) \\ &= \text{ (by Lemma 1)} \\ &\sum_{q^{i}t-1} Q(\varphi^{t-1}, \alpha^{t-1}, \beta^{t-1}) \cdot p(\beta_{t}|\alpha^{t}, \beta^{t-1}) \cdot \sum_{\varphi^{i}(\beta^{t-1})=\alpha_{t}} \prod_{\beta'^{i-1}} p(\varphi_{t}(\beta'^{t-1}))|\varphi^{t-1}(\beta'^{t-2}), \beta'^{t-1}) \\ &= \text{ (by Lemma 1)} \\ &\sum_{q^{i}t-1} Q(\varphi^{t-1}, \alpha^{t-1}, \beta^{t-1}) \cdot p(\beta_{t}|\alpha^{t}, \beta^{t-1}) \cdot p(\alpha_{t}|\alpha^{t-1}, \beta^{t-1}) \\ &= \text{ (by Lemma 1)} \\ &= p(\beta_{t}|\alpha^{t}, \beta^{t-1}) \cdot p(\alpha_{t}|\alpha^{t-1}, \beta^{t-1}) \cdot \sum_{\varphi^{i}=1} Q(\varphi^{t-1}, \alpha^{t-1}, \beta^{t-1}) \\ &= \text{ (by induction hypothesis)} \\ p(\beta_{t}|\alpha^{t}, \beta^{t-1}) \cdot p(\alpha_{t}|\alpha^{t-1}, \beta^{t-1}) \cdot p(\alpha^{t-1}, \beta^{t-1}) \\ &= \text{ (by the chain rule)} \\ p(\alpha^{t}, \beta^{t}) \\ \\ &= \text{ (by the chain rule)} \\ \end{array}$$

**Lemma 6** (Lemma 3 in the paper). Consider two IIHSs  $\mathfrak{I}$  and  $\mathfrak{I}'$  with transition functions  $\vartheta$  and  $\vartheta'$  respectively. Given  $t \geq 2$  and two sequences  $\alpha^t$  and  $\beta^t$ , assume that both  $\mathfrak{I}(\alpha^{t-1},\beta^{t-1})$  and  $\mathfrak{I}'(\alpha^{t-1},\beta^{t-1})$  are defined, that  $d_{max}(\mathfrak{I}(\alpha^{t-1},\beta^{t-1}),\mathfrak{I}'(\alpha^{t-1},\beta^{t-1})) < p(\beta_t \mid \alpha^t,\beta^{t-1})$ , and  $\vartheta(\mathfrak{I}(\alpha^t,\beta^{t-1})) \neq \emptyset$ . Then:

*I.*  $\vartheta'(\mathfrak{I}'(\alpha^t, \beta^{t-1})) \neq \emptyset$  holds as well,

2.  $\mathfrak{I}(\alpha^t, \beta^t)$  and  $\mathfrak{I}'(\alpha^t, \beta^t)$  are both defined,  $p(\beta_t \mid \alpha^t, \beta^{t-1}) > 0$ , and

$$d_{max}(\mathfrak{I}(\alpha^{t},\beta^{t}),\mathfrak{I}'(\alpha^{t},\beta^{t})) \leq \frac{d_{max}(\mathfrak{I}(\alpha^{t-1},\beta^{t-1}),\mathfrak{I}'(\alpha^{t-1},\beta^{t-1}))}{p(\beta_{t} \mid \alpha^{t},\beta^{t-1}).}$$

Proof.

1. Assume  $\vartheta(\mathfrak{I}(\alpha^t, \beta^{t-1})) \neq \emptyset$  and, by contradiction,  $\vartheta'(\mathfrak{I}'(\alpha^t, \beta^{t-1})) = \emptyset$ . Since  $d_{max}$  is a fixed point of F, we have  $d_{max} = F(d_{max})$ , and therefore

$$\begin{aligned} d_{max}(\mathfrak{I}(\alpha^{t},\beta^{t-1}),\mathfrak{I}'(\alpha^{t},\beta^{t-1})) &= F(d_{max})(\mathfrak{I}(\alpha^{t},\beta^{t-1}),\mathfrak{I}'(\alpha^{t},\beta^{t-1})) \\ &= 1 \\ &\geq p(\beta_{t} \mid \alpha^{t},\beta^{t-1}), \end{aligned}$$

against the hypothesis.

