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PROGRAMME 1

Architectures parallèles,
bases de données,
réseaux et systèmes distribués



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Transient analysis of the M/M/1 queue

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Gerardo Rubino* and Bruno Sericola*

Programme 1 — Architectures parallèles, bases de données, réseaux
et systèmes distribués
Projet Model

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Abstract: A new approach is given to obtain the transient probabilities of the $M/M/1$ queueing system. A first step of this approach deals with the generating function of the transient probabilities of the uniformized Markov chain associated with this queue. The second step consists of the inversion of this generating function. A new analytical expression of the transient probabilities of the $M/M/1$ queue is then obtained.

Key-words: M/M/1 queue, transient analysis, uniformization, generating function

(Résumé : tsvp)

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*{Patrice.Leguesdron} {Gerardo.Rubino}{Bruno.Sericola}@irisa.fr

**INSA, 20 avenue des Buttes de Coesmes, 35043 Rennes cedex

Unité de recherche INRIA Rennes

IRISA, Campus universitaire de Beaulieu, 35042 RENNES Cedex (France)

Téléphone : (33) 99 84 71 00 – Télécopie : (33) 99 38 38 32

Analyse transitoire de la file M/M/1

Résumé : Une nouvelle approche est donnée pour obtenir les probabilités transitoires de la file $M/M/1$. Une première étape de cette approche traite de la fonction génératrice des probabilités transitoires de la chaîne de Markov uniformisée associée à cette file. La deuxième étape consiste à inverser cette fonction génératrice. Une nouvelle expression analytique des probabilités transitoires de la file $M/M/1$ est ainsi obtenue.

Mots-clé : File M/M/1, analyse transitoire, uniformisation, fonction génératrice

1 Introduction

Several papers have been proposed for the study of the transient behaviour of the $M/M/1$ queueing system. In most of these papers, the derived expressions are complicated by the fact that they often refer to Bessel functions. These expressions are generally obtained by a combined use of generating functions and Laplace transforms ([1], [2], [3], [4], [5]). The use of Bessel functions leads to complex numerical solutions. An overview of these methods can be found in [6]. Recently, in [7], the authors have proposed an approach using only Taylor series but it leads also to numerical problems since the coefficients of these series are alternatively positive and negative.

In this paper, we propose a new method based on the uniformization technique and on generating functions. In the following section, we give an expression of the generating function of the transient probabilities of the uniformized Markov chain associated with the Markov process describing the $M/M/1$ queue. In Section 3 this generating function is inverted and we obtain an analytical expression of the transient probabilities of the uniformized Markov chain which leads to an analytical expression for the transient probabilities of the $M/M/1$ queue. This expression is quite simple in the case where the queue is initially empty. When the queue is not initially empty, the expression obtained is reduced to the previous one (Theorem 1(b) of [8]) by adding a term corresponding to the transient probabilities of the associated absorbing process obtained by making state 0 absorbing. In deriving this last expression, we obtain also a simple expression of the probability to go from state i to state j in n transitions on the absorbing process. The last section is devoted to some conclusions.

2 Generating functions and uniformized Markov chain

Consider the classical $M/M/1$ queue with arrival rate λ and service rate μ . We denote by $X = (X_t)_{t \geq 0}$ the continuous time Markov process associated with the state of the queue. The infinitesimal generator of the process X is denoted by A . The non zero entries of this matrix are

$$A_{0,0} = -\lambda \text{ and for any } i \geq 1, \quad A_{i,i-1} = \mu, \quad A_{i,i} = -(\lambda + \mu), \quad A_{i,i+1} = \lambda.$$

