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# A Monte Carlo estimation of the mean residence time in cells surrounded by thin layers

Antoine Lejay\*

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#### Abstract

We present a new Monte Carlo method to estimate the mean-residence time of a diffusive particle in a domain surrounded by a thin layer of low diffusivity. Through a homogenization technique, the layer is identified with a membrane. The simulations use a stochastic process called the snapping out Brownian motion the density of which matches suitable transmission conditions at the membrane. We provide a benchmark test which is a simplified form of a real-life problem coming from brain imaging techniques. We also provide a new algorithm to adaptively estimate the exponential rate of the tail of the distribution function of the probability to be in the domain using Monte Carlo simulations.

**Keywords:** semi-permeable membrane, elastic Brownian motion, snapping out Brownian motion, thin layer problem, mean-residence time, Monte Carlo methods, homogenization, first eigenvalue estimation.

## **1** Introduction

Understanding the dynamic of a diffusive process in a medium with interfaces is of great importance both for modelling and simulation. An interface could be a permeable or a semipermeable barrier. It arises as a limit of a sharp change in the properties of the underlying material. In geophysics, the inclusion of rocks, fissures, ..., leads to such interfaces, just to cite one among the many possible domains of applications (see, *e.g.*, [2, 12]).

We consider here a medium with a "thin layer" of low diffusivity seen in the limit [19] as a semi-permeable barrier the particle has difficulty to pass through. The membrane surrounding living cells is an example of such a layer. A brain imaging techniques such as *diffusion Magnetic Resonance Imaging* (dMRI) record the mean square displacement of particles of water [6]. Reconstructing the brain activity is done by estimating the diffusion coefficient through an inverse problem. Understanding the diffusive behavior of particles in living tissue, where thin layers are frequently present, is then of essential for applying Monte Carlo methods [16].

In brain imaging, the *mean residence time* (MRT) in the cells is an important macroscopic parameter related to the rate of convergence toward equilibrium. It could serves as an input for simplified, homogenized models such as double porosity models [3, 4, 15, 18]. The MRT is also commonly used in pharmacokinetics and geophysics [14, 22].

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We propose a way to estimate the MRT in media with thin layers from Monte Carlo simulations using continuous processes instead of random walks as in [18]. We simulate paths of a stochastic process the density of which solves the PDE giving the concentration of the species of interest. To avoid the large computational cost induced by the layers, which forces to use a small time step around it, we replace the process by a simpler one called the *snapping out Brownian motion* [9]. This is equivalent to replacing the thin layer by a semi-permeable membrane. We then show how to estimate the MRT from the Monte Carlo simulation. By the ways, we improve the estimation procedure of a first eigenvalue problem proposed in [11]. Finally, we provide a one-dimensional benchmark test which shows the effectiveness of our approach unless the permeability of the membrane is too low.

### **2** A Monte Carlo estimation of the mean residence time

The concentration C(t,x) of the diffusive species (water, ...) follows an equation of the type

$$\partial_t C(t,x) = \nabla \cdot (D(x)\nabla C(t,x)), \ t > 0, \ x \in \Omega, \ d = 1,2,3$$
(1)

in a periodic cell  $\Omega$  with a scalar diffusivity. We consider here a simplified situation where  $\Omega$  contains an interior part  $\Omega_{intra}$  separated from an exterior part  $\Omega_{extra}$  by a layer  $\Omega_{layer}$  of constant width  $\epsilon > 0$ . The surface separating  $\Omega_{layer}$  and  $\Omega_{intra}$  (resp.  $\Omega_{extra}$ ) is denoted by  $\Gamma_{-}$  (resp.  $\Gamma_{+}$ ). We assume that  $\Gamma_{\pm}$  are of class  $\mathscr{C}^{2}$ .

For  $D_0, \mu > 0$ , the diffusivity is  $D(x) = D_0$  for  $x \in \Omega_{\text{intra}} \cup \Omega_{\text{extra}}$ , while  $D(x) = D_{\text{layer}} = \mu \epsilon$  for  $x \in \Omega_{\text{laver}}$ .

Under the above conditions, (1) is well posed and has a unique solution for any initial condition which is bounded and measurable.

#### 2.1 The transmission condition

Since *D* is constant except on the interfaces  $\Gamma_{-}$  and  $\Gamma_{+}$ , we recast (1) into a *transmission* problem.

