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Volume Laws for Boxed Plane Partitions and Area Laws for Ferrers Diagrams

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We asymptotically analyse the volume random variables of general, symmetric and cyclically symmetric plane partitions fitting inside a box. We consider the respective symmetry class equipped with the uniform distribution. We also prove area limit laws for two ensembles of Ferrers diagrams. Most limit laws are Gaussian.

Keywords: plane partition, partition, Ferrers diagram, limit law, combinatorial probability

1 Introduction

A plane partition fitting inside an (r, s, t)-box is an $r \times s$ -array Π of non-negative integers $p_{i,j} \leq t$ with weakly decreasing rows and columns. It can be visualised as a pile of unit cubes in the box $\mathcal{B}(r, s, t) :=$ $[0, r] \times [0, s] \times [0, t]$ "flushed into the corner", see figure 1 and Bressoud (1999). The figure also illustrates the connection to tilings of a hexagon of side lengths r, s and t by lozenges. Yet another interpretation is viewing such a pile of cubes as an order ideal in the product of three finite chains (total orders) with the respective lengths r, s and t. In the following we mean by "plane partition" one which fits inside an (r, s, t)-box. The volume of a plane partition is the sum of its parts or the number of unit cubes in the pile or the cardinality of the order ideal, respectively. If one side length, say t of the bounding box is one, such a pile can be regarded as Ferrers diagram fitting inside an $r \times s$ -rectangle of area equal to the volume. Let $\mathcal{PP}(r,s,t)$ denote the set of all plane partitions fitting inside $\mathcal{B}(r,s,t)$. A plane partition $\Pi \in \mathcal{PP}(r,r,t)$ is called *symmetric* if the corresponding pile of cubes is symmetric about x = y. To put it differently the cube (i, j, k) belongs to the pile, if and only if (j, i, k) does. The respective tiling is symmetric w.r.t. a vertical axis. A plane partition $\Pi \in \mathcal{PP}(r, r, r)$ is called *cyclically symmetric* if the corresponding tiling is invariant under a rotation about $2\pi/3$, i.e. the cube (i, j, k) belongs to the pile if and only if (k, i, j) and (j, k, i) do. Call these subsets SPP(r, t) and CSPP(r) respectively. Several authors discussed properties of randomly chosen plane partitions, for example in Wilson (2001) diagonal sums are considered. In Cohn et al. (1998) an "Arctic Circle Theorem" for tilings of a large hexagon is conjectured, which states that a typical tiling is "unordered" inside a certain ellipse and "periodic" or "frozen" outside. The shape of the "frozen region" is proved in Johansson (2002).

It is possible to compute volume limit laws, as the generating functions counting plane partitions by

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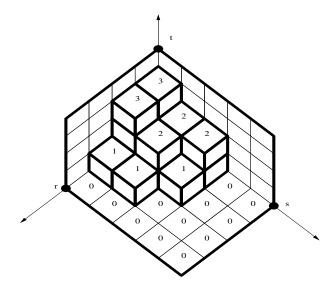


Fig. 1: A plane partition in a (4, 5, 4)-box

volume are available for the above classes. For the following formulas we refer to Bressoud (1999). The volume generating function of $\mathcal{PP}(r, s, t)$ is

$$\prod_{i=1}^{r} \prod_{j=1}^{s} \prod_{k=1}^{t} \frac{1 - q^{i+j+k-1}}{1 - q^{i+j+k-2}}.$$
(1.1)

Symmetric plane partitions in $\mathcal{B}(r,r,t)$ have the volume generating function

$$\left(\prod_{i=1}^{r} \prod_{k=1}^{t} \frac{1 - q^{2i+k-1}}{1 - q^{2i+k-2}}\right) \left(\prod_{1 \le i < j \le r} \prod_{k=1}^{t} \frac{1 - q^{2+2(i+j+k-2)}}{1 - q^{2(i+j+k-2)}}\right). \tag{1.2}$$

