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# Beyond the delta method

Antoine Lejay\*      Sara Mazzonetto†

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## Abstract

We give an asymptotic development of the maximum likelihood estimator (MLE), or any other estimator defined implicitly, in a way which involves the limiting behavior of the score and its higher-order derivatives. This development, which is explicitly computable, gives some insights about the non-asymptotic behavior of the renormalized MLE and its departure from its limit. We highlight that the results hold whenever the score and its derivative converge, including to non Gaussian limits.

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**Keywords:** parametric estimation; Taylor expansion; hypothesis testing; delta method; maximum likelihood estimation; moments method; Implicit Function Theorem.

## 1 Introduction

The traditional approach of dealing with parametric estimation consists in defining an estimator as the root or minimizer of a function and then to study its asymptotic properties as the size of the observations increases. This is the case for the Maximum Likelihood Estimation (MLE, [35]) and its variants (pseudo- or quasi- likelihood estimation), or the Generalized Method of Moments which are used in a broad

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variety of problems [25, 34]. A typical statement is that the estimator is consistent and asymptotically (mixed) normal: The properly renormalized error between the estimator and the unknown value converges in distribution to a (mixed) Gaussian random variable.

The limiting distribution may be used to quantify the quality of the estimation, and to construct for example confidence intervals. As pointed by D.A. Sprott [40, 41], the quality of the approximation depends more on the distance between the law of the estimator and the limiting distribution than on the size itself. Various tools have been introduced to improve the estimator for small or average sample size or to quantify its departure from its limiting distribution. Among them, let us cite Edgeworth expansions [38], Berry-Esséen inequality [12, 33, 44], Stein method [3, 5], jackknife [17], bootstrap [19], “magic formula” [8], Bartlett corrections [9, 15], and so on. Some of these corrections use considerations on the density of the estimator, others are pointwise ones.

We look for an expansion of the univariate root  $\theta(n)$  of  $F_n(\theta(n)) = 0$  — think of the MLE — of the form

$$\theta(n) = \theta + H(n, G_n)$$

where  $\theta$  is the true, unknown parameter,  $G_n$  is asymptotical pivotal converging to some  $G$  (think of the standard Gaussian random variable coming from a Central Limit Theorem) and  $x \mapsto H(n, x)$  is a non-linear, random function of the kind

$$H(n, x) = \sum_{k=1}^p \alpha_k(n) \varphi(n)^k x^k + O(\varphi(n)^{p+1}). \quad (1)$$

In the latter expansion,  $\varphi(n)$  decreases to 0 as  $n \rightarrow \infty$  and

$$\{\alpha_k(n)\}_{k=1}^p \xrightarrow[n \rightarrow \infty]{\text{law or probability}} \{\alpha_k\}_{k=1}^p \quad (2)$$

for some  $\alpha_k$ 's. Of course, the  $\alpha_k$ 's and the  $\alpha_k(n)$ 's may depend on  $\theta$ . The expansions may have different forms for some particular values of  $\theta$  for which some of the  $\alpha_k$  vanishes (which we call a *phase transition*). We highlight that our approach is pointwise, and the order  $p$  is related to the regularity of  $F_n$ . Note that the expansion of  $\theta(n)$  is such that

$$\varphi(n)^{-1}(\theta(n) - \theta) \xrightarrow[n \rightarrow \infty]{\text{dist.}} c_1 G.$$

The main result is provided in Theorem 3 in Section 4: An expansion of  $\theta(n)$  as just described.

The family of coefficients  $\{\alpha_k(n)\}_{k=1}^p$  in (1) is not unique. However, in Section 3, given a function  $F_n$  and knowing the asymptotic behavior of  $F_n$  and its high

order derivatives  $D^k F_n$ , we give a procedure to compute automatically the  $\alpha_k(n)$ 's satisfying (2) from  $F_n$  and its derivatives.

Our approach is based on an asymptotic implicit function theorem (see Theorem 1 below), inspired from perturbative approaches. Performing asymptotic expansions has a long history with a broad range of applications, notably to study perturbations of (partial) differential equations and dynamical systems [16, 27, 29] or in change of scale analysis [10]. Although in one case, we recover the Lagrange inversion formula for power series [21, 22, 39], our starting point is different and seems not to have been considered in a general form but mostly for particular cases. Moreover, we also consider functions that are not necessarily analytic. Our approach seems to be new because it assumes weak general assumptions on the regularity of  $F_n$  and takes into account different possible asymptotic behaviors of  $D^k F_n$  (for statistical applications, think that  $D^k F_n$  may converge thanks to the Law of Large Numbers or to the Central Limit Theorem, hence with different rates).

It is a sort of generalization of the Delta method [42] in the sense that knowing the asymptotic behavior of the score and its derivative, we get the asymptotic behavior of a function of it, in particular the estimator. Besides, we do not restrict ourselves to a first order linearization.

In Section 5.1 we study the effects on our procedure of change of variable for reparametrization of the space of parameters. Related to the problem of reparametrization, changing  $F_n$  to  $F_n/\Phi$  for a suitable  $\Phi$  does not change the root  $\theta(n)$  but leads to a different expansion in (1) which is sometimes more convenient (see Section 5.2 and Section 6.2).

We illustrate our results through several known examples: exponential family (Section 6.1) and parametric estimation for stochastic processes (Section 6.4). We highlight that we are not restricted to the MLE but we may consider as well the Generalized Method of Moment, pseudo- or quasi-likelihood estimators, and other variants (see *e.g.*, [34]). Besides, we are not restricted to independent samples nor to asymptotic normality.

Our procedure can also be used to study the lack of normality of the estimator under a fixed sample size  $n$  when  $G$  is Gaussian. The lack of normality may come from the fact that  $G_n$  has not reached the Gaussian regime and/or by the non-linearity of  $H(n, \cdot)$ . We develop this in Example 6 on the exponential distribution and in Example 8 on the Binomial distribution.

In [31], we apply our procedure to the Skew Brownian motion to obtain new expansions and the asymptotic behaviors of the MLE estimator of the skewness parameter. In this context, we recover the expansion given in [32], and give an alternative one. This is due to the fact that we have generalized the procedure to

make a finer expansion for some specific asymptotics of the  $D^k F_n$ . We refer to this situation as *phase transition*, which we explain in this paper on the binomial model (see Section 6.3).

Before introducing Theorem 3, in the next two sections, we present some preliminary analytical results in which no randomness is involved. However, to introduce some notations, we may refer to the desired probabilistic properties.

## 2 The asymptotic implicit function theorem

Let  $\Theta, \mathbb{I}$  be subsets of  $\mathbb{R}$ . We assume that  $\Theta$  contains an open interval around 0 and that  $(0, \epsilon) \subset \mathbb{I}$  for some  $\epsilon > 0$ . The set  $\Theta$  is the *space of parameters*,  $\mathbb{I}$  is the *space of rates*. Let  $\mathbf{S}$  be a set, which we call the *space of scales*. We consider a function  $\varphi : \mathbf{S} \rightarrow \mathbb{I}$ , which we call the *rate function*.

*Hypothesis 1.* We fix  $p \in \{1, \dots, +\infty\}$  and we consider  $\beta := \{\beta_k\}_{k=0,1,\dots,p+1} \in \mathbb{Z}^{p+2}$ . We define

$$\beta_\star := \max_{m=2,\dots,p+1} \beta_m, \quad (3)$$

$$\gamma_m := p + 1 + (\beta_m - \beta_\star) \wedge 0 \text{ for } m = 1, \dots, p. \quad (4)$$

From now on we assume Hypothesis 1. Later, we focus on two sequences

$$\beta^\frown := (1, 2, 2, \dots, 2) \text{ and } \beta^\smile := (1, 2, 1, 2, \dots, 1, 2)$$

which are relevant for statistical applications. We refer to these choices of  $\beta$  respectively as “up-flat” and “zigzag” cases.

**Notation 1.** In a systematic manner, we use sums over multi-indices satisfying some constraints. Such multi-indices are denoted by  $(k_1, \dots, k_m)$  and each  $k_i$  is an integer.

**Notation 2.** Let us consider a parametric family of functions  $\theta \mapsto F(s, \theta)$  which is of class  $\mathcal{C}^p(\Theta, \mathbb{R})$  for any  $s \in \mathbf{S}$ . We denote by  $D^m$  the derivative of order  $m$  with respect to  $\theta$ . We define

$$\{F\}_m(s) := \varphi(s)^{\beta_m} D^m F(s, 0) \text{ for } m = 0, \dots, p \text{ and } s \in \mathbf{S}.$$

The idea of the above notation is that the  $\beta_m$  are such that  $\varphi(s)^{\beta_m} D^m F(s, 0)$  converges.

We state a series expansion which approximates the roots of  $F(s, \cdot)$ .

**Theorem 1** (Approximation of roots). *Assume Hypothesis 1 holds and assume that  $\beta_1 \geq \beta_*$ . Let  $F \in \mathcal{C}^{p+1}(\Theta, \mathbb{R})^{\mathbf{S}}$  and  $\varphi$  be a rate function. Let  $s \in \mathbf{S}$  such that there exist a sequence  $\boldsymbol{\alpha}(s) = \{\alpha_k(s)\}_{k=0, \dots, p}$  and a convex neighborhood  $U(s)$  of 0 in  $\Theta$  satisfying:*

(i)  $\alpha_0(s) = 0$  and

$$\{F\}_0(s)\varphi(s)^{-\beta_0} + \sum_{m=1}^p \frac{\{F\}_m(s)}{m!} \sum_{k_1 + \dots + k_m < \gamma_m} \alpha_{k_1}(s) \cdots \alpha_{k_m}(s) \varphi(s)^{\sum_{n=1}^m k_n - \beta_m} = 0. \quad (5)$$

(ii) *There exists  $\theta(s) \in U(s)$  such that  $F(s, \theta(s)) = 0$ .*

(iii) *It holds that*

$$|\varphi(s)|^{\beta_1} |\mathbf{D}F(s, w)| \geq c(s) > 0, \text{ for all } w \in U(s)$$

and  $\max_{m \in \{2, \dots, p+1\}} \sup_{w \in U(s)} |\mathbf{D}^m F(s, w)| < +\infty.$

(iv) *It holds that  $\sum_{i=1}^p |\alpha_i(s)| |\varphi(s)|^i \in U(s)$ .*

Then

$$\left| \theta(s) - \sum_{i=1}^p \alpha_i(s) \varphi(s)^i \right| \leq K(s) \max \left\{ \sum_{k=1}^p |\alpha_k(s)|, \left( \sum_{k=1}^p |\alpha_k(s)| \right)^{p+1} \right\} |\varphi(s)|^{p+1},$$

with

$$K(s) \leq \frac{\max_{\substack{m \in \{1, \dots, p\} \\ \beta_{m+1} \leq \beta_*}} |\varphi(s)|^{\beta_1 - \beta_*} |\{F\}_m(s)| + \max_{m \in \{2, \dots, p+1\}} \sup_{w \in U(s)} |\varphi(s)|^{\beta_1} |\mathbf{D}^m F(s, w)|}{c(s)} < \infty.$$

The proof of Theorem 1 is provided in Appendix A. We also consider the case of analytic functions, in which the expansion is infinite.

**Theorem 2** (Approximation of roots, analytic case). *Assume Hypothesis 1 holds for  $p = +\infty$ . Let  $F : \mathbf{S} \times \Theta \rightarrow \mathbb{R}$  be such that  $F(s, \cdot)$  is analytic and  $\varphi$  be a rate function. Let  $s \in \mathbf{S}$  such that there exist a sequence  $\boldsymbol{\alpha}(s) = \{\alpha_k(s)\}_{k \geq 0}$  and a convex neighborhood  $U(s)$  of 0 in  $\Theta$  satisfying:*

(i)  $\alpha_0(s) = 0$  and (5) holds with  $p = +\infty$  and  $\gamma_m = +\infty$  for  $m \geq 1$ .

(ii) *There exists  $\theta(s) \in U(s)$  such that  $F(s, \theta(s)) = 0$ .*

(iii) *It holds that  $|\varphi(s)|^{\beta_1} |\mathbf{D}F(s, w)| \geq c(s) > 0$  for all  $w \in U(s)$ .*

(iv) *It holds that  $\sum_{i=1}^{+\infty} |\alpha_i(s)| |\varphi(s)|^i \in U(s)$ .*

Then

$$\theta(s) = \sum_{i=1}^{+\infty} \alpha_i(s) \varphi(s)^i.$$

### 3 Formal manipulation of expansions

In Theorem 1 and 2, the expansion of the root  $\theta(s)$  to  $F(s, \theta(s)) = 0$  is performed in term of  $\varphi(s)$  with  $s$ -dependent coefficients. This implies a lack of uniqueness in the expansion. To enforce uniqueness, we identify the coefficients by transforming series to formal ones. Although natural, this choice is arbitrary and “righteous”<sup>1</sup>.

The main idea is to decouple the scale ( $s$ , playing the role of sample size) and the rate ( $\varphi(s)$ ), considering the latter as a variable. In Section 3.1 we introduce the necessary operators to decouple and re-couple: dummization and evaluation. Next, in Section 3.2 we recall Faà di Bruno and Taylor formulae and show how dummization and evaluation behave in Taylor expansion of composition of functions. In Section 3.3, we propose a definition of sequences that satisfy Eq. (5). We call them  $(\beta, p)$ -related sequences. We provide a different presentation in Section 3.4. Next we show that there exists a unique  $(\beta, p)$ -related sequence when  $\beta$  is  $\beta^\frown$  and  $\beta^\smile$  and we provide a recursive formula for it in Sections 3.5-3.6. In the “up-flat” case, we provide different characterizations of the  $(\beta^\frown, p)$ -related sequence (see Lemma 7 and Lemma 8) that we propose also in the context of Theorem 1 in Section 3.7. To conclude, we show in Section 3.8 how an additive perturbation acts on the sequence.

#### 3.1 Dummization and evaluation

For a ring  $\mathbb{K}$  and a family of indeterminates  $\mathbf{X} := (X_1, \dots, X_n)$ ,  $\mathbb{K}[X_1, \dots, X_n]$  (resp.  $\mathbb{K}(X_1, \dots, X_n)$ ) denotes the ring of polynomials (resp. rational functions) with indeterminates  $X_1, \dots, X_n$  and coefficients in  $\mathbb{K}$ . Also, the set of formal power series is  $\mathbb{K}[[X_1, \dots, X_n]]$ .

To a finite, numerical sequence  $\mathbf{a} = (a_1, \dots, a_q)$ , we consider the “dummization” operation which consists in transforming each of the  $a_i$ ’s to an *indeterminate*  $a_i^\circ$ . Therefore  $\mathbb{R}[\mathbf{a}^\circ]$  (resp.  $\mathbb{R}(\mathbf{a}^\circ)$ ) is the ring of polynomials (resp. rational functions) whose indeterminates are the  $a_i^\circ$ .

We consider another indeterminate  $z$ , that we will use later as a placeholder for  $\varphi(s)^\circ$  or for  $x \in \mathbb{R}$ . We denote by  $\mathbb{K}[z]$  the ring of polynomials whose coefficients are in  $\mathbb{K}$ . We will be concerned with  $\mathbb{K} = \mathbb{R}(\mathbf{a}^\circ)$ . We also denote by  $\mathbb{K}[z, z^{-1}]$  the Laurent series whose coefficients are in  $\mathbb{K}$ .

The *evaluation* — denoted by  $\epsilon$  — consists transforming back  $a_i^\circ$  to  $a_i$ . It is the dual operation of the dummization. We extend  $\epsilon$  as an algebra homomorphism

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<sup>1</sup>As pointed out in [2] regarding homogenization theory, a “criminal path” leads to other ways to compute effective coefficients.

from  $\mathbb{R}[[\mathbf{a}^\bullet]]$  to  $\mathbb{R}$ . We furthermore extend  $e$  on  $\mathbb{R}(\mathbf{a}^\bullet)$  by setting  $e(P/Q) := e(P)/e(Q)$  as well as to  $\mathbb{K}[[z, z^{-1}]]$  by setting  $e(z^k) = \varphi(s)^k$  for any  $k \in \mathbb{Z}$ .

**Notation 3.** Let  $p \in \mathbb{N}$ . Given a sequence  $\boldsymbol{\delta} = \{\delta_k\}_{k=0, \dots, p} \subset \mathbb{K}$ , we denote by  $P_p(z; \boldsymbol{\delta})$  the polynomial  $P_p(z; \boldsymbol{\delta}) := \sum_{k=0}^p \delta_k z^k \in \mathbb{K}[z]$ .

**Notation 4** (Projection onto polynomials of degrees at most  $p$ ). Let  $p \in \mathbb{N}$ . We denote by  $\mathfrak{p}_p$  is the projection of a power series  $P \in \mathbb{K}[[z]]$  onto the polynomials of degree at most  $p$  by setting  $z^q = 0$  whenever  $q > p$ .

Note that we can extend the latter notation to  $p = \infty$  taking  $\mathfrak{p}_\infty$  the identity and  $P_\infty(z; \boldsymbol{\delta})$  the series  $\sum_{k=0}^\infty \delta_k z^k$  when it is well defined.

The next lemma is immediate by expanding powers.

**Lemma 1.** Let  $p \in \mathbb{N}$ . Consider three families  $\boldsymbol{\delta}^\bullet, \boldsymbol{\alpha}^\bullet, \boldsymbol{\eta}^\bullet$  with components in a field  $\mathbb{K}$ . Assume that  $\alpha_0^\bullet = 0$ . These families  $\boldsymbol{\delta}^\bullet, \boldsymbol{\alpha}^\bullet$  and  $\boldsymbol{\eta}^\bullet$  are linked by the relation

$$\eta_0^\bullet = \delta_0^\bullet \text{ and } \eta_k^\bullet = \sum_{m=1}^k \delta_m^\bullet \sum_{k_1 + \dots + k_m = k} \alpha_{k_1}^\bullet \cdots \alpha_{k_m}^\bullet \in \mathbb{R}(\boldsymbol{\delta}^\bullet) \text{ for all } k \in \{1, \dots, p\} \quad (6)$$

if and only if

$$P_p(z; \boldsymbol{\eta}^\bullet) = \mathfrak{p}_p \circ P_p(\cdot; \boldsymbol{\delta}^\bullet) \circ P_p(z; \boldsymbol{\alpha}^\bullet).$$

## 3.2 Taylor and Faà di Bruno formulae

The Taylor formula is the main tool to go back and forth between the realm of analysis and the one of algebra.

