

# Predicting survival and trial outcome in non-small cell lung cancer integrating tumor and blood markers kinetics with machine learning

Sébastien Benzekry, Mélanie Karlsen, Célestin Bigarré, Abdessamad El Kaoutari, Bruno Gomes, Martin Stern, Ales Neubert, Rene Bruno, François Mercier, Suresh Vatakuti, et al.

# ▶ To cite this version:

Sébastien Benzekry, Mélanie Karlsen, Célestin Bigarré, Abdessamad El Kaoutari, Bruno Gomes, et al.. Predicting survival and trial outcome in non-small cell lung cancer integrating tumor and blood markers kinetics with machine learning. 2023. hal-04375606

# HAL Id: hal-04375606 https://inria.hal.science/hal-04375606v1

Preprint submitted on 5 Jan 2024

**HAL** is a multi-disciplinary open access archive for the deposit and dissemination of scientific research documents, whether they are published or not. The documents may come from teaching and research institutions in France or abroad, or from public or private research centers. L'archive ouverte pluridisciplinaire **HAL**, est destinée au dépôt et à la diffusion de documents scientifiques de niveau recherche, publiés ou non, émanant des établissements d'enseignement et de recherche français ou étrangers, des laboratoires publics ou privés.



Distributed under a Creative Commons Attribution - ShareAlike 4.0 International License

# Predicting survival and trial outcome in non-small cell lung cancer integrating tumor and blood markers kinetics with machine learning

Sébastien Benzekry<sup>o</sup> <sup>1</sup> <sup>[]</sup>, Mélanie Karlsen<sup>1</sup>, Célestin Bigarré<sup>o</sup> <sup>1</sup>, Abdessamad El
 Kaoutari<sup>o</sup> <sup>1</sup>, Bruno Gomes<sup>2</sup>, Martin Stern<sup>3</sup>, Ales Neubert<sup>4</sup>, Rene Bruno<sup>5</sup>, François

<sup>8</sup> Mercier<sup>6</sup>, Suresh Vatakuti<sup>7</sup>, Peter Curle<sup>8</sup>, Candice Jamois<sup>9</sup>

# For correspondence: sebastien.benzekry@inria.fr

Present address: COMPO team, Pharmacy faculty 27 Bd Jean Moulin 13385 Marseille, FRANCE

Data availability: Qualified researchers may request access to individual patient level data through the clinical study data request platform (https://vivli.org/). Further details on Roche's criteria for eligible studies are available here (https://vivli.org/ members/ourmembers/). For further details on Roche's Global Policy on the Sharing of Clinical Information and how to request access to related clinical study documents, see here https://www.roche.com/ innovation/process/ clinical-trials/data-sharing/.

Funding: This work was sponsored by the Roche Pharma Research and Early Development (pRED) One-D Modeling and Simulation Digital Initiative. It also benefited from funding from ITMO Cancer AVIESAN and French Institut National du Cancer (grant #19CM148-00)

**Competing interests:** The authors declare the existence of a financial competing interest

<sup>1</sup>COMPutational pharmacology and clinical Oncology Department, Inria Sophia q Antipolis-Méditerranée, Cancer Research Center of Marseille, Inserm UMR1068, CNRS 10 UMR7258, Aix Marseille University UM105, Marseille, France; <sup>2</sup>Pharma Research and 11 Early Development, Early Development Oncology, Roche Innovation Center Basel, 12 Switzerland; <sup>3</sup>Pharma Research and Early Development, Early Development Oncology, 13 Roche Innovation Center Zurich, Switzerland; <sup>4</sup>Pharma Research and Early 14 Development, Data & Analytics, Roche Innovation Center Basel, Switzerland; <sup>5</sup>Modeling 15 and Simulation, Clinical Pharmacology, Genentech Research and Early Development, 16 Marseille France; <sup>6</sup>Modeling and Simulation, Clinical Pharmacology, Genentech 17 Research and Early Development, Roche Innovation Center Basel; <sup>7</sup>Pharma Research 18 and Early Development, Predictive Modeling and Data Analytics, Roche Innovation 19 Center Basel, Switzerland; <sup>8</sup>Inovigate, Basel, Switzerland; <sup>9</sup>Pharma Research and Early 20 Development, Translational PKPD and Clinical Pharmacology, Roche Innovation Center 21

Basel, Switzerland

# 24 Abstract

23

36

Existing survival prediction models rely only on baseline or tumor kinetics data and lack machine 25 learning integration. We introduce a novel kinetics-machine learning (kML) model that integrates 26 baseline markers, tumor kinetics and four on-treatment simple blood markers (albumin, CRP, 27 lactate dehydrogenase and neutrophils). Developed for immune-checkpoint inhibition (ICI) in 28 non-small cell lung cancer on three phase 2 trials (533 patients), kML was validated on the two 29 arms of a phase 3 trial (ICI and chemotherapy, 377 and 354 patients). It outperformed the current 30 state-of-the-art for individual predictions with a test set c-index of 0.790, a 12-months survival 31 accuracy of 78.7% and a hazard ratio of 25.2 (95% CI: 10.4 - 61.3, p < 0.0001) to identify long-term 32 survivors. Critically, kML predicted the success of the phase 3 trial using only 25 weeks of 33 on-study data (predicted HR = 0.814 (0.64 - 0.994) versus final study HR = 0.778 (0.65 - 0.931)). Our 34 model constitutes a valuable approach to support personalized medicine and drug development. 35

37 Introduction

<sup>38</sup> Lung cancer is the leading cause of cancer death worldwide<sup>1</sup>, with non-small cell lung cancer <sup>39</sup> (NSCLC) being the most prevalent type, representing 80% - 85% of case<sup>2</sup>. Immune-checkpoint in-

hibitors (ICI) (e.g., atezolizumab (ATZ)) have led to significant improvements in survival rates for pa-

tients with advanced cancers such as NSCLC<sup>3,4</sup>. However, there is still a large variability in clinical

response and progression eventually occurs in a majority of patients<sup>5</sup>. Additionally, drug devel-

<sup>43</sup> opment in immuno-oncology is highly challenging, with a 95% attrition rate<sup>6</sup>. Current approaches

<sup>44</sup> for go/no-go decisions are based on interim endpoints (e.g., progression-free survival, overall re-

45 sponse rate) that have often been found to be poor predictors of the primary endpoint of most

<sup>46</sup> clinical trials in oncology, overall survival (OS)<sup>7</sup>. This calls for better surrogate markers at interim

analyses. Altogether, there is a need for better and validated predictive models of OS for both
 personalized health care (individual predictions) and drug development (trial predictions).

Currently, PDL1 expression is the only routine biomarker used for NSCLC patients<sup>5,8</sup> despite 49 being controversial<sup>9,10</sup>. Tumor mutational burden<sup>8,11,12</sup> and transcriptomic data<sup>5,13,14</sup> have also 50 been investigated but did not reach clinical practice. Here we posit that such static and single 51 marker approach is intrinsically limited and that substantial additional predictive performances 52 could be gained by: 1) using multi-modal integrative analyses relying on a combination of markers 53 and machine learning algorithms 5,12,14,15 and 2) including dynamic markers obtained from early 54 on-treatment data<sup>15,16</sup>. The nonlinear mixed-effects (NLME) modeling approach is well suited for 55 the latter<sup>17</sup>, and tumor kinetics (TK) model-based metrics have been shown to carry significant 56 predictive value for OS in oncology, including ATZ monotherapy in advanced NSCLC<sup>18-20</sup>. The first 57 main novelty of the current study is to establish the predictive value of model-based parameters 58

of simple blood markers kinetics (BK), in addition to TK.

The second main novelty is to apply machine learning (ML) algorithms, increasingly used in biology and medicine<sup>21</sup> but only rarely for TK-OS modeling<sup>22</sup>, instead of classical survival models. Extensions of classical ML models to survival data have been proposed (e.g., random survival forests<sup>23</sup>), but their actual superiority over standard approaches remains controversial<sup>24</sup>. In addition, most ML studies to date are underpowered due to low sample sizes in both training and test sets.

<sup>66</sup> Here, we coupled the strengths of NLME modeling with ML to derive a predictive model of OS

<sup>67</sup> from baseline and on-treatment data, called kinetics-machine learning (kML, Figure 1A). We lever-

aged large training and test datasets to achieve robust results (Figure 1B). Subsequently, we tested
 the operational predictive capabilities of kML in two relevant scenarios: 1) individual prediction of

the operational predictive capabilities of KML in two relevant scenarios: 1) individual prediction
 OS and 2) prediction of the outcome of a phase 3 trial from early on-study data.

# 71 Methods

# 72 Data

For both training and external validation (testing) sets, patients from French centers were excluded for legal reasons (N = 118, not included in the numbers above). The training set comprised the FIR (NCT01846416)<sup>25</sup>, POPLAR (NCT01903993)<sup>3</sup> and BIRCH (NCT02031458)<sup>26</sup> phase 2 clinical trials. The test set was the atezolizumab arm of the OAK phase 3 trial (NCT02008227)<sup>27</sup> for individual predictions and additionally the docetaxel arm for trial predictions (Supplementary Figure 1). These

<sup>78</sup> studies were conducted in accordance with the Declaration of Helsinki after approval by institu-

tional review boards or independent ethics committees. All patients provided written informed
 consent.

The outcome considered was overall survival (OS), defined as the time between treatment start and death or last follow-up, in which case the data was right-censored. The median follow-up was

<sup>83</sup> 35.2 months (95%CI:34.5–35.7) in the training set and 26.8 months (95%CI:26.3–27.5) in the test set.

# 84 Preprocessing

- 85 Baseline data
- <sup>86</sup> The baseline data consisted of 63 variables spanning demographic and biological data, clinical infor-
- <sup>87</sup> mation and disease status (see Supplementary Figure 2–4 for a description of the main variables).
- 88 PD-L1 expression on tumor cells was measured by immunohistochemistry or quantitative poly-
- <sup>89</sup> merase chain reaction, with four possible levels (0: < 1%; 1:  $\ge$  1%; 2:  $\ge$  5% and 3:  $\ge$  50%)<sup>3</sup>. We refer
- <sup>90</sup> to the above-mentioned identifiers and references for further details on the other variables. Data
- <sup>91</sup> were measured in accordance to the studies principles.
- <sup>92</sup> Tumor and blood markers kinetics (TK and BK)
- 93 Patients with only one baseline SLD measurement and no SLD measurement during the treatment
- period were excluded (N = 110). For BK, first time points prior to treatment start were discarded.
- <sup>95</sup> Then, four exclusion rules were established to identify anomalous data points: 1) values outside
- <sup>96</sup> physiologically possible bounds, 2) duplicates, 3) values that abruptly went to an extreme out-of-
- $_{
  m 97}$  range value between two measurements, 4) only the BK value at the closest time point to treatment
- <sup>98</sup> initiation was kept. Eventually, in order to have sufficient data for Bayesian estimation with early
- <sup>99</sup> data, patients with less than three observations before cycle 5 were removed. We refer to the
- <sup>100</sup> supplementary methods for details.

# <sup>101</sup> Nonlinear mixed-effects modeling

- <sup>102</sup> Population approach
- <sup>103</sup> Statistical hierarchical nonlinear mixed-effects modeling (NLME) was used to implement a popula-
- <sup>104</sup> tion approach<sup>28</sup> for the kinetic data and parameter estimation was conducted using the Monolix
- <sup>105</sup> software<sup>29</sup> Mathematical details are given in the supplementary methods.