We denote by  $n_{-}(x)$  (resp.  $n_{+}(x)$ ) the normal derivative at a point  $x \in \Gamma_{-}$  (resp.  $x \in \Gamma_{+}$ ) which is directed toward  $\Omega_{\text{intra}}$  (resp.  $\Omega_{\text{extra}}$ ). The concentration *C* is of class  $\mathscr{C}^{1}$  on the left and right of  $\Gamma_{\pm}$ . Besides, it satisfies

$$\partial_t C(t,x) = D(x) \triangle C(t,x), t > 0, x \notin \Gamma_{\pm},$$

with

$$[n_{\pm}(x) \cdot D(x)\nabla C(t,x)]_{\Gamma_{+}} = 0 \text{ and } [C(t,x)]_{\Gamma_{+}} = 0,$$
(2)

where at a point  $x \in \Gamma_{\pm}$ ,

$$[f(x)]_{\Gamma_{\pm}} = \lim_{\delta \to 0} f(x + \delta n_{\pm}(x)) - f(x - \delta n_{\pm}(x)).$$

The condition (2) is called the *transmission condition* (see [7, §III.13, p. 224]). It specifies the continuity of the *flow*  $n_{\pm} \cdot D(x)\nabla C(t,x)$  across the interface  $\Gamma_{\pm}$  as well as the continuity of the concentration.

#### 2.2 The associated stochastic process

For the diffusivity D,  $\Omega$ ,  $\Gamma_{\pm}$ , ... above extended to  $\mathbb{R}^d$  by periodicity, let us consider the stochastic process X associated to the divergence form operator  $\mathscr{L} = \nabla \cdot (D\nabla \cdot)$ . This process is a strong Markov process with continuous paths [20]. When D is  $\mathscr{C}^1$ , the process X is solution to the Stochastic Differential Equations  $X_t = x + \int_0^t \sigma(X_s) dB_s + \int_0^t \nabla D(X_s) ds$ , where  $\sigma \sigma^T(x) = 2D(x)$ . In our case,  $\nabla D$  is not defined on  $\Gamma_{\pm}$ . Therefore, no such representation holds for X.

The probability  $p_{\text{cross}}$  that the particle moves at distance at least R during a time step  $\Delta t$  decreases exponentially fast with  $R^2/\Delta t$  (this is a consequence of the Gaussian estimates on the density transition function of X [20]). For numerical simulations, we substitute to  $\{X_{t+s}\}_{s\geq 0}$  a process  $\{Z_{t+s}\}_{s\geq 0}$  with the same distribution until it leaves the ball of radius R centered on  $X_t$ . In practice,  $R = 4\sqrt{D_0\Delta t}$  is sufficient, as the probability  $p_{\text{cross}}$  is of order  $6 \times 10^{-5}$  as discussed in [12]. For away from the interfaces, the particle could be moved using Gaussian steps, which consists only in neglecting the probability to cross them.

This explain why, up to negligible errors, we focus on the dynamic of the process only at the vicinity of the layers.

Let  $\phi$  be a  $\mathscr{C}^2$  diffeomorphism of  $\mathbb{R}^d$ . Then  $Y_t = \phi(X_t)$  is also a strong Markov process with continuous paths. Its infinitesimal generator  $\mathscr{A}$  has domain  $\text{Dom}(\mathscr{A}) = \phi(\text{Dom}(\mathscr{L}))$ . From standard computations, for  $x \notin \Gamma_{\pm}$  and  $f \in \text{Dom}(\mathscr{L})$ ,

$$\mathcal{A}f(x) = \mathcal{L}(f \circ \phi)(\phi^{-1}(x)) = D(\phi^{-1}(x)) \sum_{i,j,k=1,\dots,d} \frac{\partial^2 f}{\partial x_j \partial x_k}(x) \frac{\partial \phi_k}{\partial x_i}(x) \frac{\partial \phi_j}{\partial x_i}(x) + D(\phi^{-1}(x)) \sum_{i,j=1,\dots,d} \frac{\partial f}{\partial x_j}(x) \frac{\partial^2 \phi_j}{\partial x_i^2}(x).$$
(3)

When  $\phi$  is a translation or a rotation,  $\mathscr{A}f(x) = \mathscr{L}f(\phi(x))$ .