Here the first product corresponds to singleton orbits and the second product corresponds to doubleton orbits of $\mathcal{B}(r,r,t)$ under switching the first two coordinates. For the cyclically symmetric plane partitions we have the volume generating function

$$\left(\prod_{1 \le i < j < k \le r} \frac{1 - q^{3(i+j+k-1)}}{1 - q^{3(i+j+k-2)}}\right)^{2} \left(\prod_{1 \le i < k \le r} \frac{1 - q^{3(2i+k-1)}}{1 - q^{3(2i+k-2)}} \cdot \frac{1 - q^{3(i+2k-1)}}{1 - q^{3(i+2k-2)}}\right) \left(\prod_{i=1}^{r} \frac{1 - q^{3i-1}}{1 - q^{3i-2}}\right),\tag{1.3}$$

where the first product runs over C_3 orbits of $\mathcal{B}(r,r,r)$ with all coordinates distinct, the second runs over orbits with exactly two coordinates equal and the third corresponds to singleton orbits. The above generating functions are polynomials in q whose degree is the volume of the bounding box.

We equip $\mathcal{PP}(r, s, t)$, $\mathcal{SPP}(r, t)$ and $\mathcal{CSPP}(r)$ with the uniform distribution ("uniform fixed bounding

box ensemble"). Denote by $X=X_{rst}$ (resp. $=X_{rt},\ X_r$) the corresponding volume random variables. In the subsequent discussion we will drop subscripts as the side lengths are always r,s and t (resp. r and t,r). The probability generating function of this random variable is P(q):=G(q)/G(1) where G(q) is the respective generating function. In the $\mathcal{PP}(r,s,t)$ case this reads with the notion $\alpha_{ijk}:=i+j+k-1$

$$P(q) = \prod_{i=1}^{r} \prod_{j=1}^{s} \prod_{k=1}^{t} \frac{(\alpha_{ijk} - 1) (1 - q^{\alpha_{ijk}})}{\alpha_{ijk} (1 - q^{\alpha_{ijk} - 1})}.$$
 (1.4)

In the following $g(x)^{(N)}$ denotes the N^{th} derivative of g w.r.t. x. Since X is integer valued and bounded, moments of X of arbitrary order N exist and can be computed via

$$\mathbb{E}\left(X^{N}\right) = (-I)^{N} \left(P(e^{Ix})\right)^{(N)}\Big|_{x=0},\tag{1.5}$$

where $P(e^{Ix})$ ($I^2 = -1$) is the characteristic function of X.

2 Mean, variance and concentration properties

Denote by μ , μ_{spp} and μ_{cspp} the mean value of the volume variables on $\mathcal{PP}(r, s, t)$, $\mathcal{SPP}(r, t)$ and $\mathcal{CSPP}(r)$, respectively.

Lemma 2.1 The volume distributions are symmetric about half the volume of the bounding box, and hence the expected volumes are $\mu = rst/2$, $\mu_{spp} = r^2t/2$ and $\mu_{cspp} = r^3/2$.

Proof: Let v be the volume of the bounding box \mathcal{B} . The random variables X and v-X are seen to be equally distributed as follows. Let a plane partition of volume k be represented by a pile of k green cubes inside \mathcal{B} . Now fill up the empty space in \mathcal{B} with red cubes. The red cubes represent a plane partition of volume v-k. This construction describes a bijection between plane partitions of volume k and v-k which respects symmetry and cyclic symmetry. Now the lemma follows as $\mathbb{E}(X) = \mathbb{E}(v-X)$. \square In

order to prove the Gaussian limits we consider the characteristic functions $P(e^{Ix})$ of the random variables X. More precisely, we will study the logarithms of the characteristic functions, since sums are easier to handle than products. The following lemma enables us to compute explicit formulas for the variances of the random variables and to estimate the Taylor coefficients of the functions in question.

Lemma 2.2 For positive real numbers α , c with $\alpha > c \ge 1$, we have the expansion

$$\log\left(\frac{(\alpha-c)(1-e^{\alpha x})}{\alpha(1-e^{(\alpha-c)x})}\right) = \sum_{N\geq 1} H_{N,c}(\alpha)x^{N}.$$
 (2.1)

Here $H_{N,c}$ is a polynomial of degree N-1 in α . In particular we have $H_{1,c}(\alpha)=\frac{c}{2}$ and $H_{2,c}(\alpha)=-\frac{c}{12}\alpha+\frac{c^2}{24}$. Furthermore there is a positive constant D, such that the inequality

$$|H_{N,c}(\alpha)| \le D \cdot \alpha^{N-1} \cdot (2c)^N \tag{2.2}$$

holds for all $N \in \mathbb{N}$.