The Faà di Bruno formula is a way to compute the higher order derivatives of the composition of two functions [27]. It is a useful tool to deal with power series [1, 28] and can be seen both from the points of view of analysis [30] or of algebra [20].

**Proposition 1** (Faà di Bruno formula). Let  $U, V$  be open subsets of  $\mathbb{R}$ . Let  $f : U \rightarrow \mathbb{R}$  and  $g : V \rightarrow U$  be two functions of class  $\mathcal{C}^p$  for  $p \geq 1$ . Then  $f \circ g$  is of class  $\mathcal{C}^p(V, \mathbb{R})$  and <sup>2</sup>

$$D^\ell(f \circ g)(x) = \ell! \sum_{m=1}^{\ell} \frac{D^m f(g(x))}{m!} \sum_{k_1 + \dots + k_m = \ell} \frac{D^{k_1} g(x)}{k_1!} \cdots \frac{D^{k_m} g(x)}{k_m!}, \quad (7)$$

for  $\ell = 1, \dots, p$  and  $x \in V$ .

<sup>2</sup>We use a different convention than the usual ones on summing the indices written by regrouping the multi-index  $(k_1, \dots, k_m)$  by their respective values and counting the number of such partitions.



**Definition 1** (Formal Taylor series). Let  $f$  be a function of class  $\mathcal{C}^p(\mathbb{R}, \mathbb{R})$  for  $p = 1, \dots, +\infty$ . We set  $\delta_k := D^k f(0)/k!$ ,  $k = 0, \dots, p$  and  $\boldsymbol{\delta} := \{\delta_k\}_{k=0, \dots, p}$ . The *Taylor series operator*  $\mathcal{T}_p$  is defined as

$$\mathcal{T}_p f(x) := P_p(x; \boldsymbol{\delta}) = \sum_{k=0}^p \delta_k x^k \in \mathbb{R}[[x]]$$

that is, it gives the Taylor series of  $f$  around 0. The *formal Taylor series* of  $f$  is

$$\mathcal{T}_p^\bullet f(z) := P_p(z; \boldsymbol{\delta}^\bullet) = \sum_{k=0}^p \delta_k^\bullet z^k \in \mathbb{K}[[z]] \text{ with } \mathbb{K} = \mathbb{R}[\boldsymbol{\delta}^\bullet]. \quad (8)$$

If  $e(z) = x$  for  $f \in \mathcal{C}^{p+1}(\mathbb{R}, \mathbb{R})$ ,

$$e(\mathcal{T}_p^\bullet f(z)) = \mathcal{T}_p f(x) = f(x) - R(x)$$

with  $|R(x)| \leq |x|^{p+1} \sup_{y \in [-x, x]} |D^{p+1} f(y)| / (p+1)!$  for any  $x$  small enough.

A consequence of Lemma 1 and the Faà di Bruno formula is the following one.

**Lemma 2.** Let  $p \in \mathbb{N}$ , let  $f, g$  and  $h$  be three functions of class  $\mathcal{C}^{p+1}$  such that  $g(0) = 0$  and  $\mathcal{T}_p k(x) = e(\mathcal{T}_p^\bullet k(z))$  for  $k = f, g, h$  with  $e(z) = x$ .

Then  $\mathcal{T}_p^\bullet h = \mathfrak{p}_p(\mathcal{T}_p^\bullet f \circ \mathcal{T}_p^\bullet g)$  if and only if  $\mathcal{T}_p h = \mathcal{T}_p(f \circ g) = \mathfrak{p}_p(\mathcal{T}_p f \circ \mathcal{T}_p g)$ .

We end this section by relating the composition of analytic functions and their formal developments.

**Notation 5.** For  $r > 0$ , we denote by  $\mathcal{A}_r$  the class of analytic functions whose radius of convergence is at least  $r$ . We also set  $B_r := \{x \in \mathbb{R} \mid |x| < r\}$  the open ball of radius  $r$ .

*Hypothesis 2.* The functions  $f, g, h$  satisfy  $f \in \mathcal{A}_{r_f}$ ,  $g \in \mathcal{A}_{r_g}$ ,  $h \in \mathcal{A}_{r_h}$  with  $g(B_{r_g}) \subset B_{r_f}$ , so that  $f \circ g \in \mathcal{A}_{r_g}$ .

Note that for an analytic function with a positive radius of convergence  $r_f$ ,  $f(x) = \mathcal{T}_\infty f(x)$  whenever  $|x| < r_f$ .

**Lemma 3.** Let the same assumptions of Lemma 2, and let  $f, g, h$  as in Hypothesis 2. Then  $\mathcal{T}_\infty^\bullet h = \mathcal{T}_\infty^\bullet f \circ \mathcal{T}_\infty^\bullet g$  if and only if  $h = f \circ g$ .

*Remark 1.* Let  $p \in \mathbb{N}$  and  $\boldsymbol{\delta} = \{\delta_k\}_{k=0}^\infty \subset \mathbb{R}$  such that  $P_\infty(\cdot; \boldsymbol{\delta}) \in \mathcal{A}_r$  for some  $r > 0$ . Then  $P_p(x; \boldsymbol{\delta}) = \mathfrak{p}_p P_\infty(x; \boldsymbol{\delta}) = \mathcal{T}_p P_\infty(x; \boldsymbol{\delta})$  and  $P_p(x; \boldsymbol{\delta}) = \mathcal{T}_p P_p(x; \boldsymbol{\delta})$ . In particular if  $f \in \mathcal{A}_r$  for some  $r > 0$ , then  $\mathcal{T}_p f = \mathfrak{p}_p \mathcal{T}_\infty f = \mathcal{T}_p \mathcal{T}_\infty f$ .

### 3.3 Constructing $(\beta, p)$ -related sequences

The goal of this section is, for  $s \in \mathbf{S}$  fixed, to find a sequence  $\alpha(s)$  satisfying (5). The idea is to consider the coefficients  $\alpha_m(s)^\bullet$  and  $\{F\}_m(s)$  as dummy variables, then identify the coefficients  $\alpha_m(s)^\bullet$  by the dummy version of Eq. (5), and finally transfer them back to their numerical values.

**Definition 2** ( $(\beta, p)$ -related sequences). Let  $p$  and  $\beta$  as in Hypothesis 1. Let  $\delta = \{\delta_k\}_{k=0, \dots, p}$  and  $\alpha = \{\alpha_k\}_{k=0, \dots, p}$ . We say that  $\alpha$  is  $(\beta, p)$ -related to  $\delta$  if there exist  $\alpha_k^\bullet$ ,  $k = 0, \dots, p$  which are rational functions in  $\mathbb{K} := \mathbb{R}(\delta^\bullet)$  such that

- (i) For  $k = 0, \dots, p$ ,  $\alpha_k = e(\alpha_k^\bullet)$  with  $\alpha_0^\bullet = 0$ .
- (ii) In  $\mathbb{K}[[z, z^{-1}]]$ ,

$$\delta_0^\bullet z^{-\beta_0} + \sum_{m=1}^p \delta_m^\bullet \sum_{k_1 + \dots + k_m < \gamma_m} \alpha_{k_1}^\bullet \cdots \alpha_{k_m}^\bullet z^{\sum_{n=1}^m k_n - \beta_m} = 0, \quad (9)$$

with  $\gamma_m$  defined by (4), meaning that all the coefficients in  $\mathbb{K}$  vanish for each of the powers of  $z$ .

*Remark 2.* If  $\alpha$  is a  $(\beta, p)$ -related sequence to  $\delta$ , then  $\alpha$  it is also  $(\beta', p)$ -related sequence to  $\delta$  for  $\beta' = \{\beta_k + \gamma\}_{k=0, \dots, p+1}$  for any  $\gamma \in \mathbb{Z}$ .

*Remark 3.* If  $\alpha$  is a  $(\beta, p)$ -related sequence to  $\delta$ , then  $\alpha$  it is also  $(\beta, p)$ -related sequence to  $\delta_\lambda$  for  $\delta_\lambda = \{\lambda \delta_k\}_{k=0, \dots, p}$  for any  $\lambda \in \mathbb{R}$ .

**Example 1** (Non-existence/non-uniqueness). We consider  $p = 2$ ,  $\beta_0 = 0$ ,  $\beta_1 = 1$ ,  $\beta_2 = 2$  and  $\beta_3 \in \{1, 2\}$ . Then  $\gamma_m = 1 + \beta_m$  for  $m = 1, 2$ . Eq. (9) becomes

$$\delta_0^\bullet + \delta_1^\bullet \alpha_1^\bullet + \delta_2^\bullet (\alpha_1^\bullet)^2 = 0.$$

The possibility to find  $\alpha_1^\bullet \in \mathbb{R}$  depends on the relative values of  $\delta$ . Even if we are able to do so, we are free to choose  $\alpha_2^\bullet$  and the choice of  $\alpha$  which is  $((0, 1, 2), 2)$ -related to  $\delta$  is not unique.

**Example 2** (Non-existence/non-uniqueness). We consider  $p = 2$ ,  $\beta = (0, 1, 1, 2)$ . If  $\delta_2^\bullet \neq 0$ , then  $\gamma_m = 1 + \beta_m$  for  $m = 1, 2$ . Eq. (9) becomes

$$\delta_0^\bullet + \delta_1^\bullet \alpha_1^\bullet + \delta_2^\bullet (\alpha_1^\bullet)^2 z = 0,$$

so that this could not be solved unless  $\delta_0^\bullet = 0$  and  $\alpha_1^\bullet = 0$ . In the latter case, we are free to choose  $\alpha_2^\bullet$ .

The next lemma relates the definition above to the assumption (i) in Theorem 1.

**Lemma 4.** Let  $F \in \mathcal{C}^p(\Theta, \mathbb{R})^{\mathbf{S}}$  and fix  $s \in \mathbf{S}$ . Let  $\alpha(s)$  be a  $(\beta, p)$ -related sequence to  $\delta(s)$  with  $\delta_k(s) = \{F\}_k(s)/k!$ ,  $k = 0, \dots, p$ . Define the evaluation map so that  $e(z) = \varphi(s)$ . Then  $\alpha(s)$  satisfies (5).

We study below two examples of  $\beta$  for which  $\alpha$  is unique when using Definition 2. However, given a sequence  $\alpha(s)$  satisfying (5), it is possible to change  $\alpha_1(s)$  by a term of order  $\varphi(s)$  so that (5) is still satisfied. Therefore, there are an infinite number of solutions to (5). Our construction is a way to impose a *unique choice* of  $\alpha(s)$  satisfying (5).

### 3.4 Scrambling operator

We provide here an alternative formulation of equation (9) of Definition 2.

**Definition 3** (Scrambling operator). Given  $P_p(z; \delta) = \sum_{k=0}^p \delta_k z^k \in \mathbb{K}[[z]]$  and  $\beta \in \mathbb{R}^{p+1}$ , the *scrambling operator* is

$$\mathfrak{s}_\beta(P_p(z; \delta)) := \sum_{k=0}^p \delta_k z^{k-\beta_k} \in \mathbb{K}[[z, z^{-1}]].$$

With Definition 3, Notation 3, and Notation 4, (9) is equivalent to

$$\mathfrak{s}_\beta \circ \mathfrak{p}_p \circ P_p(\cdot; \delta^\bullet) \circ P_p(z; \alpha^\bullet) = 0 \text{ with } \alpha_0^\bullet = 0. \quad (10)$$

Under the assumptions of Lemma 4, (9) (and so (10)) rewrites

$$\mathfrak{s}_\beta \circ \mathfrak{p}_p \circ \mathcal{T}_p^\bullet F(s, P_p(z; \alpha(s)^\bullet)) = 0 \text{ in } \mathbb{R}(\delta^\bullet(s))[[z, z^{-1}]]. \quad (11)$$

We now study  $(\beta, p)$ -related sequences for some particular choices of  $\beta$ , which are fitted for statistical applications.

### 3.5 The case “zig-zag”

Our first case of interest is  $\beta^\sim = (1, 2, 1, 2, \dots)$ . This is justified below in Section 6.3. Eq. (9) becomes

$$\begin{aligned} \delta_0^\bullet z + \sum_{\substack{m \geq 0 \\ 2m+1 \leq p}} \delta_{2m+1}^\bullet \sum_{k_1 + \dots + k_{2m+1} \leq p} \alpha_{k_1}^\bullet \dots \alpha_{k_{2m+1}}^\bullet z^{k_1 + \dots + k_{2m+1}} \\ + \sum_{\substack{m \geq 1 \\ 2m \leq p}} \delta_{2m}^\bullet \sum_{k_1 + \dots + k_{2m} < p} \alpha_{k_1}^\bullet \dots \alpha_{k_{2m}}^\bullet z^{k_1 + \dots + k_{2m} + 1} = 0. \end{aligned}$$

And it can be rewritten using the scrambling operator, which in this case is

$$\mathfrak{s}_{\beta^{\sim}} \left( \sum_{k=0}^{+\infty} \delta_k^{\bullet} z^k \right) = \frac{1}{z^2} \left( \sum_{k=0}^{+\infty} \delta_{2k+1}^{\bullet} z^{2k+1} + z \sum_{k=0}^{+\infty} \delta_{2k}^{\bullet} z^{2k} \right).$$

**Lemma 5.** *Assume that  $\delta_1 \neq 0$ . Then  $\alpha$  is  $(\beta^{\sim}, p)$ -related to  $\delta$  if and only if  $\alpha_1^{\bullet} = -\delta_0^{\bullet}/\delta_1^{\bullet}$ ,  $\alpha_{2q}^{\bullet} = 0$  for  $2q \leq p$  and*

$$\begin{aligned} \alpha_{2q+1}^{\bullet} = & - \sum_{\substack{m=2 \\ m \text{ even}}}^{2q} \frac{\delta_m^{\bullet}}{\delta_1^{\bullet}} \sum_{k_1+\dots+k_m=2q} \alpha_{k_1}^{\bullet} \cdots \alpha_{k_m}^{\bullet} \\ & - \sum_{\substack{m=3 \\ m \text{ odd}}}^{2q+1} \frac{\delta_m^{\bullet}}{\delta_1^{\bullet}} \sum_{k_1+\dots+k_m=2q+1} \alpha_{k_1}^{\bullet} \cdots \alpha_{k_m}^{\bullet} \text{ for } 2q+1 \leq p. \end{aligned}$$

*In particular, there exists one and only one sequence  $\alpha$  which is  $(\beta^{\sim}, p)$ -related to  $\delta$ .*

For  $p = 5$  with  $\widehat{\delta}_k = -\delta_k/\delta_1$  and  $\alpha_1 = \widehat{\delta}_0$ ,

$$\begin{aligned} \alpha_2 &= 0, & \alpha_3 &= \widehat{\delta}_3 \widehat{\delta}_0^3 + \widehat{\delta}_2 \widehat{\delta}_0^2, \\ \alpha_4 &= 0, & \alpha_5 &= (\widehat{\delta}_5 + 3\widehat{\delta}_3^2) \widehat{\delta}_0^5 + (5\widehat{\delta}_2 \widehat{\delta}_3 + \widehat{\delta}_4) \widehat{\delta}_0^4 + 2\widehat{\delta}_2^2 \widehat{\delta}_0^3. \end{aligned}$$

Note that  $\alpha_{2k+1}$  is not homogeneous in  $\widehat{\delta}_0^{2k+1}$  but contains terms in  $\widehat{\delta}_0^j$  for  $j \in \{k+1, \dots, 2k+1\}$ .

*Remark 4.* If  $\alpha$  is  $(\beta^{\sim}, \infty)$ -related to  $\delta$ , then  $\{\alpha_k\}_{k=0}^p$  is  $(\beta^{\sim}, p)$ -related to  $\{\delta_k\}_{k=0}^p$ .

**Lemma 6.** *We consider a sequence  $\alpha$  which is  $(\beta^{\sim}, +\infty)$ -related to  $\delta$ . Let  $p \in \{1, 2, \dots, \infty\}$ . We set*

$$f_{\text{odd}}(x) := \sum_{\substack{k \geq 0 \\ 2k+1 \leq p}} \delta_{2k+1} x^{2k+1} \quad \text{and} \quad f_{\text{even}}(x) := \sum_{\substack{k \geq 1 \\ 2m \leq p}} \delta_{2k} x^{2k}$$

and  $g(x) := P_p(x; \alpha)$ . Assume in addition that  $f_{\text{odd}}, f_{\text{even}} \in \mathcal{A}_{r_f}$  and  $g \in \mathcal{A}_{r_g}$  with  $g(B_{r_g}) \subset B_{r_f}$  (necessary if  $p = \infty$ ). Then

$$\mathcal{T}_p(f_{\text{odd}}(g(x)) + x f_{\text{even}}(g(x))) = -\delta_0 x \text{ for } x \in B_{r_f}. \quad (12)$$

Besides, if  $\delta_0 \delta_1 \neq 0$ , then there exists an analytic function  $h$  that solves (12) (with  $h$  instead of  $g$ ) and therefore  $\alpha_k := D^k h(0)/k!$  for any  $k \geq 1$ .