For every $i, j \in \mathbb{N}$ and $t \geq 0$, we denote by $P_{i,j}(t)$ the probability to have j customers in the queue (including the one being served) at time t given that there are i customers in the queue at time 0. The classical following result holds:

$$P_{i,j}(t) = \mathbb{P}(X_t = j \mid X_0 = i) = \left(e^{At} \right)_{i,j}.$$

To compute these transient probabilities, we first consider the uniformized discrete time Markov chain (see for instance [9]) associated with process X . The transition probability matrix of this Markov chain is denoted by P . We then have the following relation between matrices A and P :

$$P = I + \frac{1}{\lambda + \mu} A,$$

where I denotes the identity matrix, that is $I_{i,j} = 1$ if $i = j$ and 0 otherwise. This relation allows us to write the matrix e^{At} as

$$e^{At} = \sum_{k=0}^{+\infty} e^{-(\lambda+\mu)t} \frac{(\lambda + \mu)^k t^k}{k!} P^k. \quad (1)$$

If p and q are defined by

$$p \stackrel{\text{def}}{=} \frac{\lambda}{\lambda + \mu} \quad \text{and} \quad q \stackrel{\text{def}}{=} \frac{\mu}{\lambda + \mu},$$

the non zero entries of matrix P are

$$P_{0,0} = q, \quad \forall i \geq 1 \quad P_{i,i-1} = q \quad \text{and} \quad \forall i \geq 0 \quad P_{i,i+1} = p.$$

In what follows, we focus on the calculus of P^k using generating functions.

Let us consider the functions M defined on $\mathbb{N} \times \mathbb{N}$ with complex values. As usual, we see them as infinite matrices and denote by $M_{i,j}$ the value $M(i, j)$. With each M we associate the value

$$\nu(M) \stackrel{\text{def}}{=} \sup_i \sum_j |M_{i,j}|$$

and we denote by \mathcal{M} the set of infinite matrices M such that $\nu(M) < +\infty$. It is well known that ν is a norm on the set \mathcal{M} and that (\mathcal{M}, ν) is a Banach

algebra [10]. The *potential kernel* of the matrix $M \in \mathcal{M}$, denoted by Φ_M , is the complex function defined by

$$\Phi_M(z) \stackrel{\text{def}}{=} \sum_{k=0}^{+\infty} M^k z^k$$

for all z verifying $|z| < 1/\nu(M)$. It can be seen as the generating function of the sequence of the powers of matrix M . It is well known that Φ_M is the only solution to the matrix equation in X

$$\forall |z| < 1/\nu(M), \quad X(z) = I + zMX(z). \quad (2)$$

More generally, for every matrix H , $H\Phi_M$ is the only solution to the matrix equation in X

$$\forall |z| < 1/\nu(M), \quad X(z) = H + zX(z)M. \quad (3)$$

Another useful property of the potential kernel is given below as a lemma:

Lemma 2.1 *For any functions M and N of \mathcal{M} we have*

$$\Phi_{M+N}(z) = \Phi_M(z) + z\Phi_M(z)N\Phi_{M+N}(z) \quad (4)$$

$$= \Phi_M(z) + z\Phi_{M+N}(z)N\Phi_M(z) \quad (5)$$

for all z such that $|z| < \inf(1/\nu(M), 1/\nu(N), 1/\nu(M + N))$.

Proof. These relations follow from

$$\begin{aligned} \forall k \geq 1 \quad (M + N)^k &= M^k + \sum_{j=0}^{k-1} (M + N)^j N M^{k-j-1} \\ &= M^k + \sum_{j=0}^{k-1} M^j N (M + N)^{k-j-1} \end{aligned}$$

which can be easily checked by induction on k . □

Consider the potential kernel of P , that is

$$\Phi_P(z) = \sum_{k=0}^{+\infty} P^k z^k.$$

We are going to find an explicit expression of Φ_P and we shall inverse it to obtain an explicit expression of P^k . We now need some notation and a lemma. Let V , W and R be the matrices defined by

$$V_{i,j} = I_{i+1,j} \ , \ W_{i,j} = I_{i,j+1} \ , \ R_{i,j} = I_{i,0}I_{0,j} \ \forall i, j \in \mathbb{N}.$$