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Proof: Define the function g(t) by

$$g(t) = \log\left(\frac{e^t - 1}{t}\right). \tag{2.3}$$

Then the lhs of (2.1) is easily seen to be equal to $g(\alpha x) - g((\alpha - c)x)$. We have the following series expansion for g(t):

$$g(t) = \log\left(\frac{e^t - 1}{t}\right) = \log\left(1 + \frac{1}{2}t + \frac{1}{3!}t^2 + \ldots\right) = \sum_{N > 1} b_N t^N.$$
 (2.4)

The numbers b_N can be expressed in terms of Bernoulli numbers. For our purposes it suffices to observe that the singularities of g of smallest modulus are $\pm 2\pi I$, so the numbers $|b_N|$ decay asymptotically like $(2\pi)^{-N}$ and hence are bounded by some constant D. The N^{th} coefficient in the Taylor expansion of $g(\alpha x) - g((\alpha - c)x)$ about x = 0 is in fact a polynomial of degree N - 1, namely

$$H_{N,c}(\alpha) = b_N \cdot (\alpha^N - (\alpha - c)^N) = b_N \cdot \sum_{k=1}^N \binom{N}{k} (-1)^{k+1} c^k \alpha^{N-k}.$$
 (2.5)

As $\alpha > c \ge 1$ it can be estimated as follows:

$$\left| b_N \cdot \sum_{k=1}^{N} (-1)^{k+1} \binom{N}{k} c^k \alpha^{N-k} \right| \le D \cdot \alpha^{N-1} \cdot c^N \cdot \sum_{k=1}^{N} \binom{N}{k} \le D \cdot \alpha^{N-1} \cdot c^N \cdot 2^N. \tag{2.6}$$

This finishes the proof of Lemma 2.2. \Box Now we can easily compute the variances. Denote by σ ,

 σ_{spp} and σ_{cspp} the standard deviation of the volume variables on $\mathcal{PP}(r,s,t)$, $\mathcal{SPP}(r,t)$ and $\mathcal{CSPP}(r)$, respectively. The formula for σ^2 is already shown in Wilson (2001).

Lemma 2.3 We have

1.
$$\sigma^2 = \frac{1}{12} rst(r+s+t)$$
,

2.
$$\sigma_{spp}^2 = \frac{1}{3}tr^3 + \frac{1}{6}t^2r^2 - \frac{1}{12}t^2r + \frac{1}{6}tr^2 - \frac{1}{3}tr$$
 and

3.
$$\sigma_{cspp}^2 = \frac{3}{4}r^4 - \frac{1}{2}r^2$$
.

Proof: Recall that the variance of a random variable Y can be obtained as the $\mathbb{V}(Y) = -\log\left(P(e^{Ix})\right)''\Big|_{x=0}$, where P(q) is the probability generating function of Y. In the $\mathcal{PP}(r,s,t)$ case we apply this to P(q) as in (1.4) and obtain

$$-\log(P(e^{Ix}))''\Big|_{x=0} = -\sum_{i=1}^{r} \sum_{j=1}^{s} \sum_{k=1}^{t} \log\left(\frac{(\alpha_{ijk} - 1)(1 - e^{\alpha_{ijk}Ix})}{\alpha_{ijk}(1 - e^{(\alpha_{ijk} - 1)Ix})}\right)''\Big|_{x=0},$$
(2.7)

where $\alpha_{ijk}=i+j+k-1$. According to Lemma 2.2, each summand on the rhs is equal to

$$2!H_{2,1}(\alpha_{ijk}) = -\frac{1}{6}\alpha_{ijk} + \frac{1}{12},\tag{2.8}$$

and a straightforward calculation now yields

$$\sum_{i=1}^{r} \sum_{j=1}^{s} \sum_{k=1}^{t} \left(\frac{1}{6} (i+j+k-1) - \frac{1}{12} \right) = \frac{1}{12} \left(r^2 s t + r s^2 t + r s t^2 \right) = \frac{1}{12} r s t (r+s+t). \tag{2.9}$$