*Proof.* We restrict to the case  $p = \infty$ , see Remark 4. Let us introduce a new parameter  $\nu$  and we define

$$\check{f}(x, \nu) := f_{\text{odd}}(x) + \nu f_{\text{even}}(x)$$

and the analogous  $\check{f}^{\bullet}(z, \nu)$  with  $\delta^{\bullet}$  and  $g^{\bullet}(z) := P_{\infty}(z; \alpha^{\bullet})$ . As in Lemma 1, if  $\alpha^{\bullet}, \gamma^{\bullet} \in \mathbb{R}(\delta^{\bullet})$  and  $\delta^{\bullet}$  are linked by the relation

$$\begin{aligned} \sum_{m \geq 0} \delta_{2m+1}^{\bullet} \sum_{k_1, \dots, k_{2m+1} \geq 0} \alpha_{k_1}^{\bullet} \cdots \alpha_{k_{2m+1}}^{\bullet} z^{k_1 + \dots + k_{2m+1}} \\ + \nu^{\bullet} \sum_{m \geq 1} \delta_{2m}^{\bullet} \sum_{k_1, \dots, k_{2m} \geq 0} \alpha_{k_1}^{\bullet} \cdots \alpha_{k_{2m}}^{\bullet} z^{k_1 + \dots + k_{2m}} \\ = \sum_{m \geq 0} \gamma_{2m+1}^{\bullet} z^{2m+1} + \nu^{\bullet} \sum_{m \geq 1} \gamma_{2m}^{\bullet} z^{2m} \end{aligned} \quad (13)$$

in  $\mathbb{R}(\delta^{\bullet})[[z, \nu^{\bullet}]]$  for some indeterminate  $\nu^{\bullet}$ , then

$$\check{f}^{\bullet}(g^{\bullet}(x), \nu^{\bullet}) = \sum_{m \geq 0} \gamma_{2m+1}^{\bullet} x^{2m+1} + \nu^{\bullet} \sum_{m \geq 1} \gamma_{2m}^{\bullet} x^{2m}$$

If  $\alpha$  is  $(\beta^{\sim}, +\infty)$ -related to  $\delta$ , then (13) holds for  $\nu^{\bullet} = z$ ,  $\gamma_k = 0$  for  $k = 0, 2, 3, \dots$  and  $\gamma_1 = -\delta_0$  with  $\gamma_k := e(\gamma^{\bullet})$ . This proves (12).

We set  $q(y, x) = \delta_0 x + \check{f}(y, x)$ . If  $\delta_1 \neq 0$ , then  $D_y q(y, 0) \neq 0$ . From the analytic implicit function theorem (see *e.g.*, [1]), there exists a unique analytic function  $h$  defined in a neighborhood of 0 such that  $q(h(x), x) = 0$ , meaning that  $h$  solves  $f_{\text{odd}}(h(x)) + x f_{\text{even}}(h(x)) = -\delta_0 x$  and so (12). By uniqueness in a neighborhood of 0,  $h = g$ .  $\square$

### 3.6 The case “up-flat”

We consider the case  $\beta^{\leftarrow} = (1, 2, \dots, 2)$ , which we call the case “up-flat”. With Remark 2, (9) in the definition of  $(\beta^{\leftarrow}, p)$ -related sequences (Definition 2) can be replaced by

$$\delta_0^{\bullet} z + \sum_{m=1}^p \delta_m^{\bullet} \sum_{k_1 + \dots + k_m \leq p} \alpha_{k_1}^{\bullet} \cdots \alpha_{k_m}^{\bullet} z^{k_1 + \dots + k_m} = 0. \quad (14)$$

We now give an explicit, recursively computable expressions for the  $\alpha_k$ 's from the  $\delta_k$ . Other ways to express the  $\alpha_k$ 's are possible, for example using Bell polynomials. See also [13, 43] for considerations on the efficiency of numerical implementations.

**Lemma 7.** Assume that  $\delta_1 \neq 0$ . Then  $\alpha$  is  $(\beta^\leftarrow, p)$ -related to  $\delta$  if and only if

$$\alpha_1^\bullet = \frac{-\delta_0^\bullet}{\delta_1^\bullet} \text{ and } \alpha_q^\bullet = - \sum_{m=2}^q \frac{\delta_m^\bullet}{\delta_1^\bullet} \sum_{k_1+\dots+k_m=q} \alpha_{k_1}^\bullet \cdots \alpha_{k_m}^\bullet \text{ for } 2 \leq q \leq p. \quad (15)$$

In particular, there exists one and only one sequence  $\alpha$  which is  $(\beta^\leftarrow, p)$ -related to  $\delta$ . Besides, each of the  $\alpha_q^\bullet$  is a polynomial expression in  $\delta_m^\bullet/\delta_1^\bullet$ ,  $m = 0, \dots, q$  whose term in  $\delta_0^\bullet/\delta_1^\bullet$  has degree  $q$ .

*Proof.* We rewrite (14) as

$$\delta_0^\bullet z + \sum_{q=1}^p \sum_{m=1}^p \delta_m^\bullet \left( \sum_{k_1+\dots+k_m=q} \alpha_{k_1}^\bullet \cdots \alpha_{k_m}^\bullet \right) z^q = 0. \quad (16)$$

We identify the of  $\alpha_k^\bullet$ 's according to the power of  $z$ . In particular,  $\delta_0^\bullet + \delta_1^\bullet \alpha_1^\bullet = 0$  so that  $\alpha_1^\bullet = -\delta_0^\bullet/\delta_1^\bullet$ . Since  $\alpha_0^\bullet = 0$ , for  $2 \leq q \leq p$ ,

$$0 = \sum_{m=1}^p \delta_m^\bullet \sum_{k_1+\dots+k_m=q} \alpha_{k_1}^\bullet \cdots \alpha_{k_m}^\bullet = \delta_1^\bullet \alpha_q^\bullet + \sum_{m=2}^q \delta_m^\bullet \sum_{\substack{1 \leq k_1, \dots, k_m < q \\ k_1+\dots+k_m=q}} \alpha_{k_1}^\bullet \cdots \alpha_{k_m}^\bullet.$$

This proves that a  $(\beta^\leftarrow, p)$ -related sequence may be constructed for any  $\delta$  with  $\delta_1 \neq 0$ . The converse is proved by the same argument of identifying the coefficients in the power of  $z$ .

The form of the expression and the degree of the term  $\delta_0^\bullet/\delta_1^\bullet$  is proved by a direct induction.  $\square$

For  $p = 5$ , with  $\widehat{\delta}_k = -\delta_k/\delta_1$  and  $\alpha_1 = \widehat{\delta}_0$ , we have

$$\begin{aligned} \alpha_2 &= \widehat{\delta}_2 \widehat{\delta}_0^2, & \alpha_3 &= (\widehat{\delta}_3 + 2\widehat{\delta}_2^2) \widehat{\delta}_0^3, \\ \alpha_4 &= (\widehat{\delta}_4 + 5\widehat{\delta}_2 \widehat{\delta}_3 + 5\widehat{\delta}_2^3) \widehat{\delta}_0^4, & \alpha_5 &= (\widehat{\delta}_5 + 6\widehat{\delta}_2 \widehat{\delta}_4 + 21\widehat{\delta}_2^2 \widehat{\delta}_3 + 14\widehat{\delta}_2^4 + 3\widehat{\delta}_3^2) \widehat{\delta}_0^5. \end{aligned}$$

We now consider an alternative way to compute the coefficients in terms of the inverse of an analytic function. Let us consider the scrambling operator and (10), which was an alternative expression to (9), in this case becomes the following alternative expression to (14):

$$0 = \mathfrak{s}_{\beta^\leftarrow -2} \circ \mathfrak{p}_p \circ P_p(\cdot; \delta^\bullet) \circ P_p(z; \alpha^\bullet) = \mathfrak{p}_p \circ P_p(\cdot; \delta^\bullet) \circ P_p(z; \alpha^\bullet) - \delta_0^\bullet + \delta_0^\bullet z.$$

Therefore, (9) in Definition 2 (represented here by (14)) can be replaced by

$$\mathfrak{p}_p \circ f^\bullet \circ P_p(z; \alpha^\bullet) = -\delta_0^\bullet z \quad (17)$$

with  $f^\bullet(z) := P_p(z; \delta^\bullet) - \delta_0^\bullet = \sum_{k=1}^p \delta_k^\bullet z^k$ .

*Remark 5.* To prove the equivalence of (14) and (17) we could also use Lemma 1, if  $p \in \mathbb{N}$ . Indeed (14) is equivalent to (6) in Lemma 1 with  $\eta_0 = \delta_0$ ,  $\eta_1 = -\delta_0$  and  $\eta_k = 0$  for  $k \geq 2$  (let us recall that  $\alpha_0 = 0$  by definition).

We then see that to get  $\alpha$ , we compute the compositional inverse of a power series. Hence, our problem is related to the Lagrange inversion formula [22]. Another consequence is that one may look at other ways than series expansions to compute approximations of the root, for example using continued fractions or close form expressions.

The next lemma, Lemma 8, ensures that identifying the inverse of a function constructed as a series/polynomial of coefficients  $\delta_k$ 's yields the coefficients  $\alpha_k$ 's.

**Lemma 8.** *Let  $p \in \{1, 2, \dots, \infty\}$ ,  $\alpha$  be  $(\beta^\curvearrowright, p)$ -related to  $\delta$ , and  $f(x) := \sum_{k=1}^p \delta_k x^k = P_p(x; \delta) - \delta_0$  and  $g(x) := P_p(x; \alpha)$ . If  $p = \infty$ , assume in addition Hypothesis 2. Recall that  $g(0) = 0$ . Then*

$$\mathcal{T}_p(f \circ g(x)) = \mathfrak{p}_p(f \circ g(x)) = -\delta_0 x.$$

*If in addition  $\delta_0 \neq 0$  and  $\delta_1 \neq 0$  then  $f$  is invertible on an open neighborhood of 0 and*

$$\alpha_k = (-\delta_0)^k \mathbb{D}^k f^{-1}(0)/k! \quad \text{for any } k \geq 1.$$

*Proof.* The first part is a direct consequence of the combination of dummization-evaluation, Eq. (17) (see Lemmas 2-3). The second part of the statement follows from the Inverse Function Theorem for analytic functions [1] which one may applies when  $\delta_1 = \mathbb{D}f(0) \neq 0$ . The proof is completed, by uniqueness shown in Lemma 7.  $\square$

The next corollary follows from the fact that  $\mathcal{T}_p(f \circ g) = \mathfrak{p}_p(\mathcal{T}_p f \circ \mathcal{T}_p g)$ .

**Corollary 1.** *Let  $p \in \{1, 2, \dots, \infty\}$ ,  $\alpha$  be  $(\beta^\curvearrowright, p)$ -related to  $\delta$ , and let  $f, g$  functions of class  $\mathcal{C}^p$  (if  $p = \infty$ , assume in addition Hypothesis 2) such that  $\mathcal{T}_p f(x) = \sum_{k=1}^p \delta_k x^k = P_p(x; \delta) - \delta_0$  and  $\mathcal{T}_p g(x) = P_p(x; \alpha)$ . Then*

$$\mathcal{T}_p(f \circ g(x)) = -\delta_0 x.$$

*Remark 6.* Let  $p \neq \infty$ . If  $\alpha$  is  $(\beta^\curvearrowright, \infty)$ -related to  $\delta$ , then  $\{\alpha_k\}_{k=0}^p$  is  $(\beta^\curvearrowright, p)$ -related to  $\{\delta_k\}_{k=0}^p$ .

The above remark ensures that we can take always  $p = +\infty$ . For instance taking  $\delta_k = 0$  for  $k$  above a given threshold, say  $p$ , and we associate with  $\delta$  an infinite sequence  $\alpha$ .

**Example 3.** Let  $\delta_k = (-1)^k/\theta^{k-1}$  (resp.  $\delta_k = 1/\theta^{k-1}$ ) for  $k \geq 1$  and some  $\theta \in \mathbb{R}$ . Then for  $|x| < |\theta|$ ,

$$f(x) := \sum_{k=1}^{\infty} \delta_k x^k = -\frac{\theta x}{\theta + x} \quad (\text{resp. } f(x) = \frac{\theta x}{\theta - x}).$$

Since  $f^{-1}(x) = f(x)$  (resp.  $f^{-1}(x) = -f(-x) = \frac{\theta x}{\theta + x}$ ) then  $\alpha_k = \frac{\delta_0^k}{\theta^{k-1}}$  (resp.  $\alpha_k = -\frac{\delta_0^k}{\theta^{k-1}}$ ) for  $k \geq 1$ .

**Example 4.** Let  $\psi$  be an analytic function around 0 with  $\psi'(0) \neq 0$  (so that  $\psi$  is invertible around 0) and set

$$\delta_k := -\frac{1}{k!} \frac{D^k \psi(0)}{\psi'(0)} \text{ for } k \geq 1$$

as well as

$$f(z) := \sum_{k \geq 1} \delta_k z^k = \sum_{k \geq 1} \frac{-1}{k!} \frac{D^k \psi(0)}{\psi'(0)} z^k = \frac{\psi(0) - \psi(z)}{\psi'(0)}.$$

Then  $f^{-1}(z) = \psi^{-1}(\psi(0) - z\psi'(0))$ . By Faà di Bruno formula (7)  $D^k f^{-1}(z) = D^k \psi^{-1}(\psi(0) - z\psi'(0))\psi'(0)^k$  and

$$\alpha_k = \frac{(-\psi'(0)\delta_0)^k D^k \psi^{-1}(\psi(0))}{k!} = \frac{\psi(0)^k D^k \psi^{-1}(\psi(0))}{k!}.$$

We now give a non trivial example.

**Example 5.** Consider  $\theta \notin \{0, 1\}$ . Let us assume that  $\delta_1 = -1$  and  $\delta_k = (-\theta)^k$ ,  $k \geq 2$ . Then for  $|z| < 1/|\theta|$ ,

$$f(z) = -z + \sum_{k \geq 2} (-\theta)^k z^k = \frac{1}{1 + \theta z} - 1 + (\theta - 1)z.$$

To compute the inverse  $f^{-1}$ , we are looking for the solution  $h(z)$  to the quadratic equation

$$\frac{1}{1 + \theta h(z)} + (\theta - 1)h(z) = 1 - z, \quad \text{for any } z \text{ small enough.}$$

We assume  $\theta \in (0, 1)$ . The possible solutions are

$$h_{\pm}(z) = \frac{1 - \theta z \pm \sqrt{\theta^2 z^2 - 4\theta^2 z + 2\theta z + 1}}{2\theta(\theta - 1)},$$

where the quantity in the square root is always positive. But  $f(0) = 0$  hence  $f^{-1} = h_-$ . One could then check that  $h_-$  has the same series expansion in  $z$  as the one in Lemma 7. It is also possible to express  $h_-$  using a continued fraction.



**Corollary 2.** *Let  $r > 0$  and let  $\{f_n\}_n$  be a sequence of functions in  $\mathcal{A}_r$  and  $f \in \mathcal{A}_r$  be such that  $f_n$  converge uniformly to  $f$ . Then  $\delta_k^n := D^k f_n(0)/k!$  converges to  $\delta_k$  for any  $k \geq 1$ . Besides, if  $\delta_1 \neq 0$ , then for  $n$  large enough, the corresponding sequence  $\alpha^n$  which is  $(\beta^\frown, +\infty)$ -related to  $\delta^n = \{\delta_k^n\}_{k \geq 0}$  converges to  $\alpha$ , the sequence which is  $(\beta^\frown, +\infty)$ -related to  $\delta = \{\delta_k\}_{k \geq 0}$ .*

### 3.7 $(\beta, p)$ -related sequences in the context of Theorem 1

We have already seen in Lemma 4 how to derive a sequence  $\alpha(s)$  satisfying (5): by finding a  $(\beta, p)$ -related sequence to  $\delta(s)$  with  $\delta_k(s) = \{F\}_k(s)/k!$ ,  $k = 0, \dots, p$ . Hence, when  $\beta = \beta^\frown$  (resp.  $\beta = \beta^\smile$ ), Lemma 7 (resp. Lemma 5) provides a recursive expression of the desired sequence. We now apply Lemma 8 to provide a different expression in the case  $\beta = \beta^\frown$  in this context.

**Proposition 2.** *Let  $F$  be a function of class  $\mathcal{C}^{p+1}(\Theta, \mathbb{R})^{\mathbb{S}}$ . Take  $\theta_0 \in \Theta$  and set  $\{F\}_k := \varphi(s)^2 D^k F(s, \theta_0)$  for  $k \geq 1$  and  $\{F\}_0 := \varphi(s) F(s, \theta_0)$ . Assume that Hypotheses (ii)-(iv) of Theorem 1 holds, so that in particular,  $F(s, \cdot)$  is invertible around  $\theta_0$ . Define*

$$\mathcal{E}_{\theta_0} F(s, x) := F^{-1}\left(s, F(s, \theta_0)\left(1 - \frac{x}{\varphi(s)}\right)\right) - \theta_0.$$

Let  $\alpha_0 := 0$  and  $\alpha_k$  be the  $k$ -th Taylor coefficient of  $\mathcal{E}_{\theta_0} F(s, \cdot)$  for  $k \geq 1$ , that is

$$\alpha_k(s) := \frac{D^k \mathcal{E}_{\theta_0} F(s, 0)}{k!} = \frac{D^k F^{-1}(s, F(s, \theta_0))}{k!} \left( -\frac{\{F\}_0(s)}{\{F\}_1(s)} DF(s, \theta_0) \right)^k.$$

Then  $\{\alpha_k(s)\}_{k=0}^p$  is  $(\beta, p)$ -related to  $\{-\frac{1}{k!} \frac{\{F\}_k(s)}{\{F\}_1(s)}\}_{k=0}^p$ , it satisfies equation (5) (i.e. Item (i) of Theorem 1), and

$$\begin{aligned} \theta(s) &= \theta_0 + \mathcal{E}_{\theta_0} F(s, \varphi(s)) \\ &= \theta_0 + \mathcal{I}_p \mathcal{E}_{\theta_0} F(s, \varphi(s)) + O(\varphi(s)^{p+1}) = \theta_0 + \sum_{k=1}^p \alpha_k(s) \varphi(s)^k + O(\varphi(s)^{p+1}). \end{aligned}$$

*Proof.* By construction, for any  $x$  small enough,

$$\frac{F(s, \theta_0 + \mathcal{E}_{\theta_0} F(s, x)) - F(s, \theta_0)}{DF(s, \theta_0)} = \frac{-F(s, \theta_0)}{DF(s, \theta_0)} \frac{x}{\varphi(s)}. \quad (18)$$

In particular,  $\mathcal{E}_{\theta_0} F(s, 0) = 0$  and  $F(s, \theta_0 + \mathcal{E}_{\theta_0} F(s, \varphi(s))) = 0$  so that the root  $\theta(s)$  to  $F(s, \theta) = 0$  is  $\theta(s) = \theta_0 + \mathcal{E}_{\theta_0} F(s, \varphi(s))$ .