2. If  $\vartheta(\mathfrak{I}(\alpha^t,\beta^{t-1})) \neq \emptyset$ , then, by the first point of this lemma,  $\vartheta'(\mathfrak{I}'(\alpha^t,\beta^{t-1})) \neq \emptyset$ holds as well, and therefore both  $\mathfrak{I}(\alpha^t,\beta^t)$  and  $\mathfrak{I}'(\alpha^t,\beta^t)$  are defined. The hypothesis  $d_{max}(\mathfrak{I}(\alpha^{t-1},\beta^{t-1}),\mathfrak{I}'(\alpha^{t-1},\beta^{t-1})) < p(\beta_t \mid \alpha^t,\beta^{t-1})$  ensures that  $p(\beta_t \mid \alpha^t,\beta^{t-1}) < 0$ . Let us now prove the bound on  $d_{max}(\mathfrak{I}(\alpha^t,\beta^t),\mathfrak{I}'(\alpha^t,\beta^t))$ . By definition of  $\Phi$ , we have

$$\Phi(d_{max})(\mathfrak{I}(\alpha^{t-1},\beta^{t-1}),\mathfrak{I}'(\alpha^{t-1},\beta^{t-1})) \geq d_{max}(\mathfrak{I}(\alpha^{t},\beta^{t-1}),\mathfrak{I}'(\alpha^{t},\beta^{t-1})).$$

Since  $d_{max} = \Phi(d_{max})$ , we have

$$d_{max}(\mathfrak{I}(\alpha^{t-1},\beta^{t-1}),\mathfrak{I}'(\alpha^{t-1},\beta^{t-1})) \ge d_{max}(\mathfrak{I}(\alpha^{t},\beta^{t-1}),\mathfrak{I}'(\alpha^{t},\beta^{t-1})).$$
(9)

By definition of  $\Phi$  and of the Kantorovich metric, we have

$$\Phi(d_{max})(\mathfrak{I}(\alpha^{t},\beta^{t-1}),\mathfrak{I}'(\alpha^{t},\beta^{t-1})) \geq p(\beta_{t} \mid \alpha^{t},\beta^{t-1}) \cdot \\ d_{max}(\mathfrak{I}(\alpha^{t},\beta^{t}),\mathfrak{I}'(\alpha^{t},\beta^{t})).$$

Using again  $d_{max} = \Phi(d_{max})$ , we get

$$d_{max}(\mathfrak{I}(\alpha^{t},\beta^{t-1}),\mathfrak{I}'(\alpha^{t},\beta^{t-1})) \geq p(\beta_{t} \mid \alpha^{t},\beta^{t-1}) \cdot d_{max}(\mathfrak{I}(\alpha^{t},\beta^{t}),\mathfrak{I}'(\alpha^{t},\beta^{t})),$$

which, together with (9), allows us to conclude.

**Lemma 7** (Lemma 4 in the paper). Consider two IIHSs  $\mathfrak{I}$  and  $\mathfrak{I}'$ , and let  $p(\cdot | \cdot, \cdot)$ and  $p'(\cdot | \cdot, \cdot)$  be their distributions on the output nodes. Given T > 0, and two sequences  $\alpha^T$  and  $\beta^T$ , assume that  $p(\beta_t | \alpha^t, \beta^{t-1}) > 0$  for every t < T. Let  $m = \min_{1 \le t < T} p(\beta_t | \alpha^t, \beta^{t-1})$  and let  $\epsilon \in (0, m^{T-1})$ . Assume  $d(\mathfrak{I}, \mathfrak{I}') < \epsilon$ . Then, for every  $t \le T$ , we have

$$p(\beta_t \mid \alpha^t, \beta^{t-1}) - p'(\beta_t \mid \alpha^t, \beta^{t-1}) < \frac{\epsilon}{m^{T-1}}.$$

