# High Resolution Crystal Structures Leverage Protein Binding Affinity Predictions 

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# High Resolution Crystal Structures Leverage Protein Binding Affinity Predictions 

Simon Marillet and Pierre Boudinot and Frédéric Cazals

## RESEARCH

# High Resolution Crystal Structures Leverage Protein Binding Affinity Predictions 

Simon Marillet ${ }^{*}$ and Pierre Boudinot ${ }^{\|}$and Frédéric Cazals 围<br>Project-Team Algorithms-Biology-Structure<br>Research Report $\mathrm{n}^{\circ} 8733$ - May 2015 - 40 pages


#### Abstract

Predicting protein binding affinities from structural data has remained elusive, a difficulty owing to the variety of protein binding modes. Using the structure-affinity-benchmark (SAB, 144 cases with bound/unbound crystal structures and experimental affinity measurements), prediction has been undertaken either by fitting a model using a handfull of pre-defined variables, or by training a complex model from a large pool of parameters (typically hundreds). The former route unnecessarily restricts the model space, while the latter is prone to overfitting. We design models in a third tier, using twelve variables describing enthalpic and entropic variations upon binding, and a model selection procedure identifying the best sparse model built from a subset of these variables. Using these models, we report three main results. First, we present models yielding a marked improvement of affinity predictions. For the whole dataset, we present a model predicting $K_{d}$ within one and two orders of magnitude for $48 \%$ and $79 \%$ of cases, respectively. These statistics jump to $62 \%$ and $89 \%$ respectively, for the subset of the SAB consisting of high resolution structures. Second, we show that these performances owe to a new parameter encoding interface morphology and packing properties of interface atoms. Third, we argue that interface flexibility and prediction hardness do not correlate, and that for flexible cases, a performance matching that of the whole SAB can be achieved. Overall, our work suggests that the affinity prediction problem could be partly solved using databases of high resolution complexes whose affinity is known.


Key-words: Binding affinity prediction, protein flexibility, atomic packing, high resolution crystallography, linear regression

[^0]```
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# L'utilisation de structures cristallographiques à haute résolution améliore la prédiction d'affinité de complexes protéine - protéine 

Résumé : La prédiction d'affinité de liaison entre deux protéines à partir de données structurales reste difficile, en raison de la variété des modes d'appariement de deux protéines. À partir des données du structure-affinity-benchmark (SAB, 144 entrées comprenant les structures liées et non liées, ainsi que des mesures d'affinité expérimentales), la prédiction a été abordée soit en ajustant un modèle utilisant un petit nombre de variables prédéfinies, soit en entrainant un modèle complexe à partir d'un ensemble de paramètres de grande taille. Alors que la première stratégie restreint inutilement l'espace des paramètres, la seconde est encline au sur-apprentissage.

Ce travail propose des modèles dans un troisième registre, en utilisant douze variables décrivant les variations d'enthalpie et d'entropie intervenant lors de l'appariement, et une stratégie de sélection de modèle permettant d'identifier les meilleurs modèles parcimonieux construits à partir d'un sous-ensemble de ces variables. En utilisant ces modèles, nous rapportons ici trois résultats principaux. Premièrement, nous présentons des modèles permettant une nette amélioration des prédictions. Pour le jeux de données SAB complet, nous présentons un modèle capable de prédire le $K_{d}$ à un et deux ordres de grandeur près pour respectivement $48 \%$ et $79 \%$ des complexes. Ces statistiques passent à respectivement $62 \%$ et $89 \%$ pour les structures à haute résolution du SAB . Deuxièmement, nous expliquons que ces performances sont dues à un nouveau paramètre codant pour la morphologie de l'interface et les propriétés de packing des atomes interfaciaux. Troisièmement, nous montrons que la flexibilité de l'interface et la difficulté à prédire l'affinité ne sont pas corrélées, et que, pour les cas flexibles, nos modèles exhibent une performance égale à celle obtenue sur le SAB complet. Plus généralement, notre travail suggère que le problème de prédiction de l'affinité pourrait être en partie résolu par l'utilisation de bases de données de complexes à haute résolution dont l'affinité serait connue.

