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APPROXIMATE COUNTING: A MARTINGALE APPROACH

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ABSTRACT

Approximate counting is a probabilistic algorithm for keeping track of large numbers of events by means of a counter of limited range. In this paper we present an analysis of this algorithm using the elementary theory of martingales. The methods are also applicable to the analysis of the counter which occurs in the exponential back off protocol.

RESUME

"Approximate counting" est un algorithme probabiliste utilisé pour compter un grand nombre d'évènements avec un compteur de capacité bornée. Dans cet article, nous présentons une analyse de cet algorithme en faisant appel à la théorie élémentaire des martingales. La méthode s'applique également à l'étude du compteur associé au protocole "exponential back off".

Walter A. Rosenkrantz, professeur invité dans le projet MEVAL, depuis juillet 1985, sous la direction de Guy Fayolle.

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I - INTRODUCTION

In the paper "Counting Large Numbers of Events in Small registers", Morris [Mo 1978] proposed a probabilistic algorithm for keeping track of a large number of events M with an n bit binary counter where typically M $>> 2^n-1$ = largest integer that can be represented by the counter. This method of counting has been dubbed "Approximate counting" by Flajolet [Fl 1985] who reformulated the problem in terms of a discrete time Markov chain b(t) with state space

$$I^{+} = \{0, 1, 2, ...\}$$

and transition function given by

$$P(b(t+1) = i+1 \mid b(t) = i) = 2^{-i}, \quad i \ge 0$$

$$(1.1)$$

$$P(b(t+1) = i \mid b(t) = i) = 1-2^{-i}.$$

The process b(t) arises naturally when one "counts" an event with probability $2^{-b(t)}$ and does not record it with probability $1-2^{-b(t)}$, where b(t) = current count in the register. How well does b(t) track t?

Morris [Mo 1978] has noted that



(1.2)
$$\begin{cases} a) & \mathbb{E}(2^{b(t)}) = t+2 \\ b) & \sigma^2(2^{b(t)}) = t(t+1)/2 \end{cases}$$

and Flajolet [Fl 1985] has given a proof of (1.2).

Thus $2^{b(t)}$ -2 is an unbiased estimator of t. In addition Flajolet has shown, Theorem 1 of [Fl 1985], that

(1.3)
$$a_1(t) \le E(b(t)) - \log_2 t \le a_2(t)$$

where $a_i(t)$ are small and bounded as $t \rightarrow \infty$.

The proof of (1.3) given in [Fl 1985] is not simple since it uses Mellin transforms and other refined techniques from the theory of functions of a complex variable. It is the purpose of this paper to derive (1.2) and (1.3) by means of a more elementary method, at least more elementary to probabilists!, using only the simplest ideas of the theory of martingales. See chapter VI of [Ka-Ta 1975] for a more comprehensive account. More precisely, we shall prove that (1.3) holds with

$$a_1(t) = \log_2 (2t^{-1} + \log_2)$$

and

$$a_2(t) = \log_2 (1 + 2t^{-1}).$$

In addition we derive a new recursive formula for $E(2^{lb(t)})$, $l \ge 1$.

Interestingly by enough the same process b(t), but with b(1) = 0, occurs in a recent paper by [GGMM 1985] where b(t) represents the back off counter occuring in the exponential back off protocol (EBO) when the channel is always jammed; in fact, this is what first stimulated my interest in this problem - see [Ro 1984]. The problem here is to show that

$$\lim_{t\to\infty}\sup t E(2^{-b(t)}) < \infty.$$

In fact the following result was proved in [GGMM 1985]

(1.4)
$$1/8 \le t E(2^{-b(t)}) \le 9$$
, for all $t \ge 1$.

By a simple martingale argument we are able to obtain the sharper lower bound

(1.5)
$$1 \le t E(2^{-b(t)})$$

and with a little more effort we are able to show

(1.6)
$$t E(2^{-b(t)}) \le 4.96898, t \ge 1.$$

In addition [GGMM 1985] conjectured that

(1.7)
$$E(2^{-b(t)}) \sim c t^{-1}$$
 where $c = 1/\log 2$.

