

# **Dealing with missing data in model-based clustering through a MNAR model**

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### Dealing with missing data in model-based clustering through a MNAR model

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### Take home message

- **1** The missing data pattern may convey some information on clustering
- **2** Embed the missingness mechanism directly within the clustering modeling step

#### **1** Introduction

2 A model-based MNAR clustering approach

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#### **5** Concluding remarks

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### Missing data: an inevitable event

The larger the datasets, the more missing data may appear. . .

Two traditional solutions (for obtaining a filled dataset)

- Discard individuals with missing data: expect to add variance into analysis
- Impute missing data: expect to add bias modeling into analysis

#### General guidelines

- Obtaining a non-missing dataset is not the final goal
- Missing data management should take into account the initial analysis target

### Our analysis target: model-based clustering

Embed missing data management into this paradigm. . .

### Missing data: notations

- $y = \{y_1, \ldots, y_n\}$ : full dataset with *n* individuals
- $\mathbf{y}_i = (y_i^1, \dots, y_i^d) \in \mathbb{R}^d$ : full individual  $i \in \{1, \dots, n\}$
- $\mathbf{c} = \{c_1, \ldots, c_n\}$ : pattern of missing data for the full dataset
- $\mathbf{c}_i = (c_i^1, \ldots, c_i^d) \in \{0, 1\}^d$ : pattern of missing data for individual  $i \in \{1, \ldots, n\}$

$$
c_i^j = 1 \Leftrightarrow y_i^j \text{ is missing}
$$

- $o_i = \{j : c_i^j = 0\}$ : the observed variables indexes for individual *i*  $\mathbf{y}_{i}^{\mathsf{o}_{i}}$ : the observed variables values for individual i
- $\mathsf{y}^\mathsf{o} = \{ \mathsf{y}_1^{\mathsf{o}_1}, \dots, \mathsf{y}_n^{\mathsf{o}_n} \}$ : the observed values in  $\mathsf{y}$
- $m_i = \{j : c_i^j = 1\}$ : the missing variables indexes for individual *i*
- $\mathbf{y}_i^{m_j}$ : the missing variables values for individual  $i$
- $\mathsf{y}^m = \{\mathsf{y}_1^{m_1}, \ldots, \mathsf{y}_n^{m_n}\}$ : the missing values in  $\mathsf{y}$

 $\mathbf{y} = \{ \mathbf{y}^o, \mathbf{y}^m \}$  is the full dataset with its observed and missing parts

### Missing data: typology of the missing mechanisms

**Missing completely at random (MCAR):** 

$$
P(\mathbf{c}|\mathbf{y};\boldsymbol{\psi}) = P(\mathbf{c};\boldsymbol{\psi}) \quad \forall \mathbf{y}
$$

**Missing at random (MAR):** 

$$
P(\mathbf{c}|\mathbf{y};\boldsymbol{\psi}) = P(\mathbf{c}|\mathbf{y}^{\circ};\boldsymbol{\psi}) \quad \forall \mathbf{y}^m
$$

**Missing not at random (MNAR): the mechanism is not MCAR nor MAR** 

### Clustering: model-based approach

- **Partition with K clusters:**  $z = (z_1, \ldots, z_n)$  where
	- $z_i = (z_i^1, \ldots, z_i^K) \in \{0, 1\}^K$  $z_i^k = 1$  if  $y_i$  belongs to cluster  $k$ ,  $z_i^k = 0$  otherwise
- Gaussian mixture:  $y_1, \ldots, y_n$  are i.i.d. from the mixture

$$
f(\mathbf{y}_i;\boldsymbol{\theta}) = \sum_{k=1}^K \pi_k \phi_k(\mathbf{y}_i;\theta_k)
$$

where:

 $\pi_k = P(z_i^k = 1)$  $\phi_k(.;\theta_k):$  d-variate Gaussian pdf with mean vector and covariance matrix  $\theta_k = (\mu_k, \Sigma_k)$ , d-multinomial pdf with  $\theta_k = p_k$  probabilities vector or mixed pdf.  $\mathbf{\theta} = (\pi_1, \ldots, \pi_K, \theta_1, \ldots, \theta_K)$ 

#### Question we address in this work

Which distribution  $P(c|y, z; \psi)$  to propose in this clustering context?