*Proof.* Observe that, for every t < T,  $\mathfrak{I}(\alpha^t, \beta^t)$  must be defined, and, by repeatedly applying Lemma 3(1), we get that also  $\mathfrak{I}'(\alpha^t, \beta^t)$  is defined. By definition of  $\varphi$ , and of the Kantorovich metric, we have

$$p(\beta_t \mid \alpha^t, \beta^{t-1}) - p'(\beta_t \mid \alpha^t, \beta^{t-1}) \le \Phi(d_{max})(\mathfrak{I}(\alpha^{t-1}, \beta^{t-1}), \mathfrak{I}'(\alpha^{t-1}, \beta^{t-1})),$$

and since  $d_{max}$  is a fixed point of  $\Phi$ , we get

$$p(\beta_t \mid \alpha^t, \beta^{t-1}) - p'(\beta_t \mid \alpha^t, \beta^{t-1}) \le d_{max}(\mathfrak{I}(\alpha^{t-1}, \beta^{t-1}), \mathfrak{I}'(\alpha^{t-1}, \beta^{t-1})).$$
(10)

By applying t - 1 times Lemma 3(2), from (10) we get

$$p(\beta_t \mid \alpha^t, \beta^{t-1}) - p'(\beta_t \mid \alpha^t, \beta^{t-1}) \le \frac{d_{max}(\mathfrak{I}(\alpha^0, \beta^0), \mathfrak{I}'(\alpha^0, \beta^0))}{m^{t-1}}$$
$$= \frac{d(\mathfrak{I}, \mathfrak{I}')}{m^{t-1}}$$
$$\le \frac{d(\mathfrak{I}, \mathfrak{I}')}{m^{T-1}}$$
$$< \frac{\epsilon}{m^{T-1}}$$

**Theorem 7** (Theorem 4 in the paper). Consider two normalized IIHSs  $\mathfrak{I}$  and  $\mathfrak{I}'$ , and fix a T > 0. For every  $\epsilon > 0$  there exists  $\nu > 0$  such that if  $d(\mathfrak{I}, \mathfrak{I}') < \nu$  then  $|C_T(\mathfrak{I}) - C_T(\mathfrak{I}')| < \epsilon$ .

*Proof.* Consider two normalized IIHSs  $\mathcal{I}$  and  $\mathcal{I}'$  and choose  $T, \epsilon > 0$ . Observe that

$$|C_T(\mathfrak{I}) - C_T(\mathfrak{I}')| = |\max_{p_F(\cdot)} \frac{1}{T} I(A^T \to B^T) - \max_{p_F(\cdot)} \frac{1}{T} I(A'^T \to B'^T)|$$
  