Lemma 2.2 *Let α and β be any complex numbers. Let X be the infinite matrix defined by*

$$X_{i,j} = \alpha^i \beta^j$$

and Y the infinite matrix defined by

$$Y = \sum_{k=0}^{+\infty} W^k X V^k.$$

Then matrix Y satisfies

$$(\alpha W + \beta V)Y = (1 + \alpha\beta)Y - I. \quad (6)$$

Proof. Observe that Y is well defined for every $i, j \in \mathbb{N}$ as

$$Y_{i,j} = \begin{cases} \beta^{j-i} s_i & \text{if } i \leq j \\ \alpha^{i-j} s_j & \text{if } i \geq j \end{cases}$$

where for every $m \geq 0$,

$$\begin{aligned} s_m &= \sum_{l=0}^m (\alpha\beta)^l \\ &= \frac{1 - (\alpha\beta)^{m+1}}{1 - \alpha\beta} \text{ if } \alpha\beta \neq 1. \end{aligned}$$

To prove Relation (6), we denote by L its left hand side, that is, $L = (\alpha W + \beta V)Y$. For $i > j$, we have

$$\begin{aligned} L_{i,j} &= \alpha Y_{i-1,j} + \beta Y_{i+1,j} \text{ (with the convention } Y_{-1,j} = 0) \\ &= \alpha \alpha^{i-j-1} s_j + \beta \alpha^{i-j+1} s_j \\ &= (1 + \alpha\beta) \alpha^{i-j} s_j \\ &= (1 + \alpha\beta) Y_{i,j}. \end{aligned}$$

The case $i < j$ is similar to the case $i > j$. On the diagonal we have

$$\begin{aligned}
 L_{i,i} &= \alpha Y_{i-1,i} + \beta Y_{i+1,i} \\
 &= \alpha \beta s_{i-1} + \alpha \beta s_i \\
 &= s_i - 1 + \alpha \beta s_i \\
 &= (1 + \alpha \beta) s_i - 1 \\
 &= (1 + \alpha \beta) Y_{i,i} - 1.
 \end{aligned}$$

□

The next theorem gives an expression of the potential kernel $\Phi_P(z)$ of matrix P . First let us recall that the Catalan's numbers are defined for all positive integer n by

$$c_n \stackrel{\text{def}}{=} \binom{2n}{n} \frac{1}{n+1}.$$

The generating function of the sequence of Catalan's numbers,

$$C(z) \stackrel{\text{def}}{=} \sum_{n=0}^{+\infty} c_n z^n, \quad (7)$$

converges for $|z| \leq 1/4$ and we have

$$C(z) = \frac{1 - \sqrt{1 - 4z}}{2z}.$$

Remark that this complex function verifies

$$C(z) = 1 + zC^2(z). \quad (8)$$

Theorem 2.3 *Let $|z| < 1$ and $\eta(z) \stackrel{\text{def}}{=} C(pqz^2)$. Let $X(z)$ be the infinite matrix defined by*

$$X_{ij}(z) \stackrel{\text{def}}{=} (qz\eta(z))^i (pz\eta(z))^j$$

and $Y(z)$ the infinite matrix defined as in Lemma 2.2 by

$$Y(z) = \sum_{k=0}^{+\infty} W^k X(z) V^k.$$

Then, we have

$$\forall z \text{ such that } |z| < \frac{1}{2}, \quad \Phi_P(z) = \eta(z)Y(z) + \frac{qz\eta(z)}{1 - qz\eta(z)}\eta(z)X(z). \quad (9)$$