Mots-clés : Prédiction d'affinité de liaison, flexibilité des protéines, packing atomique, cristallographie à haute résolution, régression linéaire

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## 1 Introduction

### 1.1 Estimating Binding Affinities

Deciphering the dynamics of protein - protein interactions is a major challenge for functional genomics, as they determine almost all processes in living organisms. If structural models of complexes shed light on interactions at the atomic level, the formation of a complex and its stability are explained by its binding affinity (affinity for short). Estimating affinities is thus a central step while modeling biological systems, to eventually unravel the hidden complexity of the interactome [4]. But such estimates are also key to exert exogenous control on biological systems in general and in medicine in particular, where the importance of designing drugs [12, 24, therapeutic peptides [44, or high affinity antibodies [34] cannot be overstated.

Affinities measured by dissociation constants $\left(K_{d}\right)$ span 11 orders of magnitude, a range illustrating the diversity of biological processes and the various binding modes inherent to them. From an experimental standpoint, affinities can be measured by various techniques, including ITC, SPR, and titration by fluorescence, with free energy typical errors in the range $0.1-0.25$ $\mathrm{kcal} / \mathrm{mol}$ [24, 32, 14]. While such errors modestly impact $K_{d}$ (factor of 1.52 for $0.25 \mathrm{kcal} / \mathrm{mol}$ ), experimental conditions and in particular concentration, temperature, ionic strength, or pH may trigger important changes, up to $2.3 \mathrm{kcal} / \mathrm{mol}$ (factor of 48 on $K_{d}$ ) [32].

From a modeling perspective, the estimation of affinities relies on structure based modeling, to bridge the gap between 3D atomic coordinates and thermodynamics. More precisely, consider two species A and B forming a complex C . The aforementioned dissociation constant $K_{d}$ is defined by $K_{d}=[A][B] /[C]$, and the corresponding dissociation free energy $\Delta G_{d}$, in the $c^{\circ}=1 M$ standard state) satisfies

$$
\begin{equation*}
\Delta G_{d}=-R T \ln K_{d} / c^{\circ}=\Delta H-T \Delta S \tag{1}
\end{equation*}
$$

This equation shows that $\Delta G_{d}$ has two components coding the enthalpic and entropic changes upon binding, to be estimated from atomic coordinates. It also illustrates enthalpy - entropy compensation phenomenon [37, 18, which stipulates that a favorable enthalpic change upon association is accompanied by an entropic penalty. In fact, affinity enhancement may have an entropic origin, since a limited entropic loss may be associated with preconfigurations of specific binding sites [43, 12, 47].

In theory, estimating a dissociation free energy can be done using free energy calculations methods such as thermodynamic integration, umbrella sampling, or potential of mean forces [22, 13]. While in principle highly accurate, these methods are extremely demanding in terms of sampling, at the expense of high computational requirements to generate appropriate sampling. They are not suitable to large scale studies, which motivated the development of estimation methods focusing on relevant phenomena. For this reason, $\Delta G_{d}$ are generally modeled by a generic equation with terms accounting for the (variation of the) potential energy and entropy of the solute (partners and complex), as well as a solvation term 24. Modeling enthalpic changes requires approximating the internal energy of the system. This may be done using classical force fields such as CHARMM [6], AMBER [16] or GROMOS [27], which incorporate terms accounting for van der Waals interactions, electrostatic interactions, as well as bonded interactions. One may also use phenomenological functions modeling relevant phenomena [31, including interfacial properties [1], biophysical properties such as salt bridges and hydrogen bonds or cavities [21], conservation of a.a. [26], or hot spots which may account for a large fraction of the interaction energy 40. The solvation terms include both energetic and entropic terms, the latter referring to the loss of degrees of freedom incurred by water molecules surrounding the solute, usually estimated by weighted surface area terms [19]. Finally, the entropic variation of the
partners includes translational and rotational entropy, conformational entropy (e.g., rotameric states), and vibrational entropy. Coming up with reliable estimates for entropic change poses major challenges, yet such computations are indispensable, since as discussed above, affinity enhancement may have an entropic origin.