$$\leq \frac{1}{T} \max_{p_F(\cdot)} |I(A^T \to B^T) - I(A'^T \to B'^T)|$$

Since the directed information  $I(A^T \to B^T)$  is defined by means of arithmetic operations and logarithms on the joint probabilities  $p(\alpha^t, \beta^t)$  and on the conditional probabilities  $p(\alpha^t, \beta^t), p(\alpha^t, \beta^{t-1})$ , which in turn can be obtained by means of arithmetic operations from the probabilities  $p(\beta_t \mid \alpha^t, \beta^{t-1})$  and  $p_F(\varphi^t)$ , we have that  $I(A^T \to B^T)$  is a continuous functions of the distributions  $p(\beta_t \mid \alpha^t, \beta^{t-1})$  and  $p_F(\varphi^t)$ , for every  $t \leq T$ . Let  $p(\beta_t \mid \alpha^t, \beta^{t-1}), p'(\beta_t \mid \alpha^t, \beta^{t-1})$  be the distributions on the output nodes of  $\mathfrak{I}$  and  $\mathfrak{I}', \mathfrak{g}^{t-1}) = 0$ , then we redefine the distributions in all the output nodes of the subtree rooted in  $\mathfrak{I}(\alpha^t, \beta^t)$  so that they coincide with the distribution of the corresponding nodes of in  $\mathfrak{I}'$ , and analogously for  $p'(\beta_t \mid \alpha^t, \beta^{t-1})$ . Note that this transformation does not change the directed information, because the subtree rooted in  $\mathfrak{I}(\alpha^t, \beta^t)$  does not contribute to it, due to the fact that it depends the probability of reaching any of its nodes is 0. The continuity of  $I(A^T \to B^T)$  implies that there exists  $\epsilon' > 0$  such that, if  $|p(\beta_t \mid \alpha^t, \beta^{t-1}) - p'(\beta_t \mid \alpha^t, \beta^{t-1})| < \epsilon'$  for all  $t \leq T$  and all sequences  $\alpha^t, \beta^t$ , then, for any  $p_F(\varphi^t)$ , we have  $|I(A^T \to B^T) - I(A'^T \to B'^T)| < \epsilon$ . The result then follows from Lemma 4, by choosing

$$\nu = \epsilon' \cdot \min(\min_{\substack{1 \le t < T \\ p(\beta_t \mid \alpha^t, \beta^{t-1}) > 0}} p(\beta_t \mid \alpha^t, \beta^{t-1}), \min_{\substack{1 \le t < T \\ p'(\beta_t \mid \alpha^t, \beta^{t-1}) > 0}} p'(\beta_t \mid \alpha^t, \beta^{t-1})).$$

# **C:** Computing the reactors and the leakage for the website example

We present here more details about the cocaine auction example in Section 6.

The next step is to construct all the possible reaction functions  $\{f_t(\beta^{t-1})\}_{t=1}^T$ . As seen in Section 4.1, the reaction functions are the correspondent to the encoder in the channel. They take the feedback story and decide how the world is going to react to this situation. For this example, Table 8 shows the reaction functions for each time t.