We set

$$\delta_k(s) := \frac{-D^k F(s, \theta_0)}{k! DF(s, \theta_0)} \text{ for } k = 1, \dots, p, \quad \delta_0(s) := -\frac{\{F\}_0(s)}{\{F\}_1(s)}, \text{ and } \boldsymbol{\delta} := \{\delta_k\}_k.$$

Then, inspired by Example 4, we set

$$\psi(s, z) = \frac{F(s, \theta_0) - F(s, \theta_0 + z)}{DF(s, \theta_0)}.$$

Hence  $\mathcal{T}_p \psi(s, z) = \sum_{k=1}^p \delta_k(s) z^k$  which is analytic. From (18),

$$\psi(s, \mathcal{E}_{\theta_0} F(s, x)) = -\delta_0(s)x.$$

Note that

$$\mathfrak{p}_p(\mathcal{T}_p \psi(s, \mathcal{T}_p \mathcal{E}_{\theta_0} F(s, x))) = \psi(s, \mathcal{E}_{\theta_0} F(s, x)) = -\delta_0(s)x.$$

This ensures by Eq. (17), together with Lemmas 2-3, that the coefficients  $\boldsymbol{\alpha}(s)$  of the Taylor expansion of  $\mathcal{E}_{\theta_0} F(s, \cdot)$  around 0 are  $(\boldsymbol{\beta}^{\leftarrow}, p)$ -related to  $\boldsymbol{\delta}(s)$ . Hypothesis (i) is satisfied, so that the conclusion of Theorem 1 holds.  $\square$

### 3.8 Perturbation of the roots and $(\boldsymbol{\beta}, p)$ -related sequences

Let  $\boldsymbol{\beta} \in \{\boldsymbol{\beta}^{\leftarrow}, \boldsymbol{\beta}^{\sim}\}$ . Let  $\Lambda : \mathbf{S} \rightarrow \mathbb{R}$  and set  $\{\Lambda\}_0(s) := \varphi(s)^{\beta_0} \Lambda(s)$ . Let us consider solving

$$F(s, \theta_\Lambda(s)) = \Lambda(s).$$

We define

$$\{F_\Lambda\}_k(s) := \begin{cases} \{F\}_0(s) - \{\Lambda\}_0(s) & \text{for } k = 0, \\ \{F\}_k(s) & \text{for } k > 0 \end{cases}$$

as well as the corresponding  $\delta_{\Lambda k}(s)$  and  $\alpha_{\Lambda k}(s)$ .

**Proposition 3.** *For each  $k \geq 1$ , the coefficients  $\alpha_{\Lambda k}(s)$  are polynomial expressions in  $\{F\}_\ell(s)/\{F\}_1(s)$ ,  $\ell \leq k$  and  $\{\Lambda\}_0(s)/\{F\}_1(s)$ . Moreover, if  $\{\Lambda\}_0/\{F\}_1$  converges to 0 with  $s$  and provided that  $\alpha_k$  converges, then  $\alpha_{\Lambda k}$  and  $\alpha_k$  have the same limit for each  $k \geq 0$ .*

*Proof.* The proof is immediate once it has been noticed by  $\delta_{\Lambda k} = \delta_k$  for  $k > 1$  and  $\delta_{\Lambda 0} = \delta_0 - \{\Lambda\}_0$ . The expression of  $\alpha_{\Lambda k}(s)$  in terms of  $\{F\}_\ell(s)/\{F\}_1(s)$  and  $\Lambda(s)/\{F\}_1(s)$  are obtained by the Newton formula.  $\square$

## 4 The stochastic asymptotic implicit function theorem

In this section we consider the inference of a parameter  $\theta_0 \in \mathbb{R}$  and we assume that  $\Theta$  contains an open interval around  $\theta_0$ .

To deal with convergence, we need some topological assumptions. We consider that  $(\mathbf{S}, \prec)$  is a directed set that is a set with a partial order  $\prec$  such that for any  $s, s'$  in  $\mathbf{S}$ , there exists  $r \in \mathbf{S}$  such that  $s \prec r$  and  $s' \prec r$ . Typically, we consider  $\mathbf{S} = \mathbb{R}$  or  $\mathbf{S} = \mathbb{N}$  (sample size) yet this framework allows consider partitions and pairs (sample size and discretization step).

*Hypothesis 3.* The rate function  $\varphi : \mathbf{S} \rightarrow \mathbb{I} \cap (0, 1]$  is monotone decreasing, that is  $\varphi(s) \leq \varphi(s')$  whenever  $s' \prec s$ .

Note that, since  $(\mathbf{S}, \prec)$  is a directed set, the rate function is a *net*.

**Notation 6.** Let  $F \in \mathcal{C}^p(\Theta, \mathbb{R})^{\mathbf{S}}$  such that  $DF(s, \theta_0) \neq 0$  then, analogously to Notation 2, let for all  $k \in \{0, \dots, p\}$ ,

$$\{F\}_k(s) := \varphi(s)^{\beta_k} D^k F(s, \theta_0).$$

Let  $\delta^F$  be the family of  $(\mathbb{R}^{p+1})^{\mathbf{S}}$  defined by

$$\delta^F(s) := \left\{ -\frac{1}{k!} \frac{\{F\}_k(s)}{\{F\}_1(s)} \right\}_{k=0, \dots, p}.$$

For the expression of the  $(\beta^\frown, p)$ -related (resp.  $(\beta^\smile, p)$ -related) sequence to  $\delta^F(s)$  we refer to Lemma 7 (resp. Lemma 5). In the case “up-flat” the sequence can be given also by Proposition 2.

**Proposition 4.** *We consider a family of random functions  $(F(s, \theta))_{(s, \theta) \in \mathbf{S} \times \Theta}$  on a probability space  $(\Omega, \mathcal{F}, \mathbb{P})$  as well as  $p$  as in Hypothesis 1. Assume that for all  $s \in \mathbf{S}$  and  $\mathbb{P}$ -almost all  $\omega \in \Omega$  the function  $\theta \mapsto F(s, \theta)$  is of class  $\mathcal{C}^p$  at  $\theta_0$ , and there exists a random vector  $\boldsymbol{\mu} = \{\mu_k\}_{k=0}^p$  such that*

$$(\{F\}_0(s), \dots, \{F\}_p(s)) \xrightarrow[s \rightarrow \infty]{\text{dist.}} \boldsymbol{\mu}.$$

For every  $s \in \mathbf{S}$  let  $\boldsymbol{\alpha}(s)$  the  $(\beta, p)$ -related sequence to  $\delta^F(s)$ . Then

- (i)  $\delta^F(s) \xrightarrow[s \rightarrow \infty]{\text{dist.}} \boldsymbol{\delta} := \left\{ -\frac{1}{m!} \frac{\mu_m}{\mu_1} \right\}_{m=0, \dots, p}$ .
- (ii)  $\boldsymbol{\alpha}(s) \xrightarrow[s \rightarrow \infty]{\text{dist.}} \boldsymbol{\alpha}$  where  $\boldsymbol{\alpha}$  is the  $(\beta, p)$ -related sequence to  $\boldsymbol{\delta}$ .

**Definition 4** (Boundedness in probability). A family of random variables  $\{X_s\}_{s \in \mathbf{S}}$  is said to be *bounded in probability* if for any  $\epsilon > 0$ , there exist  $K \geq 0$  and  $s \in \mathbf{S}$  such that  $\sup_{s < s'} \mathbb{P}[|X_{s'}| \geq K] \leq \epsilon$ .

*Hypothesis 4.* We consider a family of random functions  $(F(s, \theta))_{(s, \theta) \in \mathbf{S} \times \Theta}$  on a probability space  $(\Omega, \mathcal{F}, \mathbb{P})$  as well as  $\beta \in \{\beta^{\smile}, \beta^{\smile\smile}\}$  and  $p$  as in Hypothesis 1 and the rate function as in Hypothesis 3. Assume that there exists an open neighborhood  $U \subseteq \Theta$  of  $\theta_0$  such that for all  $s \in \mathbf{S}$  and  $\mathbb{P}$ -almost all  $\omega \in \Omega$ :

- (i)  $\theta \mapsto F(s, \theta)$  is of class  $\mathcal{C}^{p+1}$  in  $U$ .
- (ii) There exists  $\theta(s) \in U$  such that  $F(s, \theta(s)) = 0$ .
- (iii)  $DF(s, \theta) \neq 0$  for all  $\theta \in U \setminus \{\theta_0\}$  and  $\{1/\{F\}_1(r)\}_{r \in \mathbf{S}}$  is bounded in probability.
- (iv) The family  $\{N_r(F)\}_{r \in \mathbf{S}}$  is bounded in probability with

$$N_s(F) := \sup_{1 \leq k \leq p+1} \sup_{\theta \in U} |\varphi(s)^{\beta_1} D^k F(s, \theta)|.$$

The following proposition is a direct consequences of Theorem 1.

**Theorem 3.** Assume Hypothesis 4 and assume that the net  $\varphi$  in Hypothesis 3 converges to 0. Assume that there exists a random vector  $\boldsymbol{\mu} = \{\mu_k\}_{k=0}^p$  such that

$$(\{F\}_0(s), \dots, \{F\}_p(s)) \xrightarrow[s \rightarrow \infty]{dist.} \boldsymbol{\mu},$$

and for every  $s \in \mathbf{S}$  let  $\boldsymbol{\alpha}(s)$  the  $(\beta, p)$ -related sequence to  $\delta^F(s)$ . Let  $\boldsymbol{\alpha}$  and  $\boldsymbol{\delta}$  as in Proposition 4. Then, using Notation 3,

- (i)  $P_p(\varphi(s); \boldsymbol{\alpha}(s)) = \sum_{k=0}^p \alpha_k(s) \varphi(s)^k$  converges to 0 in probability.
- (ii)  $|\theta(s) - \theta_0 - P_p(\varphi(s); \boldsymbol{\alpha}(s))| \varphi(s)^{-(p+1)}$  is bounded in probability.
- (iii)  $\varphi(s)^{-(k+1)} (P_p(\varphi(s); \boldsymbol{\alpha}(s)) - P_k(\varphi(s); \boldsymbol{\alpha}(s))) \xrightarrow[s \rightarrow \infty]{dist.} \alpha_{k+1}$  and

$$\varphi(s)^{-(k+1)} (\theta(s) - \theta_k(s)) \xrightarrow[s \rightarrow \infty]{dist.} \alpha_{k+1}$$

with  $\theta_k(s) := \theta_0 + P_k(\varphi(s); \boldsymbol{\alpha}(s))$  for  $k \in \{0, \dots, p-1\}$  (with  $P_0 \equiv 0$ ). In particular if  $k = 0$

$$\varphi(s)^{-1} (\theta(s) - \theta) \xrightarrow[s \rightarrow \infty]{dist.} \alpha_1.$$

- (iv)  $\theta(s) - \theta_p(s, \infty)$  converges to 0 in probability with rate  $\varphi(s) (\alpha_1(s) - \alpha_1)$  where  $\theta_p(s, \infty) := \theta_0 + P_p(s, \boldsymbol{\alpha}) = \theta_0 + \sum_{k=1}^p \alpha_k \varphi(s)^k$ .

The previous result exhibits two approximations of  $\theta(s)$ , i.e.,

$$\theta_p(s) := \theta_0 + P_p(\varphi(s), \boldsymbol{\alpha}(s)) \text{ and } \theta_p(s, \infty) := \theta_0 + P_p(\varphi(s), \boldsymbol{\alpha}).$$

In Section 6 we consider applications in which the estimator  $\theta(s)$  of  $\theta_0 \in \Theta$  (assume  $\Theta$  is open) is a zero of  $\theta \mapsto F(s, \theta) := g(\theta)T(s) - f(\theta)$  where  $f, g \in \mathcal{C}^p(\Theta)$  and  $T(s)$  is a sample statistic. Indeed this is often the case for MLE estimators for independent observations with distributions belonging to an exponential family or for the method of moments.

We aim at highlighting the following facts:

(F1) When  $F$  is the score, the result above shows the importance of the hidden variable  $\alpha_1(s) = -\{F\}_0(s)/\{F\}_1(s)$  to assess the quality of the MLE. This variable  $\alpha_1(s)$  depends on  $F$  and its derivative. The other variables  $\alpha_k(s)$  are polynomial expressions in  $\alpha_1(s)$  and  $-\{F\}_k(s)/k!\{F\}_1(s)$  for  $k \geq 1$ . In the case “up-flat”, owing to the homogeneity in  $\alpha_1(s)$  granted by Lemma 7,

$$\theta(s) = \theta_0 + \sum_{k=1}^p \xi_k(s) \alpha_1(s)^k \varphi(s)^k + O(\varphi(s)^{p+1}),$$

where the coefficients  $\xi_k(s)$ ,  $k \geq 2$  are polynomial expressions in the terms  $-\{F\}_\ell(s)/\ell!\{F\}_1(s)$ ,  $1 \leq \ell \leq k$ , and  $\xi_1(s) = 1$ . We shorten this as

$$\theta(s) = \theta_0 + H(s, \alpha_1(s)\varphi(s)) + O(\varphi(s)^{p+1}), \quad (19)$$

where  $H(s, \cdot)$  is a non-linear (polynomial or analytic) random function with  $H(s, x) = x + O(x^2)$ .

In many practical situations, the term  $\{F\}_0(s)/\{F\}_1(s) = -\alpha_1(s)$  fluctuates asymptotically as a (mixed) Gaussian random variable, while the coefficients of the expansion of  $H(s, \cdot)$  converges in probability to constants.

A similar result holds for the case “zig-zag” ( $\beta = \beta^\sim$ ), excepted that the expression is no longer polynomial in  $\alpha_1(s)\varphi(s)$ , but still polynomial in  $\alpha_1(s)$  and  $\varphi(s)$  (granted by Lemma 5). Some of the coefficients of  $H$  converges in probability to (mixed) Gaussian distributions.

(F2) Considering  $\theta_p(s, \infty)$  instead of  $\theta_p(s)$  consists in replacing  $\alpha_1(s)$  and  $H(s, \cdot)$  by their limits. It gives then a “proxy” of the behavior of the MLE which is useful when the convergence of  $\alpha_k(s)$ ’s toward their limits is faster than the one of  $\theta_p(s)$ . This happens with the Skew Brownian motion [31], or with the exponential distribution as seen below in Example 6.

(F3) In the examples in Section 6 below,  $\alpha_1(s)^2\{F\}_1(s)$  (alternatively  $\alpha_1(s)^2\mu_1$  and  $\{F\}_0(s)^2/\{F\}_1(s)$  and  $\{F\}_0(s)^2/\mu_1$ ) appears to be asymptotically “pivotal” for  $s$  large enough, in the sense that its distribution no longer depends on the parameter. We could say it is the *Wald’s statistics*. Hence, we could change (19) to express the MLE as a non-linear transform of one of these pivotal quantities. A further

direction of research will be to use this to design confidence intervals or hypothesis tests that takes the lack of normality into account.

(F4) The upper error bound in Theorem 1 (non-random version of Theorem 3) also depends on  $\alpha_1(s)$ . Therefore, when  $\alpha_1(s)$  deviates from 0, the truncated expansions  $\theta_p(s)$  are not always accurate.

(F5) In some situations, for  $k \geq 1$ ,  $\delta_k(s) := -\{F\}_k(s)/k!\{F\}_1(s)$  is both deterministic and constant in  $s$  (see Section 6.2). In the case “up-flat”

$$\theta(s) = \theta_0 + H\left(-\frac{\{F\}_0(s)}{\{F\}_1(s)}\varphi(s)\right) = \theta_0 + H(\alpha_1(s)\varphi(s))$$

and  $\theta(s, \infty) = \theta_0 + H(\alpha_1\varphi(s))$

for a deterministic, polynomial (or analytic function  $H$ ). The randomness in the MLE comes only from the one of  $\{F\}_0(s)$  that usually converges thanks to a CLT. We have seen that  $\varphi(s)^{-1}(\theta(s) - \theta_0)$  converges to  $\alpha_1$ , and often in applications  $\alpha_1$  has a Gaussian distribution  $\mathcal{N}(0, \sigma^2)$ .

Now, we quantify the distance of the distribution function of

$$\varepsilon(s) := \varphi(s)^{-1}(\theta_p(s) - \theta_0) = \varphi(s)^{-1}H(\alpha_1\varphi(s))$$

from the one of its limiting distribution. Let us consider the Kolmogorov-Smirnov distance: For any two random variables  $X_1, X_2$ , let

$$\Delta(X_1; X_2) := \sup_{x \in \mathbb{R}} |\mathbb{P}[X_1 \leq x] - \mathbb{P}[X_2 \leq x]|. \quad (20)$$

Whenever  $H$  is invertible with inverse  $H^{-1}$ , we have

$$\begin{aligned} \Delta(\varepsilon(s); \alpha_1) &\leq \Delta(\varepsilon(s); \varphi(s)^{-1}H(\alpha_1\varphi(s))) + \Delta(\varphi(s)^{-1}H(\alpha_1\varphi(s)); \alpha_1) \\ &= \Delta(\alpha_1(s); \alpha_1) + \Delta(\varphi(s)^{-1}H(\alpha_1\varphi(s)); \alpha_1). \end{aligned}$$

The distance  $\Delta(\varepsilon(s); \alpha_1)$  is then given by the competition of two terms  $\Delta(\alpha_1(s); \alpha_1)$  and  $\Delta(\varphi(s)^{-1}H(\alpha_1\varphi(s)); \alpha_1)$ . The quantity  $\Delta(\alpha_1(s); \alpha_1)$  is the distance of  $\varepsilon(s)$  to a possibly non linear transformation of  $\alpha_1\varphi(s)$ . Indeed in the case we are considering  $H$  is deterministic and does not even depend on the sample size. The fact that  $\Delta(\alpha_1(s); \alpha_1) \approx 0$  often follows from Berry-Esséen result and it rewrites: for any  $x \in \mathbb{R}$ ,

$$\mathbb{P}[\varepsilon(s) \leq x] \approx \mathbb{P}\left[\alpha_1 \leq \frac{H^{-1}(x\varphi(s))}{\varphi(s)}\right].$$

The term  $\Delta(\varphi(s)^{-1}H(\alpha_1\varphi(s)); \alpha_1)$  concerns the error introduced by the non-linearity. We show in Example 6 that the effect of the latter error might be predominant.