For large scale protein binding affinity studies, prediction models may be classified into two classes (see also the supplemental section 6.3). The first class consists of models using a small number of variables aiming at explaining intuitively important components of the affinity. Based on the observed correlation between the buried surface area (BSA) at the interface and binding affinity [15], a model splitting the BSA into polar and apolar components was first proposed [29]. A refinement of BSA models with a term coding the depth of interface atoms, called the Voronoi shelling order, was proposed [5], yielding improvements in particular for rigid cases. The previous models focusing on interfacial properties only, terms coding the percentage of charged and polar a.a. on the interacting surface (NIS) were introduced in [33], and their connexion with solvent dynamics investigated in [49. Finally, a model also taking into account the iRMSD, namely the root-mean-square displacement of the $\mathrm{C} \alpha$ atoms of interfacial residues between the bound and unbound states, was recently proposed 30.

The second class consists of models using machine learning techniques to select the most relevant features amidst a large pool of parameters. In [39], a binding affinity predictor based upon four machine learning classifiers is proposed. These classifiers were trained on 57 complexes (with high confidence on affinity), so as to select features amidst 200 candidates. In a nearby vein, a scoring function based model using statistical potential, for a total of 1092 parameters, was proposed in [51]. In a similar spirit, yet using a smaller set of features targeting various aspects of protein structures (H bonds, vdW interactions, cavities, iRMSD, dihedral angles, hot spots, a.a. propensities, electrostatics), various linear models were tested in [21]. Importantly, using a large number of variables helps to provide a detailed account of chemical properties of a.a. and atoms. Yet parameterizing such complex models is prone to overfitting, especially given the scarcity of data at hand, so that performances on external datasets are often limited.

Apart from the diversity of the models themselves, previous work may be distinguished using two aspects. First, different subsets of the SAB were used. Second, various statistical methodologies were used to assess the prediction performances. In particular, three types of cross validation were used, namely leave one out, four-fold, and five-fold. In doing so, the model is trained on a portion of the data, and the prediction performances are assessed by computing the correlation between the predicted affinities and the measured ones. Yet, while cross validation asymptotically yields consistent estimates [28], performances on datasets of small size should be interpreted with care [25], and checks on external datasets are called for [38]. In any case, a common finding of all these studies is that flexible cases of the SAB are the most difficult ones to deal with.

### 1.2 Contributions

In this work, we make a stride towards a better understanding of three core questions related to binding affinity predictions. The first one relates to the variables and models best suited to perform such predictions. We introduce sparse models relying on 12 variables aiming at capturing enthalpic and entropic changes upon binding. These parameters describe surface areas, packing properties, and their variations at the atomic level, and solely exploiting a partition of atoms into polar and nonpolar. Using these variables, we identify specific models for subsets of the SAB considered by previous studies, whose performances match or outperform those previously published, in particular for flexible and high resolution cases. Each specific model is also challenged on its non-specific dataset, to highlight the relevance of its variable in handling features specific from these datasets. In particular, this analysis singles out a novel parameter,
coding the morphology and the packing properties of the interface, namely properties reminiscent of enthalpy and entropy.

The second question relates to a key difficulty in predicting affinities, namely flexibility. In previous work, flexible cases have been described as the most challenging ones. Using our models, we show that flexibility and prediction hardness do not correlate, and that for flexible cases, a performance almost matching that of the whole SAB can be achieved.