$\beta^0$	$f_{1(1)}$	$f_{1(2)}$	$f_{1(3)}$
Ø	Candlemaker	Scarface	$a_*$

$\beta^1$	$f_{2(1)}(\beta^1)$	$f_{2(2)}(\beta^{1})$	$f_{2(3)}(\beta^1)$	$f_{2(4)}(\beta^1)$	$f_{2(5)}(\beta^{1})$	$f_{2(6)}(\beta^1)$	$f_{2(7)}(\beta^{1})$	$f_{2(8)}(\beta^1)$	$f_{2(9)}(\beta^{1})$
inc+	Candlemaker	Candlemaker	Candlemaker	Candlemaker	Candlemaker	Candlemaker	Candlemaker	Candlemaker	Candlemaker
inc++	Candlemaker	Candlemaker	Candlemaker	Scarface	Scarface	Scarface	$a_*$	$a_*$	$a_*$
$b_*$	Candlemaker	Scarface	$a_*$	Candlemaker	Scarface	a *	Candlemaker	Scarface	$a_*$
$\beta^1$	$f_{2(10)}(\beta^1)$	$f_{2(11)}(\beta^1)$	$f_{2(12)}(\beta^1)$	$f_{2(13)}(\beta^1)$	$f_{2(14)}(\beta^1)$	$f_{2(15)}(\beta^1)$	$f_{2(16)}(\beta^1)$	$f_{2(17)}(\beta^1)$	$f_{2(18)}(\beta^1)$
inc+	Scarface	Scarface	Scarface	Scarface	Scarface	Scarface	Scarface	Scarface	Scarface
inc++	Candlemaker	Candlemaker	Candlemaker	Scarface	Scarface	Scarface	$a_*$	a *	a *
$b_*$	Candlemaker	Conforme	-	<i>c</i>					
		scarjace	$a_*$	Candlemaker	Scarface	$a_*$	Candlemaker	Scarface	$a_*$
$\beta^1$	$f_{2(19)}(\beta^1)$	$f_{2(20)}(\beta^{1})$	$a_*$ $f_{2(21)}(\beta^1)$	$f_{2(22)}(\beta^1)$	$f_{2(23)}(\beta^1)$	$\frac{a_*}{f_{2(24)}(\beta^1)}$	$f_{2(25)}(\beta^1)$	$f_{2(26)}(\beta^1)$	$a_*$ $f_{2(27)}(\beta^1)$
$\beta^1$ inc+	$f_{2(19)}(\beta^1)$	$\frac{f_{2(20)}(\beta^1)}{a_*}$	$a_*$ $f_{2(21)}(\beta^1)$ $a_*$	$f_{2(22)}(\beta^1)$	$\frac{Scarface}{f_{2(23)}(\beta^1)}$ $a_*$	$\frac{a_{*}}{f_{2(24)}(\beta^{1})}$	$f_{2(25)}(\beta^1)$ $a_*$	$\frac{Scarface}{f_{2(26)}(\beta^1)}$ $a_*$	$\frac{a_*}{f_{2(27)}(\beta^1)}$
$\beta^1$ inc+ inc++	$\begin{array}{c} f_{2(19)}(\beta^{1})\\ \hline a_{*}\\ \hline Candlemaker \end{array}$	$f_{2(20)}(\beta^1)$ $a_*$ Candlemaker	$a_*$ $f_{2(21)}(\beta^1)$ $a_*$ Candlemaker	$f_{2(22)}(\beta^1)$ $a_*$ Scarface	$f_{2(23)}(\beta^1)$ $a_*$ Scarface	$a_*$ $f_{2(24)}(\beta^1)$ $a_*$ Scarface	$\frac{f_{2(25)}(\beta^1)}{a_*}$	$\frac{Scarface}{f_{2(26)}(\beta^1)}$ $\frac{a_*}{a_*}$	$a_*$ $f_{2(27)}(\beta^1)$ $a_*$ $a_*$

(a) All 3 reaction functions  $\varphi_1$ 

(b) All 27 reaction functions  $\varphi_2(\beta^1)$ 

Table 8. Reaction functions for the cocaine auction example.

Now we need to define the reactor, i.e., the reaction functions stochastic kernel. Corollary 1 shows that we can do so by using the following equations:

$$\begin{split} p(\varphi_1) &= p(\alpha_1 | \alpha^0, \beta^0) = p(\alpha_1) \\ p(\varphi_t | \varphi^{t-1}) &= \prod_{\beta^{t-1}} p(\varphi_t(\beta^{t-1}) | \varphi^{t-1}(\beta^{t-2}), \beta^{t-1}), \quad 2 \leq t \leq T \end{split}$$

For instance,  $p(f_{1(1)})=p({\rm Canclemaker})=p_1.$  In the same way,  $p(f_{1(2)})=p({\it Scarface})=p_2$  and  $p(f_{1(3)})=p(a_*)=p_3.$ 

Let us take as an example the calculation of  $p(f_{2(6)}|f_{1(3)})$ :

$$\begin{split} p(f_{2(6)}|f_{1(1)}) &= \prod_{\beta^{1}} p(f_{2(6)}(\beta^{1})|\varphi_{1(1)},\beta^{1}) \\ &= p(f_{2(6)}(inc+)|Candlemaker,inc+) \cdot p(f_{2(6)}(inc++)|Candlemaker,inc++) \\ &\quad p(f_{2(6)}(b_{*})|Candlemaker,b_{*}) \\ &= p(Candlemaker|Candlemaker,inc+) \cdot p(Scarface|Candlemaker,inc++) \\ &\quad p(a_{*}|Candlemaker,b_{*}) \\ &= p_{9} \cdot p_{13} \cdot 1 \\ &= p_{9}p_{13} \end{split}$$
(11)