Proof. Let us denote by Q the matrix $pV + qW$. We have

$$P = Q + qR.$$

For $|z| < 1$, we have $|pqz^2| < 1/4$ (observe that since $p + q = 1$, we have $pq \leq 1/4$) and so $\eta(z)$ is well defined and verifies $|\eta(z)| < 2$. Moreover, writing Relation (7) with $z = pq$ and after some algebra, we obtain that $|pz\eta(z)| < 1$ and $|qz\eta(z)| < 1$, and so

$$\nu(X(z)) = \frac{1}{1 - p|z||\eta(z)|}.$$

Applying the previous lemma, in which we set $\alpha = qz\eta(z)$ and $\beta = pz\eta(z)$, we obtain

$$z\eta(z)QY(z) = (1 + pqz^2\eta^2(z))Y(z) - I.$$

Relation (8) implies that $\eta(z) = 1 + pqz^2\eta^2(z)$, so we have

$$\eta(z)Y(z) = I + zQ\eta(z)Y(z).$$

Now since $\nu(Q) = 1$, we can apply Relation (2) and we obtain

$$\forall z \text{ such that } |z| < 1, \quad \eta(z)Y(z) = \Phi_Q(z).$$

From Relation (4) in which we set $M = Q$ and $N = qR$, we have

$$\forall z \text{ such that } |z| < 1, \quad \Phi_P(z) = \Phi_Q(z) + qz\Phi_P(z)R\Phi_Q(z),$$

that is, since $\Phi_Q(z) = \eta(z)Y(z)$,

$$\forall z \text{ such that } |z| < 1, \quad \Phi_P(z) = \eta(z)Y(z) + z\Phi_P(z)qR\eta(z)Y(z).$$

To prove Relation (9), we fix z such that $|z| < 1/2$ and we apply Relation (3) in which we set $H(z) = \eta(z)Y(z)$ and $M(z) = qR\eta(z)Y(z)$. We must verify that for $|z| < 1/2$, we have $|z| < 1/\nu(M(z))$. By definition of $Y(z)$ and since $RW = 0$, we have $RY(z) = RX(z)$. This leads to

$$\begin{aligned} \forall z \text{ such that } |z| < \frac{1}{2}, \quad \nu(M(z)) &= \nu(qR\eta(z)Y(z)) \\ &= q|\eta(z)|\nu(RY(z)) \\ &= q|\eta(z)|\nu(RX(z)) \end{aligned}$$

$$\begin{aligned}
 &= q|\eta(z)|\nu(X(z)) \\
 &= \frac{q|\eta(z)|}{1-p|z||\eta(z)|} \\
 &< \frac{2q}{1-p} = 2.
 \end{aligned}$$

So,

$$|z| < \frac{1}{2} \implies \nu(M(z)) < 2 \implies |z| < 1/\nu(M(z)).$$

We can now apply Relation (3) to obtain

$$\forall z \text{ such that } |z| < \frac{1}{2}, \quad \Phi_P(z) = \eta(z)Y(z) \sum_{k=0}^{+\infty} (qz\eta(z))^k (RY(z))^k$$

or (since $RY(z) = RX(z)$),

$$\forall z \text{ such that } |z| < \frac{1}{2}, \quad \Phi_P(z) = \eta(z)Y(z) \sum_{k=0}^{+\infty} (qz\eta(z))^k (RX(z))^k.$$

Observing that $RX(z)R = R$ and so that for every $k \geq 1$, we have $(RX(z))^k = RX(z)$, we can write

$$\forall z \text{ such that } |z| < \frac{1}{2}, \quad \Phi_P(z) = \eta(z)Y(z) + \eta(z)Y(z) \sum_{k=1}^{+\infty} (qz\eta(z))^k RX(z)$$

that is, since $|qz\eta(z)| < 1$,

$$\forall z \text{ such that } |z| < \frac{1}{2}, \quad \Phi_P(z) = \eta(z)Y(z) + \frac{qz\eta(z)}{1-qz\eta(z)}\eta(z)Y(z)RX(z).$$

By definition of $Y(z)$ and since $VR = 0$, we have $Y(z)R = X(z)R$. Furthermore, it is easy to check that $X(z)RX(z) = X(z)$. Finally,

$$\forall z \text{ such that } |z| < \frac{1}{2}, \quad \Phi_P(z) = \eta(z)Y(z) + \frac{qz\eta(z)}{1-qz\eta(z)}\eta(z)X(z)$$

which completes the proof. □

3 Transient probabilities

In this section, we focus on Relation (9) to obtain an expression of $(P^n)_{i,j}$, the transient probabilities of the uniformized Markov chain of the $M/M/1$ queue. These discrete transient probabilities will lead by means of Formula (1) to the transient probabilities of the $M/M/1$ queue. In what follows, we always work in the complex domain $|z| < 1/2$.