# 106 Structural models

<sup>107</sup> Following previous work, the TK structural model was assumed to be the sum of two exponen-<sup>108</sup> tials<sup>19,30</sup>:

$$y_{j}^{i} = \begin{cases} y_{0}^{i} e^{\mathsf{KG}^{i}t} & t \leq 0 \\ y_{0}^{i} (e^{-\mathsf{KS}^{i}t} + e^{\mathsf{KG}^{i}t} - 1) & t > 0 \end{cases}$$

where t = 0 corresponds to treatment initiation and  $y_0$ , KG and KS are three parameters, represent-

ing respectively the baseline value, growth and shrinkage rates. This model was also considered for BK, together with three other models: constant  $(y_i^i = \alpha^i, \forall j)$ , linear  $(y_i^i = \alpha^i + \beta^i t_i^i, \forall j)$  and hy-

perbolic  $(y_j^i = p^i + \frac{e^{l^i}(q^i - p^i)}{t_j^i + e^{l^i}})^{31}$ . Quantitative comparison of goodness-of-fit between models was assessed using the corrected Bayesian information criterion<sup>32</sup>.

# <sup>114</sup> Identification of individual model-based parameters

The population parameters identified on the training set were used to define prior distributions of 115 the TK and BK model parameters. These "training" priors were used for Bayesian estimation (maxi-116 mum a posteriori estimate) of the individual TK and BK model parameters, not only for the training 117 set but also for the test sets, in order to avoid leakage. To focus on the pure kinetic parameters, 118 the model-estimated baseline parameters were not kept. We additionally considered the ratio of 119 the model-predicted value at cycle 3 day 1 to the model-estimated baseline parameter. Altogether, 120 there were three individual parameters for each marker:  $X_{KG}$ ,  $X_{KS}$  and  $X_{ratio}$  for X = TK, CRP, LDH 121 and neutrophils; and albumin<sub>p</sub>, albumin<sub>l</sub> and albumin<sub>ratio</sub> for albumin. 122

# 123 Truncated data: individual-level

- 124 Individual-level truncated datasets were derived from the longitudinal TK and BK data by keeping
- data only up to: cycle 3 day 1 (C3D1, 1.5 months), C5D1 (3 months) and C10D1 (6.75 months). New
- $_{\tt 126}$   $\,$  training priors were estimated from each CXD1 training set. The resulting TK and BK truncated
- model parameter Y for marker X at cycle *i* were denoted by  $X_{Y,i}$  (e.g.,  $Idh_{KG,5}$ ).

- 128 Truncated data: study-level for trial prediction
- 129 Study-level truncated datasets were defined at the following on-study landmark times It after study
- initiation (first patient recruited): lt = 10, 25 and 60 weeks.
- Only the patients enrolled before this time and their data collected up to *lt* was used. Note
- that here t = 0 corresponds to study initiation and thus patients in these datasets have varying
- <sup>133</sup> follow-up duration (from 0 to *lt*), in contrast to individual-level truncated datasets.

# 134 Machine learning

- 135 Data preparation
- <sup>136</sup> Missing values (1.6% total, maximum 12% in one variable) were imputed with the median for nu-
- 137 meric variables and mode for categorical variables, learned on the training set, even when applied
- to the test set. All numeric variables were centered and scaled. Means and standard deviations
- <sup>139</sup> were learned on the train and carried to the test set.

# 140 Models

Model elaboration and development was performed exclusively on the training set, using 10 folds 141 cross-validation for predictive performances evaluation. Due to censoring in the data, survival 142 models were used: proportional hazards Cox regression<sup>33</sup>, extreme gradient boosting (XGB) with 143 either Cox or accelerated failure time (AFT) models<sup>34</sup> and random survival forests (RSF)<sup>23</sup>. Nested 144 cross-validation with inner bagging in each 10-fold cross-validation outer loop was used to evaluate 145 the benefit of tuning the hyperparameters<sup>35</sup>. Improvement of the performances was negligible 146 with hyperparameter tuning (Supplementary Figure 5). Therefore, we used the default values of 147 the hyperparameters. For the final RSF model: number of trees ntree = 500, number of variables 148 to possibly split at each node mtry = 5, minumum size of terminal node *nodesize* = 15, number of 149 random splits for splitting a variable nsplit = 10. 150

151 Evaluation

Predictive performances were assessed for either discrimination (c-index and classification met-152 rics at horizon times  $\tau$ ), calibration (calibration curves) or stratification (dichotomized KM survival 153 curves). For each individual, the RSF model gives two prediction outputs: a scalar value termed 154 "mortality" that we will refer to as "ML score", and time-dependent predicted survival curves<sup>23</sup>. The 155 former was used to compute the c-index using the rcorr.cens function of the mmisc R package<sup>36,37</sup>. 156 For prediction of survival at a horizon time  $\tau$ , we used the latter to compute model-predicted prob-157 abilities of death at  $\tau$ . Unless otherwise specified,  $\tau = 12$  months. Survival-adapted metrics of pre-158 dictive performance were used for sensitivity, specificity, area under the receiver-operator curve 159 (ROC AUC) and negative and positive predictive value (NPV and PPV) to account for censoring<sup>38,39</sup>. 160 For computation of accuracy, censored patients before  $\tau$  were discarded (N = 17/396 in the test 161 set at 12 months). The optimal cut-points used for individual OS predictions on the test set were 162 defined as the Kaplan-Meier estimated survival probability in the training set at  $\tau$  (0.257 at 6 months. 163 0.437 at 12 months. 0.634 at 24 months). 164

For patient stratification (dichotomized KM curves), the ML score was used, with models trained on the training set and predicted on the test set. In order to assess stratification abilities to capture the 20% of long-term survivors, cut-points were set at the 20<sup>th</sup> percentiles for each variable/score evaluated. This cut-point arbitrary definition was also motivated by the aim to ensure fair comparison between multiple parameters on the same data. Significance of differences in KM curves was established using the logrank test, and hazard ratios were computed using proportional hazards Cox regression.

- 172 Variable selection and minimal signature
- <sup>173</sup> Variable selection was performed only for the BSL data. The method was based on two steps: 1)
- <sup>174</sup> sorting the variables using least absolute shrinkage and selection operator (LASSO)<sup>40</sup> and 2) build-
- ing RSF incremental models including increasing numbers of variables. LASSO sorting was defined

- as taking the coefficients gradually becoming non-zero during likelihood maximization when the
- 177 regularization parameter decreases. The minimal signature was defined as the minimal set of vari-
- ables able to achieve a c-index larger than 0.75 and an AUC larger than 0.8, with the addition of 4
- <sup>179</sup> well-established prognosis.
- 180 Survival simulations and computation of predicted HRs
- <sup>181</sup> For each patient *i*, one output of the kML model is a survival curve  $S^{i}(t)$ . This gives the cumulative
- distribution function  $1 S^{i}(t)$  of the random variable  $T^{i}$  of the time to death for patient *i*, which
- was used to simulate 100 replicates of  $T^i$ . Pooling all patients together, we thus obtained 100 repli-
- cates of  $\{T^{i, ATZ}, T^{j, DTX}\}$  for *i* and *j* being the patient indices within the ATZ and docetaxel arms,
- respectively. Each replicate then led to 1) a predicted survival curve in each arm and 2) a Cox pro-
- portional hazard HR between the two arms. Taking the mean and the 5<sup>th</sup> and 95<sup>th</sup> percentiles over all replicates vielded the reported point estimate and corresponding 95% prediction interval. The
- all replicates yielded the reported point estimate and corresponding 95% prediction in
   same procedure was used for study-truncated data.
- 189 Data Availability
- <sup>190</sup> Qualified researchers may request access to individual patient level data through the clinical study
- <sup>191</sup> data request platform (https://vivli.org/). Further details on Roche's criteria for eligible studies are
- <sup>192</sup> available here (https://vivli.org/members/ourmembers/). For further details on Roche's Global Policy <sup>193</sup> on the Sharing of Clinical Information and how to request access to related clinical study docu-
- <sup>193</sup> ments, see here https://www.roche.com/innovation/process/clinical-trials/data-sharing/.
- 195 Code availability
- Algorithms used for data analysis are all publicly available from the indicated libraries and refer ences in the Methods section.

# 198 Results

# 199 Data

- The data consisted of advanced NSCLC patients enrolled in ATZ trials (N = 1936, Figure 1B and Supplementary Figure 1). Three ATZ phase 2 trials were pooled into a training dataset<sup>3,25,26</sup> (N = 862). The external validation (test) set comprised data from the ATZ arm (N = 553) of the OAK phase 3 trial<sup>27</sup>. For trial outcome prediction the docetaxel arm (N = 521) was added as an additional test
- 204 set.

Variables comprised baseline (pre-treatment) and longitudinal (on-treatment) data (Figure 1A). 205 The former included: patients and disease characteristics (p = 63 variables, 43 numeric and 20 categorical, denoted BSL) and transcriptomic ("RNAseg", p = 58,311 transcripts) data. The latter 207 included: longitudinal investigator-assessed sum of largest diameters (SLD) of lesions as per the 208 RECIST criteria<sup>41</sup> denoted by tumor kinetics (TK k = 5.473/3.015 time points in the train/test sets re-209 spectively, median 5/4 data points per patient, range 2/2 - 24/20); and longitudinal measurements 210 of four blood markers (albumin, C-reactive protein (CRP), lactate dehydrogenase (LDH) and neu-211 trophils), denoted together as blood markers kinetics (BK, k = 60, 779/38, 460 data points, median 212 11-7-11-11/9-9-9-10 data points per patient, range 3-3-3-3/3-3-3-3 --60-63-63-63-78/82-47-77-89 for 213 albumin-CRP-LDH-neutrophils in the train/test sets, respectively). 214

# <sup>215</sup> Nonlinear mixed-effects modeling (NLME) of longitudinal markers

- <sup>216</sup> We first developed NLME models for the longitudinal data (Figure 1B). The TK structural model was
- the sum of an increasing and a decreasing exponential function (double exponential model)<sup>30</sup>. It
- was able to accurately describe the training data with no goodness-of-fit misspecification (Figure
- <sup>219</sup> 2A and Supplementary Figure 6). Population parameters were estimated with good accuracy (all
- relative standard errors smaller than 9%, Table 1).

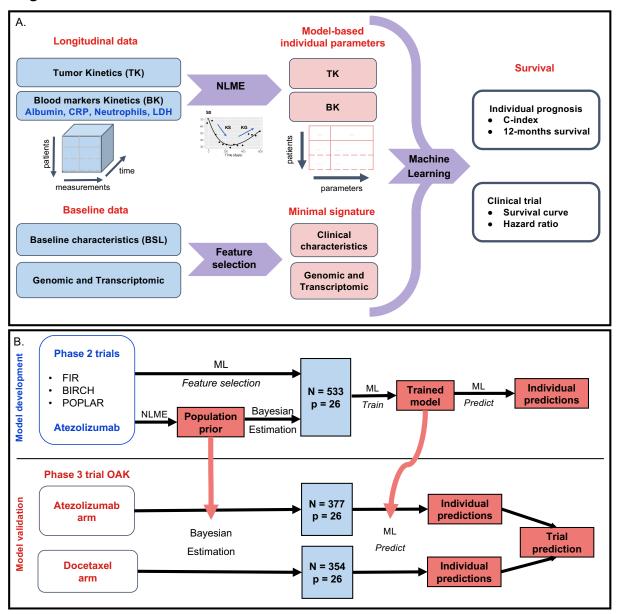


Figure 1: Schematic of the kML framework

Figure 1. Study schematic A. Baseline and longitudinal data were combined into a machine learning algorithm in order to predict individual survival prognosis. Longitudinal data were modelled using nonlinear mixed-effects modelling, whereas machine learning-based feature selection was applied to the baseline data to derive a minimal signature. Tumor kinetics and biological kinetics parameters were combined with the minimal signature to predict survival. Predictive performances were assessed using survival metrics (c-index and survival at horizon times).
B. Algorithm used to develop the model on the train data and carry it to the test set for external validation. Each step — preprocess, learning of the Bayesian priors, dimensionality reduction, feature selection, choice, tuning and training of the machine learning algorithm — were calibrated on the training set and then applied to the test set.