Note that some reaction functions can have probability 0, which is consistent with probabilistic automaton. For instance:

$$\begin{split} p(f_{2(25)}|f_{1(3)}) &= \prod_{\beta^1} p(f_{2(4)}(\beta^1)|\varphi_{1(3)},\beta^1) \\ &= p(f_{2(4)}(\textit{inc+})|a_*,\textit{inc+}) \cdot p(f_{2(4)}(\textit{inc++})|a_*,\textit{inc++}) \cdot p(f_{2(4)}(b_*)|a_*,b_*) \\ &= p(b_*|a_*,\textit{inc+}) \cdot p(b_*|a_*,\textit{inc++}) \cdot p(\textit{Candlemaker}|a_*,b_*) \\ &= 1 \cdot 1 \cdot 0 \\ &= 0 \end{split}$$
(12)

#### Calculating the information leakage for the cocaine auction example

Let us now calculate the information leakage for this cocaine auction example using the concepts from Section 5. We are going to analyze three different scenarios:

- **Example a:** There is feedback, but the probability of an observable (in general) does not depend on the history of secrets. In the auction protocol, it corresponds to a scenario where the probability of one of the mob members to bid can depend on the increment imposed by the seller, but the history of who has previously bidded in the past has no influence on the choice of increments by the seller during the coming turns. In other words, the server cannot use the information of who is bidding to change his strategy of defining the new increments. That situation corresponds to the original description of the protocol in [15], where the seller does not have access to the identity of the bidder, for the sake of anonymity preservation. In general, we have that  $p(\beta_t | \alpha^t, \beta^{t-1}) = p(\beta_t | \beta^{t-1})$  for every  $1 \le t \le T$ . However, there is an exception: if there is no bidder, case modelled by the secret being  $a_*$ , then the auction is finished, which is signalized by the observable  $b_*$ .
- **Example b:** The most general case, no assumption is made to restrict the model. The presence of feedback allows the probability of the guy bidding to depend on the increment on the price. For instance, if *Candlemaker* is richer than *Scarface*, it is more likely that the later bids if the increment in the price is *inc*++ instead of *inc*+. Also, the probability of an observable can depend on the history of secrets, i.e., in

general  $p(\beta_t | \alpha^t, \beta^{t-1}) \neq p(\beta_t | \beta^{t-1})$  for  $1 \leq t \leq T$ . This scenario can represent a situation where the seller is corrupted and can use his information to affect the outcome of the auction. As an example, suppose that the seller is a friend of *Scarface*'s and he wants to help him in the auction. One way of doing so is to check who was the winner of the last bidding round. Whenever the winner is *Scarface*, the seller chooses for increment the small value *inc+*, hoping that it will give a good chance for *Scarface* to bid in the next round. On the other hand, whenever the seller detects that the winner is *Scarface*, he chooses for the next increment the greater value *inc++*, hoping that it will minimize the chances of *Candlemaker* to bid in the next round (and therefore maximizing the chances of the auction to end having *Scarface* as the final winner).

**Example c:** There is no feedback. In the cocaine auction, we can have the (maybe unrealistic) situation in which the increment added to the bid has no influence on the probability of *Candlemaker* or *Scarface* being the bidder. Matematically, we have that  $p(\alpha_t | \alpha^{t-1}, \beta^{t-1}) = p(\alpha_t | \alpha^{t-1})$  for every  $1 \le t \le T$ . However, like in Example b, we do not impose any restriction to  $p(\beta_t | \alpha^t, \beta^{t-1})$ .

For each scenario we need to attribute values to the probabilities in the protocol tree in Figure 4. The probabilities for each example are listed in Table 9.

Table 10 shows a comparison between some relevant values on the three cases.