In fact Flajolet has shown [F1 1986] that (1.7) must be modified to take into account bounded fluctuations $\omega(t)$ of small amplitude i.e. he shows that

(1.8)
$$E(2^{-b(t)}) = c t^{-1} + \omega(t)$$
.

The proof of (1.8) is, as is to be expected, quite delicate. Using only very simple tools we are able to prove the following weak form of conjecture (1.7)

(1.9)
$$\begin{cases} a) & \lim \sup_{t \to \infty} \sum_{s=1}^{t-1} E(2^{-b(s)})/\log t \le c \\ t \to \infty & s=1 \end{cases}$$

$$t-1$$

$$b) & \lim \inf_{t \to \infty} \sum_{s=1}^{t-1} E(2^{-b(s)})/\log t \ge 1.$$

These results are obtained in part 3.

Acknowledgments: I wish to thank my colleagues F. Baccelli, G. Fayolle, P. Flajolet and P. Robert here at INRIA for several helpful discussions. In particular the proof of the lower bound (2.9) is due to P. Robert.

11 - A METHOD FOR CONSTRUCTING MARTINGALES ASSOCIATED WITH A MARKOV CHAIN

The following discussion assumes the reader is familiar with chapter VI of Karlin-Taylor. Let y(t) be a Markov chain with countable state space $I^{\dagger} = \{0, 1, 2, \ldots\}$ and transition matrix P = (P(i,j)). Let $\mathcal{J}(t) = \mathcal{J}(y(u), 0 \le u \le t)$ denote the smallest σ -field generated by the random variables y(u), $0 \le u \le t$. Intuitively $\mathcal{J}(t)$ is the past up until time t. We recall the

Definition: we say that
$$x(t)$$
 is a martingale with respect to (t) if (i) $x(t)$ is $\mathcal{J}(t)$ measurable (2.1)
$$(ii) \quad E(|x(t)|) < \infty \text{ and } (iii) \quad E(x(t+1) | \mathcal{J}(t)) = x(t).$$

In order to construct martingales associated with the Markov chain y(t) we make use of the operators P and A acting on the function space

$$D = \{f \mid f : I^{+} \rightarrow R, \sum_{j=0}^{\infty} P(i,j) \mid f(j) \mid < \infty \}$$

Definition: Let f & D

(i)
$$Pf(i) \stackrel{\sim}{\Delta} \stackrel{\sim}{:} P(i,j) f(j)$$

(2.2) (ii)
$$Af(i) \stackrel{\frown}{\Delta} Pf(i) - f(i)$$

(iii)
$$P | f|$$
 (i) $\Delta \sum_{j=0}^{\infty} P(i,j) | f|$ (j).

Note that $E\{f(y(t+1)) - f(y(t)) | \mathcal{F}(t)\} = E\{f(y(t+1)) - f(y(t)) | y(t)\} = Af(y(t))$ and using this fact it is easy to establish the following

Lemma : Suppose $P[f](i) < \infty$ and set

(2.3)
$$x(t) = f(y(t)) - \sum_{s=0}^{t-1} Af(y(s)), t \ge 1 \text{ and } x(0) = f(y(0)).$$

Then $\{x(t), \mathcal{J}(t)\}\$ is a martingale.

For future reference we note the following special case. Suppose $Af(i) \equiv c$ all $i \geq 0$. Then f(y(t)) - ct is a martingale.

Lemma (2.3) yields a novel and computationally simple proof of Flajolet's proposition 0 as well as a recursive formula for the higher moments of $2^{b(t)}$.

Proposition:

(2.4)

Let
$$b(0) = 1$$
. Then $E(2^{b(t)}) = t+2$ and $\sigma^2(2^{b(t)}) = t(t+1)/2$.

Proof. We begin by noting that

(2.5)
$$Ar(j) = 2^{-j} [r(j+1) - r(j)].$$

Let $f_{\ell}(j) = 2^{\ell j}$ where $\ell = 0, 1, 2, \ldots$ An easy calculation shows that

(2.6)
$$Ar_g(j) = (2^{g-1}) r_{g-1}(j)$$
.