Note that we can also consider

$$\theta(s) = \theta_0 + H\left(-\frac{\{F\}_0(s)}{\{F\}_1(s)}\varphi(s)\right) + O(\varphi(s)^{p+1})$$

and  $\theta_p(s, \infty) = \theta_0 + H(\alpha_1\varphi(s)),$

or dealing when  $H$  depends on  $s$  and is possibly random by using the limit of its coefficients as an approximation.

(F6) Sometimes a suitable change of variable, as in the following Section 5.1, provides an expression of  $\theta(s)$  instead of an approximation (see Example 8 below).

(F7) There are also situations which exhibit some *boundary layer*, meaning that the coefficients  $\alpha_k$  are themselves high and  $\theta_p$  are not suitable approximation of the MLE. This is the case for the binomial distribution with a parameter  $\theta_0$  close to 0 or 1. See Example 7.

(F8) For specific values of the parameter to estimate, sometimes several choices of  $\beta$  are possible (note that  $\beta_k^\wedge \leq \beta_k^\frown$ ), leading to different finite approximation of the estimator. We call this phenomena phase transition. See section 6.3 for an example.

## 5 Reparametrization and Changes of scale

### 5.1 Reparametrization through changes of variables

In the context of statistical estimation, it is sometimes convenient to perform some change of variable on the space of parameters.

Let us consider a one-to-one map  $\Phi$  from the space of parameters  $\Theta$  to the space  $\Theta^\dagger$  with inverse  $\Psi$ . We assume that  $\Phi \in \mathcal{C}^{p+1}$ . For the sake of simplicity, we assume that 0 belongs to the interiors of both  $\Theta$  and  $\Theta^\dagger$ . Moreover, we assume that  $\Phi(0) = 0$ .

**Notation 7** (Pullback). For a function  $f : \Theta \rightarrow \mathbb{R}$ , we set  $\Psi^*f := f \circ \Psi$  so that  $\Psi^*f$  is a function from  $\Theta^\dagger$  to  $\mathbb{R}$ . The operator  $\Psi^*$  is the *pullback*. For  $f : \mathbf{S} \times \Theta \rightarrow \mathbb{R}$ , we extend the pullback to  $\Psi^*f(s, \eta) = f(s, \Psi(\eta))$  for  $(s, \theta) \in \mathbf{S} \times \Theta^\dagger$ . For a function  $F \in \mathcal{C}^{p+1}(\Theta, \mathbb{R})^{\mathbf{S}}$ , we define  $F^\dagger(s, \eta) := \Psi^*F(s, \eta) = F(s, \Psi(\eta))$  which is a function in  $\mathcal{C}^{p+1}(\Theta^\dagger, \mathbb{R})^{\mathbf{S}}$ . We also set  $\{F^\dagger\}_k(s) := \varphi(s)^{\beta_k} D^k F^\dagger(s, 0)$  for  $k = 0, \dots, p$ .

We focus on the case “up-flat” ( $\beta = \beta^\frown$ ). From the Faà di Bruno formula (Proposition 1), we see that  $\{\{F^\dagger\}_k(s)\}_{k=0, \dots, p}$  converges whenever  $\{\{F\}_k(s)\}_{k=0, \dots, p}$

does. This is not necessarily true in the case “zig-zag”. Indeed, let  $\varphi(s) \in (-1, 1)$ , then in the “zig-zag” case,  $\varphi(s)^{-1}\{F\}_1(s)$  may diverge and so

$$\{F^\dagger\}_2(s) = \{F\}_2(s)(D\Psi(0))^2 + \varphi(s)^{-1}\{F\}_1(s)D^2\Psi(0)$$

does not converge neither. The same happens for  $\{F^\dagger\}_{2k}(s)$ , by Faà di Bruno formula (7) (in Proposition 1). So, starting from “zig-zag” case, taking  $F^\dagger$  we bump into the “up-flat” case. More generally, if  $\varphi(s) \in (-1, 1)$ , starting from  $\beta = (\beta_0, \beta_1, \beta_2, \dots, \beta_p)$ , we bump into  $(\beta_0, \beta_1, \beta_2^*, \beta_3^*, \dots, \beta_p^*)$  where  $\beta_k^* := \max_{n \in \{1, \dots, k\}} \beta_n$ . This latter kind of coefficient, as  $\beta^\frown$ , is “stable” under such changes of variables.

**Proposition 5.** Fix  $s \in \mathbf{S}$ . Let us define  $\delta = \{\delta_k\}_{k=0, \dots, p}$  and  $\delta^\dagger = \{\delta_k^\dagger\}_{k=0, \dots, p}$  by

$$\delta_k(s) := \frac{\{F\}_k(s)}{k!} \text{ and } \delta_k^\dagger(s) := \frac{\{F^\dagger\}_k(s)}{k!} \text{ for } k = 0, 1, \dots, p.$$

Let  $\alpha(s)$  (resp.  $\alpha^\dagger(s)$ ) be the unique  $(\beta^\frown, p)$ -related sequence to  $\delta(s)$  (resp.  $\delta^\dagger(s)$ ) given in Lemma 7. Then

$$\alpha_q(s) = \sum_{m=1}^q \frac{D^m \Psi(0)}{m!} \sum_{k_1 + \dots + k_m = q} \alpha_{k_1}^\dagger(s) \cdots \alpha_{k_m}^\dagger(s) \text{ for } q = 1, \dots, p, \quad (21)$$

or, in other words,

$$\sum_{k=1}^p \alpha_k(s) \bullet z^k := \mathfrak{p}_p \circ \mathcal{T}_p^\bullet \Psi \left( \sum_{k=1}^p \alpha_k^\dagger(s) \bullet z^k \right),$$

where  $\mathcal{T}_p^\bullet \Psi \in \mathbb{R}(\delta^\bullet(s))[[z]]$  is the formal Taylor expansion of  $\Psi$  (see Definition 1).

The meaning of this proposition is that expanding the root  $\theta^\dagger(s)$  of  $F^\dagger(s, \theta^\dagger(s)) = 0$  in terms of  $\varphi(s)$  using our procedure of Lemma 7 and then extending  $\Psi(\theta^\dagger(s))$  using a Taylor series leads to the same expansion as the one given computing the root  $\theta(s)$  of  $F(s, \theta(s)) = 0$  using Lemma 7. This is summarized in the diagram of Figure 1.

*Proof.* In this proof let  $F_\bullet := \mathcal{T}_p^\bullet F$  and  $F_\bullet^\dagger := \mathcal{T}_p^\bullet F^\dagger$  from Definition 1. By (11) and Lemma 7, the sequence  $\alpha(s) \bullet \in \mathbb{R}(\delta^\bullet(s))^p$  corresponding to  $\alpha(s)$  is the unique solution to

$$\mathfrak{s}_{\beta^\frown} \circ \mathfrak{p}_p \circ F_\bullet(s, P(z)) = 0 \text{ with } P(z) := \sum_{k=1}^p \alpha_k(s) \bullet z^k \text{ and } \alpha_k(s) \bullet \in \mathbb{R}(\delta(s)^\bullet). \quad (22)$$



$$\begin{array}{ccc}
(F, \theta) & \xrightarrow{\text{expansion}} & \sum_{k=1}^p \alpha_k^\bullet z^k & \xrightarrow{\text{evaluation}} & \theta_0 + \sum_{k=1}^p \alpha_k \varphi^k \\
\downarrow \Psi^* \times \Phi & & \uparrow \mathfrak{p}_p \circ \mathcal{T}_p^\bullet \Psi & & \\
(F^\dagger, \theta^\dagger) & \xrightarrow{\text{expansion}} & \sum_{k=1}^p \alpha_k^\dagger z^k & \xrightarrow{\text{evaluation}} & \theta_0^\dagger + \sum_{k=1}^p \alpha_k^\dagger \varphi^k
\end{array}$$

Figure 1: Change of variable.

Similarly,  $\alpha^\dagger(s)^\bullet$  is the unique solution to

$$\mathfrak{s}_{\beta^\frown} \circ \mathfrak{p}_p \circ F_\bullet^\dagger(s, P^\dagger(z)) = 0 \text{ with } P^\dagger(z) := \sum_{k=1}^p \alpha_k^\dagger(s)^\bullet z^k = 0$$

and  $F_\bullet^\dagger(z) = \sum_{k=0}^p \delta_k^\dagger(s)^\bullet z^k \in \mathbb{R}[\delta^\dagger(s)^\bullet][[z]]$ .

The  $\alpha_k^\dagger(s)^\bullet$ 's are rational functions in  $\mathbb{R}(\delta^\dagger(s)^\bullet)$ . From the very definition,  $F^\dagger(s, \eta) := F(s, \Psi(\eta))$ . By the Faà di Bruno formula (Proposition 1),  $\{F^\dagger\}_k(s)$  is a linear combination of the elements  $\{F\}_\ell(s)$ ,  $\ell = 1, \dots, k$  and the derivatives of  $\Psi$ . Therefore, the  $\alpha_k^\dagger(s)^\bullet$ 's and  $\delta_k^\dagger(s)^\bullet$ 's can be seen as rational functions in  $\mathbb{R}(\delta(s)^\bullet)$ . Thus  $F_\bullet^\dagger(z)$  and  $\mathcal{T}_p^\bullet \Psi \circ P^\dagger(z)$  can be seen as elements of  $\mathbb{R}[\delta(s)^\bullet][[z]]$ .

By Lemma 2,

$$\mathfrak{p}_p \circ F_\bullet^\dagger = F_\bullet^\dagger = \mathcal{T}_p^\bullet(\Psi^* F) = \mathfrak{p}_p \circ (\mathcal{T}_p^\bullet \Psi)^* F_\bullet.$$

Since  $\mathcal{T}_p^\bullet \Psi \circ P^\dagger(z) \in \mathbb{R}[\delta(s)^\bullet][[z]]$  and

$$0 = \mathfrak{s}_{\beta^\frown} \circ \mathfrak{p}_p \circ F_\bullet^\dagger(s, P^\dagger(z)) = \mathfrak{s}_{\beta^\frown} \circ \mathfrak{p}_p \circ F_\bullet(s, \mathcal{T}_p^\bullet \Psi \circ P^\dagger(z)).$$

Hence  $\mathcal{T}_p^\bullet \Psi \circ P^\dagger(z)$  solves (22), which is known to have a unique solution,  $P(z)$ . Thus  $P(z) = \mathcal{T}_p^\bullet \Psi \circ P^\dagger(z)$ . The result follows by an application of the evaluation map.  $\square$

## 5.2 Change of scale

We now introduce another change of variable, called *change of scale*, leading to a different expansion. The impact of this change of scale technique will be seen in Section 6.2 to obtain a “natural” expansion in the context of exponential families.

Let us take a one-to-one function  $\Phi \in \mathcal{C}^{p+1}(\Theta, \Theta^\dagger)$  whose derivative  $\Phi'$  never vanishes. Such a function  $\Phi$  is seen as a *scale*.

**Notation 8.** For a differentiable function  $f : \mathbb{R} \rightarrow \mathbb{R}$ , we write  $D_\Phi f = Df/D\Phi$ , where  $D$  is the derivative operator.

Since we solve  $F(s, \theta) = 0$ , a natural operation consists in setting

$$\check{F}(s, \theta) := \Phi^\sharp F(s, \theta) := \frac{F(s, \theta)}{D\Phi(\theta)}$$

and to solve  $\check{F}(s, \theta) = 0$ . The root  $\theta(s)$  is the same for  $F$  and  $\check{F}$ . However, as discussed below, the expansion could be different.

In the case of the MLE, as  $F(s, \cdot)$  takes scalar values, there exists a function  $G \in \mathcal{C}^{p+2}(\Theta, \mathbb{R})^{\mathbf{S}}$  such that  $DG(s, \cdot) = F(s, \cdot)$ . Considering  $\check{F}(s, \theta)$  corresponds to a change of variable at the level of the function  $G$ . Indeed, with Notation 8,

$$D_\Phi G(s, \theta) = \frac{DG(s, \theta)}{D\Phi(\theta)} = \check{F}(s, \theta) \text{ in } \mathcal{C}^{p+1}(\Theta, \mathbb{R})^{\mathbf{S}}.$$

We set  $G^\dagger(s, x) := G(s, \Psi(x))$  and  $F^\dagger(s, x) := F(s, \Psi(x))$  so that from standard computations,

$$D(G^\dagger) = (D_\Phi G)^\dagger \text{ and } F^\dagger = D_\Psi(G^\dagger).$$

Therefore,

$$\check{F}^\dagger := (\check{F})^\dagger = D(G^\dagger) \text{ so that } \check{F} = \Phi^* D(\Psi^* G).$$

We denote by  $\check{\alpha}(s)$  and  $\check{\alpha}^\dagger(s)$  the unique  $(\mathcal{B}^\leftarrow, p)$ -related sequences to  $\{\{\check{F}\}_k(s)/k!\}_{k=0, \dots, p}$  and  $\{\{\check{F}^\dagger\}_k(s)/k!\}_{k=0, \dots, p}$  defined by Lemma 7.

Proposition 5 may be applied to compute  $\check{\alpha}^\dagger(s)$  and then  $\check{\alpha}(s)$ , instead of computing the derivative of  $\check{F}$ . We summarize the various relationships in Figure 2.

Are  $P_{\check{\alpha}}$  and  $P_\alpha$  the same?

Let us reduce to the case  $\Phi(0) = 0$ . If  $\Phi(x) = \lambda x$  for some  $\lambda \in \mathbb{R} \setminus \{0\}$ , then the two expansions coincide by the fact that  $F = DG = \lambda D_\Phi G$  and Remark 3.

Note that if  $\alpha_1 \neq 0$  and  $\check{\alpha}_1 \neq 0$  then they are different unless  $\Phi''(0) = 0$  since

$$\check{\alpha}_1(s) = \frac{\alpha_1(s)}{1 + \frac{\Phi''(0)}{\Phi'(0)} \alpha_1(s) \varphi(s)^{\beta_1 - \beta_0}}.$$

Let us observe that  $\check{\alpha}_1$  and  $\alpha_1$  coincide asymptotically as  $s \rightarrow \infty$ , when  $\varphi(s)$  vanish and  $\alpha_1(s)$  converges in distribution.

$$\begin{array}{ccccccc}
G^\dagger & \xrightarrow{D_\Psi} & (F^\dagger, \theta^\dagger) & \xrightarrow{\text{expansion}} & \sum_{k=1}^p \alpha_k^\dagger \bullet z^k & \xrightarrow{\text{evaluation}} & \theta^\dagger + \sum_{k=1}^p \alpha_k^\dagger \varphi^k \\
\uparrow \Psi^* & & \uparrow \Psi^* \times \Phi & & \downarrow \mathfrak{p}_p \circ \mathcal{T}_p^\bullet \Psi & & \\
G & \xrightarrow{D} & (F, \theta) & \xrightarrow{\text{expansion}} & \sum_{k=1}^p \alpha_k^\bullet z^k & \xrightarrow{\text{evaluation}} & \theta + \sum_{k=1}^p \alpha_k \varphi^k \\
\downarrow \text{Id} & & \downarrow \Phi^\# \times \text{Id} & & & & \\
G & \xrightarrow{D_\Phi} & (\check{F}, \theta) & \xrightarrow{\text{expansion}} & \sum_{k=1}^p \check{\alpha}_k^\bullet z^k & \xrightarrow{\text{evaluation}} & \theta + \sum_{k=1}^p \check{\alpha}_k \varphi^k \\
\downarrow \Psi^* & & \downarrow \Psi^* \times \Phi & & \uparrow \mathfrak{p}_p \circ \mathcal{T}_p^\bullet \Psi & & \\
G^\dagger & \xrightarrow{D} & (\check{F}^\dagger, \theta^\dagger) & \xrightarrow{\text{expansion}} & \sum_{k=1}^p \check{\alpha}_k^\dagger \bullet \varphi^k & \xrightarrow{\text{evaluation}} & \theta^\dagger + \sum_{k=1}^p \check{\alpha}_k^\dagger \varphi^k
\end{array}$$

Figure 2: Change of scales and change of variable.

## 6 An application to parametric models

### 6.1 Exponential family

In this section, we consider the class univariate models from the exponential family of laws. This is a broad class of random variables [14] with a wide range of applications in statistical theory.

Let  $\Theta$  be an open interval and let  $n$  independent copies  $X_1, \dots, X_n$  with joint distribution under  $\mathbb{P}_\theta$  of the form

$$p(x_1, \dots, x_n, \theta) = \left( \prod_{j=1}^n h(x_j) \right) \exp \left( w(\theta) \sum_{j=1}^n T(x_j) - nA(\theta) \right) \quad (23)$$

for suitable functions  $T : \mathbb{R} \rightarrow \mathbb{R}$ ,  $h : \mathbb{R} \rightarrow \mathbb{R}$  and  $w, A \in \mathcal{C}^{p+1}(\Theta)$ ,  $p \geq 1$ . Hence we are considering an univariate parametric model from an exponential family of laws. The exponential family covers a wide variety of models [7].

Therefore the MLE  $\theta(n)$ , assume it exists, is the zero of the score function

$$\theta \mapsto F(n, \theta) := n(w'(\theta)T_n - A'(\theta)) \quad \text{where} \quad T_n := \frac{1}{n} \sum_{j=1}^n T(X_j).$$

We now apply the results obtained in the previous section to obtain an expansion of the MLE estimator  $\theta(n)$ . Before to do so, we would like to make some remarks.