TK: tumor kinetics; BK: blood markers kinetics; ML: machine learning; NLME: nonlinear mixed-effects modelling

Table 1. Parameters from nonlinear mixed-effects modeling of tumor and blood marker kinetics

	К		CR	P	LD	Н	Neutro	phils	Albu	umin	
$KG_{pop}$ (week <sup>-1</sup> )	0.004 92	(6.80)	0.008 14	( 9.38)	0.00238	(10.48)	0.004 36	( 8.69)	p <sub>pop</sub> (g/l)	29.4	(3.82)
$KG_{pop}$ (week <sup>-1</sup> )	0.00778	(8.22)	0.0137	(14.14)	0.001 84	(13.73)	0.000 987	(21.16)	$I_{pop}$ (log (day))	8.09	(2.74)
ω <sub>KG</sub>	1.36	(3.80)	1.61	( 4.25)	1.55	( 5.36)	1.41	( 4.48)	$\omega_p$	0.476	(7.48)
$\omega_{KS}$	1.41	(4.66)	1.81	( 6.29)	1.92	( 5.34)	2.46	( 5.82)	$\omega_l$	0.359	(6.42)
error <sup>1</sup>	6.82	(1.15)	0.559	( 1.23)	0.138	( 0.79)	0.207	( 0.82)	error <sup>1</sup>	0.0549	(0.77)

Parameter value (relative standard error (%)). TK: constant error, others: proportional error. CRP : C-reactive protein; LDH : lactate deshydrogenase.

To analyze the BK data, we first investigated whether significant kinetic patterns could be ob-221 served beyond random noise (due to, e.g., measurement errors, see raw data in Supplementary 222 Figures 7–10). The latter was considered as the null hypothesis, described by a constant model. 223 It was tested against three alternative empiric models: linear, hyperbolic (monotonous but non-224 linear and saturating) and double-exponential (nonlinear and non-monotonous). For all four BKs. 225 we found significant kinetics compared with the constant model, as shown by lower corrected Bayesian information criterion and relative error between model fits and data (Supplementary Fig-227 ure 11). The best descriptive models were hyperbolic for albumin and double-exponential for the 228 other BKs. Individual fits to patient kinetics with the best models showed substantial descriptive 220 power (Figure 2A), which was confirmed by data versus model fits plots (Supplementary Figures 230 12–15). Parametric identifiability of population parameters was excellent for all models (Table 1). 231

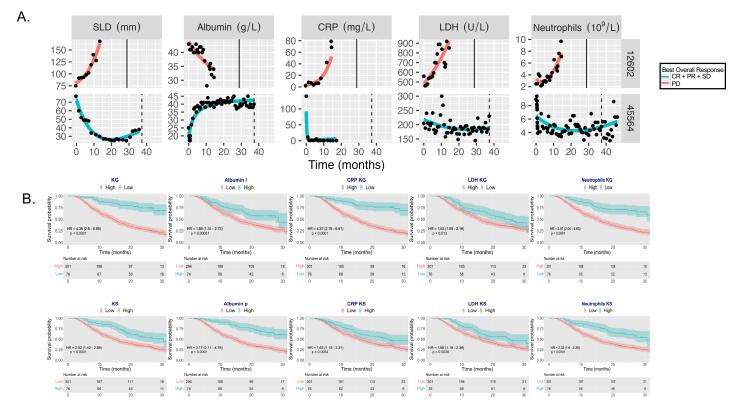
We further assessed the stratification value of the individual model-based kinetic marker for OS prognosis (Figure 2B). The TK parameter KG (growth rate) exhibited good stratifying ability (HR = 4.39 (2.8 – 6.89)), which was similar to the CRP<sub>KG</sub> parameter (HR = 4.37 (2.76 – 6.91)). Ranked by HR importance; (controlled by the 20<sup>th</sup> percentile definition of the cut-point, see methods), the following four best parameters were albumin<sub>p</sub> (HR = 3.17 (2.11 – 4.78)), neutrophils<sub>KG</sub> (HR = 3.07 (2.04 – 4.63)), neutrophils<sub>KS</sub> (HR = 2.33 (1.6 – 3.39)) and TK<sub>KS</sub> (HR = 2.02 (1.42 – 2.89)). All kinetic parameters carried substantial prognostic power (p < 0.0001, log rank test).

For TK and BKs we complemented the initial model parameters with an additional metric that
 was considered valuable for early prediction: the model-predicted ratio of change over baseline at
 cycle 3 day 1.

# 242 Survival prediction using kinetics-machine learning (kML): model development

Four feature sets resulted from the analysis above: BSL, RNAseq, TK and BK (Figure 1A). The development of a kinetics-machine learning (kML) comprised two main steps: choice of the algorithm and derivation of a minimal signature (Figure 1B). The first was achieved by benchmarking four models that used all variables (p = 119, N = 553). The random survival forest (RSF) model found to exhibit the best performances (Supplementary Figure 5) and was thus selected. Notably, we found significantly better predictive performances of RSF over a classical Cox proportional hazard regression model (p = 0.0006).

Feature selection on BSL variables was performed building incremental RSF models based on 250 LASSO importance-sorted variables (Figure 3A). The model using all of them achieved the best 251 score. Nevertheless, keeping in mind the objective to ultimately support decision making and pa-252 tient stratification, a minimal (11 features), near-optimal, set of BSL variables was selected and 253 denoted mBSL. It was defined as the first seven variables reaching the plateau (CRP, heart rate, neu-254 trophils to lymphocytes ratio, neutrophils, lymphocytes to leukocytes ratio, liver metastases and 255 ECOG score), complemented with four variables with established prognostic or predictive value 256 and available in routine care: PD-L1 expression (50% cut-off)<sup>3</sup>, hemoglobin<sup>42</sup>, SLD<sup>22</sup> and LDH<sup>43,44</sup>. 257 Applying stringent criteria to the RNAseq data (see supplementary methods), we selected 167 258 transcripts as candidates for final variable selection using Bolasso regression model to identify the 250



# Figure 2: Goodness-of-fit metrics and plots of dynamic BK models

**Figure 2. Goodness-of-fit metrics and plots of dynamic BK models A.** Representative individual fits for the TK and BK best empirical models showing non-trivial kinetic parameters well captured by the dynamic models. Survival is indicated by a vertical line (solid = death, dashed = censored). **B.** Stratified Kaplan-Meier curves at the 20<sup>th</sup> percentile level on the test set, for TK and BK model-based parameters. Missing values were removed in this univariable analysis, explaining the difference of initial number of patients for albumin that had 9 patients in this case. CR: complete response; PR: partial response; SD: stable disease; PD: progressive disease.

#### Table 2. Contingency table for OS prediction at 12 months

	TRUTH									
		Al	ive (0)	de	ead (1)	Total				
MODEL	Alive (-) Dead (+)	182		30		212	(58.7%)			
WODEL	Dead (+)	48		101		149	(41.3%)			
	Total	230	(63.7%)	131	(36.3%)	361				

Note: 16/377 censored patients with survival time  $\leq$  12 months removed for computation of accuray. sensitivity, specificity, PPV and NPV don't correspond exactly to the numbers because they are computed from KM estimate, thus adjusting for censoring bias.

<sup>260</sup> optimal set of predictors<sup>45</sup>. Finally, we ended up with 52 RNAseq variables that corresponded to <sup>261</sup> the highest average c-index of 0.64.

We then compared the cross-validated c-index of each feature set on the train data (Figure 262 3B). Because of negligible discrimination performances (*c-index* = 0.62 + 0.050) and non-systematic 263 availability of those data, the RNAseg set was removed from the model. The selected set of clinical 264 data at baseline (mBSL) exhibited moderate discrimination performances (*c-index* = 0.710 + 0.038). 265 which was slightly outperformed by the TK set (*c-index* = 0.723 + 0.025). Interestingly, the BK set 266 significantly outperformed both baseline clinical and TK (*c-index* =  $0.793 \pm 0.038$ , p = 0.0004 and 0.0005267 respectively. Student's t-test). Jointly, mBSL, TK and BK performed significantly better than any 268 feature set alone (*c-index* = 0.824 + 0.050, p = 0.00007, 0.0002 and 0.055), as well as any combination 269 of two sets among the three (mBSL + TK: c-index = 0.77 + 0.026, mBSL + BK: c-index = 0.81 + 0.027. 270 TK + BK: *c-index* = 0.80 + 0.049). The resulting model combining mBSL. TK and BK was denoted kML 271 (kinetics-machine learning). 272 During cross-validation on the training set, kML exhibited excellent predictive performances 273

across multiple metrics, with minimal between-folds variability ( $AUC = 0.919 \pm 0.056$ , accuracy = 0.873 \pm 0.052, Figure 3C).

## 276 External validation

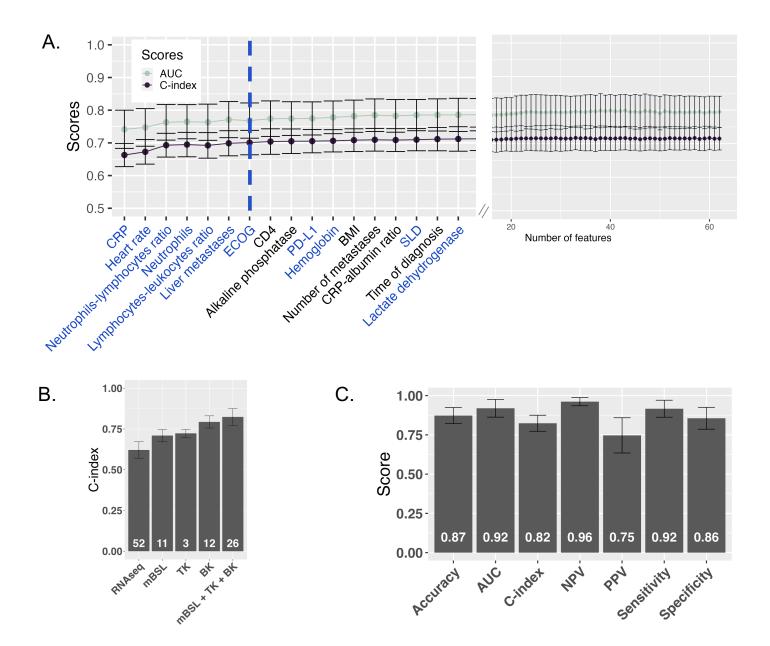
The predictive performance of the final kML model (mBSL, TK and BK) was assessed on the ATZ 277 test set (377 patients). At the population level, the model-predicted survival curve was in excellent 278 agreement with the observed data (Figure 4A). Notably, the prediction interval from the model 279 was narrow, indicating high precision. At the individual level, consistent with the cross-validation 280 results, substantial discrimination performances were observed (*c-index* = 0.790, accuracy and AUC 281 for 12-months survival probability 0.787 and 0.874, respectively. Figure 4B), All classification metrics 282 for prediction of survival at 12 months were high (> 0.78), except PPV, indicating worse ability to 283 predict death than survival. Although smaller, they were similar to the cross-validation results. 284 In addition, calibration curves revealed good performance, at multiple horizon times (Figure 285 4C). Model-predicted probabilities were concordant with the observed KM estimates of the survival 286

probabilities, over the entire range of the binned predicted probabilities. This is further illustrated
by the contingency Table 2. For instance, among 212 patients predicted to be alive at 12 months,
182 (85.8%) were actually alive. Predictive AUC was good at other horizon times (0.846 and 0.910 at
6 and 24 months, respectively, Supplementary Figure 16). However, PPV and sensitivity were very

low at 6 months.