Recall that the layer  $\Omega_{\text{layer}}$  is surrounded by two  $\mathscr{C}^2$ -interfaces  $\Gamma_-$  and  $\Gamma_+$  at distance  $\epsilon$ . We use an argument similar to the one of [2] to reduce the problem to a "nearly 1-dimensional" one.

**Proposition 1.** At the vicinity of any point  $\hat{x}$  in  $\Gamma_-$ , there exists a  $\mathscr{C}^2$  diffeomorphism  $\phi$  such that  $\phi(\Gamma_-)$  and  $\phi(\Gamma_+)$  are the hyperplanes orthogonal to (1, 0, ..., 0) passing through  $\phi(\hat{x}) = 0$  and  $(\epsilon, 0, ..., 0)$ .

*Proof.* Up to compositions with a rotation and a translation, we may assume that our reference point is  $\hat{x} = 0$  and  $\Gamma_{-} = U \cap \{(\omega(x'), x') | x' \in \mathbb{R}^{d-1}\}$  for a smooth map  $\omega : \mathbb{R}^{d-1} \to \mathbb{R}$  and a neighborhood U of 0. Let  $d : \mathbb{R}^d \to \mathbb{R}_+$  be the distance between a point and  $\Gamma_-$ . At the vicinity of  $\Gamma_-$ , the distance is a well defined function with the same regularity as  $\Gamma_-$  [5]. In addition, there exists a unique point y in  $\Gamma_-$  such that d(x) = |x - y|. Let  $n(x) = \nabla \omega(x')$  be the unit vector normal to  $\Gamma_-$  oriented toward the exterior of the cell at the point  $x = (\omega(x'), x')$ . Write  $\psi(d, x') = x + dn(x)$  with  $x = (\omega(x'), x')$  for  $d \in \mathbb{R}$  and  $x' \in \mathbb{R}^{d-1}$ . Thus,  $\phi$  is locally a diffeomorphism and  $\phi^{-1}$  maps  $\Gamma_- \cap V$  to  $\{(0, x') | x' \in \mathbb{R}^{d-1}\} \cap U$  as well as  $\Gamma_+ \cap V$  to  $\{(\epsilon, x') | x' \in \mathbb{R}^{d-1}\} \cap U$  for two neighborhoods V and U of 0. This function  $\phi$  is the diffeomorphism we are looking at.

Using this, we focus on the component which is normal to the interface  $\Gamma_{\pm}$ . We are then is the same situation as the one described in [10].

**Corollary 1.** Using the diffeomorphism  $\phi$  given above at the vicinity of a point  $\hat{x}$  of the layer, we set  $X^{\phi} = \phi(X)$ . The first component  $(X^{\phi})^1$  of  $X^{\phi}$  is a diffusion process. Write  $d(x) = D(\phi^{-1}(x, 0, ..., 0))$  for  $x \in \mathbb{R}$ . The infinitesimal generator of  $(X^{\phi})^1$  is  $\mathscr{P} = \nabla(d\nabla \cdot)$ .

#### 2.3 The thin layer approximation

The particles have difficulties in traversing the thin layer which is then seen as a semi-permeable barrier. Thanks to a homogenization result [19], our diffusion equation could be replaced by a simpler one where the layer enclosed between  $\Gamma_+$  and  $\Gamma_-$  is reduced to a single interface  $\Gamma$ . The condition at this single interface  $\Gamma$  is

$$\mu(C(t,x+) - C(t,x-)) = D_0 n_+(x) \cdot \nabla C(t,x) \text{ and } [n_+(x)C(t,x)]_{\Gamma} = 0, \ x \in \Gamma,$$
(4)

where *C* is the solution to the PDE

$$\partial_t C(t, x) = D_0 \triangle C(t, x) \text{ for } x \notin \Gamma \text{ and } t > 0.$$
 (5)

This approximation then simplifies theoretical or numerical computations in modelling [3, 4, 21].

In dimension 1, a similar analysis can be carried at the level of the stochastic process. Therefore, X may be replaced by a simpler process Y, called a *SNapping Out Diffusion* (SNOD), living in a medium where the layers are replaced by semi-permeable barriers of vanishing width which we call *membranes*. The coefficient  $\mu$  in (4) is called the *membrane permeability*.

The state space of such a process is then a disjoint union of close intervals, where each finite endpoint corresponds to one side of a membrane. The density transition function p(t,x,y) of X is solution to (4)-(5) in the variables (t,x).