First  $\theta(n)$  solves

$$T_n = \frac{A'(\theta)}{w'(\theta)} = D_w A(\theta)$$

where Notation 8 is used.

Next, if on one hand  $\mathbb{E}_\theta[F(n, \theta)] = nw'(\theta)\mathbb{E}_\theta[T_n] - nA'(\theta)$ , on the other hand (if the model is regular, in particular  $\partial_\theta^j p(x, \theta)$  is integrable and the following integrals are finite and equal  $\int \partial_\theta^j p(x, \theta) dx = \partial_\theta^j \int p(x, \theta) dx$ , for  $j = 1, 2$ ) it holds that  $\mathbb{E}_\theta[F(n, \theta)] = 0$ . We deduce that  $\mathbb{E}_\theta[T_n] = D_w A(\theta)$  for all  $\theta \in \Theta$ . Thus the MLE estimator  $\theta(n)$  solves

$$\mathbb{E}_\theta[T_n] = T_n.$$

If  $T(X_1)$  is a square integrable random variable then the strong LLN and the CLT ensure respectively that  $T_n$  converges  $\mathbb{P}_\theta$  a.s. to  $\mathbb{E}_\theta[T(X_1)]$  and, under  $\mathbb{P}_\theta$ ,

$$G_n := \frac{T_n - \mathbb{E}_\theta[T_n]}{\sqrt{\text{Var}[T_n]}} = \sqrt{n} \frac{T_n - D_w A(\theta)}{\sqrt{D_w^2 A(\theta)}} \quad (24)$$

converges in law to a standard Gaussian random variable, say  $G$ .

Moreover, if the model is regular, then

$$\text{Var}[T_n] = \frac{1}{n} \frac{A''(\theta)w'(\theta) - A'(\theta)w''(\theta)}{(w'(\theta))^3} = \frac{1}{n} D_w^2 A(\theta).$$

The Fisher information of the model satisfies  $\mathcal{F}(\theta) = D_w^2 A(\theta)(w'(\theta))^2$ .

Taking  $\mathbf{S} = \mathbb{N}$ ,  $\varphi(s) = \frac{1}{\sqrt{s}}$ , and  $\boldsymbol{\beta} = \boldsymbol{\beta}^\curvearrowright$ ,

$$\begin{aligned} \{F\}_0(n) &= \sqrt{n} (w'(\theta)T_n - A'(\theta)) = \sqrt{\mathcal{F}(\theta)} G_n \xrightarrow[n \rightarrow \infty]{\text{dist.}} \sqrt{\mathcal{F}(\theta)} G =: \mu_0, \\ \{F\}_k(n) &= w^{(k+1)}(\theta)T_n - A^{(k+1)}(\theta) \xrightarrow[n \rightarrow \infty]{\mathbb{P}_\theta} w^{(k+1)}(\theta)D_w A(\theta) - A^{(k+1)}(\theta) =: \mu_k. \end{aligned}$$

By Proposition 4, we get the coefficients  $\delta_k(n) := -\{F\}_k(n)/k!\{F\}_1(n)$ ,  $k \geq 0$  of  $\boldsymbol{\delta}(n)$  as well as the coefficients  $\delta_k := -\mu_k/k!\mu_1$  of its limit  $\boldsymbol{\delta}$ . Theorem 3 (with  $\theta_0 = \theta$ ) provides the two approximations of the MLE estimator  $\theta(n)$  of  $\theta$

$$\theta_p(n) := \theta + \sum_{k=1}^p \alpha_k(n)n^{-k/2} \quad \text{and} \quad \theta_p(n, \infty) := \theta + \sum_{k=1}^p \alpha_k n^{-k/2}$$

with  $\alpha_k(n)$  and  $\alpha_k$  the  $(\boldsymbol{\beta}^\curvearrowright, p)$ -related sequence to  $\boldsymbol{\delta}(n)$  and  $\boldsymbol{\delta}$  respectively. Since  $\mu_1 = -\mathcal{F}(\theta)$ ,

$$\alpha_1(n) = -\frac{\{F\}_0(n)}{\{F\}_1(n)} \xrightarrow[n \rightarrow \infty]{\text{dist.}} -\frac{\mu_0}{\mu_1} = \frac{G}{\sqrt{\mathcal{F}(\theta)}}.$$

Let us explicit the expansions with  $p = 3$ :

$$\theta_3(s) = \theta + \frac{\alpha_1(n)}{\sqrt{n}} + \delta_2(n) \left( \frac{\alpha_1(n)}{\sqrt{n}} \right)^2 + (2\delta_2(n)^2 + \delta_3(n)) \left( \frac{\alpha_1(n)}{\sqrt{n}} \right)^3, \quad (25)$$

and

$$\theta_3(n, \infty) = \theta - \frac{G}{\sqrt{n\mathcal{F}(\theta)}} - \frac{\mu_2}{2\mu_1} \left( \frac{G}{\sqrt{n\mathcal{F}(\theta)}} \right)^2 + \left( \frac{1}{2} \frac{\mu_2^2}{\mu_1^2} - \frac{\mu_3}{6\mu_1} \right) \left( \frac{G}{\sqrt{n\mathcal{F}(\theta)}} \right)^3.$$

Let us provide some examples of exponential families of laws to which the above procedure applies.

**Example 6** (Exponential distribution). We consider the Exponential distribution. The parameters is  $\theta > 0$ . The related functions are

$$T(x) = x, \quad w(\theta) = -\theta, \quad A(\theta) = -\log(\theta),$$

so that  $D_w A(\theta) = \frac{1}{\theta}$ ,  $D_w^2 A(\theta) = \mathcal{F}(\theta) = \frac{1}{\theta^2}$ ,  $\{F\}_k = \frac{(-1)^k k!}{\theta^{k+1}}$ ,  $k \geq 1$ .

In this case the MLE is  $\theta(n) = 1/T_n$ . The two approximations  $\theta_p(n)$  and  $\theta_p(n, \infty)$  in Theorem 3 are (using computations similar to the ones of Example 3)

$$\theta_p(n) = \theta \sum_{k=0}^p \frac{(-G_n)^k}{n^{k/2}} \quad \text{and} \quad \theta_p(n, \infty) = \theta \sum_{k=0}^p \frac{(-G)^k}{n^{k/2}}$$

with  $G_n = \sqrt{n}(\theta T_n - 1)$  (See (24)) and  $G \sim \mathcal{N}(0, 1)$ . With  $p = +\infty$  since all the involved functions are analytic<sup>3</sup>,

$$\theta_\infty(n) = \frac{\theta}{1 + \frac{G_n}{\sqrt{n}}} = \mathcal{H} \left( \frac{G_n}{\sqrt{n}} \right) = \theta + H \left( \frac{G_n}{\sqrt{n}} \right)$$

with

$$\mathcal{H}(x) := \frac{\theta}{1+x} \quad \text{and} \quad H(x) = \frac{-\theta x}{1+x}.$$

Remark that  $G_n > -\sqrt{n}$  almost surely, while  $\theta_\infty(n, \infty) = \mathcal{H}(G/\sqrt{n})$  on the event  $\{G > -\sqrt{n}\}$ . We then get the approximation of the distribution function of

$$\varepsilon(n) := \sqrt{n}(\theta_\infty(n) - \theta) = \sqrt{n} \left( \frac{1}{T_n} - \theta \right) = \sqrt{n} H \left( \frac{G_n}{\sqrt{n}} \right) \approx \sqrt{n} H \left( \frac{G}{\sqrt{n}} \right)$$

---

<sup>3</sup>The result may be found directly by using the relation between  $\theta(n)$ ,  $T_n$  and  $G_n$ .

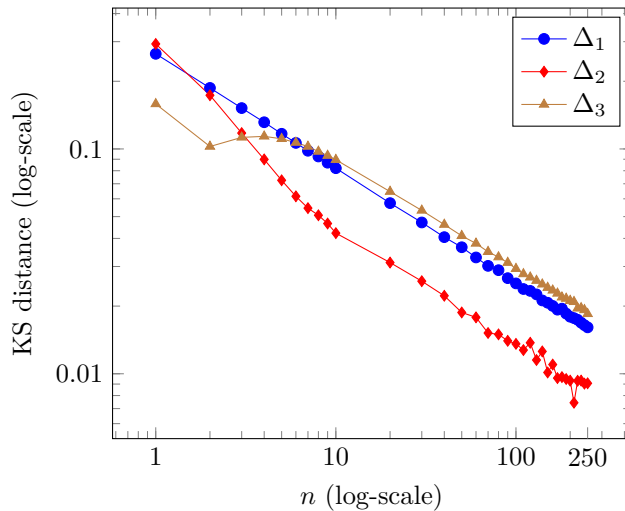


Figure 3: Kolmogorov-Smirnov distances defined in (26) for  $\theta = 1$ .

as

$$\mathbb{P}[\varepsilon(n) \leq x] \approx \mathbb{P}[\sqrt{n}(\theta_\infty(n, \infty) - \theta) \leq x] = \mathbb{P}\left[G \leq \frac{-x}{\theta + \frac{x}{\sqrt{n}}}\right],$$

using the Gaussian approximation of  $G_n$  by  $G$ .

With this example, we illustrate Fact (F5) in Section 4. We write  $\alpha_1 = \theta G$  with  $G \sim \mathcal{N}(0, 1)$ , as the Fisher information  $\mathcal{J}(\theta)$  is  $1/\theta^2$ . In Figure 3, we represent for several values of the sample size  $n$  the following Kolmogorov-Smirnov distances<sup>4</sup>,

$$\begin{aligned} \Delta_1 &:= \Delta(\varepsilon(n), \theta G), \quad \Delta_2 := \Delta(\varepsilon(n), \sqrt{n}H(G/\sqrt{n})) \\ \text{and } \Delta_3 &:= \Delta(\sqrt{n}H(G/\sqrt{n}), \theta G) \end{aligned} \quad (26)$$

with  $\Delta$  defined in (20). These distances are computed using  $10^6$  samples of  $T_n$ . The Kolmogorov-Smirnov statistics allows one to asserts that the Monte Carlo error, of order  $10^{-3}$ , is small in front of the  $\Delta_i$ . Since  $H$  is invertible,  $\Delta_2 = \Delta(G_n, G)$ , up to neglecting  $\{G \leq -\sqrt{n}\}$ . We then see that  $\Delta(\varepsilon(n), \sqrt{n}H(G/\sqrt{n}))$  is around half  $\Delta(\varepsilon(n), \theta G)$ , meaning that when  $n$  is small, it is far more satisfactory to use  $\sqrt{n}H(G/\sqrt{n})$  than  $\theta G$  to construct for example confidence intervals or tests for the value  $\theta$ .

**Example 7** (Binomial distribution). Let  $\theta \in (0, 1)$  be the parameter of a binomial model  $(N, \theta)$  with  $N \in \mathbb{N}$  fixed ( $N = 1$  corresponds to the Bernoulli model). Only

<sup>4</sup>We actually neglect the events  $\{G \leq -\sqrt{n}\}$  which is exponentially small.

the following functions differ from the previous case: The related functions are

$$T(x) = x, \quad w(\theta) = \log \frac{\theta}{1-\theta}, \quad A(\theta) = -N \log(1-\theta)$$

$$\text{so that } D_w A(\theta) = N\theta, \quad D_w^2 A(\theta) = N\theta(1-\theta), \quad \mathcal{F}(\theta) = \frac{N}{\theta(1-\theta)}.$$

In this case the MLE is the empirical mean divided by  $N$ . The derivative of the score and their a.s. limit (LGN): for  $m \geq 1$ ,

$$\begin{aligned} \{F\}_m(n) &= n^{1-\beta_{m/2}} m! \left( (-1)^m \frac{T_n}{\theta^{m+1}} - \frac{(N-T_n)}{(1-\theta)^{m+1}} \right), \\ \{F\}_m(\infty) &:= \mu_m = -m! N \left( \frac{(-1)^{m+1}}{\theta^m} + \frac{1}{(1-\theta)^m} \right). \end{aligned}$$

Besides, the CLT ensures that

$$\{F\}_0(n) = \sqrt{n} \frac{T_n - N\theta}{\theta(1-\theta)} = G_n \sqrt{\mathcal{F}(\theta)} \xrightarrow[n \rightarrow \infty]{\text{dist.}} \{F\}_0(\infty) := \sqrt{\mathcal{F}(\theta)} G.$$

And  $\{F\}_1(\infty) = -\mathcal{F}(\theta)$ .

In particular the two approximations  $\theta_p(n)$  and  $\theta_p(n, \infty)$  (for  $p = 2$ ) in Theorem 3 are

$$\theta_2(n) = \theta + \frac{G_n}{\sqrt{n}} \frac{\sqrt{N\theta^3(1-\theta)^3}}{T_n(1-2\theta) + N\theta^2} + \frac{G_n^2}{n} N\theta^2(1-\theta)^2 \frac{(\theta^3 + (1-\theta)^3)T_n - \theta^3 N}{((1-2\theta)T_n + N\theta^2)^3}.$$

and

$$\theta_2(n, \infty) = \theta + \frac{G\sqrt{\theta(1-\theta)}}{\sqrt{n \cdot N}} + \frac{G^2}{n} \frac{1-2\theta}{N}.$$

To illustrate a *boundary layer* effect, let us consider  $\theta_3(n, \infty)$  in the previous example:

$$\theta_3(n, \infty) = \theta + \frac{G\sqrt{\theta(1-\theta)}}{\sqrt{n \cdot N}} + \frac{G^2}{n} \frac{1-2\theta}{N} + \frac{G^3}{(n \cdot N)^{3/2}} \frac{1-5\theta+5\theta^2}{\sqrt{\theta(1-\theta)}}.$$

If  $\theta$  is close to 0 or to 1 then the coefficient of  $\frac{1}{\sqrt{n}}$  goes to 0 and the other ones give their maximal contribution. In particular the coefficient of  $\frac{1}{n^{3/2}}$  goes to infinity as  $\theta$  goes to 0 or to 1. Actually the expansion we used is not adapted when  $\theta$  is close to 0 or 1. This is a *boundary layer*. More precisely, assume for simplicity  $N = 1$ ,  $n = 10^m$ , and  $\theta = 10^{-t}$  (or  $\theta = 1 - 10^{-t}$ ) for  $m, t \in \mathbb{N}$ :  $\theta_3(n, \infty) - \theta = 10^{-(m+t)/2} G + 10^{-m} G^2 + 10^{-(3m-t)/2} G^3$ .

If  $t = m$  all terms provide a contribution of the same order  $10^{-m}$ .

If  $t > m$ , say  $t = 2m$  then  $\theta(n) - \theta \approx 10^{-3m/2} G + 10^{-m} G^2 + 10^{-m/2} G^3$  and the dominant term is the last one.

*Remark 7.* Note that if  $\theta$  is close to  $1/2$  then the coefficient of  $\frac{1}{\sqrt{n}}$  in  $\theta_3(n)$  gives its maximal contribution, the one of  $\frac{1}{n}$  is close to 0, and the coefficient of  $\frac{1}{n^{3/2}}$  gives its maximal contribution (i.e.  $\frac{-1^n}{2\sqrt{N}}G^3$ ). Moreover the latter term becomes 0 for  $\theta = \frac{5 \pm \sqrt{5}}{10} \in (0, 1)$ . The case  $\theta = 1/2$  allows for a different choice of  $\beta$ :  $\beta = \beta^\sim$ . This is shown in Section 6.3.

## 6.2 Using the natural parameters for the exponential family

We saw in Sections 5.1 and 5.2 two ways to consider another parametrization for the space of parameters. Such parametrizations could lead to simpler expressions. In the case of the exponential family, let us discuss the use of the “natural scale”. For this, we use the change of variable  $\Phi := w$  (with inverse  $\Psi$ ), where the notations of (23) are in use.

We consider that  $A$  and  $w$  are of class  $\mathcal{C}^{p+1}$  (or are analytic, in which case  $p = +\infty$ ). The likelihood and  $\check{F}$  are

$$\mathcal{L}_n(\theta) = n(w(\theta)T_n - A(\theta)) \text{ and } \check{F}(n, \theta) = n(T_n - D_w A(\theta)).$$

The MLE  $\theta(n)$  solves  $\theta(n) = D_w A^{-1}(T_n)$  that can be rewritten as

$$\theta(n) = \mathcal{H}_{\theta_0} \left( \frac{G_n}{\sqrt{n}} \right) \text{ with } \mathcal{H}_{\theta_0}(x) := D_w A^{-1} \left( D_w A(\theta_0) + x \sqrt{D_w^2 A(\theta_0)} \right)$$

with  $G_n$  defined by (24). Proposition 2 shows that

$$\theta_p(n) = \mathcal{I}_p \mathcal{H}_{\theta_0}(G_n/\sqrt{n})$$

that is

$$\theta_p(n) = \theta_0 + \sum_{k=1}^p \check{\alpha}_k(n) n^{-k/2} = \theta_0 + \sum_{k=1}^p \omega_k \left( \frac{G_n}{\sqrt{n}} \right)^k \quad (27)$$

with

$$\check{\alpha}_k(n) := G_n^k \omega_k \text{ and } \omega_k := \frac{D^k(D_w A^{-1})(D_w A(\theta_0))}{k!} (D_w^2 A(\theta_0))^{k/2}.$$

In particular, the  $\omega_k$ 's are deterministic and do not depend on  $n$ . They are the Taylor coefficients of  $\mathcal{H}_{\theta_0}$ .

**Example 8** (Binomial distribution). For the binomial distribution (see Example 7), this expression simplifies for all  $p \in \{1, \dots, \infty\}$  as

$$\theta_p(n) = \theta + \frac{\sqrt{\theta(1-\theta)}}{\sqrt{N \cdot n}} G_n = T_n$$



which is the empirical mean. Unlike for the exponential distribution, the distance between  $\sqrt{n}(\theta_p(n) - \theta)$  and the normal distribution  $\mathcal{N}(0, \theta(1 - \theta)/N)$  comes solely from the distance between  $G_n$  and the distribution  $\mathcal{N}(0, 1)$ .

Note that in Example 6 the change of variable does not provide new approximations for the MLE because the model is basically already in natural scale.