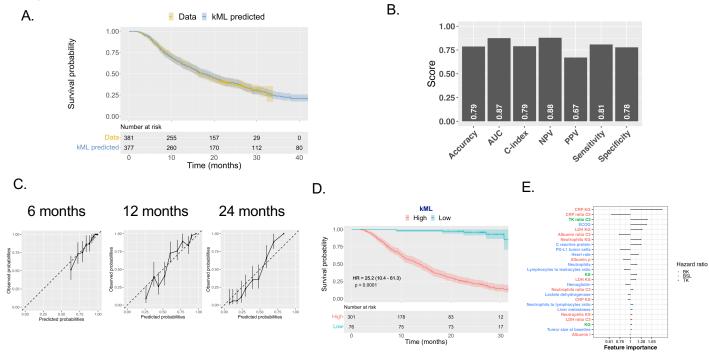
Notably, the kML mortality score derived from the model and learned on the training set was able to accurately stratify OS in the test set (HR = 25.2 (10.4-61.3), p < 0.0001, Figure 4D), indicating excellent ability to identify the 20% of long-term survivors. It outperformed all single kinetic markers (Figure 2C).

<sup>296</sup> Variables importance was assessed by running a post-hoc multivariable Cox regression (Figure



# Figure 3: Minimal baseline (mBSL) signature and kinetics-ML (kML) model

**Figure 3. Minimal baseline (mBSL) signature and kinetics-ML (kML) model A.** Cross-validated (CV) performance scores on the training set (c-index and AUC, mean ± standard deviation) for incremental random survival forest (RSF) models using an increasing number of baseline clinical and biological variables sorted by LASSO importance. The dashed blue line shows the minimal number of variables reaching the plateau. Blue-colored variables correspond to the minimal clinical signature (mBSL). **B.**Comparative CV c-indices of RSF models based either on RNAseq, mBSL, TK, BK and mBSL + TK + BK (final model, kML) variables showing increased predictive performances over baseline when using model-based parameters of kinetic markers. Numbers on the bars indicate the number of variables. **C.** CV performances of the kML model for discrimination (c-index) and classification (survival prediction at 12-months OS).



# Figure 4: Predictive performances of kML on the ATZ test set

**Figure 4. Predictive performances of kML on the ATZ test set A.** Comparison of the population-level survival curves between the data (KM estimator) and the model prediction. **B.** Scores of discrimination metrics. Classification metrics were computed for prediction of OS at 12 months. **C.** Calibration curves at 6, 12 and 24 months, showing the observed survival probabilities (with KM 95% confidence interval) versus the predicted ones in 10 bins corresponding to the model-predicted survival probability deciles. Dashed line is the identity. **D.** Dichotomized KM survival curves based on the ML model-predicted score (high versus low), at the 20<sup>th</sup> percentile cut-off. **E.** Variables importance (multivariable hazard ratios) in the full time-course kML model.

<sup>297</sup> 4F). Interestingly, the top two variables were BKs ( $CRP_{KG}$  and CRP ratio C3). In addition, TK and <sup>298</sup> BK made up for six out of the seven top important features and were found more important than <sup>299</sup> PD-L1.

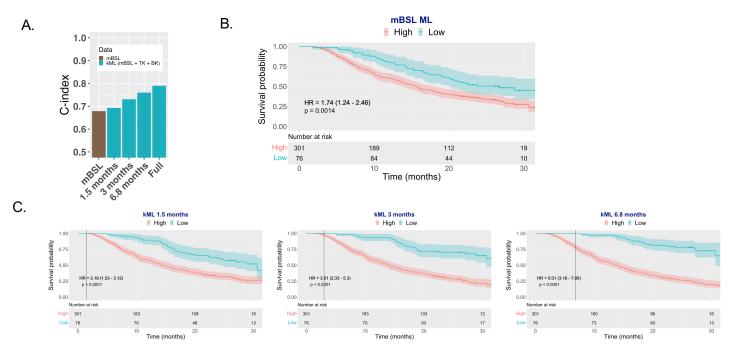
Given the large sample size of our data, we further assessed the model performances when trained on smaller data sets (Supplementary Figure 17). The learning curve revealed that approximately 200 patients were necessary to reach similar performance to the ones obtained with the full training set (N = 533), for both cross-validation and external validation on the test set (*c-index* =  $0.82 \pm 0.056$  vs *c-index* =  $0.82 \pm 0.050$  in cross-validation, 0.78 vs 0.79 on the test set, models trained with 200 vs 533 patients, respectively). Trained with only 60 patients, kML reached already good performances (*c-index* =  $0.76 \pm 0.15$  and 0.74 in cross-validation and test, respectively).

Together, these results demonstrate important predictive performances of overall survival following ATZ treatment using the kML model.

# <sup>309</sup> Application to individual survival prognosis from early on-treatment data

Results above required full on-treatment time-course data to compute TK and BK markers, thus cannot be used to make early predictions. To investigate the operational applicability of our method-

- ology, data from the test set were truncated at the beginning of treatment cycles number 3, 5 and
- <sup>313</sup> 10, respectively corresponding to 1.5, 3 and 6.75 months. We found that integrating longer on-
- treatment data in kML, the predictive performances steadily increased (Figure 5A and Supplemen-
- tary Figure 18). Using the baseline variables only (mBSL), the stratification ability was significant but moderate (HR = 1.74 (1.24 - 2.46), p = 0.0014, Figure 5B). In contrast, kML exhibited increasing
- statification ability from data at 1.5 months (HR = 2.19 (1.53 3.12), p < 0.001), 3 months (HR = 3.51



# Figure 5: Individual-level predictions of kML from cycle-truncated data

**Figure 5. Predictive value of kML from cycle-truncated data A.** Predictive power (c-index) of ML models using baseline (BSL) or truncated data at 1.5, 3 and 6.8 months as well as the full time-course. **B.** Stratified KM survival curves using a RSF model trained on the minimal baseline (mBSL) variables. **C.** Stratified KM survival curves using kML from 1.5 months (2 cycles), 3 months (4 cycles) and 6.8 months (9 cycles) truncated data. Truncation time is indicated by the vertical line.

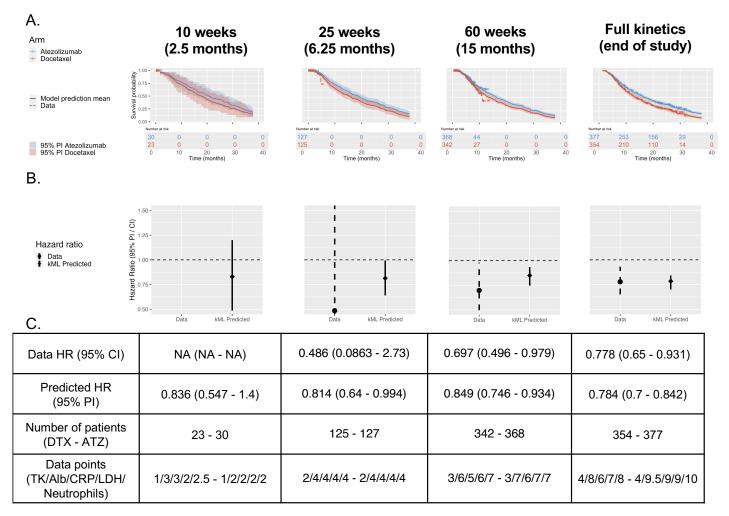
TK: tumor kinetics; BK: biological kinetics; LDH: lactate dehydrogenase; CRP: C-reactive protein.

- $_{318}$  (2.33 5.3), p < 0.0001) and 6.8 months (HR = 5.01 (3.16 7.95), p < 0.0001), see Figure 5C.
- <sup>319</sup> Further investigation of the predictive performances of individual kinetic markers revealed
- that TK parameters were the most informative at 6 weeks (1.5 months, first imaging assessment).
- Adding BKs to TKs brought additional predictive value starting at 3 months, and BKs outperformed
- <sup>322</sup> TK from 6.75 months on (Supplementary Figure 19A). Among BKs, neutrophils kinetics appeared to
- <sub>323</sub> be the most predictive, followed by CRP, albumin and LDH. However, the combined BK signature
- <sup>324</sup> outperformed each individual BK, indicating that their collective predictive capabilities were not <sup>325</sup> driven by any single biomarker alone.

Interestingly, the most important variable at 1.5 months was a kinetic one, TK ratio C3 with following variables being from mBSL (e.g., liver metastases, PDL1 and ECOG). When more ontreatment variables become available, this shifted to TK and BK (TK ratio C3, TK<sub>KS</sub>, TK<sub>KG</sub>, CRP<sub>KG</sub>, LDH<sub>KG</sub>), see Supplementary Figure 19B.

# 330 Application to clinical trial outcome prediction from early on-study data

The kML model can also be applied for the prediction of the outcome of a clinical trial (survival 331 curves and associated hazard ratio), from early on-study data. We performed on-study runca-332 tions on the test set based on a number of weeks after the date of the first patient recruited (see 333 methods). Here, we applied the model to predict not only patients receiving ATZ, but also doc-334 etaxel (Figure 1B). Predictions of the kML model applied to each arm yielded very accurate results 335 when using data from the entire study (predicted HR = 0.784 (0.7 - 0.842)), versus data HR = 0.778336 (0.65 – 0.931), Figure 6A–B). Notably, the model prediction intervals were narrower than the data 337 Kaplan-Meier confidence intervals, probably because the kML-trained model incorporates the in-338 formation from the three phase 2 trials. Using only early data, the model was already able to detect 330



# Figure 6: Use of kML for early prediction of the outcome of a clinical trial

Figure 6. Use of kML for early-prediction of the outcome of a clinical trial A. Survival curves model-based predictions and prediction intervals versus actual data from on-study data at multiple horizon times after study initiation. Note that the model is able to predict full survival curves even if based on early kinetics. B. Compared data and kML-predicted hazard ratios. C. Description of hazard ratios, number of patients and number of data points available in each arm, at the landmark on-study time points.

PI: prediction interval, CI: confidence interval, DTX: docetaxel arm, ATZ: atezolizumab arm.

- a (non-significant) tendency at 10 weeks, with only 23 and 30 patients in each arm, and very short 340
- follow-up. Starting from data available at 25 weeks (6.25 months), the model correctly predicted 341
- a positive outcome of the study, with a 95% prediction interval of the HR below 1. Of note, the 342
- available data at this time (dashed lines, Figure 6A and red HR CIs in Figure 6B) was far from being 343
- conclusive. The model prediction was stable from 25 weeks on whereas the OS data only exhib-344
- ited significant HR starting from 60 weeks and required more than 300 patients in each arm to be 345 conclusive. 346

#### Discussion 347

- Blood markers from hematology and biochemistry are routinely collected during clinical care or 348
- drug trials. They are cost-effective and easily obtained both before and during treatment. There 349
- is limited exploration regarding the predictive capabilities of the kinetics of such data. Combining 350
- BSL variables with on-treatment data (TK and BK), we addressed this guestion using a novel hy-351
- brid NLME-ML methodology. The resulted kML model demonstrated excellent predictive perfor-352

mances for OS in two aspects: 1) patient-level predictions (discrimination, calibration and patient

<sup>354</sup> stratification) and 2) trial-level predictions. The kML model outperformed current state-of-the-art

methods based on either baseline or on-treatment data alone, utilizing only routine clinical infor-

mation, with a c-index of 0.79 and an accuracy of 78% for prediction of 12-month survival, on the

1357 test dataset. Overall, kML incorporates 26 features, out of which 15 features require monitoring

five quantities over time (tumor size, albumin, CRP, LDH and neutrophils).