The third one pertains to the quality of predictions. For the whole dataset, we present a model predicting $K_{d}$ within one and two orders of magnitude for $48 \%$ and $79 \%$ of cases, respectively. These statistics jump to $62 \%$ and $89 \%$ respectively, for the subset of the SAB consisting of high resolution structures, a marked improvement over previous work, also stressing the dependence of energies on atomic details.

## 2 Estimating Affinities: Datasets and Parameters

### 2.1 Datasets from the Structure Affinity Benchmark

We use the structure-affinity benchmark [32] (SAB, denoted SAB-A), providing 144 cases with crystal structures for the partners and the complex, as well as an experimentally measured dissociation free energy $\Delta G_{d}^{e x p_{i}}$. Following previous work, we extract seven datasets using a flexibility criterion, and one dataset of high resolution structures. These datasets are (supplemental Fig. 33: $\mathrm{SAB}-\mathrm{R}_{1.0}, \mathrm{SAB}-\mathrm{R}_{1.1}$ and $\mathrm{SAB}-\mathrm{R}_{1.5}$, three datasets consisting of rather rigid cases; $\mathrm{SAB}-\mathrm{F}_{1}$ and SAB- $\mathrm{F}_{1.5}$, two datasets consisting of rather flexible cases; SAB-I, a dataset consisting of intermediate cases; and SAB-A-HR, 37 high resolution entries (resolution $\leq 2.5$ ) 21]. We also ruled out two cases with more than $20 \%$ atoms missing in the bound versus unbound forms, and three cases with an upper bound on the affinity rather than a proper value.

### 2.2 Parameters involved in Affinity Prediction Models

In the sequel, having presented key geometric constructions associated with solvent accessible models of the partners and of the complex, we define parameters meant to capture information on enthalpic and entropic contributions associated with complex formation (Fig. 11 and Table 1).

### 2.2.1 Key Geometric Constructions

Surface areas. The solvent accessible surface area (SASA for short) of a solvent accessible model is the sum of the surface areas exposed by the individual atoms. Upon complex formation, the buried surface area (BSA) is the surface area of the partners buried at the interface, namely the SASA lost by the individual atoms. This quantity has long been known as the simplest and most descriptive parameter of specific protein interfaces [1].

Voronoi interfaces and their shelling order (SO). In describing a protein - protein interface, various parameters are of interest beyond the mere list of atoms, namely its shape (e.g. elongated vs isotropic), its partition into a core and a rim, its curvature, or its number of patches. A parameter free Voronoi interface model encapsulating all these parameters into a single construction, the $\alpha$-complex derived from the Voronoi (power) diagram of the atoms, has been proposed [10, 35]. In a nutshell, define the restriction of an atom as the intersection between its ball in the solvent accessible model and its cell in the Voronoi diagram. The Voronoi interface identifies pairs of neighboring restrictions, such that each pair involves either two different
partners or a partner and the interfacial solvent. The atoms found in at least one such pair are denoted $\mathcal{I}$ and their complement $\mathcal{I}^{C}$. This Voronoi-based model was instrumental to show that the interface may involve atoms which do not lose solvent accessibility, and also to stress the role of water mediated contacts [10. We note in passing that the exposed atoms in the set $\mathcal{I}^{C}$ form the non interacting surface (NIS) [33.

Consider the BSA, and more specifically the atoms of one partner contributing to the BSA. The exposed surface of the atoms contributing to the BSA define a binding patch (patch for short) [5]. The shelling order (SO) of an atom from a patch is its least distance, counted in integer steps, to the nearest atom from the NIS. That is, the atoms on the border of the patch have a SO of 1 and the remaining ones have a $S O>1$ (Fig. 1(B)). Thus, the SO generalizes core-rim models [31], since the rim corresponds to $S O=1$, and the core to $S O>1$.