For $i, j \in \mathbb{N}$, $Y_{i,j}(z)$ is given in the proof of Lemma 2.2 in which we set $\alpha = qz\eta(z)$ and $\beta = pz\eta(z)$, that is

$$Y_{i,j}(z) = \begin{cases} (pz\eta(z))^{j-i} \sum_{k=0}^i p^k q^k z^{2k} \eta(z)^{2k} & \text{if } i \leq j \\ (qz\eta(z))^{i-j} \sum_{k=0}^j p^k q^k z^{2k} \eta(z)^{2k} & \text{if } i \geq j \end{cases}$$

In what follows, we consider only the case $i \leq j$. The case $i \geq j$ is similar since, as mentioned in [8, page 325], we have for every $i, j \geq 0$,

$$P_{i,j}(t) = \left(\frac{\lambda}{\mu}\right)^{j-i} P_{j,i}(t).$$

So, let $i, j \in \mathbb{N}$ such that $i \leq j$. By definition of $X(z)$ we obtain

$$(\Phi_P(z))_{i,j} = (pz\eta(z))^{j-i} \sum_{k=0}^i p^k q^k z^{2k} \eta(z)^{2k+1} + p^j q^{i+1} z^{i+j+1} \eta(z)^{i+j+1} \frac{\eta(z)}{1 - qz\eta(z)}. \quad (10)$$

In the next subsection, we first consider the case $i = 0$, that is, the case when the queue is initially empty.

3.1 The queue is initially empty

The next theorem gives an expression of the transient probabilities of the uniformized Markov chain associated with X , the process describing the behaviour of the $M/M/1$ queue. To prove it we need the following lemma which gives an analytical expression of the powers of $\eta(z)$.

Lemma 3.1 For every $k \geq 1$ and for every z such that $|z| \leq 1/4$, we have

$$C^k(z) = \sum_{n=0}^{+\infty} s(k, n)z^n,$$

where

$$s(k, n) = k \frac{(2n + k - 1)!}{n!(n + k)!}.$$

Proof. (see for instance [11, page 154]) □

Theorem 3.2 For every $j \geq 0$,

$$(P^n)_{0,j} = \begin{cases} 0 & \text{if } n < j \\ \frac{p^j}{q^j} \sum_{k=0}^{\lfloor \frac{n-j}{2} \rfloor} s(n + 1 - 2k, k) p^k q^{n-k} & \text{if } n \geq j. \end{cases}$$

where $\lfloor x \rfloor$ denotes the largest integer not greater than the real number x .

Proof. By setting $i = 0$ in Relation (10), we obtain

$$(\Phi_P(z))_{0,j} = (pz\eta(z))^j \frac{\eta(z)}{1 - qz\eta(z)}. \quad (11)$$

This relation can be written as

$$(\Phi_P(z))_{0,j} = \frac{p^j}{q^{j+1}z} \sum_{k=j+1}^{+\infty} [qz\eta(z)]^k.$$

Using Lemma 3.1, this gives

$$[qz\eta(z)]^k = [qzC(pqz^2)]^k = \sum_{n=0}^{+\infty} s(k, n) p^n q^{n+k} z^{2n+k}.$$