358

Regarding baseline markers, the predictive value of PD-L1 expression, commonly used in clini-359 cal care, is controversial<sup>9,10</sup>. Previous studies reported an AUC for durable response of 0.601 and 360 a PFS HR of 1.90 (PD-L1 > 1% vs  $0\%)^8$ . Baseline tumor mutational burden showed similar predic-361 tive value initially  $(AUC = 0.646)^{11}$ , but led to disappointing results in a recent prospective studv<sup>46</sup>. 362 Baseline blood counts were previously reported to predict overall survival<sup>43,47-49</sup> and treatment 363 response (AUC = 0.74)<sup>42</sup>. The ROPRO score, derived from a large pan-cancer cohort and incorpo-36/ rating baseline clinical and biological data (27 variables) achieved a c-index of 0.69 and a 3-months 365 AUC of 0.743 for prediction of survival in the OAK clinical trial<sup>50</sup>. Here, we confirmed these find-366 ings and established a minimal signature of such data composed of only 11 variables (CRP, heart 367 rate, neutrophils to lymphocytes ratio, neutrophils, lymphocytes to leukocytes ratio, liver metas-368 tases, ECOG, PD-L1 > 50%, hemoglobin, SLD and LDH), vet with similar predictive performances 360 (*c-index* = 0.678) and significant stratification ability (HR = 1.74, p = 0.0014). Altogether, our kML 370 model demonstrated substantially better predictive performances than these baseline models. 371

We further confirmed the established predictive value of TK model-based parameters<sup>19,20</sup>. Blood-372 or serum-derived longitudinal markers kinetics have to date rarely been modeled. Gavrilov et al. 373 proposed to model NLR kinetics and demonstrated improved OS predictions over TK alone<sup>31</sup>. Here 374 we extended to four BKs: albumin, CRP, LDH and neutrophils. This choice was not only motivated 375 by observed statistical associations, but also from biological considerations. Albumin is associated 376 with nutritional status (cachexic state) and is known to evolve with time in responders. CRP is a 377 marker of systemic inflammation<sup>44</sup>. Increased CRP, decreased albumin level, and increased CR-378 P/albumin ratio have been reported to be associated with poor survival<sup>51</sup>. Neutrophils play a role 379 in inflammation by promoting a favorable microenvironment for cancer cell growth and spread. 380 and activation of carcinogenic signaling pathways<sup>52</sup>. Elevated LDH levels are a marker of cancer 381 cells turnover rate, and LDH has a potential role for prediction of potential invisible metastases<sup>44</sup> 382 We found that all these markers had non-trivial on-treatment kinetics. However, data fits were not 383 perfect, possibly due to the simplicity and empiric nature of the models we used. Further mecha-384 nistic modeling of the joint kinetics of BKs and TK could bring relevant biological information and 385 vield more accurate predictive parameters. We found that all four BKs were contributive to the 386 model and that, combined, they outperformed TK performances. 387

We analyzed the RNAseq data using standard methods and found only negligible predictive performances. Such result could be explained by the fact that the tissue of origin that was used was heterogeneous across the patients (primary tumor or metastasis), was limited to a local area of the tumor, and could come from tissue sampled long before treatment initiation. Given that our main objective was to derive a predictive model from markers available in routine practice, we excluded it from our minimal signature. A refined analysis, especially focusing on immune-based signatures, could improve our results<sup>5</sup>.

Machine learning models, although increasingly used in pharmacological studies —including 395 recently for TK-OS modeling and variable selection<sup>22,53</sup> —have vet rarely been rigorously compared 396 to classical statistical models<sup>24</sup>. Here, such comparison revealed significantly better performance 397 of the nonlinear random survival forest RSF model compared to the linear proportional hazards 398 Cox model. In our approach, we did not use the propagation of standard statistical quantification 390 of the parameters' estimates uncertainty to evaluate the accuracy of the model predictions. Rather, 400 we relied on the RSF-outputted individual survival curves to sample virtual individuals and compute 401 prediction intervals. 402

A drawback of classical TK–OS studies is that they make use of the full observed kinetics to

<sup>404</sup> predict overall survival, which can lead to time-dependent covariate bias<sup>54</sup> and limit their practical

405 applicability at bedside. We used individual-truncated data sets and found that kML was already

<sup>406</sup> improving predictions over mBSL using data at 1.5 months, which corresponds to the first imag-

ing assessment of the treatment effect. At later times, stratification abilities increased to highly significant levels (e.g., HR = 5 at 6.8 months).

A strength of our study is that we relied on well-curated data with high number of patients from
 clinical trials. However, when extrapolating to other settings — earlier trial phases, real-world data
 — limited number of patients might be available. Yet, we found that using only 60 patients to train
 kML was sufficient to reach near-optimal performances.

Not only kML has value for personalized health care, but it also revealed useful for prediction 413 of a phase 3 trial using early on-study data. Our model was able to predict the study's positive 414 outcome with data at  $\sim 6$  months, versus 10 months using TK only<sup>19</sup>. Relying on the data alone. 415 such positive outcome was only detectable at 15 months. These results could have important im-416 plications for drug development as they could inform earlier on go/no-go decisions. Consequently, 417 this could allow to detect futility more easily and more rapidly during clinical trials, allowing to 418 avoid treating patients with an inefficient investigational treatment and to reassign funds and en-419 ergy to other researches. Of note, in a recent evaluation based on resampling the first-line NSCLC 420 ATZ study IMpower150 to mimic small, short follow up early Phase Ib studies, TK model-based 421 metrics had better operating characteristics to predict Phase III success compared with RECIST 422 endpoints ORR and PES<sup>55</sup>. Extension of such results with the addition of BKs is thus a promising 423 line of research. In addition, kML, trained on ATZ data, yielded excellent predictive abilities for the 424 docetaxel (control) arm. This suggests that the relationships between TK / BK and OS might be 425 drug-independent. In turn, this opens future perspectives in terms of testing kML on drugs with 426 different mechanism of action, or combinations. 427

Further avenues of research comprise the development of integrative models from advanced 428 multi-modal data such as the one collected during the PIONeeR clinical study (NCT03833440)<sup>56,57</sup> 429 that include guantitative image analysis from multiplex immune-histochemistry, genomic and tran-430 scriptomic data, biological and clinical markers. In addition, mechanistic modeling of quantitative 431 and physiologically meaningful longitudinal data (immune-monitoring, vasculo-monitoring, circu-432 lating DNA<sup>12,58-60</sup>, soluble factors<sup>61</sup>, pharmacokinetics, TK and a large number of BKs from either 433 hematology or biochemistry) payes the way to an improved understanding and prediction of mech-434 anisms of relapse to  $ICI^{62}$ . Furthermore, the predictive abilities of kML — at the both individual and 435 study levels — should be evaluated in model-based prospective trials<sup>63</sup>. 436

In conclusion, our study shows that integrating model-based on-treatment dynamic data from
 routine biological markers shows great promise for both personalized health care and early pre diction of the outcome of clinical trials during drug development.

# 440 References

- Bray, F., Ferlay, J., *et al.* Global cancer statistics 2018: GLOBOCAN estimates of incidence and mortality worldwide for 36 cancers in 185 countries. *CA: A Cancer Journal for Clinicians*, 394– 424. ISSN: 1542-4863. doi:10.3322/caac.21492 (2018).
- Duma, N., Santana-Davila, R. & Molina, J. R. Non–Small Cell Lung Cancer: Epidemiology, Screen ing, Diagnosis, and Treatment. *Mayo Clinic Proceedings*, 1623–1640. ISSN: 0025-6196, 1942 5546. doi:10.1016/j.mayocp.2019.01.013 (2019).
- Fehrenbacher, L., Spira, A., *et al.* Atezolizumab versus docetaxel for patients with previously
   treated non-small-cell lung cancer (POPLAR): a multicentre, open-label, phase 2 randomised
   controlled trial. *The Lancet*, 1837–1846. ISSN: 0140-6736, 1474-547X. doi:10.1016/S0140-6736(16)
- controlled trial. *The Lancet*, 1837–1846. ISSN: 0140-6736, 1474-547X. doi:10.1016/S0140-6
   00587-0 (2016).

- Grant, M. J., Herbst, R. S. & Goldberg, S. B. Selecting the optimal immunotherapy regimen in driver-negative metastatic NSCLC. *Nature Reviews Clinical Oncology*, 625–644. ISSN: 1759-4782. doi:10.1038/s41571-021-00520-1 (2021).
- Camidge, D. R., Doebele, R. C. & Kerr, K. M. Comparing and contrasting predictive biomarkers
   for immunotherapy and targeted therapy of NSCLC. *Nature Reviews Clinical Oncology*, 341–355.
   ISSN: 1759-4782. doi:10.1038/s41571-019-0173-9 (2019).
- 6. Hutchinson, L. & Kirk, R. High drug attrition rates—where are we going wrong? *Nature Reviews Clinical Oncology*, 189–190. ISSN: 1759-4782. doi:10.1038/nrclinonc.2011.34 (2011).
- Hua, T., Gao, Y., Zhang, R., Wei, Y. & Chen, F. Validating ORR and PFS as surrogate endpoints in phase II and III clinical trials for NSCLC patients: difference exists in the strength of surrogacy in various trial settings. *BMC Cancer*, 1022. ISSN: 1471-2407. doi:10.1186/s12885-022-10046-z
  (2022).
- Rizvi, H., Sanchez-Vega, F., *et al.* Molecular Determinants of Response to Anti–Programmed
   Cell Death (PD)-1 and Anti–Programmed Death-Ligand 1 (PD-L1) Blockade in Patients With
   Non–Small-Cell Lung Cancer Profiled With Targeted Next-Generation Sequencing. *Journal of Clinical Oncology*. doi:10.1200/JCO.2017.75.3384 (2018).
- 9. Doroshow, D. B., Bhalla, S., *et al.* PD-L1 as a biomarker of response to immune-checkpoint inhibitors. *Nature Reviews Clinical Oncology*, 345–362. ISSN: 1759-4782. doi:10.1038/s41571-021-00473-5 (2021).
- So, W. V., Dejardin, D., Rossmann, E. & Charo, J. Predictive biomarkers for PD-1/PD-L1 check point inhibitor response in NSCLC: an analysis of clinical trial and real-world data. *Journal for Immunotherapy of Cancer*, e006464. ISSN: 2051-1426. doi:10.1136/jitc-2022-006464 (2023).
- Hellmann, M. D., Ciuleanu, T.-E., *et al.* Nivolumab plus Ipilimumab in Lung Cancer with a
  High Tumor Mutational Burden. *New England Journal of Medicine*, 2093–2104. doi:10.1056 /
  NEJMoa1801946 (2018).
- Gandara, D. R., Paul, S. M., *et al.* Blood-based tumor mutational burden as a predictor of clinical benefit in non-small-cell lung cancer patients treated with atezolizumab. *Nature Medicine*,
  1441–1448. ISSN: 1546-170X. doi:10.1038/s41591-018-0134-3 (2018).
- 479 13. Cristescu, R., Mogg, R., *et al.* Pan-tumor genomic biomarkers for PD-1 checkpoint blockade–
   480 based immunotherapy. *Science*, eaar3593. doi:10.1126/science.aar3593 (2018).
- 14. Sankar, K., Ye, J. C., *et al.* The role of biomarkers in personalized immunotherapy. *Biomarker Research*, 32. ISSN: 2050-7771. doi:10.1186/s40364-022-00378-0 (2022).
- Acosta, J. N., Falcone, G. J., Rajpurkar, P. & Topol, E. J. Multimodal biomedical Al. *Nature Medicine*,
   1773–1784. ISSN: 1546-170X. doi:10.1038/s41591-022-01981-2 (2022).
- Kurtz, D. M., Esfahani, M. S., *et al.* Dynamic Risk Profiling Using Serial Tumor Biomarkers for
  Personalized Outcome Prediction. *Cell*, 699–713.e19. ISSN: 1097-4172. doi:10.1016/j.cell.2019.
  06.011 (2019).
- Bonate, P. L. *Pharmacokinetic-Pharmacodynamic Modeling and Simulation* 2nd ed. 2011. ISBN:
   978-1-4419-9484-4 (Springer-Verlag New York Inc., New York, 2011).
- Claret, L., Girard, P., *et al.* Model-based prediction of phase III overall survival in colorectal
   cancer on the basis of phase II tumor dynamics. *J Clin Oncol*, 4103–4108. doi:10.1200/JCO.
   2008.21.0807 (2009).
- Claret, L., Jin, J. Y., *et al.* A Model of Overall Survival Predicts Treatment Outcomes with Ate zolizumab versus Chemotherapy in Non-Small Cell Lung Cancer Based on Early Tumor Kinet ics. *Clin Cancer Res*, 3292–3298. doi:10.1158/1078-0432.CCR-17-3662 (2018).
- Benzekry et al. 2024 | Dynamical modeling coupled to machine learning for anti-PDL1 treatment in NSCLC

- <sup>496</sup> 20. Chan, P., Marchand, M., *et al.* Prediction of overall survival in patients across solid tumors fol-
- lowing atezolizumab treatments: A tumor growth inhibition-overall survival modeling frame work. *CPT: pharmacometrics & systems pharmacology,* 1171–1182. ISSN: 2163-8306. doi:10.1002/

<sup>499</sup> psp4.12686 (2021).