Atomic packing properties. Early models to assess atomic packing properties resorted to the volume of Voronoi cells [23], preferably using the power diagram of the atoms instead of the Euclidean Voronoi diagram [3], since different atomic radii are accommodated. However, the Voronoi cell of an atom located on the convex hull of the protein (or complex) is unbounded. To avoid boundary effects, we focus in the sequel on the aforementioned atomic restrictions, whose volume can be computed accurately [9]. That is, denoting volume_bound (a) (resp. volume_unbound $(a))$ the volume of the Voronoi restriction of an atom $a$ in the bound form (resp. unbound form), the difference between these quantities defines the volume variation of this atom (Eq. 77).

### 2.2.2 Partners: Enthalpic Contributions

Local interactions. The BSA alone does not account for the interface geometry, as the same surface area may be obtained for by morphologies as diverse as a perfectly isotropic patch, or a long and skinny patch, letting alone curvature. The obliviousness to interface morphology is intuitively detrimental, since morphology relates to the cooperativity of phenomena inherent to non-bonded interactions. To take into account such morphological features, a weighted average of atomic shelling orders, called the internal path length (IPL) was defined from the shelling order [5] ${ }^{1}$. The IPL has been shown to improve the analysis of correlations between interface morphology against conserved residues and interfacial solvent dynamics [5].

In terms of binding energies, a limitation of IPL is that the SO of an atom does not not account for the atomic environment of this atom-that is two atoms with identical SO may be located in a dense an loose environments respectively. This is detrimental since a dense packing is likely to favor local interactions, in particular van der Waals interactions. Since a packed interface is more likely to result in a high affinity, the shelling order is weighted by the inverse of the volume, yielding the inverse volume-weighted internal path length (Eq. (8)).

### 2.2.3 Partners: Entropic Contributions

Assessing entropic variations requires taking several components into account, in particular configurational entropy and vibrational entropy. Large conformational changes yielding structured elements correspond to entropic penalties, and can be assessed using the interface root mean square deviation (iRMSD). In the sequel, we refine this measure using atomic packing properties.

[^1]Packing properties. A closely packed environment yields favorable interactions by increasing the number of neighbors. But it also entails an entropic penalty for that atom, illustrating the classical enthalpy - entropy compensation, which holds in particular for biological systems involving weak interactions [18, 14]. We therefore use our atomic volumes and their variations upon binding (Eq. (7)) to model both the interaction energy and the entropic changes upon binding.

To model entropic changes, we resort to volume variations. We do so by considering four categories of atoms. For interface atoms, we define two groups, those found on the $\operatorname{rim}(\mathcal{I}, S O=$ $1)$, retaining solvent accessibility, and the reaming ones ( $\mathcal{I}, S O>1$ ). Likewise, for the set of non interface atoms, we distinguish between those retaining solvent accessibility ( $\mathcal{I}^{C}$ and SASA $>0$ in the complex), and those which do not ( $\mathcal{I}^{C}$ and SASA $=0$ in the complex). Adding up volume variations for these four categories of atoms yields the following four Sum of Volumes Differences ( $S V D$ ) parameters, namely SVD_SO1 $(\mathcal{I}, S O(a)=1$; Eq. (9p), SVD_SOGT1 $(\mathcal{I}, S O(a)>1$; Eq. 10) $)$, SVD_NI_B $\left(\mathcal{I}^{C}, \operatorname{SASA}(a)=0 ;\right.$ Eq. 11) $), \operatorname{SVD\_ NI\_ E~}\left(\mathcal{I}^{C}, \operatorname{SASA}(a)>0\right.$; Eq. 12 ).

### 2.2.4 Solvent Interactions and Electrostatics

The interaction between a protein molecule and water molecules is complex. In particular, the exposition to the solvent of non polar groups hinders the ability of water molecules to engage into hydrogen bonding, yielding an entropic loss for such water molecules. To account for these effects, we use the fractions of charged and polar a.a. on the non interacting surface [33], respectively denoted NIS ${ }^{\text {polar }}$ (Eq. 13) and NIS ${ }^{\text {charged }}$ (Eq. 14). We also use the variation of these quantities to account for conformational changes upon binding, yielding the quantities $\Delta$ NIS $^{\text {polar }}$ (Eq. 15p) and $\Delta$ NIS $^{\text {charged }}$ (Eq. 16).