In particular $Af_1(j) \equiv 1$ and therefore, by lemma (2.3), we have

$$f_1(b(t)) - \sum_{s=0}^{l-1} 1 - 2^{b(t)} - t$$

is a martingale. Consequently

$$E(f_1(b(t)) - t) = f_1(b(0)) = f_1(1) = 2$$
 i.e.

(2.7)
$$E(f_1(b(t)) = E(2^{b(t)}) = t+2.$$

Lemma (2.3) and formula (2.6) can now be used to derive a recursive formula for the moments $E(2^{lb(t)})$, $l \ge 2$. To illustrate these ideas consider the case l = 2. Since $Af_2(j) = 3f_1(j)$ lemma (2.3) and (2.7) yield

$$E(f_{2}(b(t)) - \sum_{s=0}^{t+1} 3f_{1}(b(s)) = f_{2}(b(0)) = 4$$

$$E(2^{2b(t)}) = E(f_2(b(t)) = 4 + 3 \sum_{s=0}^{t-1} E(f_1(b(s)) = 4 + 3 \sum_{s=0}^{t-1} (s+2) = 4 + 3 \cdot \frac{(t-1)t}{2} + 6 \cdot t = \frac{3}{2} t(t+3) + 4.$$

MOW

$$\sigma^2(2^{b(t)}) = E(2^{2b(t)}) - E(2^{b(t)})^2 = \frac{3}{2} t(t+3) + 4 - (t+2)^2 = \frac{t(t+1)}{2}.$$

: 3.

1. 1. E.

Using (2.3) we can give a recurrence formula for the higher moments of $2^{b(t)}$. More precisely let $M_{\ell}(t) = E(2^{\ell b(t)})$, $\ell = 0, 1, \ldots$ Then lemma (2.3) and (2.6) yield

(2.8)
$$\mathbb{H}_{\ell}(t) = 2^{\ell} + (2^{\ell-1}) \sum_{s=0}^{t-1} \mathbb{H}_{\ell-1}(s).$$

Note. We believe formula (2.8) to be new.

The formula $E(2^{b(t)})$ = t+2 together with Jensen's inequality applied to the concave function $\phi(t)$ = $\log_2 t$ yield

$$E(b(t)) \le \log_2(t+2) = \log_2 t + \log(1+2t^{-1}).$$

We now proceed to derive a lower bound for $E(2^{b(t)})$ using an argument suggested to the author by P. Robert (INRIA).

An easy calcuulation shows that

$$E(b(t+1) - b(t)) = E(2^{-b(t)}) \ge 2^{-E(b(t))}$$

and therefore

$$2^{\mathbb{E}(b(t))}$$
. $\mathbb{E}(b(t+1) - b(t)) \ge 1$.

Thus

$$\begin{array}{ll} t-1 & & \\ \Sigma & 2^{E(b(s))} & E(b(s+1) - b(s)) \ge t. \\ s=0 & & \end{array}$$

Set $\phi(s) = E(b(s))$; so $\phi(0) = 1$. The preceeding expression can be rewritten as $\frac{t^{-1}}{\sum_{s=0}^{\infty} 2^{\phi(s)}} (\phi(s+1) - \phi(s)) \ge t.$

On the other hand

$$t \leq \sum_{s=0}^{t-1} 2^{\phi(s)} \left[\phi(s+1) - \phi(s) \right] \leq \int_{1}^{t} 2^{s} ds = e^{(\log 2)\phi(t)} - e^{\log 2} / \log 2 \quad i.e.$$

$$2^{\phi(t)} \geq 2 + (\log 2).t$$

and this implies

$$E(b(t)) = \phi(t) \ge \log_2 (t[\log 2 + 2t^{-1}]) = \log_2 t + \log_2 [\log 2 + 2t^{-1}].$$

Summing up then we're shown that

(2.9)
$$\log_2 t + \log_2 (2t^{-1} + \log_2) \le E(b(t)) \le \log_2 t + \log_2 (1 + 2t^{-1}).$$

Clearly this implies $\lim_{t\to\infty} \frac{E(b(t))}{\log_2 t} = 1$ and also that

(2.10)
$$\log_2(\log 2) \le E(b(t)) - \log_2 t \le 1/t$$
.