- Benzekry, S. Artificial intelligence and mechanistic modeling for clinical decision making in
   oncology. *Clinical Pharmacology & Therapeutics*, 471–486. ISSN: 1532-6535. doi:10.1002/cpt.
   1951 (2020).
- Chan, P., Zhou, X., *et al.* Application of Machine Learning for Tumor Growth Inhibition Overall
   Survival Modeling Platform. *CPT: pharmacometrics & systems pharmacology*, 59–66. ISSN: 2163 8306. doi:10.1002/psp4.12576 (2021).
- Ishwaran, H., Kogalur, U. B., Blackstone, E. H. & Lauer, M. S. Random survival forests. *The Annals of Applied Statistics*, 841–860. ISSN: 1932-6157, 1941-7330. doi:10.1214/08-AOAS169
   (2008).
- <sup>509</sup> 24. Christodoulou, E., Ma, J., *et al.* A systematic review shows no performance benefit of machine
   <sup>510</sup> learning over logistic regression for clinical prediction models. *Journal of Clinical Epidemiology*,
   <sup>511</sup> 12–22. ISSN: 0895-4356. doi:10.1016/j.jclinepi.2019.02.004 (2019).
- Spigel, D. R., Chaft, J. E., *et al.* FIR: Efficacy, Safety, and Biomarker Analysis of a Phase II Open Label Study of Atezolizumab in PD-L1–Selected Patients With NSCLC. *Journal of Thoracic On- cology*, 1733–1742. ISSN: 1556-0864. doi:10.1016/j.jtho.2018.05.004 (2018).
- Peters, S., Gettinger, S., *et al.* Phase II Trial of Atezolizumab As First-Line or Subsequent Therapy for Patients With Programmed Death-Ligand 1-Selected Advanced Non-Small-Cell Lung Cancer (BIRCH). *Journal of Clinical Oncology: Official Journal of the American Society of Clinical Oncology*, 2781–2789. ISSN: 1527-7755. doi:10.1200/JCO.2016.71.9476 (2017).
- Rittmeyer, A., Barlesi, F., *et al.* Atezolizumab versus docetaxel in patients with previously treated non-small-cell lung cancer (OAK): a phase 3, open-label, multicentre randomised controlled trial. *The Lancet*, 255–265. ISSN: 0140-6736, 1474-547X. doi:10.1016/S0140-6736(16)32517-X
   (2017).
- Lavielle, M. *Mixed Effects Models for the Population Approach* ISBN: 1-4822-2650-2 (CRC Press, 2014).
- 29. Lixoft. Monolix version 2020R1. Antony, France, 2020.
- Stein, W. D., Figg, W. D., *et al.* Tumor growth rates derived from data for patients in a clinical trial correlate strongly with patient survival: a novel strategy for evaluation of clinical trial data.
   *The Oncologist*, 1046–1054. doi:10.1634/theoncologist.2008-0075 (2008).
- 31. Gavrilov, S., Zhudenkov, K., *et al.* Longitudinal Tumor Size and Neutrophil-to-Lymphocyte Ratio
   Are Prognostic Biomarkers for Overall Survival in Patients With Advanced Non-Small Cell Lung
   Cancer Treated With Durvalumab. *CPT: pharmacometrics & systems pharmacology*, 67–74. ISSN:
   2163-8306. doi:10.1002/psp4.12578 (2021).
- <sup>533</sup> 32. Delattre, M., Lavielle, M. & Poursat, M.-A. A note on BIC in mixed-effects models. *Electronic* Journal of Statistics, 456–475. ISSN: 1935-7524, 1935-7524. doi:10.1214/14-EJS890 (2014).
- S35 33. Cox, D. R. Regression Models and Life-Tables. *Journal of the Royal Statistical Society. Series B* (*Methodological*), 187–220. ISSN: 0035-9246 (1972).
- S4. Chen, T. & Guestrin, C. XGBoost: A Scalable Tree Boosting System in Proceedings of the 22nd
   ACM SIGKDD International Conference on Knowledge Discovery and Data Mining (Association for
   Computing Machinery, San Francisco, California, USA, 2016), 785–794. ISBN: 9781450342322.
   doi:10.1145/2939672.2939785.
- <sup>541</sup> 35. Cawley, G. C. & Talbot, N. L. C. On over-fitting in model selection and subsequent selection <sup>542</sup> bias in performance evaluation. *Journal of Machine Learning Research*, 2079–2107 (2010).

- 36. Harrell, F. E., Califf, R. M., Pryor, D. B., Lee, K. L. & Rosati, R. A. Evaluating the yield of medical
   tests. *JAMA*, 2543–2546. ISSN: 0098-7484 (1982).
- <sup>545</sup> 37. Harrell, F. E. Hmisc: Harrell miscellaneous. *R package*. https://CRAN.R-project.org/package=
   <sup>546</sup> Hmisc (2022).
- Heagerty, P. J., Lumley, T. & Pepe, M. S. Time-dependent ROC curves for censored survival data and a diagnostic marker. *Biometrics*, 337–344. ISSN: 0006-341X. doi:10.1111/j.0006-341x.
   2000.00337.x (2000).
- Heagerty, P. J. & Saha-Chaudhuri, p. b. P. survivalROC: Time-dependent ROC curve estimation
   from censored survival data tech. rep. (2013). https://CRAN.R-project.org/package=survivalROC.
- Tibshirani, R. Regression Shrinkage and Selection via the Lasso. *Journal of the Royal Statistical Society. Series B (Methodological)*, 267–288. ISSN: 0035-9246 (1996).
- 41. Eisenhauer, E. A., Therasse, P., *et al.* New response evaluation criteria in solid tumours: Revised
   RECIST guideline (version 1.1). *European Journal of Cancer. Response assessment in solid tumours* (*RECIST*): Version 1.1 and supporting papers 228–247. ISSN: 0959-8049. doi:10.1016/j.ejca.2008.
   10.026 (2009).
- Benzekry, S., Grangeon, M., *et al.* Machine Learning for Prediction of Immunotherapy Efficacy
   in Non-Small Cell Lung Cancer from Simple Clinical and Biological Data. *Cancers*, 6210. doi:10.
   3390/cancers13246210 (2021).
- 43. Havel, J. J., Chowell, D. & Chan, T. A. The evolving landscape of biomarkers for checkpoint in hibitor immunotherapy. *Nature Reviews Cancer*, 133–150. ISSN: 1474-1768. doi:10.1038/s41568 019-0116-x (2019).
- <sup>564</sup> 44. Blank, C. U., Haanen, J. B., Ribas, A. & Schumacher, T. N. The "cancer immunogram". *Science*, <sup>565</sup> 658–660. doi:10.1126/science.aaf2834 (2016).
- Bach, F. R. *Bolasso: model consistent Lasso estimation through the bootstrap* in (Association for Computing Machinery, New York, NY, USA, 2008), 33–40. ISBN: 978-1-60558-205-4. doi:10.
   1145/1390156.1390161.
- Peters, S., Dziadziuszko, R., *et al.* Atezolizumab versus chemotherapy in advanced or metastatic
   NSCLC with high blood-based tumor mutational burden: primary analysis of BFAST cohort C
   randomized phase 3 trial. *Nature Medicine*, 1831–1839. ISSN: 1546-170X. doi:10.1038/s41591 022-01933-w (2022).
- Soyano, A. E., Dholaria, B., *et al.* Peripheral blood biomarkers correlate with outcomes in advanced non-small cell lung Cancer patients treated with anti-PD-1 antibodies. *Journal for Immunotherapy of Cancer*, 129. ISSN: 2051-1426. doi:10.1186/s40425-018-0447-2 (2018).
- <sup>576</sup> 48. Diem, S., Schmid, S., *et al.* Neutrophil-to-Lymphocyte ratio (NLR) and Platelet-to-Lymphocyte ratio (PLR) as prognostic markers in patients with non-small cell lung cancer (NSCLC) treated
   <sup>578</sup> with nivolumab. *Lung Cancer (Amsterdam, Netherlands)*, 176–181. ISSN: 1872-8332. doi:10.1016/
   <sup>579</sup> j.lungcan.2017.07.024 (2017).
- Peng, L., Wang, Y., *et al.* Peripheral blood markers predictive of outcome and immune-related adverse events in advanced non-small cell lung cancer treated with PD-1 inhibitors. *Cancer immunology, immunotherapy: Cll*, 1813–1822. ISSN: 1432-0851. doi:10.1007/s00262-020-02585 w (2020).
- Becker, T., Weberpals, J., *et al.* An enhanced prognostic score for overall survival of patients
   with cancer derived from a large real-world cohort. *Annals of Oncology*, 1561–1568. ISSN: 0923 7534. doi:10.1016/j.annonc.2020.07.013 (2020).
- 587 51. Yang, J.-R., Xu, J.-Y., *et al.* Post-diagnostic C-reactive protein and albumin predict survival in
   588 Chinese patients with non-small cell lung cancer: a prospective cohort study. *Scientific Reports*,
   599 8143. ISSN: 2045-2322. doi:10.1038/s41598-019-44653-x (2019).