It follows that for every $j \geq 0$,

$$(\Phi_P(z))_{0,j} = \frac{p^j}{q^{j+1}z} \sum_{k=j+1}^{+\infty} \sum_{n=0}^{+\infty} s(k, n) p^n q^{n+k} z^{2n+k}$$

$$\begin{aligned}
&= \frac{p^j}{q^{j+1}z} \sum_{n=0}^{+\infty} \sum_{k=j+1}^{+\infty} s(k, n) p^n q^{n+k} z^{2n+k} \\
&= \frac{p^j}{q^{j+1}z} \sum_{n=0}^{+\infty} \sum_{k=2n+j+1}^{+\infty} s(k - 2n, n) p^n q^{k-n} z^k \\
&= \frac{p^j}{q^{j+1}z} \sum_{k=j+1}^{+\infty} \sum_{n=0}^{\lfloor \frac{k-(j+1)}{2} \rfloor} s(k - 2n, n) p^n q^{k-n} z^k \\
&= \frac{p^j}{q^j} \sum_{k=j}^{+\infty} \sum_{n=0}^{\lfloor \frac{k-j}{2} \rfloor} s(k + 1 - 2n, n) p^n q^{k-n} z^k
\end{aligned}$$

and so we obtain the desired expression for $(P^n)_{0,j}$. \square

The transient probabilities of the $M/M/1$ queue given that the queue is empty at time $t = 0$ are then

$$P_{0,j}(t) = \frac{p^j}{q^j} \sum_{n=j}^{+\infty} e^{-(\lambda+\mu)t} \frac{(\lambda + \mu)^n t^n}{n!} \sum_{k=0}^{\lfloor \frac{n-j}{2} \rfloor} s(n + 1 - 2k, k) p^k q^{n-k} \quad (12)$$

that is,

$$P_{0,j}(t) = \frac{p^j}{q^j} \sum_{n=j}^{+\infty} e^{-(\lambda+\mu)t} \frac{(\lambda + \mu)^n t^n}{n!} \sum_{k=0}^{\lfloor \frac{n-j}{2} \rfloor} \frac{n + 1 - 2k}{n + 1} \binom{n + 1}{k} p^k q^{n-k}. \quad (13)$$

If $j = 0$, we obtain from Theorem 3.2,

$$(P^n)_{0,0} = \sum_{k=0}^{\lfloor \frac{n}{2} \rfloor} s(n + 1 - 2k, k) p^k q^{n-k},$$

but a simpler expression can be obtained using the expression of $(\Phi_P(z))_{0,0}$ as follows. If we set $j = 0$ in Relation (11), we have

$$(\Phi_P(z))_{0,0} = \frac{\eta(z)}{1 - qz\eta(z)}$$

and the relation $\eta(z) = 1 + pqz^2\eta^2(z)$ leads to

$$(\Phi_P(z))_{0,0} = \frac{1 - pz\eta(z)}{1 - z}$$

which gives

$$(\Phi_P(z))_{0,0} = \sum_{n=0}^{+\infty} \left(1 - \sum_{k=0}^{\lfloor \frac{n-1}{2} \rfloor} \binom{2k}{k} \frac{p^{k+1}q^k}{k+1} \right) z^n$$

with the convention that when an index decreases in a sum the value of that sum is 0. With this convention, we have

$$(P^n)_{0,0} = 1 - \sum_{k=0}^{\lfloor \frac{n-1}{2} \rfloor} \binom{2k}{k} \frac{p^{k+1}q^k}{k+1}.$$

An interpretation of this result is the following. Let NBP denote the number of customers served during a busy period. Using the distribution of NBP (see for instance [12, page 65]), that is, for $k \geq 1$

$$\mathbb{P}(NBP = k) = \binom{2k-2}{k-1} \frac{p^{k-1}q^k}{k},$$

we obtain

$$(P^n)_{0,0} = 1 - \frac{p}{q} \mathbb{P}(NBP \leq n).$$

It is shown in [13] that the right-hand side of this last relation is the probability that the n th customer will find the queue empty on its arrival.