### 6.3 A phase transition of the approximation

The possibility to change scale (*i.e.*,  $\beta = \beta^\sim$ ) in Theorem 3 sometimes leads to a phase transition in the approximation of the MLE estimator. This happens when the distribution of the score is symmetric with respect to a parameter  $\theta_0 \in \Theta$ .

Let us consider the case of the binomial model in Example 7. If the true parameter  $\theta = \frac{1}{2}$  then for all  $m \in \mathbb{N}$  even, when  $\beta = \beta^\sim$ ,

$$\{F\}_m(n) = \frac{m!}{\sqrt{n}\theta^{2m}} \{F\}_0(n) \xrightarrow[n \rightarrow \infty]{\mathbb{P}_\theta} 0$$

If  $\beta = \beta^\sim$  then the latter coefficients for  $m$  even are

$$\{F\}_m(n) = \frac{m!}{\theta^m} \{F\}_0(n) \xrightarrow[n \rightarrow \infty]{\text{dist.}} \mu_m := \frac{m!}{\theta^m} \mu_0 = \frac{m!}{\theta^m} \sqrt{\mathcal{J}(\theta)} G$$

and for  $m$  odd,  $\{F\}_m(n) = -\frac{m!N}{\theta^{m+1}} = \frac{m!}{\theta^{m-1}} \{F\}_1(n)$ .

Applying Theorem 3 with  $\beta = \beta^\sim$  and  $\theta_0 = \theta = \frac{1}{2}$  provides new approximations of the MLE  $\theta(n)$  (see Lemma 5). The expansion for  $n \in \mathbb{N}$ ,  $p \in \{1, \dots, \infty\}$ , coincides with the one obtained in the previous section:

$$\theta_p(n) = \theta + \frac{\alpha_1(n)}{\sqrt{n}}.$$

Indeed all coefficients  $\alpha_k(n) = 0$  for  $k \geq 3$ . By Lemma 5 it suffices to prove that  $\alpha_{2k+1} = 0$  for  $k \geq 1$  and this can be shown by induction using the fact that  $\{F\}_{2k}(n) = \frac{(2k)!}{\theta^{2k}} \{F\}_0(n)$  and  $\{F\}_{2k+1}(n) = \frac{(2k+1)!}{\theta^{2k}} \{F\}_1(n)$ .

In the case of the MLE estimator of the skewness parameter of Skew Brownian motion, such a phase transition has been noticed in [32] and made explicit in [31].

### 6.4 An application to statistics of diffusions

Let us consider the inference of the parameter  $\theta$  in the Stochastic Differential Equations (SDE)

$$dX_t^\theta = \theta \cdot b(X_t^\theta) dt + \sigma dB_t, \quad t \geq 0.$$

for a Brownian motion  $B$ , a function  $b$  regular enough, and  $\sigma \in (0, +\infty)$ .

For this, we follow the methodology of [11], which is based on the martingale approximations of the log-likelihood.

We fix a time step  $\Delta > 0$ . We assume we observe a realization  $(X_0, \dots, X_n)$  of the vector  $(X_0^\theta, X_1^\theta, \dots, X_n^\theta)$  where  $X_i^\theta$  is a shorthand for  $X_{i\Delta}^\theta$ ,  $i = 0, \dots, n$ .

The unknown parameter  $\theta$  is estimated by  $\theta(n)$  that solves  $F_n(\theta(n)) = 0$  with

$$F_n(\theta) := \sum_{i=1}^n \frac{b(X_{i-1})}{\sigma^2} (X_i - M(X_{i-1}, \theta)) \quad (28)$$

where

$$M(x, \theta) := \mathbb{E}[X_\Delta^\theta \mid X_0^\theta = x].$$

The asymptotic of  $\theta(n)$  is studied when  $n \rightarrow \infty$  while  $\Delta$  is kept fixed (long time and fixed time step).

By a straightforward computation,

$$D_\theta^k F_n(\theta) := - \sum_{i=1}^n \frac{b(X_{i-1})}{\sigma^2} D_\theta^k M(X_{i-1}, \theta).$$

*Hypothesis 5.* The process  $X^\theta$  is ergodic. Its invariant measure is its renormalized speed measure

$$\mu_\theta(dx) = \frac{1}{A(\theta)} \exp\left(2\theta \int_0^x \frac{b(y)}{\sigma^2} dy\right) dx,$$

where  $A(\theta)$  is a normalizing constant such that  $\int_{\mathbb{R}} \mu_\theta(dx) = 1$ .

We denote by  $(t, x, y) \mapsto p_\theta(t, x, y)$  the density transition function of  $X^\theta$ . For  $f : \mathbb{R}^2 \rightarrow \mathbb{R}$ , we write

$$Q^\theta(f) := \int_{\mathbb{R}} \int_{\mathbb{R}} f(x, y) p_\theta(\Delta, x, y) \mu_\theta(dx).$$

The following proposition is a consequence of ergodicity.

**Proposition 6** (See *e.g.*, [11, Lemma 3.1]). *Let  $f : \mathbb{R}^2 \rightarrow \mathbb{R}$  be a measurable function such that  $Q^\theta(f^2) < +\infty$ . Then*

$$\frac{1}{n} \sum_{i=1}^n f(X_{i-1}^\theta, X_i^\theta) \xrightarrow[n \rightarrow \infty]{prob.} Q^\theta(f)$$

in  $L^2(\mathbb{P}_\theta)$ . If  $Q^\theta(f) = 0$ , then

$$\frac{1}{\sqrt{n}} \sum_{i=1}^n f(X_{i-1}^\theta, X_i^\theta) \xrightarrow[n \rightarrow \infty]{\text{dist.}} G \text{ with } G \sim \mathcal{N}(0, Q^\theta(f^2))$$

under  $\mathbb{P}_\theta$ .

Hence, for  $k \geq 1$ , owing to the definition of the invariant measure, under  $\mathbb{P}_\theta$ , replacing the observations  $\mathbf{X}_i$  by the corresponding random variable  $X_i^\theta$ ,

$$\frac{-D_\theta^k F_n(\theta)}{n} = \frac{1}{n} \sum_{i=1}^n \frac{b(X_{i-1}^\theta)}{\sigma^2} D_\theta^k M(X_{i-1}^\theta, \theta) \xrightarrow[n \rightarrow \infty]{\text{prob.}} \int_{\mathbb{R}} \frac{b(x)}{\sigma^2} D_\theta^k M(x, \theta) \mu_\theta(dx),$$

while, if  $\int_{\mathbb{R}} \frac{b(x)}{\sigma^2} D_\theta^k M(x, \theta) \mu_\theta(dx) = 0$ ,

$$\frac{1}{\sqrt{n}} D_\theta^k F_n(\theta) = \frac{1}{\sqrt{n}} \sum_{i=1}^n \frac{b(X_{i-1}^\theta)}{\sigma^2} D_\theta^k M(X_{i-1}^\theta, \theta) \xrightarrow[n \rightarrow \infty]{\text{dist.}} \mathcal{N}(0, c^2)$$

with

$$c^2 = \int_{\mathbb{R}} \int_{\mathbb{R}} p(\Delta, x, y) \frac{b(x)^2}{\sigma^4} (y - M(x, \theta))^2 \mu_\theta(dx) dy.$$

We note that the quantity  $M(x, \theta)$  cannot be always computed in an explicit manner, because

$$M(x, \theta) = \mathbb{E}_\theta[X_\Delta^\theta | X_0^\theta = x] = x + \theta \int_0^\Delta \mathbb{E}_\theta[b(X_s^\theta) | X_0^\theta = x] ds. \quad (29)$$

**Example 9** (Ornstein-Uhlenbeck process). Let us consider that  $b(x) = -x$  and  $\theta > 0$  so that  $X^\theta$  is the Ornstein-Uhlenbeck process. From (29),

$$M(x, \theta) = x \exp(-\Delta\theta) \text{ and } \mu_\theta(dx) = \frac{\sqrt{\theta}}{\sqrt{\pi\sigma^2}} \exp\left(-\frac{\theta x^2}{\sigma^2}\right) dx$$

which is a centered Gaussian law of variance  $\sigma^2/(2\theta)$ . Therefore,  $D^k M(x, \theta) = x(-\Delta)^k \exp(-\Delta\theta)$ . We are then in the case ‘‘up-flat’’. We consider the change of variable  $\eta := \Phi(\theta) := e^{-\Delta\theta}$  (see Section 5.1) and we get

$$F_n^\dagger(\eta) = \sum_{i=1}^n \frac{-X_{i-1}}{\sigma^2} (X_i - X_{i-1}\eta).$$

It holds that  $D_\eta F_n^\dagger = \sum_{i=1}^n \frac{1}{\sigma^2} (X_{i-1})^2$  and  $D_\eta^k F_n^\dagger = 0$ . By Theorem 3,

$$\eta(n) = \eta - \frac{F_n^\dagger(\eta)}{D_\eta F_n^\dagger(\eta)} = \frac{\sum_{i=1}^n X_{i-1} X_i}{\sum_{i=1}^n (X_{i-1})^2}$$

so that

$$\theta(n) = \frac{-1}{\Delta} \log \frac{\sum_{i=1}^n \mathbf{X}_{i-1} \mathbf{X}_i}{\sum_{i=1}^n (\mathbf{X}_{i-1})^2}.$$

We have the same formula as the one in [11, Example 2.1, p. 21], which is obtained by direct means (note that we use  $b(x) = -x$  instead of  $b(x) = x$  as in [11]).

**Example 10** (Using an approximation of  $M(x, \theta)$ ). When no close form solution to  $M(x, \theta)$  (see (29)), we consider that we have a smooth approximation  $M_\delta(x, \theta)$  of  $M(x, \theta)$  depending on a parameter  $\delta$ . We set

$$\epsilon_\delta(\theta) := \sum_{i=1}^n \frac{b(\mathbf{X}_{i-1})}{\sigma^2} (M_\delta(\mathbf{X}_{i-1}, \theta) - M(\mathbf{X}_{i-1}, \theta))$$

and define  $F_{\delta,n}$  similarly to  $F_n(\theta)$  in (28) with  $M_\delta$  instead of  $M$ . The solution  $\theta_{\delta,n}$  to  $F_{\delta,n}(\theta_{\delta,n}) = 0$  also solves

$$F_n(\theta_{\delta,n}) = \epsilon_\delta(\theta_{\delta,n}). \quad (30)$$

Assume that we are in the case “up-flat” and that  $\epsilon_\delta(\theta_{\delta,n})/\sqrt{n}$  converges to 0 as  $(\delta, n)$  converges to  $(0, +\infty)$  as a net (this could imply some relationship between the rate of convergence of  $\delta$  to 0 and the one of  $n$  to  $+\infty$ ). Proposition 7 allows one to conclude that  $\theta_{\delta,n}$  and  $\theta(n)$  have the same asymptotic behavior.

## 7 Conclusion

We have given an expansion of the solution  $\theta(n)$  to a problem  $F_n(\theta(n)) = 0$  in a way which involves the asymptotic behavior of  $F_n$  and its derivatives. This covers (quasi-, pseudo-) maximum likelihood estimation and Generalized Method of Moments. We could also consider errors in the observations. Remark that  $n$  is a quantity related to the asymptotic behavior (*e.g.*, the sample size). Our hypotheses are rather weak so that we may consider independent samples as well as dependent samples such as the ones arising in statistics of diffusion or chronological series. With the inference in view, we have an expansion of type  $\theta(n) = \theta_0 + G_n \varphi(n) + E_n(G_n \varphi(n)) + O(\varphi(n)^{p+1})$  where  $G_n$  converges in distribution to some random variable  $G$ ,  $\varphi(n)$  is a rate (say  $\varphi(n) = 1/\sqrt{n}$ ) and  $E_n$  is a non-linear, random function with  $E_n(x) = O(x^2)$ . In the cases considered, the convergence of  $G_n$  depends on the one of  $F_n$  and its first derivative. The function  $E_n$  has a Taylor expansion whose coefficients have an asymptotic behavior that depends on the derivatives of  $F_n$ .

Our expansion can be used to access semi-asymptotic behaviors of the estimator, as we have illustrated in a toy example in Section 6.1. The distribution of  $\varphi(n)(\theta(n) -$

$\theta_0$ ) is close to the one of  $G$ . Its departure from  $G$  is due to the distance between the distributions of  $G_n$  and the one of  $G$ , as well as the behavior of  $E_n$ . We illustrated this in Fact (F5) (Section 4) and in Example 6. So, Berry-Esséen type results seems accessible as well because (in the cases we specify) the extension we establish is a non-linear transformation of an asymptotically pivotal quantity (see for instance [36, 37]). Note that Berry-Esséen results on the convergence of the latter to its limit may be combined with our expansion (as in the case of i.i.d. exponential random variable). Similar approaches should also apply to Wasserstein distance (see *e.g.*, [4, 6]).

Let us mention that the main results of this document (in particular Theorem 1) can be easily extended also for multidimensional parameters thanks to multidimensional Inverse Function Theorem (see, *e.g.*, [1]). It should also hold true in an infinite dimensional setting, which should be useful to deal with semiparametric models although there are some fine technical points to deal with (see the discussion in [23, 24, 42]). In the multidimensional setting, a Taylor expansion is used in [26] to compute the first two moments of functions of estimators. This might be combined easily with our approach to obtain higher moments. We expect to compute cumulants [18] as well in a wide variety of cases, and to identify some corrections to enforce the quality of the estimation [42].

## A Asymptotic inversion

The goal of the section is proving Theorem 1. The proof of Theorem 2 is similar, yet simpler as there is no need to control the remainder.

### A.1 A Taylor formula with remainder

In this section, we basically consider the  $p$ -th order Taylor polynomial of  $z \mapsto F(s, P_p(z; \boldsymbol{\alpha}))$  around 0 and a convenient expression for the remainder.

*Hypothesis 6.* Let  $p \in \mathbb{N} \setminus \{0\}$ . Let  $\boldsymbol{\alpha} := \{\alpha_k\}_{k=0, \dots, p} \in \mathbb{R}^{p+1}$  with  $\alpha_0 = 0$ . With  $P_m$  defined by

$$P_m(z; \boldsymbol{\alpha}) := \sum_{k=0}^m \alpha_k z^k \text{ for } m = 0, 1, \dots, p \text{ and } z \in \mathbb{R}, \quad (31)$$

we assume that  $P_m(\mathbb{I}; \boldsymbol{\alpha}) \subseteq \Theta$  for all  $m \in \{1, \dots, p\}$ .

For the sake of simplicity, in this section, we consider a real-valued function  $F$  on  $\Theta$ , but we could consider a real-valued function on  $\mathbf{S} \times \Theta$ .

**Notation 9.** For  $m = 1, \dots, p$ ,  $z \in \mathbb{R}$ , and  $F \in \mathcal{C}^p(\Theta, \mathbb{R})$  we define

$$R_m(z; \boldsymbol{\alpha}) := \sum_{k_1 + \dots + k_m \leq p} \alpha_{k_1} \cdots \alpha_{k_m} z^{k_1 + \dots + k_m} \\ \times \int_0^1 \int_0^{\tau_1} \cdots \int_0^{\tau_{m-1}} \left( D^m F(\tau_m P_{p-(k_1 + \dots + k_{m-1})}(z; \boldsymbol{\alpha})) \right. \\ \left. - D^m F(\tau_m P_{p-(k_1 + \dots + k_m)}(z; \boldsymbol{\alpha})) \right) d\tau_m \cdots d\tau_1. \quad (32)$$

**Notation 10.** We define for  $z \in \mathbb{I}$ ,

$$P_p^*(z; \boldsymbol{\alpha}) := \sum_{k=1}^p |\alpha_k| \cdot |z|^k, \\ J(z; \boldsymbol{\alpha}) := [-P_p^*(z; \boldsymbol{\alpha}), P_p^*(z; \boldsymbol{\alpha})] \cap \Theta, \\ \text{and } |F|_m(z; \boldsymbol{\alpha}) := \sup_{w \in J(z; \boldsymbol{\alpha})} |D^m F(w)|. \quad (33)$$

**Lemma 9.** Assume Hypothesis 6. Let  $F \in \mathcal{C}^{p+1}(\Theta, \mathbb{R})$ . Then for all  $z \in \mathbb{I}$ ,

$$F(P_p(z; \boldsymbol{\alpha})) - F(0) \\ = \sum_{m=1}^p \frac{1}{m!} D^m F(0) \sum_{k_1 + \dots + k_m \leq p} \alpha_{k_1} \cdots \alpha_{k_m} z^{k_1 + \dots + k_m} + \sum_{m=1}^p R_m(z; \boldsymbol{\alpha}). \quad (34)$$

Besides, for  $m = 1, \dots, p$ ,

$$|R_m(z; \boldsymbol{\alpha})| \leq |F|_{m+1}(z; \boldsymbol{\alpha}) P_p^*(1; \boldsymbol{\alpha})^{m+1} |z|^{p+1} \text{ for } |z| \leq 1. \quad (35)$$

*Proof.* For simplicity we do not specify the dependence on the sequence  $\boldsymbol{\alpha}$ . By the fundamental theorem of calculus,

$$F(P_p(z)) - F(0) = \int_0^1 DF(\tau P_p(z)) \cdot P_p(z) d\tau \\ = DF(0) P_p(z) + \int_0^1 (DF(\tau_1 P_p(z)) - DF(0)) \cdot P_p(z) d\tau_1. \quad (36)$$

Adding and removing another suitable term yields

$$F(P_p(z)) - F(0) \\ = DF(0) P_p(z) + \sum_{k_1=1}^p \alpha_{k_1} z^{k_1} \int_0^1 \left( DF(\tau_1 P_{p-k_1}(z)) - DF(0) \right) d\tau_1 \\ + \sum_{k_1=1}^p \alpha_{k_1}(s) z^{k_1} \int_0^1 \left( DF(\tau_1 P_p(z)) - DF(\tau_1 P_{p-k_1}(z)) \right) d\tau_1. \quad (37)$$

Note that  $P_{p-k_1}(z) \in J(z; \boldsymbol{\alpha})$  for any  $k_1 \in \{1, \dots, p\}$  with  $J(z; \boldsymbol{\alpha})$  defined in Notation 10. The Lipschitz constant of  $DF$  on  $J(z; \boldsymbol{\alpha})$  is the the sup-norm of  $D^2F$  on  $J(z; \boldsymbol{\alpha})$ , that is  $|F|_2(z)$  in Notation 10. Then

$$\begin{aligned} & \sum_{k_1=1}^p |\alpha_{k_1}| \cdot |z|^{k_1} \cdot |DF(\tau_1 P_p(z)) - DF(\tau_1 P_{p-k_1}(z))| \\ & \leq |F|_2(z) \sum_{k_1=1}^p \sum_{k_2=p-k_1+1}^p |\alpha_{k_1} \alpha_{k_2}| \cdot |z|^{k_1+k_2} \leq |F|_2(z) \sum_{\substack{1 \leq k_1, k_2 \leq p \\ k_1+k_2 \geq p+1}} |\alpha_{k_1} \alpha_{k_2}| \cdot |z|^{k_1+k_2}. \end{aligned}$$

This justifies our definition of  $R_1(z)$  in Notation 9 as  $|R_1(z)| = O(|z|^{p+1})$  for  $z$  close to 0. Besides,

$$\sum_{\substack{1 \leq k_1, k_2 \leq p \\ k_1+k_2 \geq p+1}} |\alpha_{k_1} \alpha_{k_2}| \leq \left( \sum_{k=1}^p |\alpha_k| \right)^2 \leq P_p^*(1)^2.$$

This proves (35) for  $R_1$ .