When a SNOD reaches a membrane, a local time clock drives a switching in which the process is replaced randomly on one or the other side of the membrane.

Using the Lamperti's type transform presented below in Proposition 2 to reduce its diffusivity to 1/2 (the one of the Brownian motion), a SNOD living in  $\mathbb{R}$  with one membrane at 0 becomes a *SNapping Out Brownian motion* (SNOB) [1, 9], a process which depends on a single parameter  $\kappa > 0$  related to  $D_0$  and  $\mu$ .

A SNOB behaves like an elastic Brownian motion on  $(-\infty, 0-] \cup [0+,\infty)$ . When its local time at  $0\pm$  goes above an independent exponential random variable of mean  $1/\kappa$ , the process is replaced with probability 1/2 to either at 0+ or 0-. It starts afresh with the same dynamics. The behavior at 0 models a semi-permeable barrier the particle goes through after a random time, in a Markovian manner.

The infinitesimal generator of the SNOB is  $\mathscr{S} = \frac{1}{2} \triangle$ . The domain of  $\mathscr{S}$  contains as a core the functions satisfying the transmission conditions

$$\nabla u(0+) = \nabla u(0-)$$
 and  $\kappa(u(0+) - u(0-)) = 2\nabla u(0)$ 

and which are bounded and of class  $\mathscr{C}^2$  elsewhere with a bounded first and second order derivatives.

**Proposition 2** ([8]). Let Y be a SNOD in a one-dimensional media of constant diffusivity  $D_0$  with one membrane at 0 of permeability  $\mu$ . With  $\psi(x) = y/\sqrt{2D_0}$ , the process  $Y^{\psi} = \psi(Y)$  is a SNOB of parameter  $\kappa = 2^{3/2} \mu/\sqrt{D_0}$  when restricted to the vicinity of 0.

When f is a bounded, continuous function and Z a SNOB of parameter  $\kappa$ , it was shown in [9] that

$$\mathbb{E}_{x}[f(Z_{t})] = \mathbb{E}_{x}\left[\frac{1 + e^{-\kappa L_{t}^{0}(B)}}{2}f(|B_{t}|)\right] + \mathbb{E}_{x}\left[\frac{1 - e^{-\kappa L_{t}^{0}(B)}}{2}f(-|B_{t}|)\right] \text{ for } x \ge 0,$$
(6)

where *B* is a Brownian motion and  $L_t^0(B)$  its symmetric local time at 0.

Using the representation (6) of the semi-group and some explicit formula for the simulation of the local time [13], a simulation algorithm could be given for the SNOB which is presented in [8, 9].

The tangential components behave like a Brownian motion and are then easily simulated.

#### 2.4 A Monte Carlo estimator of the mean residence time

The mean residence time (MRT) is an important macroscopic quantity that could be used to construct two-scale models such as in [3].

In the periodic medium  $\Omega$ , the operator  $\check{\mathscr{L}} = \nabla(D\nabla \cdot)$  with domain  $\text{Dom}(\check{\mathscr{L}}) = \{f \in L^2(\Omega) | \check{\mathscr{L}}f \in L^2(\Omega)\}$  is self-adjoint with a compact resolvent. In particular, it admits a spectral decomposition  $0 \leq \lambda_0 < \lambda_1 < \cdots$  with eigenfunctions  $\eta_0, \eta_1, \ldots$  chosen orthonormal for the  $L^2(\Omega)$  scalar product (regarding the sign of the eigenvalues, we use the convention  $\mathscr{L}\eta_k = -\lambda_k\eta_k$ ). The first eigenvalue is  $\lambda_0 = 0$  and its associated eigenfunction is  $\eta_0(x) = 1/\sqrt{V_{\text{per}}}$ , there  $V_{\text{per}}$  is the volume of the periodic cell  $\Omega$ .

**Proposition 3** ([3, 8]). The mean residence time  $\tau$  in the cell satisfies the approximation

$$\frac{1}{\tau} \approx \frac{V_{\text{intra}}}{V_{\text{per}}} \lambda_1, \tag{7}$$

where  $V_{intra}$  is the volume of interior part of the cell.

Determining the first eigenvalue  $\lambda_1$  is then crucial for estimating the MRT. We explain now its relationship with the process  $\check{X}$  defined as the projection of X on the periodic cell  $\Omega$ .