Using now the expression of $(P^n)_{0,0}$, we get

$$P_{0,0}(t) = \sum_{n=0}^{+\infty} e^{-(\lambda+\mu)t} \frac{(\lambda+\mu)^n t^n}{n!} \left(1 - \sum_{k=0}^{\lfloor \frac{n-1}{2} \rfloor} \binom{2k}{k} \frac{p^{k+1}q^k}{k+1} \right). \quad (14)$$

If BP denotes the busy period of the M/M/1 queueing system, it is known that

$$P_{0,0}(t) = 1 - \frac{\lambda}{\mu} \mathbb{P}(BP \leq t)$$

(see [1, page 153].) Relation (14) gives then an expression of the busy period distribution:

$$\mathbb{P}(BP \leq t) = \sum_{n=1}^{+\infty} e^{-(\lambda+\mu)t} \frac{(\lambda+\mu)^n t^n}{n!} \sum_{k=0}^{\lfloor \frac{n-1}{2} \rfloor} \binom{2k}{k} \frac{p^k q^{k+1}}{k+1}.$$

This expression was also obtained and studied in [14].

3.1.1 Limiting behaviour of the transient probabilities

We can verify that the well known limiting behaviour of the $M/M/1$ queue is obtained when t tends to infinity. Let ρ be defined as $\rho = \lambda/\mu = p/q$ if $\lambda < \mu$ and $\rho \equiv 1$ otherwise.

Using Theorem 3.2 for $j = 0$ and the previous relation giving $(P^n)_{0,0}$ we obtain the following relation:

$$1 - \sum_{k=0}^{\lfloor \frac{n-1}{2} \rfloor} \binom{2k}{k} \frac{p^{k+1}q^k}{k+1} = \sum_{k=0}^{\lfloor \frac{n}{2} \rfloor} s(n+1-2k, k) p^k q^{n-k}.$$

If $\pi_{0,j}$ denotes the limit of $P_{0,j}(t)$ when t grows to infinity, we have

$$\begin{aligned} \pi_{0,j} &= \lim_{t \rightarrow +\infty} P_{0,j}(t) = \lim_{n \rightarrow +\infty} (P^n)_{0,j} \\ &= \frac{p^j}{q^j} \lim_{n \rightarrow +\infty} \sum_{k=0}^{\lfloor \frac{n-j}{2} \rfloor} s(n+1-2k, k) p^k q^{n-k} \\ &= \frac{p^j}{q^j} \lim_{n \rightarrow +\infty} \sum_{k=0}^{\lfloor \frac{n}{2} \rfloor} s(n+1-2k, k) p^k q^{n-k} \\ &= \frac{p^j}{q^j} \lim_{n \rightarrow +\infty} \left[1 - \sum_{k=0}^{\lfloor \frac{n-1}{2} \rfloor} \binom{2k}{k} \frac{p^{k+1}q^k}{k+1} \right] \\ &= \frac{p^j}{q^j} (1 - pC(pq)) \\ &= \rho^j (1 - \rho) \end{aligned}$$

which is the well known formula for the stationary behaviour of the $M/M/1$ queue.

3.2 The queue is not initially empty

Using some results obtained in [8] (equations (3), (6), (10) and Theorem 1(b)), we can write $P_{i,j}(t)$, for $i \leq j$, as follows:

$$P_{i,j}(t) = e^{-(\lambda+\mu)t} \left(\frac{\lambda}{\mu} \right)^{(j-i)/2} \left[I_{j-i}(2t\sqrt{\lambda\mu}) - I_{i+j+2}(2t\sqrt{\lambda\mu}) \right] + \left(\frac{\mu}{\lambda} \right)^{i+1} P_{0,i+j+1}(t) \quad (15)$$

where I_k denotes the modified Bessel function of the first kind, that is,

$$I_k(x) = \sum_{m=0}^{+\infty} \frac{(x/2)^{2m+k}}{m!(k+m)!}.$$

So, the general case reduces to the case of the queue initially empty by means of Relation (15). However, it may be interesting to show how to derive (15) from the generating function expression. Combining Relations (10) and (11), we can write for all $0 \leq i \leq j$,