III - AN IMPROVED ESTIMATE FOR E(2-b(t))

In this part of the paper we assume b(1) = 0 and repeating the same argument of proposition (2.4) we see that

$$f_1(b(t)) - \sum_{s=1}^{t-1} Af_1(b(s))$$

is a martingale with

$$E(f_1(b(t)) - \sum_{s=1}^{t-1} Af_1(b(s)) = f_1(b(1)) = 1.$$

Consequently $E(2^{b(t)}) = t$. We now apply Jensen's inequality with $\phi(x) = x^{-1}$ to deduce the much sharper lower bound

(3.1)
$$E(2^{-b(t)}) \ge 1/t$$
.

To obtain the upper bound

(3.2)
$$\mathbb{E}(2^{-b(t)}) \le 4.1/t$$
.

We derive an exponential bound on the first passage time $\tau_y = \inf\{t: b(t) \ge y\}$. This leads to a sharper estimate than the one obtained by [GCMM 1985] who used Chebychev's inequality; otherwise the two proofs are the same.

Lemma: For any
$$\lambda \le \lambda(y) = -\log(1 - 2^{-y-1})$$
(3.3)
$$y-1$$

$$P(\tau_y \ge t) \le \exp(-\lambda t + \lambda) \prod_{i=1}^{y-1} (1 - (1 - \exp(-\lambda))2^i)^{-1}.$$

<u>Proof.</u> Let y_i be the random variable with geometric distribution given by $P(y_i = j) = 2^{-\frac{1}{2}(1-2^{-\frac{1}{2}})^{\frac{1}{2}-1}}$, $j \ge 1$. Thus its moment generating function

$$E(e^{\lambda y}) = (1 - (1 - \exp(-\lambda))2^{1})^{-1}.$$

Clearly $\tau_1 = 1$ and for y > 1 we have

Since the y_i , $1 \le i \le y-1$, are mutually independent we see at once that

$$M_{y}(\lambda) = E(\exp(\lambda \tau_{y})) = \exp(\lambda) \prod_{i=1}^{y-1} E(\exp(\lambda y_{i}))$$

$$= \exp(\lambda) \prod_{i=1}^{y-1} (1 - (1 - \exp(-\lambda))2^{i})^{-1}.$$

By Chebychev's inequality $\exp(\lambda t)$ $P(\tau_y \ge t) \le M_y(\lambda)$ which is precisely lemma (3.3). Note that a sufficient condition for

$$M_y(\lambda) < \infty$$
 is $\sup_{1 \le i \le y-1} 2^i (1 - \exp(-\lambda)) < 1$.

In particular if we set $\lambda = \lambda(y)$ where

(3.4)
$$\lambda(y) = -\log(1-2^{-y})$$

then $M_y(\lambda) < \infty$ for $\lambda \le \lambda(y)$ and

(3.5)
$$\log(1 - (1 - \exp(-\lambda))2^{i}) \ge - (2\log 2)(1 - \exp(-\lambda))2^{i}$$
,

where we've used the inequality

$$log(1-x) \ge -(2log2)x$$
 on $0 \le x \le \frac{1}{2}$.

Our next step is to set $\lambda = \lambda(y)$ in lemma (3.3) and then take the log of both sides which leads to

$$\log P(\tau_{y} \ge t) \le -\lambda(y)(t-1) + (2\log 2) \sum_{i=1}^{y-1} 2^{i-y}$$

$$\le -\lambda(y)t + (2\log 2)(1-2^{-y}) + \lambda(y).$$

Consequently

$$P(\tau_v \ge t) \le c(y) \exp(-\lambda(y)t)$$

where

$$c(y) = \exp{\lambda(y) + (2\log 2)(1 - 2^{-y})} \le 4, y \ge 1.$$

We have thus derived the exponential bound

(3.6)
$$P(\tau_{v} \ge t) \le 4 \exp(-\lambda(y)t).$$

We now proceed to estimate $E(2^{-b(t)})$. Set $m = \lfloor \log_2 t \rfloor$ so $(\log_2 t)-1 < m \le \log_2 t$, consequently