The other variances are calculated in an analogous way with suitable choices of α and c.

we can investigate concentration properties of the families of volume random variables, i.e. we study the quotients of standard deviation and mean when the box gets large. According to Lemma 2.1 and Lemma 2.3 these quotients tend to zero if at least two side lengths of the bounding box tend to infinity. The latter is in particular satisfied when $r \to \infty$ in the $\mathcal{SPP}(r,t)$ and $\mathcal{CSPP}(r)$ case. In the general case of $\mathcal{PP}(r,s,t)$ we have

$$\left(\frac{\sigma}{\mu}\right)^2 = \frac{1}{3}\left(\frac{1}{st} + \frac{1}{rt} + \frac{1}{rs}\right). \tag{2.10}$$

If say r and s are unbounded, the right hand side of (2.10) tends to zero for $r, s \to \infty$ and the family is concentrated to the mean, i.e. for every $\delta > 0$ we have

$$\mathbb{P}\left(1 - \delta \le \frac{X}{\mathbb{E}(X)} \le 1 + \delta\right) \longrightarrow 1 \qquad (r, s \to \infty)$$

On the other hand, if two coordinates are fixed, say r, s, then the quotient is bounded away from zero by $\frac{1}{3rs}$ for $t \to \infty$. These families are not concentrated. Similar considerations hold in the case of $\mathcal{SPP}(r,t)$ when r is fixed and $t \to \infty$. In the next section we investigate the limit laws in these two cases.

3 Limit laws

We first give a result for the concentrated families.

Proposition 3.1 If at least two of r, s, t tend to infinity, the family

$$\left(Y := \frac{X - \mu}{\sigma}\right)$$

of normalised volume random variables on $\mathcal{PP}(r,s,t)$ converges in distribution to a standard normal distributed random variable. The same statement holds for the volume variables on $\mathcal{SPP}(r,t)$ if at least $r \to \infty$ and on $\mathcal{CSPP}(r)$ if $r \to \infty$.

Proof: For $\mathcal{PP}(r,s,t)$. Let $r,s\to\infty$ and t vary arbitrarily. The characteristic function of Y is $\phi(x):=e^{-\mu Ix/\sigma}P\left(e^{Ix/\sigma}\right)$, with P as in eq. (1.4). We prove that $\log\left(\phi(x)\right)\to-\frac{x^2}{2}$ for $x\in\mathbb{R}$ if $r,s\to\infty$. Then by Levy's continuity theorem (cf. Feller (1971)) the assertion follows.

By Lemma 2.1 the random variables Y and -Y have the same distribution. So the characteristic function is real valued and the Taylor coefficients of ϕ and $\log(\phi)$ of odd order vanish. By Lemma 2.2 we have the Taylor expansion

$$\log\left(\phi(x)\right) = -\frac{x^2}{2} + \sum_{N>2} \sum_{i=1}^{r} \sum_{j=1}^{s} \sum_{k=1}^{t} H_{2N,1}(i+j+k-1) \left(\frac{xI}{\sigma}\right)^{2N}$$
(3.1)

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According to the estimate (2.2) we can bound the modulus of the $2N^{th}$ summand, $N \geq 2$, by

$$D\sum_{i=1}^{r}\sum_{j=1}^{s}\sum_{k=1}^{t}2^{2N}(i+j+k-1)^{2N-1}\left(\frac{|x|}{\sigma}\right)^{2N} \le 4^{N}D|x|^{2N}\frac{rst(r+s+t)^{2N-1}}{\sigma^{2N}}.$$
 (3.2)

Plugging the explicit expression for σ^2 of Lemma 2.3 into the rhs of (3.2) we obtain the estimate for the $2N^{th}$ summand, $N \ge 2$, in the expansion (3.1)

$$\left| \sum_{i=1}^{r} \sum_{j=1}^{s} \sum_{k=1}^{t} H_{2N,1}(i+j+k-1) \left(\frac{xI}{\sigma} \right)^{2N} \right| \le D|x|^{2N} \cdot 48^{N} \left(\frac{1}{rs} + \frac{1}{rt} + \frac{1}{st} \right)^{N-1}. \tag{3.3}$$