The density transition function p(t, x, y) with respect to the Lebesgue measure of X admits the spectral expansion

$$p(t,x,y) = \sum_{k=0}^{+\infty} e^{-\lambda_k t} \eta_k(x) \eta_k(y)$$

Since  $\eta_0(x)$  is constant and for a fixed point  $x \in \Omega$ ,

$$p(t) = \int_{\Omega_{\text{intra}}} p(t, x, y) \, \mathrm{d}y = \frac{V_{\text{intra}}}{V_{\text{per}}} + C\eta_1(x)e^{-\lambda_1 t} + o(\exp(-\lambda_1 t)) \tag{8}$$

with  $C = \int_{\Omega_{\text{intra}}} \eta_1(y) dy$ .

As t increases, p(t) converges to the ratio  $p_{\infty} = V_{intra}/V_{per}$ . Thus,  $-\lambda_1$  is actually the exponential rate of convergence toward equilibrium.

On the other hand,  $p(t) = \mathbb{P}_x[\dot{X}_t \in \Omega_{intra}]$ . By simulating a large number N of paths of the stochastic process  $\check{X}$ , we estimate p(t) at some given times  $\{t_k\}_{k=1,...,M}$  using the empirical estimators

$$p_N(t_k) = \frac{1}{N} \sum_{i=1}^N \mathbb{1}_{\check{X}_{t_k}^{(i)} \in \Omega_{\text{intra}}},\tag{9}$$

where  $\check{X}^{(1)}, \ldots, \check{X}^{(N)}$  are N independent samples of the paths  $\check{X}$ .

*Remark* 1. For practical computation, we use the projection on  $\Omega$  of the SNOD approximation  $\check{Y}$  of  $\check{X}$ . A similar analysis may be carried on its infinitesimal generator.

From the knowledge of  $\{p_N(t_k)\}_{k=1,\dots,M}$ , we estimate  $\lambda_1$  through a linear regression on  $\{\log(p_N(t_k) - p_\infty)\}_{k=1,\dots,M}$ . However, the first and last times  $t_1$  and  $t_M$  shall be carefully chosen. For this, we use the procedure presented in Section 3.

# 3 A procedure to estimate some exponential rate from Monte Carlo measurement

Let p(t) be the probability of an event which decreases to 0 with the time *t* at some exponential rate  $\lambda$ . This is typical for first eigenvalues problems, such as (8) and the estimation of survival probability of a Brownian motion killed when it exits from a domain [11].

More precisely, we are willing to estimate  $\lambda$  under the assumption that

$$p(t) = p_{\infty} + c \exp(-\lambda t) + r(t) \text{ with } r(t) = o_{t \to \infty}(\exp(-\lambda t)), \tag{10}$$

where  $p_{\infty} \ge 0$  is known, yet c > 0,  $\lambda > 0$  and r are unknown. In particular, we ignore how fast  $r(t)\exp(-\lambda t)$  decreases to 0.

At different times  $0 < t_1 < \cdots < t_M$ , this quantity  $p(t_i)$  is estimated by  $p_N(t_i)$  from a Monte Carlo estimation with N independent samples. A Gaussian confidence interval allows one to assert that with a probability  $\alpha$ ,

$$|p(t) - p_N(t)| \le c_\alpha \frac{\sqrt{p(t)(1 - p(t))}}{\sqrt{N}} \text{ for } N \text{ large enough and } t > 0, \tag{11}$$

with, e.g.,  $c_{\alpha} = 2.57$  for  $\alpha = 99\%$ . From (10)

From (10),

$$\log(p_N(t) - p_\infty) = \log c - \lambda t + \xi(t)$$

with

$$\xi(t) = \log\left(1 + e^{\lambda t}r(t) + e^{\lambda t}(p_N(t) - p(t))\right) \approx e^{\lambda t}r(t) + e^{\lambda t}(p_N(t) - p(t)).$$

The difficulty is then to find a time interval  $[t_-, t_+]$  so that  $\xi(t)$  is small in front of  $\log c$  and  $\lambda t$ , so that a linear regression could be used.