To challenge a.a. terms with their atomic counterparts and see which ones are best suited to perform affinity predictions, we also included the atomic solvation energy from Eisenberg et al [20], describing the free energies of transfer from 1-octanol to water per surface unit $\left(\AA^{2}\right)$. The corresponding variable, ATOM_SOLV, is a weighted sum of atomic solvent accessible surface areas (Eq. 17), and may be seen as the atomic-scale counterparts of NIS ${ }^{\text {charged }}$ and NIS ${ }^{\text {polar }}$.

Finally, we include an intermediate-grained description of the non-interacting surface which consists in the atomic-wise polar area of the complex. The corresponding term, POLAR_SASA (Eq. 18)), is also a weighed sum of exposed areas.

### 2.3 Parameters Computation

To compute the atoms at interface along with their shelling order, packing and volume, we use the application sbl-vorshell-bp-ABW-atomic.exe from the Structural Biology Library (SBL) [7, see http://sbl.inria.fr. Contacts mediated by water molecules are included because crystallographic water molecules are biologically relevant [45].

To compute the solvent accessible atoms of the molecules, we use the application sbl-vorlume-pdb.exe software [9], also from the SBL [7]. In that case, water molecules are not considered since they contribute to the protein surface solvation as much the bulk solvent.

### 2.4 Statistical Methodology

Denote $\mathcal{V}$ the pool of twelve variables specified by Eq. (8) to (18) (Table 1), plus the iRMSD defined in the SAB . We now explain how to assess models predicting $\Delta G_{d}^{e x p_{i}}$ using variables from $\mathcal{V}$.

### 2.4.1 The Dataset Level

Linear models. Consider a dataset $\mathcal{D}$ from the SAB , and a set of parameters $V \subset \mathcal{V}$. The set $V$ can be used to fit a linear model denoted $M_{V}(\mathcal{D})$ aiming at predicting the $\Delta G_{d}^{\text {exp }_{i}}$ from the dataset $\mathcal{D}$.

To assess predictive performances, we perform $N_{X V}(=10000)$ repetitions of 5 -fold cross validation. Specifically, for one fold, the model is trained on $4 / 5$ of the dataset, and predictions are run on the remaining $1 / 5$ of complexes. Processing the five folds yields one repetition of the cross validation procedure, resulting in one prediction for the $\Delta G_{d}^{\exp _{i}}$ of each complex. We denote $\hat{g}_{i j}$ the prediction of the $\Delta G_{d}^{e x p_{i}}$ of the $i$ th complex during the $j$ th repetition, for $j=1, \ldots, N_{X V}$.

Considering one cross-validation repetition, we define the correlation $\operatorname{Corr}_{j}$ associated with this repetition as the correlation between the experimental values $\left\{\Delta G_{d}^{e x p_{i}}\right\}$ and the predictions $\left\{\hat{g}_{i j}\right\}$. An overall assessment of the model using the $N_{X V}$ repetitions is obtained by the following median of correlations (see also the supplemental section 6.4.3):

$$
\begin{equation*}
C[V, \mathcal{D}]=\operatorname{median}_{j} \operatorname{Corr}_{j} \tag{2}
\end{equation*}
$$

For a complex, the prediction $\hat{g}_{i}$ is defined as the median across repetitions i.e.

$$
\begin{equation*}
\hat{g}_{i}=\operatorname{median}_{j} \hat{g}_{i j} \tag{3}
\end{equation*}
$$