We could then iterate the above decomposition in (37) by replacing  $F$  by  $DF$  for the second term in the right-hand side of the above integral. We obtain

$$\begin{aligned} & DF(\tau_1 P_{p-k_1}(z)) - DF(0) \\ & = \tau_1 D^2F(0) P_{p-k_1}(z) + \sum_{k_2 \leq p-k_1} \tau_1 \alpha_{k_2} z^{k_2} \int_0^1 \left( D^2F(\tau_2 \tau_1 P_{p-k_1-k_2}(z)) - D^2F(0) \right) d\tau_2 \\ & \quad + \sum_{k_2 \leq p-k_1} \tau_1 \alpha_{k_2} z^{k_2} \int_0^1 \left( D^2F(\tau_1 \tau_2 P_{p-k_1}(z)) - D^2F(\tau_2 \tau_1 P_{p-k_1-k_2}(z)) \right) d\tau_2. \quad (38) \end{aligned}$$

Using a change of variable  $\tau'_2 = \tau_1 \tau_2$  allows one to transform  $\tau_1 \int_0^1 f(\tau_1 \tau_2) d\tau_2$  to  $\int_0^{\tau_1} f(\tau_2) d\tau_2$  for any suitable function  $f$ .

Integrating (38) between 0 and 1 and using  $p$  times the development (37) to the successive derivatives of  $F$  leads to (34).

The remainder's terms  $R_k$  are controlled in a similar way as for  $R_1$ .  $\square$

## A.2 Application of the Taylor formula

In this section, we use Notations 2, 9-10.

**Proposition 7.** *Assume Hypothesis 1. Let  $F \in \mathcal{C}^{p+1}(\Theta, \mathbb{R})^{\mathbf{S}}$  and  $s \in \mathbf{S}$  such that  $\{F\}_1(s) \neq 0$ . Let  $\boldsymbol{\alpha}(s) \in \mathbb{R}^{p+1}$  such that (5) is satisfied. Assume that  $P_m(\varphi(s); \boldsymbol{\alpha}(s)) \in \Theta$  for all  $m \in \{1, \dots, p\}$ . Then*

$$\begin{aligned} |F(s, P_p(\varphi(s); \boldsymbol{\alpha}(s)))| &\leq \sum_{m \in \{2, \dots, p+1\}} P_p^*(1; \boldsymbol{\alpha}(s))^m |F|_m(\varphi(s); \boldsymbol{\alpha}(s)) |\varphi(s)|^{p+1} \\ &\quad + \sum_{\substack{m \in \{1, \dots, p\} \\ \beta_m + 1 \leq \beta_\star}} P_p^*(1; \boldsymbol{\alpha}(s))^m |\{F\}_m(s)| \cdot |\varphi(s)|^{\gamma_m - \beta_m}. \end{aligned}$$

*Proof.* In the proof we replace any dependence on  $(\varphi(s); \boldsymbol{\alpha}(s))$  by simply  $(s')$ . Using Notation 2, we rewrite (34) in Lemma 9 as

$$\begin{aligned} F(s, P_p(s')) - \{F\}_0(s) \varphi(s)^{-\beta_0} \\ = \sum_{m=1}^p \frac{\{F\}_m(s)}{m!} \sum_{k_1 + \dots + k_m \leq p} \alpha_{k_1}(s) \cdots \alpha_{k_m}(s) \varphi(s)^{k_1 + \dots + k_m - \beta_m} + \sum_{m=1}^p R_m(s'). \end{aligned}$$

For  $m = 1, \dots, p$ , the control (35) implies that

$$|R_m(s')| \leq |F|_{m+1}(s') P_p^*(1; \boldsymbol{\alpha}(s))^{m+1} |\varphi(s)|^{p+1} \text{ since } |\varphi(s)| \leq 1$$

for all  $m \in \{1, \dots, p\}$ . Using the definition of  $\gamma_m$  in (4), we split

$$\begin{aligned} \{(k_1, \dots, k_m) \mid 1 \leq k_1 + \dots + k_m \leq p\} \\ = \{(k_1, \dots, k_m) \mid 1 \leq k_1 + \dots + k_m < \gamma_m\} \\ \cup \{(k_1, \dots, k_m) \mid \gamma_m \leq k_1 + \dots + k_m \leq p\}. \end{aligned}$$

Therefore, the convention that the sum is null if the index set is empty, yields for  $m \in \{1, \dots, p\}$  that if  $\beta_m + 1 \leq \beta_\star$  then

$$\begin{aligned} \left| \sum_{\gamma_m \leq k_1 + \dots + k_m \leq p} \alpha_{k_1}(s) \cdots \alpha_{k_m}(s) |\varphi(s)|^{k_1 + \dots + k_m} \right| \\ \leq \sum_{\gamma_m \leq k_1 + \dots + k_m \leq p} |\alpha_{k_1}(s) \cdots \alpha_{k_m}(s)| \cdot |\varphi(s)|^{\gamma_m} \leq P_p^*(1; \boldsymbol{\alpha}(s))^m |\varphi(s)|^{\gamma_m}. \end{aligned}$$

Using (5) we obtain the result.  $\square$

### A.3 Proof of Theorem 1 on the approximation of roots

*Proof of Theorem 1.* Since  $DF(s, \cdot)$  is continuous on  $U(s)$ , The fact that  $|DF(s, \cdot)|$  has a positive lower bound on  $U(s)$  (Condition (iii)) implies that  $F(s, \cdot)$  is invertible



on  $U(s)$  by the Inverse Function Theorem [1]. Besides, for any  $\sigma, \sigma' \in F(s, U(s))$ ,

$$\begin{aligned} |DF^{-1}(s, \sigma) - DF^{-1}(s, \sigma')| &\leq \sup_{r \in [\sigma, \sigma']} \frac{|\sigma - \sigma'|}{|DF(s, F^{-1}(s, \tau))|} \\ &\leq \sup_{\theta \in U(s)} \frac{|\sigma - \sigma'|}{|DF(s, \theta)|} = \sup_{\theta \in U(s)} \frac{|\varphi(s)|^{\beta_1}}{|\varphi(s)|^{\beta_1} |DF(s, \theta)|} |\sigma - \sigma'| \leq \frac{|\varphi(s)|^{\beta_1}}{c(s)} |\sigma - \sigma'|. \end{aligned}$$

Condition (iv) implies that  $J(\varphi(s)) \subset U(s)$ , where  $J$  is defined in Notation 10.

With  $g(s) := F(s, P_p(\varphi(s); \boldsymbol{\alpha}(s)))$  and since  $F(s, \theta(s)) = 0$ ,

$$|P_p(\varphi(s); \boldsymbol{\alpha}(s)) - \theta(s)| = |F^{-1}(s, g(s)) - F^{-1}(s, 0)| \leq \frac{|\varphi(s)|^{\beta_1}}{c(s)} |g(s)|.$$

The result follows from Proposition 7. □

## References

- [1] R. Abraham, J. E. Marsden, and T. Ratiu. *Manifolds, tensor analysis, and applications*. Second. Vol. 75. Applied Mathematical Sciences. Springer-Verlag, New York, 1988, pp. x+654. DOI: 10.1007/978-1-4612-1029-0.
- [2] G. Allaire, A. Lamacz, and J. Rauch. *Crime Pays; Homogenized Wave Equations for Long Times*. 2018.
- [3] A. Anastasiou and C. Ley. “Bounds for the asymptotic normality of the maximum likelihood estimator using the Delta method”. In: *Latin American Journal of Probability and Mathematical Statistics* 14.1 (2017), p. 153. DOI: 10.30757/alea.v14-09.
- [4] A. Anastasiou and C. Ley. “Bounds for the asymptotic normality of the maximum likelihood estimator using the delta method”. In: *ALEA Lat. Am. J. Probab. Math. Stat.* 14.1 (2017), pp. 153–171.
- [5] A. Anastasiou and G. Reinert. “Bounds for the normal approximation of the maximum likelihood estimator”. In: *Bernoulli* 23.1 (Feb. 2017). DOI: 10.3150/15-bej741.
- [6] A. Anastasiou and G. Reinert. “Bounds for the normal approximation of the maximum likelihood estimator”. In: *Bernoulli* 23.1 (2017), pp. 191–218. DOI: 10.3150/15-BEJ741.
- [7] O. Barndorff-Nielsen. *Information and exponential families in statistical theory*. Wiley Series in Probability and Statistics. Reprint of the 1978 original [MR0489333]. John Wiley & Sons, Ltd., Chichester, 2014, pp. x+238. DOI: 10.1002/9781118857281.

- [8] O. Barndorff-Nielsen. “On a formula for the distribution of the maximum likelihood estimator”. In: *Biometrika* 70.2 (1983), pp. 343–365. DOI: 10.1093/biomet/70.2.343.
- [9] M. S. Bartlett. “Approximate confidence intervals”. In: *Biometrika* 40 (1953), pp. 12–19. DOI: 10.1093/biomet/40.1-2.12.
- [10] A. Bensoussan, J.-L. Lions, and G. Papanicolaou. *Asymptotic analysis for periodic structures*. Providence, RI: AMS Chelsea Publishing, 2011.
- [11] B. M. Bibby and M. Sørensen. “Martingale Estimation Functions for Discretely Observed Diffusion Processes”. In: *Bernoulli* 1.1/2 (Mar. 1995), p. 17. DOI: 10.2307/3318679.
- [12] J. P. N. Bishwal. “Berry-Esseen inequalities for the discretely observed Ornstein-Uhlenbeck-gamma process”. In: *Markov Process. Related Fields* 17.1 (2011), pp. 119–150.
- [13] R. P. Brent and H. T. Kung. “Fast Algorithms for Manipulating Formal Power Series”. In: *Journal of the ACM* 25.4 (Oct. 1978), pp. 581–595. DOI: 10.1145/322092.322099.
- [14] L. D. Brown. *Fundamentals of statistical exponential families with applications in statistical decision theory*. English. Vol. 9. IMS Lect. Notes, Monogr. Ser. Hayward, CA: Institute of Mathematical Statistics, 1986.
- [15] G. M. Cordeiro and F. Cribari-Neto. *An introduction to Bartlett correction and bias reduction*. SpringerBriefs in Statistics. Springer, Heidelberg, 2014. DOI: 10.1007/978-3-642-55255-7.
- [16] R. B. Dingle. *Asymptotic expansions: Their derivation and interpretation*. London-New York: Academic Press, a subsidiary of Harcourt Brace Jovanovich Publishers. XV. 1973.
- [17] R. J. M. M. Does, L. W. G. Strijbosch, and W. Albers. “Using Jackknife Methods for Estimating the Parameter in Dilution Series”. In: *Biometrics* 44.4 (Dec. 1988), p. 1093. DOI: 10.2307/2531738.
- [18] H. Döring, S. Jansen, and K. Schubert. “The method of cumulants for the normal approximation”. English. In: *Probab. Surv.* 19 (2022), pp. 185–270. DOI: 10.1214/22-PS7.
- [19] B. Efron. “Bootstrap confidence intervals for a class of parametric problems”. In: *Biometrika* 72.1 (1985), pp. 45–58. DOI: 10.1093/biomet/72.1.45.

- [20] A. Frabetti and D. Manchon. “Five interpretations of Faà di Bruno’s formula”. In: *Faà di Bruno Hopf algebras, Dyson-Schwinger equations, and Lie-Butcher series. Based on a conference hosted by IRMA, Strasbourg, France, June 27–July 1, 2011*. Zürich: European Mathematical Society (EMS), 2015, pp. 91–147.
- [21] H. Furstenberg. “Algebraic functions over finite fields”. In: *Journal of Algebra* 7.2 (Oct. 1967), pp. 271–277. DOI: 10.1016/0021-8693(67)90061-0.
- [22] I. M. Gessel. “Lagrange inversion”. In: *J. Combin. Theory Ser. A* 144 (2016), pp. 212–249. DOI: 10.1016/j.jcta.2016.06.018.
- [23] R. D. Gill. “Non- and semi-parametric maximum likelihood estimators and the von Mises method. I”. In: *Scand. J. Statist.* 16.2 (1989). With a discussion by J. A. Wellner and J. Præstgaard and a reply by the author, pp. 97–128.
- [24] R. D. Gill and A. W. van der Vaart. “Non- and semi-parametric maximum likelihood estimators and the von Mises method. II”. In: *Scand. J. Statist.* 20.4 (1993), pp. 271–288.
- [25] A. Hall. *Generalized Method of Moments*. Oxford University Press, 2005.
- [26] J. Hurt. “Asymptotic expansions of functions of statistics”. English. In: *Apl. Mat.* 21 (1976), pp. 444–456.
- [27] W. P. Johnson. “The curious history of Faà di Bruno’s formula”. In: *Am. Math. Mon.* 109.3 (2002), pp. 217–234. DOI: 10.2307/2695352.
- [28] S. G. Johnston and J. Prochno. “Faà di Bruno’s formula and inversion of power series”. In: *Advances in Mathematics* 395 (Feb. 2022), p. 108080. DOI: 10.1016/j.aim.2021.108080.
- [29] P. A. Lagerstrom. *Matched asymptotic expansions. Ideas and techniques*. Vol. 76. New York etc.: Springer-Verlag, 1988.
- [30] R. B. Leipnik and C. E. M. Pearce. “The multivariate Faà di Bruno formula and multivariate Taylor expansions with explicit integral remainder term”. In: *The ANZIAM Journal* 48.3 (Jan. 2007), pp. 327–341. DOI: 10.1017/S1446181100003527.
- [31] A. Lejay and S. Mazzonetto. *Estimation of the rate of convergence of the MLE estimator for the SBM*. In preparation. 2022.
- [32] A. Lejay, E. Mordecki, and S. Torres. “Is a Brownian motion skew?” English. In: *Scand. J. Stat.* 41.2 (2014), pp. 346–364. DOI: 10.1111/sjos.12033.
- [33] M. N. Mishra and B. L. S. Prakasa Rao. “On a Berry-Esseen type bound for the maximum likelihood estimator of a parameter for some stochastic partial differential equations”. In: *J. Appl. Math. Stoch. Anal.* 2 (2004), pp. 109–122. DOI: 10.1155/S1048953304309019.

- [34] W. K. Newey and R. J. Smith. “Higher Order Properties of GMM and Generalized Empirical Likelihood Estimators”. In: *Econometrica* 72.1 (2004), pp. 219–255. URL: <http://www.jstor.org/stable/3598854>.
- [35] Y. Pawitan. *In all likelihood: statistical modelling and inference using likelihood*. Oxford University Press, 2001.
- [36] I. Pinelis. “Optimal-order uniform and nonuniform bounds on the rate of convergence to normality for maximum likelihood estimators”. In: *Electron. J. Stat.* 11.1 (2017), pp. 1160–1179. DOI: 10.1214/17-EJS1264.
- [37] I. Pinelis and R. Molzon. “Optimal-order bounds on the rate of convergence to normality in the multivariate delta method”. In: *Electron. J. Stat.* 10.1 (2016), pp. 1001–1063. DOI: 10.1214/16-EJS1133.
- [38] I. M. Skovgaard. “Edgeworth expansions of the distributions of maximum likelihood estimators in the general (non-i.i.d.) case”. In: *Scand. J. Statist.* 8.4 (1981), pp. 227–236.
- [39] A. D. Sokal. “A ridiculously simple and explicit implicit function theorem”. In: *Sém. Lothar. Combin.* 61A (2009/11), Art. B61Ad, 21.
- [40] D. A. Sprott. “Application of maximum likelihood methods to finite samples”. In: *Sankhyā Ser. B* 37.3 (1975), pp. 259–270.
- [41] D. A. Sprott. “Normal likelihoods and their relation to large sample theory of estimation”. In: *Biometrika* 60 (1973), pp. 457–465. DOI: 10.1093/biomet/60.3.457.
- [42] A. W. van der Vaart. *Asymptotic statistics*. Vol. 3. Cambridge Series in Statistical and Probabilistic Mathematics. Cambridge University Press, Cambridge, 1998. DOI: 10.1017/CB09780511802256.
- [43] J. van der Hoeven. “Relax, but don’t be too lazy”. In: *J. Symb. Comput.* 34.6 (2002), pp. 479–542. DOI: 10.1006/jscs.2002.0562.
- [44] B. Wang, H. Jiang, and J. Yu. “Berry-Esseen bound for parameter estimation in some time inhomogeneous diffusions and applications”. In: *Statist. Probab. Lett.* 81.8 (2011), pp. 921–929. DOI: 10.1016/j.spl.2011.04.002.