- 590 52. Bruni, D., Angell, H. K. & Galon, J. The immune contexture and Immunoscore in cancer prog 591 nosis and therapeutic efficacy. *Nature Reviews Cancer*, 662–680. ISSN: 1474-1768. doi:10.1038/
   592 541568-020-0285-7 (2020).
- 53. Liu, G., Lu, J., Lim, H. S., Jin, J. Y. & Lu, D. Applying interpretable machine learning workflow to
   evaluate exposure-response relationships for large-molecule oncology drugs. *CPT: pharmaco- metrics & systems pharmacology*, 1614–1627. ISSN: 2163-8306. doi:10.1002/psp4.12871 (2022).
- 54. Desmée, S., Mentré, F., Veyrat-Follet, C. & Guedj, J. Nonlinear Mixed-Effect Models for Prostate Specific Antigen Kinetics and Link with Survival in the Context of Metastatic Prostate Cancer:
   a Comparison by Simulation of Two-Stage and Joint Approaches. *The AAPS Journal*, 691–699.
   ISSN: 1550-7416. doi:10.1208/s12248-015-9745-5 (2015).
- 55. Bruno, R., Marchand, M., *et al.* Tumor Dynamic Model-Based Decision Support for Phase IbII
   Combination Studies: A Retrospective Assessment Based on Resampling of the Phase III Study
   IMpower150. *Clinical Cancer Research*, OF1–OF9. ISSN: 1078-0432. doi:10.1158/1078-0432.CCR 22-2323 (2023).
- Greillier, L., Monville, F., *et al.* Abstract LB120: Comprehensive biomarkers analysis to explain
   resistances to PD1-L1 ICIs: The precision immuno-oncology for advanced non-small cell lung
   cancer (PIONeeR) trial. *Cancer Research*, LB120. ISSN: 0008-5472. doi:10.1158/1538-7445.
   AM2022-LB120 (2022).
- Barlesi, F., Monville, F., *et al.* Comprehensive biomarkers (BMs) analysis to predict efficacy of
   PD1-L1 immune checkpoint inhibitors (ICIs) in combination with chemotherapy: a subgroup
   analysis of the Precision Immuno-Oncology for advanced Non-Small CEll Lung CancER (PIO NeeR) trial. *Annals of Oncology.* doi:10.1016/iotech/iotech100100 (2022).
- Assaf, Z. J. F., Zou, W., *et al.* A longitudinal circulating tumor DNA-based model associated with
   survival in metastatic non-small-cell lung cancer. *Nature Medicine*, 859–868. ISSN: 1546-170X.
   doi:10.1038/s41591-023-02226-6 (2023).
- 59. Nabet, B. Y., Esfahani, M. S., *et al.* Noninvasive Early Identification of Therapeutic Benefit from Immune Checkpoint Inhibition. *Cell*, 363–376.e13. ISSN: 1097-4172. doi:10.1016/j.cell.2020.09.
  001 (2020).
- 60. Cabel, L., Proudhon, C., *et al.* Clinical potential of circulating tumour DNA in patients receiv ing anticancer immunotherapy. *Nature Reviews. Clinical Oncology*, 639–650. ISSN: 1759-4782.
   doi:10.1038/s41571-018-0074-3 (2018).
- 61. Barrera, L., Montes-Servín, E., *et al.* Cytokine profile determined by data-mining analysis set
   into clusters of non-small-cell lung cancer patients according to prognosis. *Annals of Oncology: Official Journal of the European Society for Medical Oncology*, 428–435. ISSN: 1569-8041. doi:10.
   1093/annonc/mdu549 (2015).
- 625 62. Ciccolini, J., Benzekry, S. & Barlesi, F. Deciphering the response and resistance to immune 626 checkpoint inhibitors in lung cancer with artificial intelligence-based analysis: when PIONeeR
   627 meets QUANTIC. *British Journal of Cancer*, 1–2. ISSN: 1532-1827. doi:10.1038/s41416-020-0918-3
   628 (2020).
- 63. Ciccolini, J., Barbolosi, D., André, N., Barlesi, F. & Benzekry, S. Mechanistic Learning for Com binatorial Strategies With Immuno-oncology Drugs: Can Model-Informed Designs Help Inves tigators? *JCO Precision Oncology*, 486–491. doi:10.1200/PO.19.00381 (2020).

#### **Supplementary Figures** 632

Train + Test			141
Test - OAK GO28915	Phase 3 RCT of ATZ versus docetaxel (DTX) in patients with previously treated NSCLC	Stage IIIB or IV, previously chemo treated	55
Train			86
BIRCH GO28754	Phase 2 study of ATZ in advanced or metastatic NSCLC	Locally advanced or metastatic NSCLC (lines 1, 2 or 3)	59
POPLAR GO28753	Phase 2 randomised controlled trial of ATZ versus docetaxel in NSCLC	Locally advanced or metastatic NSCLC who failed platinum therapy (line 2)	134
FIR GO28625	Phase 2 study for the efficacy and safety of ATZ in advanced NSCLC	PD-L1 positive locally advanced or metastatic NSCLC (lines 1 and 2+)	133
Study	Description	Population	N

Four monotherapy studies of atezolizumab in advanced NSCLC. NSCLC: Non-Small Cell Lung Cancer; p = number of parameters, N: number of patients treated with atezolizumab (patients from French centers were excluded for legal reasons (N=118); In total, data from 1074 patients from OAK were used as Test set (553 from the ATZ arm, 521 from the DTX arm); PD: Pharmacodynamic; SLD: Sum of the Largest Diameters. CRP: C Reactive Protein; LDH: Lactate Dehydrogenase.

#### Supplementary Figure 1. Train and test data sets

Characteristic	Total, N = 14151	FIR, N = 1331	POPLAR, N = 1341	BIRCH, N = 5951	OAK, N = 5531	p-value <sup>2</sup>
Age	64 (57, 70)	67 (60, 73)	62 (55, 69)	65 (57, 71)	64 (57, 70)	<0.001
Sex						0.3
Female	568 (40%)	57 (43%)	47 (35%)	251 (42%)	213 (39%)	
Male	847 (60%)	76 (57%)	87 (65%)	344 (58%)	340 (61%)	
Weight	72 (61, 82)	70 (60, 83)	73 (63, 84)	72 (61, 82)	71 (60, 82)	0.3
BMI	24.9 (22.1, 28.1)	24.8 (21.9, 27.6)	25.2 (22.8, 28.7)	25.0 (22.1, 28.2)	24.7 (22.0, 28.1)	0.4
Unknown	65	8	5	30	22	
Race						<0.001
Asian	228 (16%)	6 (4.5%)	23 (17%)	77 (13%)	122 (22%)	
Others, unknown or missing	73 (5.2%)	9 (6.8%)	9 (6.7%)	23 (3.9%)	32 (5.8%)	
White	1,114 (79%)	118 (89%)	102 (76%)	495 (83%)	399 (72%)	
Smoking history						0.13
Current	171 (12%)	18 (14%)	22 (16%)	60 (10%)	71 (13%)	
Never	248 (18%)	16 (12%)	27 (20%)	102 (17%)	103 (19%)	
Previous	996 (70%)	99 (74%)	85 (63%)	433 (73%)	379 (69%)	
Characteristic	Total, N = 14151	FIR, N = 1331	POPLAR, N = 1341	BIRCH, N = 5951	OAK, N = 5531	p-value <sup>2</sup>
Heart rate	81 (71, 92)	80 (70, 94)	84 (74, 96)	80 (70, 90)	81 (72, 93)	0.049
Systolic blood pressure	122 (111, 133)	122 (113, 132)	122 (112, 131)	120 (110, 133)	123 (113, 135)	0.13

<sup>1</sup>Median (IQR); n (%) <sup>2</sup>Kruskal-Wallis rank sum test; Pearson's Chi-squared test

636

633

635

## Supplementary Figure 2. Patient characteristics: demographics and clinics

638

Characteristic	Total, N = 14151	FIR, N = 133 <sup>1</sup>	POPLAR, N = 134 <sup>1</sup>	BIRCH, N = 5951	OAK, N = 5531	p-value <sup>2</sup>
Disease type						0.4
Locally advanced	77 (5.4%)	3 (2.3%)	8 (6.0%)	32 (5.4%)	34 (6.1%)	
Metastatic	1,338 (95%)	130 (98%)	126 (94%)	563 (95%)	519 (94%)	
Line						< 0.001
≥2	1,255 (89%)	102 (77%)	134 (100%)	466 (78%)	553 (100%)	
1	160 (11%)	31 (23%)	0 (0%)	129 (22%)	0 (0%)	
Histology						< 0.001
Non-squamous	1,016 (74%)	95 (100%)	87 (65%)	427 (72%)	407 (74%)	
Squamous	361 (26%)	0 (0%)	47 (35%)	168 (28%)	146 (26%)	
Unknown	38	38	0	0	0	
Stage						< 0.001
1 .	123 (8.9%)	11 (8.7%)	4 (3.0%)	73 (12%)	35 (6.5%)	
Ш	140 (10%)	11 (8.7%)	9 (6.8%)	73 (12%)	47 (8.7%)	
III	367 (27%)	30 (24%)	37 (28%)	165 (28%)	135 (25%)	
IV	753 (54%)	75 (59%)	82 (62%)	273 (47%)	323 (60%)	
Unknown	32	6	2	11	13	
Number of metastases						0.4
1	386 (28%)	35 (27%)	34 (26%)	158 (27%)	159 (29%)	
2	655 (48%)	55 (42%)	61 (47%)	289 (50%)	250 (46%)	
3	334 (24%)	41 (31%)	34 (26%)	128 (22%)	131 (24%)	
Unknown	40	2	5	20	13	
Liver metastases	272 (19%)	29 (22%)	33 (25%)	105 (18%)	105 (19%)	0.3
Number of met. loc.						0.3
Four sites	70 (4.9%)	13 (9.8%)	6 (4.5%)	27 (4.5%)	24 (4.3%)	
One site	426 (30%)	37 (28%)	39 (29%)	178 (30%)	172 (31%)	
Three sites	264 (19%)	28 (21%)	28 (21%)	101 (17%)	107 (19%)	
Two sites	655 (46%)	55 (41%)	61 (46%)	289 (49%)	250 (45%)	
Tumor size	59 (43, 95)	59 (57, 59)	70 (43, 107)	60 (37, 96)	70 (43, 102)	0.001

639 640 <sup>1</sup>n (%); Median (IQR) <sup>2</sup>Pearson's Chi-squared test; Kruskal-Wallis rank sum test

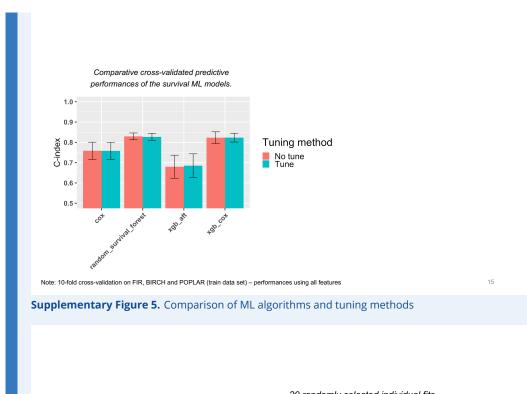
Supplementary Figure 3. Patient characteristics: disease

Characteristic	Total, N = 14151	FIR, N = 1331	POPLAR, N = 1341	BIRCH, N = 5951	OAK, N = 5531	p-value <sup>2</sup>
PD-L1 tumor cells						<0.001
0	606 (43%)	4 (3.0%)	40 (30%)	212 (36%)	350 (64%)	
1	209 (15%)	21 (16%)	56 (42%)	69 (12%)	63 (11%)	
2	381 (27%)	105 (79%)	37 (28%)	160 (27%)	79 (14%)	
3	217 (15%)	3 (2.3%)	1 (0.7%)	154 (26%)	59 (11%)	
Unknown	2	o	0	0	2	
PD-L1 immune cells						<0.001
0	252 (18%)	3 (2.3%)	40 (30%)	16 (2.7%)	193 (35%)	
1	398 (28%)	5 (3.8%)	56 (42%)	122 (21%)	215 (39%)	
2	439 (31%)	25 (19%)	19 (14%)	297 (50%)	98 (18%)	
3	322 (23%)	99 (75%)	19 (14%)	159 (27%)	45 (8.2%)	
Unknown	4	1	0	1	2	
ECOG						0.6
Status 0	505 (36%)	42 (32%)	44 (33%)	215 (36%)	204 (37%)	
Status 1 or 2	909 (64%)	90 (68%)	90 (67%)	380 (64%)	349 (63%)	
Unknown	1	1	0	0	0	

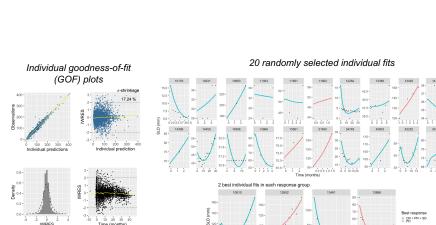
<sup>1</sup>n (%) <sup>2</sup>Pearson's Chi-squared test