Summing in (3.3) over $N \ge 2$ then yields the estimate

$$\left|\log(\phi(x)) + x^2/2\right| \le 48D|x|^2 \frac{48|x|^2 \left(\frac{1}{rs} + \frac{1}{rt} + \frac{1}{st}\right)}{1 - 48|x|^2 \left(\frac{1}{rs} + \frac{1}{rt} + \frac{1}{st}\right)}.$$
(3.4)

The rhs tends to zero for any fixed real x if r and s tend to infinity. This proves the assertion for the $\mathcal{PP}(r,s,t)$ case. The other cases are shown with similar estimates.

we consider the non-concentrated case. If r=s=1 we have a single column of unit cubes of height t. The volume clearly is uniformly distributed. So, if rs>1 the volume random variable can be viewed as sum of dependent random variables with values in $\{0,\ldots,t\}$. Now the easiest guess is that for large t the dependence vanishes and the volume of a single column is uniformly distributed. This guess is the right one as the following proposition shows.

Proposition 3.2 If r and s are fixed and t tends to infinity, the family

$$\left(Z_t := \frac{X_{rst}}{t}\right)$$

of rescaled random variables converges in distribution to the (rs)-fold convolution of the uniform distribution on [0,1]. In the SPP(r,t) case the so rescaled sequence converges in distribution to the convolution of r factors of the uniform distribution on [0,1] and r(r-1)/2 factors of the uniform distribution in [0,2].

Proof: For $\mathcal{PP}(r,s,t)$. We show that the characteristic function of Z_t converges pointwise to $\left(\frac{e^{Ix}-1}{Ix}\right)^{rs}$, which is the characteristic function of the rs-fold convolution of the uniform distribution on [0,1]. The Fourier transform of Z_t is $P(e^{Ix/t})$, with P as in (1.4). We expand the product running from 1 to t in $P(e^{Ix/t})$. All but the last term in the numerator and the first term in the denominator cancel out:

$$P(e^{Ix/t}) = \prod_{i=1}^{r} \prod_{j=1}^{s} \frac{(i+j-1)(1-e^{Ix(i+j+t-2)/t})}{(i+j+t-1)(1-e^{Ix(i+j-2)/t})}$$

The single factors are easily seen to converge to $\frac{e^{Ix}-1}{Ix}$. The $\mathcal{SPP}(r,t)$ case is worked out analogously. \Box

4 Ferrers Diagrams

A Ferrers diagram is a convex lattice polygon which contains both upper corners and the lower left corner of its smallest bounding rectangle. Here convex means, that the intersection of the polygon with any horizontal or vertical line is convex. When t=1 a boxed plane partition can be viewed as a Ferrers diagram fitting inside an $r \times s$ rectangle. Then formula (1.1) reduces to the q-binomial coefficient ${r+s\brack s}_q$. The class of Ferrers diagrams with h rows and w columns has the area generating polynomial

$$q^{h+w-1} \begin{bmatrix} h+w-2 \\ h-1 \end{bmatrix}_q \tag{4.1}$$

since at least the h+w-1 unit squares which constitute the top row and the leftmost column are contained in such a polygon. So, by Proposition 3.1 a Gaussian area limit law arises in this uniform fixed height and width ensemble, when both height and width tend to infinity.

In statistical physics (Prellberg and Owczarek (1995)) as well as in combinatorics (Bousquet-Mélou (1996)), one is also interested in area limit laws in the so-called uniform fixed-perimeter ensembles, where all Ferrers diagrams of a fixed half-perimeter are considered equally likely. For fixed $m \geq 0$ the area generating function of Ferrers diagrams of half-perimeter m+2 is obtained by summing formula (4.1) over all pairs $(h,w) \in \mathbb{N}^2$ with h+w=m+2. This can be written as

$$q^{m+1} \sum_{h=0}^{m} \begin{bmatrix} m \\ h \end{bmatrix}_{q}. \tag{4.2}$$

Observe that the index of summation in (4.2) can be interpreted as the height minus one and that the total number of such Ferrers Diagrams is 2^m . So the following probability generating function (PGF)