Likewise, the median prediction error is defined by

$$
\begin{equation*}
e_{i} \equiv \Delta \Delta G_{i}[V, \mathcal{D}]=\operatorname{median}_{j}\left(\Delta G_{d}^{\exp p_{i}}-\hat{g}_{i j}\right) \tag{4}
\end{equation*}
$$

and the median absolute prediction error by:

$$
\begin{equation*}
e_{i}^{a b s} \equiv \Delta \Delta G_{i}^{a b s}[V, \mathcal{D}]=\operatorname{median}_{j}\left(\left|\Delta G_{d}^{e x p_{i}}-\hat{g}_{i j}\right|\right) \tag{5}
\end{equation*}
$$

Using this latter value, we define the prediction ratio $p_{\delta}^{\Delta \Delta G}$ of the model for the dataset as the percentage of cases such that the dissociation free energy is off by a specified amount $\delta$ :

$$
\begin{equation*}
p_{\delta}^{\Delta \Delta G}=\% \text { cases in } \mathcal{D} \text { such that } \Delta \Delta G_{i}^{a b s}[V, \mathcal{D}] \leq \delta \tag{6}
\end{equation*}
$$

In particular, setting $\delta$ to $1.4,2.8$ and $4.2 \mathrm{kcal} / \mathrm{mol}$ in the previous equation yields cases whose $K_{d}$ is approximated within one, two and three orders of magnitude respectively.

Finally, a permutation test yields a p-value for each model 41. In a nutshell, the rationale consists of generating randomized datasets by shuffling the $\Delta G_{d}^{\text {exp }_{i}}$ values of the dataset. Then, one computes the correlation of Eq. (22) for each such dataset, from which the p-value is inferred (supplement, Algorithm 1 ).

Variable selection and model evaluation. To define parsimonious templates from the set $\mathcal{V}$ for a given dataset, we select subsets of $\mathcal{V}$ involving up to at most five variables-an upper bound dictated by the fact that beyond five variable, the performance of the best model starts to decrease (supplemental Fig. 5). This defines a pool of models $\mathcal{M}=\left\{M_{1}, \ldots, M_{1585}\right\}{ }^{2}$. As explained, each set $V$ undergoes five fold cross validation (supplement, Algorithm 2).

[^2]Model selection: specific models of a dataset $\mathcal{D}$. We wish to single out the best models from $\mathcal{M}$ to fit the $\Delta G_{d}^{\text {exp }_{i}}$ of complexes in $\mathcal{D}$. Here, we define best as models with highest $C[V, \mathcal{D}]$. To see how, observe that compare two models $M_{U}$ and $M_{V}$, a univariate two-sample test suffices to check whether the two sets of correlations (one per model) obtained for the $N_{X V}$ repetitions come from the same distribution (the null hypothesis H 0 ), or whether one dominates the other. In an analogous spirit and since we are handling a pool of models $\mathcal{M}$, we wish to identify within $\mathcal{M}$ a subset of models whose distribution cannot be distinguished from the best model. To this end, we decompose the models as $\mathcal{M}=\mathcal{M}_{1} \cup \mathcal{M}_{2}$ such that (i) the best model is in $\mathcal{M}_{1}$, (ii) in comparing two models from $\mathcal{M}_{1}$, one does not reject H0, and (iii) in comparing one model from $\mathcal{M}_{1}$ against one model from $\mathcal{M}_{2}$, one rejects H 0 . The models in $\mathcal{M}_{1}$ are called the specific models for the dataset $\mathcal{D}$. The corresponding procedure is based on the Kruskal-Wallis test (supplemental, Algorithm 3). The p-value threshold was set to $\alpha=0.01$.

### 2.4.2 The Structure Affinity Benchmark Level

Remarkably, the previous procedure yields a single specific model for each dataset (details below). We also challenge the set of variables associated with this model on the remaining seven datasets, applying for each the cross-validation procedure explained above.

We also used the eight datasets to define the best overall model. To this end, we sorted the models by correlation value $C[V, \mathcal{D}]$ for each dataset, and took the model with lowest median rank among all datasets.