642 643

Supplementary Figure 4. Patient characteristics: PD-L1 and ECOG







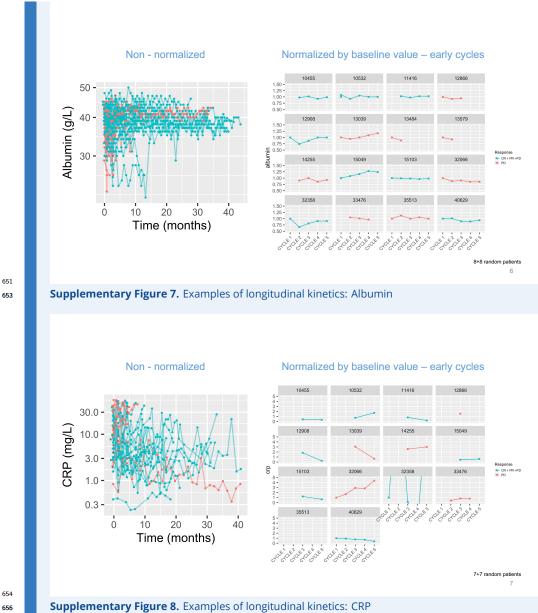
of-fit qu

Go

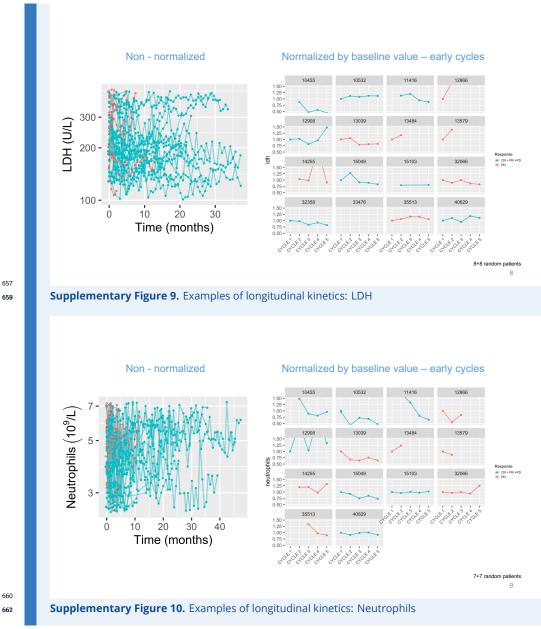
648 650

Supplementary Figure 6. TK modeling goodness-of-fit

Best response CR + PR + SD PD



653



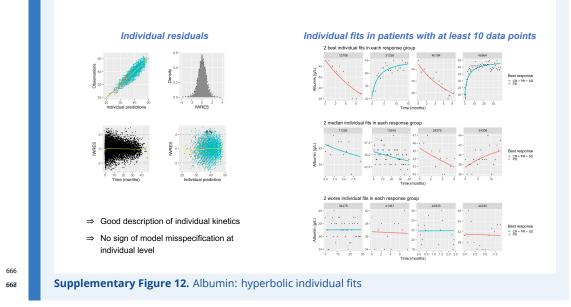
65<del>9</del>

	Albumin		CRP		LDH		Neutrophils	
Model	BICc	b	BICc	b	BICc	b	BICc	b
Double-exponential	48,395	0.058	28,764	0.21	39,886	0.56	102,449	0.14
Hyperbolic	48,007	0.056	29,712	0.22	40,915	0.62	102,943	0.14
Linear	49,436	0.063	30,020	0.23	42,462	0.70	105,193	0.17
Constant	49,724	0.065	31,332	0.25	42,982	0.74	106,249	0.18

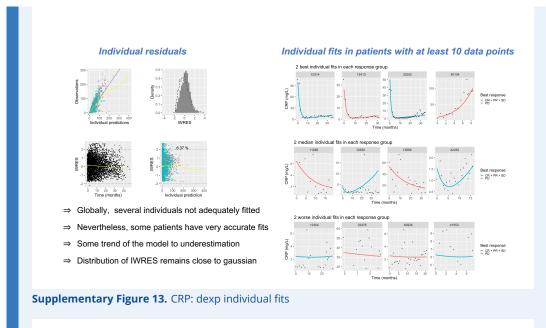
Corrected Bayesian Information Criterion (BICc) for four empirical kinetic models of BK. b : standard deviation of the proportional error model

663 665

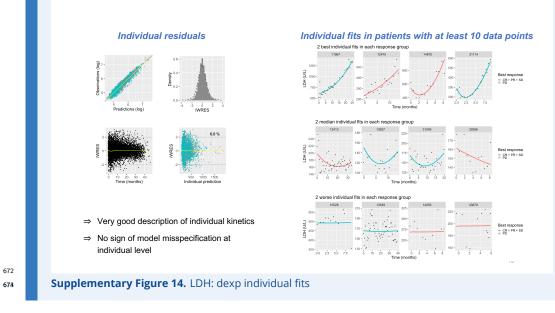




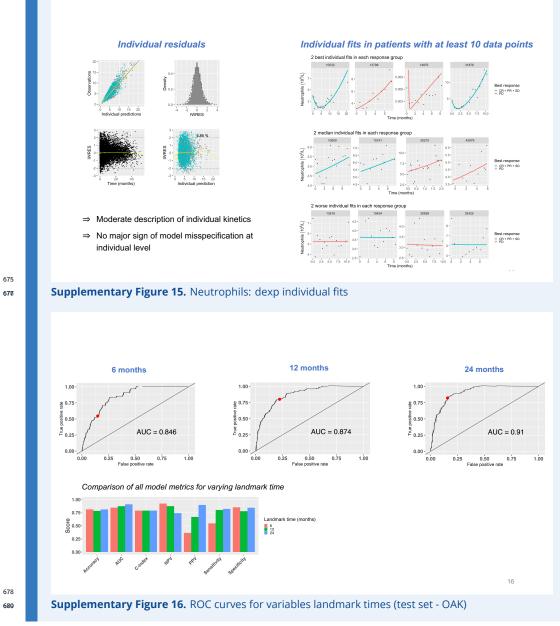
10



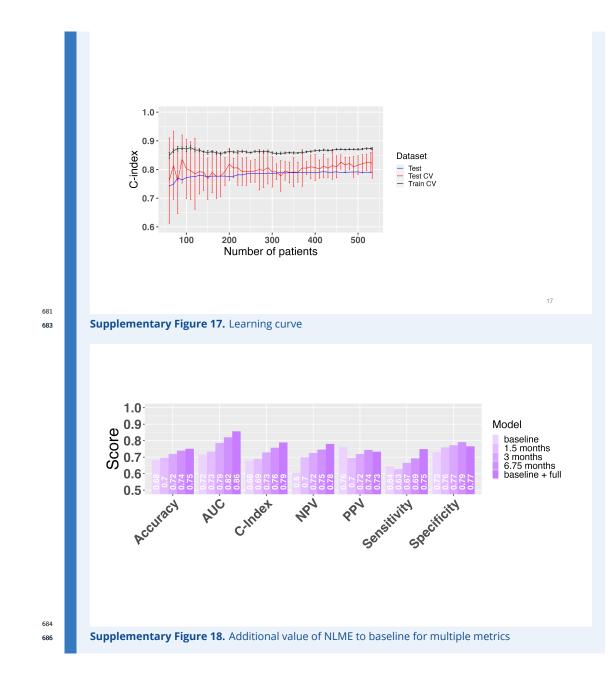




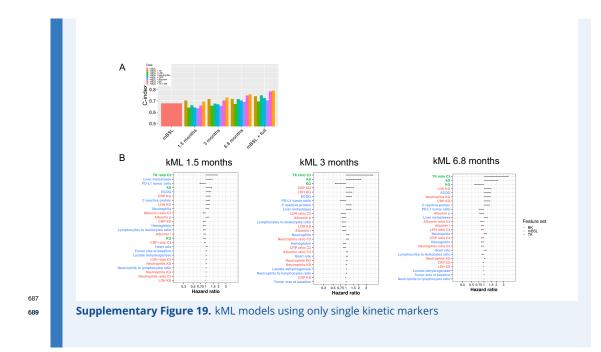
....







Benzekry et al. 2024 | Dynamical modeling coupled to machine learning for anti-PDL1 treatment in NSCLC



# Supplementary Methods

# Dimensionality reduction for RNAseq

Initial expression data from RNAseq consisted of 715 patients and 58, 311 transcripts. The first step of data filtering removed all transcripts with less than 10 read counts for all patients, then selected genes with highest variability between patients (top 15,000 transcripts most variable). Then, data were normalized using upper quartile normalization which consisted in dividing each read count by the 75<sup>th</sup> percentile of the read counts of the corresponding sample and the final expression values were  $\log_2$  transformed. Subsequently, a univariable Cox regression model was employed to statistically assess the correlations between the expression levels of the transcripts and overall survival. Bonferroni correction was used to adjust p-values from multiple univariate tests. This step was performed using the RegParallel R package. We selected transcripts with high predictive values using following criteria: adjusted log rank < 0.01 and HR < 0.85 or HR > 1.2. The remaining transcripts were used to perform a bootstrap Lasso Cox regression with cross-validation using mainly the glmnet R package. Finally, the smallest number of transcripts with best predictive model (highest C-index) was selected for further analysis.

# Rules for BK processing

- 1. Observations outside lower (LB) and upper (UB) physiological bounds were discarded using the following values, determined from discussion with a clinical oncologist: albumin,  $LB = 10 g L^{-1}$ ,  $UB = 100 g L^{-1}$ ; CRP, no LB,  $UB = 300 mg L^{-1}$ ; LDH, LB = 50 U/L, UB = 2000 U/L; neutrophils, no LB, UB = 20.
- 2. For duplicates, the first one recorded was kept.
- 3. Denoting  $BK_n$  the value of the a "BK" at time  $t_n$  for a given patient, we excluded values such that:  $BK_n \notin (BK_{n-1}, BK_{n+1})$  AND  $|BK_n BK_{n-1}| > 3 \times sd_{BK}$  AND  $|BK_n BK_{n+1}| > 3 \times sd_{BK}$ , where  $sd_{BK}$  is the standard deviation of  $\{BK_n\}_n$ , i.e. all time points for this patient.
- 4. The BK value at the closest time point to treatment initiation was kept, provided this time point was no more than 40 days before or 10 days after treatment initiation (otherwise, patient was disregarded).

# Nonlinear mixed-effects modeling

Denoting by  $\mathcal{M}(t;\theta)$  a structural dynamic model that depends on time *t* and a set of parameters  $\theta$ , longitudinal observations  $y_j^i$  in patient *i* at time  $t_j^i$  were assumed to follow the observation model

$$y_j^i = \mathcal{M}\left(t_j^i; \theta^i\right) + \varepsilon_j^i$$

where  $\varepsilon_j^i \sim \mathcal{N}\left(0, \sigma_j^i\right)$  is the gaussian-distributed error model. The latter was either constant  $(\sigma_j^i = a, \forall i, j)$  for TK or proportional  $(\sigma_j^i = b\mathcal{M}(t_j^i; \theta^i), \forall i, j)$  for BK. To describe inter-individual variability, individual parameters  $\theta^i$  were assumed to follow log-normal distributions:

$$\ln\left(\theta^{i}\right) = \ln\left(\theta_{non}\right) + \eta^{i}, \quad \eta^{i} \sim \mathcal{N}(0, \omega^{2})$$

with population-level parameters  $\theta_{pop}$  and  $\omega$ . Estimation of these was performed using the stochastic approximation of expectation maximization algorithm implemented in the Monolix software.