In Example a, since the probability of observables (in general) do not depend on the history of secrets, there is (almost) no information flowing from the input to the output, and the directed information  $I(A^T \to B^T)$  is close to zero, i.e., there leakage is low. The only reason why the leakage is not zero is because the end of an auction needs to be signalized. However, due to presence of feedback, the directed information in the other sense  $I(B^T \to A^T)$  is non-zero, so is the mutual information  $I(A^T; B^T)$ . This is an example where the mutual information does not correspond to the real information leakage, since the some of the correlation between input and output can be attributed to the feedback.

In Example b, due to feedback,  $A^T$ , the information flow from outputs to inputs  $I(B^T \to A^T)$  is not zero. In this way, the mutual information  $I(A^T; B^T)$  is higher than the directed information  $I(A^T \to B^T)$ , and the actual leakage of information is given by the latter.

In Example c, the absence of feedback implies that  $I(B^T \to A^T)$  is zero. In that case the values of  $I(A^T; B^T)$  and  $I(A^T \to B^T)$  coincide, and our model collapses to the classic model.

Probability	Example a	Example b	Example c
variable	value	value	value
$n_1$	0.7	0.7	0.7
$\frac{P_1}{n_2}$	0.2	0.2	0.2
P 2	0.1	0.1	0.1
P3 04	0.9	0.1	0.1
94 05	0.1	0.9	0.9
95 06	0.9	0.9	0.9
90 07	0.1	0.1	0.1
$\frac{q_i}{p_0}$	0.6	0.6	0.6
<i>P</i> 9 <i>n</i> 10	0.0	0.0	0.0
<i>P</i> 10	0.5	0.5	0.5
<i>p</i> <sub>11</sub>	0.1	0.1	0.1
<i>P</i> 12	0.3	0.3	0.0
<i>P</i> 13	0.5	0.5	0.5
<i>p</i> <sub>14</sub>	0.2	0.2	0.1
<i>p</i> <sub>15</sub>	0.4	0.4	0.5
<i>p</i> <sub>16</sub>	0.4	0.4	0.2
<i>p</i> <sub>17</sub>	0.2	0.2	0.5
<i>P</i> 18	0.0	0.0	0.5
<i>p</i> 19	0.5	0.5	0.2
<i>P</i> 20	0.1	0.1	0.5
<i>q</i> 22	0.4	0.1	0.1
<i>q</i> 23	0.0	0.9	0.9
<i>q</i> 24	0.7	0.9	0.9
$q_{25}$	0.5	0.1	0.1
$q_{27}$	0.2	0.1	0.1
$q_{28}$	0.8	0.9	0.9
$q_{29}$	0.1	0.9	0.9
$q_{30}$	0.9	0.1	0.1
$q_{32}$	0.4	0.1	0.1
$q_{33}$	0.6	0.9	0.9
$q_{34}$	0.7	0.9	0.9
$q_{35}$	0.3	0.1	0.1
$q_{37}$	0.2	0.1	0.1
$q_{38}$	0.8	0.9	0.9
$q_{39}$	0.1	0.9	0.9
$q_{40}$	0.9	0.1	0.1

 Table 9. Values of the probabilities in Figure 4 in 3 different examples.

Interpretation	Symbol	Example a	Example b	Example c
Input uncertainty	$H(A^T)$	2.3833	2.4891	2.3607
Reactor uncertainty	$H_R$	2.3768	2.4832	2.3607
A posteriori uncertainty	$H(A^T B^T)$	1.3683	0.0677	0.6646
Mutual information	$I(A^T; B^T) = H(A^T) - H(A^T B^T)$	1.0150	1.8214	1.6961
Leakage	$I(A^T \to B^T) = H_R - H(A^T   B^T)$	1.0085	1.8155	1.6961
Feedback information	$I(B^T \to A^T)$	0.185955	0.0060	0.0000
	Table 10 Values for the event	a		

Table	10.	Values	for the	examp	les.
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