$$E(2^{-b(t)}; b(t) \ge m) \le 2/t.$$

Here E(f; A) means integrating f over the subset A. So the problem now is to show that

$$E(2^{-b(t)}; b(t) \le m-1) \le a/t$$

and to estimate the constant a. Now

(3.7)
$$E(2^{-b(t)}; b(t) \le m-1) \le \sum_{j=0}^{m-1} 2^{-j} P(b(t) \le j).$$

But

$$P(b(t) \le j) = P(\tau_j \ge t) \le 4 \exp(-\lambda(j)t)$$

and

$$-\lambda(j) = + \log(1 - 2^{-j}) \le -2^{-j}$$

together imply

$$P(b(t) \le j) \le 4 \cdot exp(-2^{-j}t)$$

 $\le 4 \cdot exp(-2^{m-j}).$

The right hand side of (3.7) is bounded by the sum

$$S_1 = \sum_{j=0}^{m-1} 2^{-j} \exp(-2^{m-j}).$$

Since $-2^{m-j} \le -(m-j)^2$ it follows that

$$S_{1} \le 4 \sum_{j=0}^{m-1} 2^{-j} [\exp(-2)]^{m-j} = 4 \exp(-2m) \sum_{j=0}^{m-1} {\frac{\exp(2)}{2}}^{j} \le$$

$$\le (4 / [(e^{2}/2)-1]) 2^{-m} \le a/t$$

where

$$a = 8 / [(e^2/2)-1] = 2.96898.$$

Finally, putting these estimates together we obtain

(3.8)
$$E(2^{-b(t)}) \le 4.96898 / t.$$

We conclude this part of the paper by establishing a weak form of the conjecture (1.7). More precisely set $m(t) = E(2^{-b(t)})$ and apply lemma (2.3) to the function f(j) = j. Since $Af(j) = 2^{-j}$ we see at once that

$$E(b(t) - \sum_{s=1}^{t-1} 2^{-b(s)}) = 0$$

and therefore

On the other hand as we've already seen $E(2^{-b(t)}) \ge 1/t$ which implies

(3.10)
$$\sum_{s=1}^{t-1} E(2^{-b(s)}) \ge \sum_{s=1}^{t-1} 1/s \ge 1 + \int_{1}^{t-1} \frac{ds}{s} = 1 + \log(t-1).$$

Combining (3.9) and (3.10) yields

(b) $\lim_{t\to\infty}\inf \frac{t=1}{\sum_{s=1}^{\infty}E(2^{-b(s)})}$ $\geq 1.$

Note that if in fact $E(2^{-b(s)})$ ~ c/s then Σ $E(2^{-b(s)})$ ~ c log t. This is why we call (3.11) the weak form of conjecture (1.7).

EV - AN ESTIMATE FOR THE TAIL OF THE DISTRIBUTION

Using the result that $E(2^{2b(t)}) = \frac{4}{2} + (3/2)$ t(t+3) and Chebychevis inequality it is easy to see that

(4.1)
$$P(b(t) \ge 2 \log_2 t + \delta) = O(2^{-2\delta}t^{-2}), \delta \ge 0.$$

in ract

$$P(b(t) \ge 2 \log_2 t + \delta) = P(2^{2b(t)} \ge 2^{2\delta}t^4) \le \frac{4 + (3/2)t(t+3)}{2^{2\delta} \cdot t^4} \le c_1 2^{-2\delta} \cdot t^{-2}$$

where c_1 is independent of δ and t. This is in fact a considerable strengthening of Flajolet's proposition 4 wherein he shows that

$$P(b(t) = 2\log_2 t + \delta) = 0(2^{-\delta}t^{-.99}),$$

uniformly in t and $\delta \ge 0$.

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