$$(\Phi_P(z))_{i,j} = (\Phi_Q(z))_{i,j} + \frac{q^{i+1}}{p^{i+1}} (\Phi_P(z))_{0,i+j+1} \quad (16)$$

where Q is the matrix defined in the proof of Theorem 2.3 by

$$Q = pV + qW.$$

To obtain an expression of $P_{i,j}(t)$ in the general case, it remains to perform the inversion of Φ_Q . Let us consider the so-called absorbing process associated with X , obtained by making state 0 absorbing. For every i, j , $(Q^n)_{i,j}$ is the probability to go from state $i+1$ to state $j+1$ in n transitions on the absorbing process. The potential kernel, $\Phi_Q(z)$, of Q is equal to $\eta(z)Y(z)$ (see Theorem 2.3). It follows that, for $i \leq j$,

$$(\Phi_Q(z))_{i,j} = \sum_{k=0}^i p^{k+j-i} q^k z^{2k+j-i} \eta(z)^{2k+j-i+1} \quad (17)$$

which, using Lemma 3.1, can be written

$$(\Phi_Q(z))_{i,j} = \sum_{n=0}^{+\infty} \sum_{k=0}^i s(2k+j-i+1, n) p^{n+k+j-i} q^{n+k} z^{2n+2k+j-i}.$$

In the last sum, the variable change $n+k \rightarrow k$ leads to

$$(\Phi_Q(z))_{i,j} = \sum_{n=0}^{+\infty} \sum_{k=n}^{n+i} s(2k-2n+j-i+1, n) p^{k+j-i} q^k z^{2k+j-i}.$$

Exchanging first the order of summations and exchanging then the indexes n and k , we obtain

$$\begin{aligned} (\Phi_Q(z))_{i,j} &= \sum_{n=0}^{i-1} \sum_{k=0}^n s(2n-2k+j-i+1, k) p^{n+j-i} q^n z^{2n+j-i} \\ &+ \sum_{n=i}^{+\infty} \sum_{k=n-i}^n s(2n-2k+j-i+1, k) p^{n+j-i} q^n z^{2n+j-i}. \end{aligned}$$

It follows that the only non zero values of the sequence $(Q^n)_{i,j}$ are the following:

$$(Q^{2n+j-i})_{i,j} = \begin{cases} p^{n+j-i} q^n \sum_{k=0}^n s(2n-2k+j-i+1, k) & \text{if } n \leq i \\ p^{n+j-i} q^n \sum_{k=n-i}^n s(2n-2k+j-i+1, k) & \text{if } n \geq i+1. \end{cases}$$

In order to simplify this expression, we use the following result which is immediat to prove:

$$\sum_{k=0}^l s(h-2k+1, k) = \binom{h}{l} \text{ for every } l \geq 0 \text{ and } h \geq 2l$$

We obtain that

$$(Q^{2n+j-i})_{i,j} = \begin{cases} p^{n+j-i} q^n \binom{2n+j-i}{n} & \text{if } n \leq i \\ p^{n+j-i} q^n \left[\binom{2n+j-i}{n} - \binom{2n+j-i}{n-i-1} \right] & \text{if } n \geq i+1. \end{cases}$$

Putting these results into Relation (16) we obtain, after some algebra, the general expression (15).

4 Conclusion

In this paper we have derived analytical formulae giving the transient behavior of the $M/M/1$ queueing system. These formulae are quite simple, especially when the queue is initially empty. Moreover, we get a simple expression of the probability to go from state i to state j in n transitions on the absorbing process obtained by making state 0 absorbing. Further work could be the extension of this method to more general queueing systems such as the $M/M/c$ queue.

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615 rue de Jardin Botanique, BP 101, 54600 VILLERS LÈS NANCY
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