When  $t_{-}$  and  $t_{+}$  are fixed,  $c_{est}$  and  $\lambda_{est}$  are the values which minimize

$$SQ(t_{-},t_{+};c,\lambda) = \sum_{\substack{i=1,...,M\\t_{-} \le t_{i} \le t_{+}}} (\log c - \lambda t_{i} - \log(p_{N}(t_{i}) - p_{\infty}))^{2} = \sum_{\substack{i=1,...,M\\t_{-} \le t_{i} \le t_{+}}} \xi(t_{i})^{2}.$$

The residual sum of squares is

$$\operatorname{RSS}(t_{-},t_{+}) = \operatorname{SQ}(t_{-},t_{+};c_{\operatorname{est}},\lambda_{\operatorname{est}}) = \min_{c,\lambda} \operatorname{SQ}(t_{-},t_{+};c,\lambda).$$
(12)

Several criteria have been proposed in [11] to find the best time interval  $[t_-, t_+]$  for estimating  $\lambda$  through a least squares procedure. Yet they partially rely on a graphical procedure, which we want to avoid.

The Monte Carlo error can be neglected unless  $c_{\alpha}\sqrt{p(t)(1-p(t))}/\sqrt{N}$ , the magnitude of  $p_N(t) - p(t)$ , has the same order as  $p(t) - p_{\infty} \approx c \exp(-\lambda t)$ . Hence, it is useless to consider times t at which

$$c_{\alpha} \frac{\sqrt{p_N(t)(1-p_N(t))}}{\sqrt{N}} \ge \beta(p_N(t)-p_{\infty}) \text{ with say } \beta = \frac{1}{10}.$$

This criterion fixes a limit for the maximum sampling time  $t_M$  to consider. After this time, the estimations are no more relevant as being "flooded" by the Monte Carlo error.

Getting rid of r(t) is more cumbersome, as this function is generally unknown.

In the report [8], we have proposed a new methodology in which  $t_{-}$  and  $t_{+}$  are determined adaptively using the Akaike Information Criteria (AIC) [17, Example 13.6, p. 179]. This approach is suitable as the number of values of  $p_N(t_i)$  to use varies with  $t_{-}$  and  $t_{+}$ .

Since there are two parameters (c and  $\lambda$ ) to estimate, we use the customary simplified formula for the AIC

AIC
$$(t_-, t_+) = n \log(\text{RSS}(t_-, t_+)) - n \log n + 4$$
  
with  $n = \#\{t_- \le t_i \le t_+; i = 1, \dots, M\}.$ 

The procedure is: We draw randomly some pairs  $(t_-, t_+)$  with  $0 \le t_- < t_+ \le t_M$ . We select the one for which AIC $(t_-, t_+)$  is minimized.

This procedure, whose result in illustrated in Figure 1, no longer relies on a graphical interpretation as in [11].

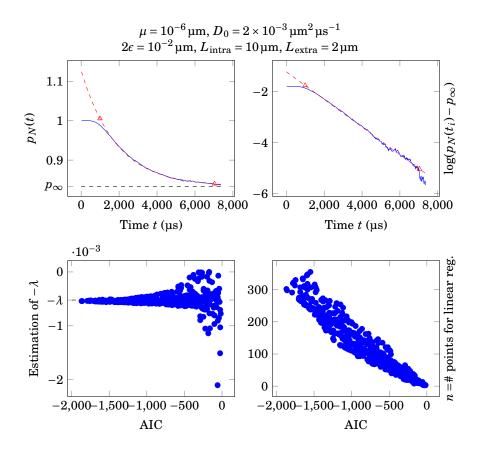


Figure 1: Estimation of  $\lambda$  using a least square procedure. The curve in red represents the estimated function  $c \exp(-\lambda t)$  from the time interval  $[t_-, t_+]$  delimited by the triangle. The two figures at the bottoms show that the values which are close to the minimum of AIC provide pretty close estimated values of  $\lambda$  and use a lot of sample times. We have used N = 100000 paths and  $t_{i+1} - t_i = 20 \,\mu\text{s}$  with the parameters of Table 1.

### 4 A computational benchmark test in dimension one

We test our simulation scheme by estimating a MRT which we compare to the numerical, deterministic estimation of an eigenvalue problem through (7). For this, we place ourselves in a one-dimensional medium, which allows explicit computations of the first positive eigenvalue  $\lambda = \lambda_1$ . We consider a cell  $\Omega_{intra}$  surrounded by two thin layers at positions  $y_1$  and  $y_2$  in a periodic medium  $\Omega$ . The set of points in the layers is  $\Omega_{layer}$ . We set  $\Omega_{extra} = \Omega \setminus (\Omega_{intra} \cup \Omega_{layer})$ . The diffusivity is  $D(x) = D_0$  for  $x \in \Omega_{intra} \cup \Omega_{extra}$ . In each layers, of width  $\epsilon$ , the diffusivity is  $D(x) = \mu \epsilon$ .