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### Logistic model: a natural and flexible candidate

$$
P(\mathbf{c}|\mathbf{y}, \mathbf{z}; \psi) = \prod_{i=1}^{n} \prod_{j=1}^{d} P(c_i^j | \mathbf{y}, \mathbf{z}; \psi)
$$

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\text{I}(\mathsf{M}) = (\alpha_0, \beta_1 \cdots \beta_k \cdots \beta_k
$$

### MNARz analysis: it depends on y through z!

$$
P(c_i^j = 1 | \mathbf{y}; \boldsymbol{\theta}, \boldsymbol{\psi}) = \sum_{k=1}^K P(c_i^j = 1 | \mathbf{y}, \mathbf{z}; \boldsymbol{\psi}) P(\mathbf{z} | \mathbf{y}; \boldsymbol{\theta})
$$

Example of a univariate Gaussian model with the three components

 $0.2N(\cdot; 0, 1) + 0.3N(\cdot; 1, 2) + 0.5N(\cdot; 2, 3)$ 

and with parameters of the logit expression:  $\alpha_0 = 1, \beta_1 = 1, \beta_2 = -1, \beta_3 = 1$ 



### MNARz analysis: pattern c gives information on partition z!

Draw Bayes error of a MNARz model with two components and 20% of missing data

 $\pi_k = 0.5$ ,  $\|\boldsymbol{\mu}_2 - \boldsymbol{\mu}_1\|$  varies,  $\Sigma_1 = \Sigma_2 = I, |\beta_2 - \beta_1|$  varies



Both  $\mu_k$  and  $\beta_k$  act on the Bayes error

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### Ignorable vs. non ignorable model

A missing mechanism is ignorable if likelihoods can be decomposed as

$$
L(\boldsymbol{\theta}, \boldsymbol{\psi}; \underbrace{\mathbf{y}^o, \mathbf{c}}_{\text{observed data}}) = L(\boldsymbol{\psi}; \mathbf{c} | \mathbf{y}^o) \times L(\boldsymbol{\theta}; \mathbf{y}^o)
$$

Some simple algebra show that this occurs when missing mechanism is not MNAR

#### Inference of θ

"If the missing mechanism is ignorable then likelihood-based inferences for  $\theta$  from  $L(\theta; y^{\circ})$  will be the same as likelihood based inference for  $\theta$  from  $L(\theta, \psi; y^{\circ}, c)$ ." ([Little and Rubin, 2002] Section 6.2)

- **MCAR** is ignorable
- **MNARZ, MNARy and MNARyz are non ignorable**

### EM algorithm: looks simple

Decomposition of  $\mathit{Q}(\boldsymbol{\theta},\boldsymbol{\psi};\widehat{\boldsymbol{\theta}},\widehat{\boldsymbol{\psi}})$ 

The expected complete log-likelihood conditional related to observed data is:

$$
E\left[L(\boldsymbol{\theta}, \boldsymbol{\psi}; \mathbf{c}, \mathbf{y}, \mathbf{z}); \widehat{\boldsymbol{\theta}}, \widehat{\boldsymbol{\psi}} | \mathbf{y}^o, \mathbf{c}\right] = Q_{\mathsf{y}}(\boldsymbol{\theta}; \widehat{\boldsymbol{\theta}}, \widehat{\boldsymbol{\psi}}) + Q_{\mathsf{c}}(\boldsymbol{\psi}; \widehat{\boldsymbol{\theta}}, \widehat{\boldsymbol{\psi}})
$$

$$
Q_{y}(\theta; \widehat{\theta}, \widehat{\psi}) = \sum_{i=1}^{n} \sum_{k=1}^{K} \tau_{i}^{k} E\left[\log(\pi_{k} \phi_{k}(\mathbf{y}_{i}; \theta_{k})) | \mathbf{y}_{i}^{o_{i}}, \mathbf{c}_{i}; \widehat{\theta}, \widehat{\psi}\right]
$$

$$
Q_{c}(\psi; \widehat{\theta}, \widehat{\psi}) = \sum_{i=1}^{n} \sum_{k=1}^{K} \tau_{i}^{k} E\left[\log(P(\mathbf{c}_{i} | \mathbf{z}_{i}^{k} = 1, \mathbf{y}_{i}; \theta, \psi)) | \mathbf{y}_{i}^{o_{i}}, \mathbf{c}_{i}; \widehat{\theta}, \widehat{\psi}\right]
$$

$$
\tau_{i}^{k} = P(z_{i}^{k} = 1 | \mathbf{c}_{i}, \mathbf{y}_{i}^{o_{i}}; \widehat{\theta}, \widehat{\psi}) = \frac{\widehat{\pi}_{k} \phi_{k}(\mathbf{y}_{i}^{o_{i}}; \theta_{k}^{o_{i}}) P(\mathbf{c}_{i} | \mathbf{z}_{i}^{k} = 1, \mathbf{y}_{i}^{o_{i}}; \widehat{\psi})}{\sum_{h=1}^{K} \widehat{\pi}_{h} \phi_{h}(\mathbf{y}_{i}^{o_{i}})^{o_{i}} P(\mathbf{c}_{i} | \mathbf{z}_{i}^{h} = 1, \mathbf{y}_{i}^{o_{i}}; \widehat{\psi})}
$$