$$2^{-m}q^{m+1}u\sum_{h=0}^{m}u^{h}\begin{bmatrix} m\\ h \end{bmatrix}_{q}$$
 (4.3)

describes the ensemble of Ferrers diagrams of half-perimeter m+2 counted by height and area. The additional height parameter allows us to use the results for height- and width-ensembles discussed above. More precisely, we condition the area variable in the perimeter-ensemble on the height variable. In what follows we will prove that the joint distribution of the (properly rescaled) height and area variables in this ensemble converge in distribution to the two-dimensional standard normal distribution. The following proposition from Fligner (1981) (see also Sethuraman (1961)) is tailor-made for this situation.

Proposition 4.1 Let X_m, Y_m be real valued random variables and let Y_m be supported in a lattice $L_m := \{\alpha_m + k\delta_m | k \in \mathbb{Z}\}$, where $\delta_m > 0$ and $\alpha_m \in \mathbb{R}$, i.e. $\mathbb{P}(Y_m \in L_m) = 1$. Suppose Y_m satisfies a local limit law μ with a density g(y) w.r.t. the Lebesgue measure on \mathbb{R} (this implies $\delta_m \to 0$), i.e. for all $y \in \mathbb{R}$ and every sequence (y_m) with $y_m \in L_m$ and $y_m \to y$ we have $\mathbb{P}(Y_m = y_m)/\delta_m \to g(y)$. Suppose further that for μ -almost all $y \in \mathbb{R}$ the conditional distributions $\mathbb{P}(X_m \in \cdot | Y_m = y_m)$ converge weakly to a measure $\nu(\cdot,y)$ ($y_m \to y, y_m \in L_m$). Then the joint distribution of (X_m,Y_m) converges weakly to the measure ν defined by

$$\nu(A \times B) := \int_{B} \nu(A, y) \mathrm{d}\mu(y)$$

for all Borel sets $A, B \subseteq \mathbb{R}$.

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For q=1 formula (4.3) is simply $2^{-m}u(1+u)^m$, the PGF of the height random variable H_m . So H_m It satisfies the assumptions of Theorem IX.14 in Flajolet and Sedgewick (2008) and hence a Gaussian local limit law arises. The mean and standard deviation of the height are easily computed to be asymptotically equal to m/2 and $\sqrt{m}/2$, respectively. So let $Y_m=2(H_m-m/2)/\sqrt{m}$. The random variable Y_m is supported in \mathbb{Z}/\sqrt{m} . Let $y\in\mathbb{R}$ and $y_m\to y$, where each $y_m\in\mathbb{Z}/\sqrt{m}$. The area random variable A_m , conditioned on the event $\{Y_m=y_m\}$, has PGF

$$q^{m+1} \binom{m}{m/2 + y_m \sqrt{m}/2}^{-1} \begin{bmatrix} m \\ m/2 + y_m \sqrt{m}/2 \end{bmatrix}_q. \tag{4.4}$$

An application of Lemma 2.1 and Lemma 2.3 with $r=m/2+y_m\sqrt{m}/2$, s=m-r and t=1 shows that the mean and the variance of the conditioned area variables are asymptotically equal to $m^2/8$ and $m^3/48$, respectively. So according to Proposition 3.1, the rescaled area variable $X_m:=(A_m-m^2/8)/\sqrt{m^3/48}$ conditioned to $\{Y_m=y_m\}$ converges weakly to a Gaussian random variable with mean zero and variance one. Now Proposition 4.1 yields

Proposition 4.2 Denote by A_m and H_m the random variables of area and height of a Ferrers diagram in the uniform fixed perimeter ensemble described by (4.3). Let

$$X_m = \frac{A_m - m^2/8}{\sqrt{m^3/48}}, \quad Y_m = \frac{H_m - m/2}{\sqrt{m}/2}.$$

Then, as $m \to \infty$, (X_m, Y_m) converges weakly to the two-dimensional standard normal-distribution. In particular, the normalised area random variables converge to the standard normal distribution.

The last result can also be obtained via the moment method, by analysing a q-difference equation satisfied by the half-perimeter and area generating function of Ferrers diagrams.

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