Using the thin layer approximation,  $\Omega_{\text{layer}}$  is just reduced to two points at the positions of the membranes replacing the layers.

Still using this approximation, we consider the process  $\check{Y}$  living in the periodic medium with two membranes at  $y_1$  and  $y_2$ , and diffusivity  $D_0$ . The process  $Z = \check{Y}/\sqrt{2D_0}$  has diffusivity 1/2 outside and inside the cell. Around each membrane,  $\check{Y}$  acts<sup>1</sup> like a SNOB with parameter  $\kappa = 2^{3/2} \mu / \sqrt{D_0}$  (Proposition 2). We simulate  $\check{Y}$  with the algorithm given in [9] to deal locally with each membrane.

According to (7), the MRT is then approximated by  $\tau \approx \lambda^{-1} V_{\text{per}} / V_{\text{intra}}$ .

We use a realistic range of parameters given in Table 1. We perform N independent

<sup>&</sup>lt;sup>1</sup>That is, by neglecting the excursions which are so large that  $\check{Y}$  crosses the other interface during the small time step

Name	Quantity	From	То
$D_0$	diffusivity in $\Omega_{intra} \cup \Omega_{extra}$	$2  imes 10^{-3}  \mu m^2  \mu s^{-1}$	$3  imes 10^{-3}  \mu m^2 \mu s^{-1}$ $10^{-6}  \mu m^2 \mu s^{-1}$
$\mu\epsilon$	diffusivity in $\Omega_{ m layer}$	$10^{-8}\mu{ m m}^2\mu{ m s}^{-1}$	$10^{-6}\mu{ m m}^2\mu{ m s}^{-1}$
$\epsilon$	layer width	0.01 µm	
$\mu = \frac{D_{\text{layer}}}{\epsilon}$	membrane's permeability	$10^{-6}\mu{ m m}\mu{ m s}^{-1}$	$10^{-4}\mu{ m m}\mu{ m s}^{-1}$
$V_{ m intra}$	intracellular size	5 µm	20 µm
$V_{ m extra}$	extracellular size	$pprox V_{ m intra}/10$	
$V_{ m per}$	medium's period	$V_{ m intra}$ + $V_{ m extra}$	
T	max. observations' time	$20000\mu s$	$50000\mu s$

Table 1: A realistic range of parameters.

simulations. With (9), we compute empirical estimations  $p_N(t_k)$  of  $p(t_k)$  given by (8) at given times  $0 < t_1 < \cdots < t_M \le T$ . We then use the procedure of Section 3 to estimate  $\lambda$ .

Being the first non-zero eigenvalue of  $\hat{\mathscr{L}}$ ,  $\lambda$  could also be computed deterministically as an eigenvalue of a Sturm-Liouville problem. Numerically, it is easily checked that the eigenvalues with the thin layer approximation are close to the ones of the original problem [8]. The process is simulated only using the thin layer approximation, which leads to faster computations. Without the thin layer approximation, the simulations could be performed using the algorithms given in [12]. Yet the thinness of the layer forces to use a very small time step, burdening the computational cost.

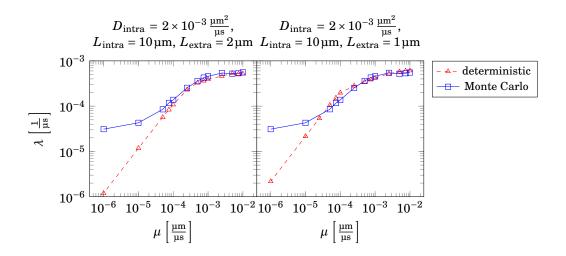


Figure 2: Estimation of  $\lambda$  by a deterministic method and a Monte Carlo one with N = 100000 particles.

Numerical results are given in Figure 2. Unless  $\mu$  is too small, the Monte Carlo results are in good agreements with the deterministic approach. A small value of  $\mu$  means that the particles have difficulties to pass through the semi-permeable barrier. This problem is related to a rare events simulation problem. This shall be subject to future investigations.

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