ш

### EM and/or SEM algorithms

MCAR (and also MAR . . . ): classical formula! . . . . . . . . . . . . . . . . . (EM , SEM)  $\tau_i^k \propto \widehat{\pi}_k \phi_k(\mathbf{y}_i^{o_i}; \theta_k^{o_i})$ 

 $\vert$  MNARz  $\vert$ : needs some new calculus but still simple  $\ldots$ ............. (EM, SEM)  $\mathcal{L}_{\mathcal{A}}$ 

$$
\tau_i^k \propto \widehat{\pi}_k \phi_k(\mathbf{y}_i^{o_i}; \theta_i^{o_i}) \prod_{j=1}^d (1 + \exp(-r_i^j \widehat{\beta}_k))^{-1} \text{ where } r_i^j = \left\{ \begin{array}{cl} 1 & \text{if } c_i^j = 1 \\ -1 & \text{otherwise} \end{array} \right.
$$

MNARy : needs approximations . (EM , SEM)

$$
P(c_i^j|\textbf{y}_i^{o_i}, \textbf{z}_i^k=1;\boldsymbol{\psi})=\left\{\begin{array}{cl} \int_{-\infty}^{+\infty}\frac{1}{1+\exp(-(\alpha_jy_i^j))}\phi_k(y_i^j;\theta_k^j)dy_i^j & \text{if } c_i^j=1\\ \\ \frac{1}{1+\exp(\alpha_jy_i^j)} & \text{otherwise} \end{array}\right.
$$

In the Gaussian case, there is no closed form [Pirjol, 2013] (same for MNARyz) But SEM is still simple in that case thanks to random drawing instead of expectation

### Link with some usual procedures!

Concatenation [Jones, 1996]: model equivalence  $MNARz^{j}(y^{\circ}, c) \Longleftrightarrow MCAR(-|y^{\circ}|c)$  )  $MNARz(y^o, c) \iff MCAR$  $\sqrt{ }$  $\overline{1}$  $\sqrt{ }$  $\left(y^{\circ}\right)$  $\left(\sum_{j=1}^d\right)$ c j . Y  $\overline{1}$ Y  $\overline{1}$  $\setminus$  $\overline{1}$ 

"All Available Cases" [Little and Rubin, 2002]: estimation equivalence In case of conditional independence between variables, whatever MCAR or MNAR\*:

Classical (S)EM  $\Longleftrightarrow$  (S)EM without estimating missing y<sup>m</sup>

. . . an opportunity to reduce the computing time

### What about model selection?

Can select between MCAR and MNAR∗ with any information criterion (BIC, ICL)

Even if the missing mechanism is ignorable for MCAR. . .

. . . need to model c to compare a MCAR and a MNAR model

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### Hospital Data

- Number of patients:  $n = 5$  146
- Number of features:  $d = 7$ 
	- Age
	- Size
	- Weight
	- Cardiac frequency
	- Hemoglobin concentration
	- Temperature
	- **Minimum Diastolic and Systolic Blood Pressure**
- Percentage of missing data: 6.4%

### ICL comparison





- **MCAR, MNARy and MNARz are equivalent until**  $K = 3$
- **MNARz** and MNARyz clearly indicate presence of an additional cluster ( $K = 4$ )

It seems to be an illustation of the effect of c through MNARz and MNARyz

### Missing Pattern



It seems that MNARz modelling leads to a missing free cluster

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#### **Summary**

- Interest to put a model on c
- Interest of the simple but meaningful model MNARz
- $\blacksquare$  Link between our models and usual methods

### Ongoing works

- Deeper analysis of the previous results with doctors...
- **Implement the proposed models/algo. in the Mixmod software**<sup>a</sup>
- Use mixed data algorithms for medical study with the same MNAR\* models

a http://www.mixmod.org

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