## 3 Results

### 3.1 Specific Models yield Enhanced Correlations. . .

Model selection yielded a single best specific model for each dataset (Table 2). These models are all statistically significant with a p-value smaller than 0.01 , except for the one associated with the dataset SAB-I, therefore omitted from subsequent analysis.

In terms of correlations between estimates and $\Delta G_{d}^{\text {exp }_{i}}$ (Supplemental table 4, 5 -fold cross validation), our specific models outperform those previous previously in $5 / 8$ cases. For two of the three remaining cases, the top correlation is provided by the complex model from [39], which we estimated to use 94 variables (see supplemental Section 6.3.1). For the remaining one, [21] provides the best results with a seven variables model. Unfortunately, the corresponding variables are not specified.

In terms of correlation values themselves, three facts emerge. First, the model specific of the high resolution model dataset yields a remarkable correlation of 0.77 . Second, for flexible datasets, satisfactory performances are observed, which is unexpected since such cases are generally considered as the most challenging ones for affinity predictions. In particular, for flexible cases characterized by an interface iRMSD larger than $1.5 \AA$, a correlation of 0.46 is obtained, a value comparable to that of the whole dataset, namely 0.48 . Finally, the best overall model, when challenged by individual datasets, shows performances comparable to those of their specific models with maximum drop in correlation of 0.6 . This is a clear assessment of its robustness.

## 3.2 . . . and Improved Predictions on a per Complex Basis

The correlation between predictions and $\Delta G_{d}^{\exp _{i}}$ provides a global performance assessment of a model for a dataset. To gain insights at the individual complex level, we use the individual predictions $\hat{g}_{i j}$. Using these individual predictions, we compute the prediction ratio (Eq. 6) for
$\delta=1.4,2.8$ and $4.2 \mathrm{kcal} / \mathrm{mol}$, respectively, yielding the fraction of cases for which $K_{d}$ is predicted within one, two and three orders of magnitude. Three striking facts emerge from Table 3 .

First, the merits of our specific models as well as those of the best overall model clearly emerge. As a quantitative measure, we collect the min and max prediction ratios for the aforementioned three values of $\delta$, yielding a three pairs min-max percentages. For our best overall model, one gets $44-57 \%, 74-86 \%$ and $91-95 \%$ within one, two and three orders of magnitude. In contrast, the intervals for [33] are $46-51 \%, 68-83 \%$ and $85-95 \%$, and those for [30] are $22-44 \%, 57-73 \%$ and $85-93 \%$. Collecting now the min and max prediction ratios of the specific models on their specific datasets, one gets $46-62 \%, 78-89 \%, 85-97 \%$. Thus, for the whole SAB, both the specific models and the best overall model yield improved performances.

Second, the prediction ratios of the model specific of the high resolution dataset turn out to be $62 \%, 89 \%$ and $97 \%$ within one, two and three orders of magnitude, an outstanding performance.

Third, concerning the flexible datasets, considered as the most challenging ones in previous studies, model 7 (dataset $\mathrm{SAB}-\mathrm{F}_{1}$ ) and model 8 (dataset $\mathrm{SAB}-\mathrm{F}_{1.5}$ ) reach performances comparable to those obtained on the whole SAB , namely $p_{1.4}^{\Delta \Delta G}$ values of $50 \%$ and $50 \%$ respectively, instead of $47 \%$. This shows that the difficulty of predicting binding affinity for flexible interfaces can be circumvented by the right choice of variables. This observation is also backed up by the lack of correlation between the interface flexibility and the prediction error (Fig. 22).

Specific cases. Inspecting extreme cases is informative (named cases, Fig. 20. The individual predictions $\hat{g}_{i}$ from Eq. 3 are provided in the supplemental Table 9 .

On the one hand, the affinity of three complexes with subpicomolar affinity (1EMV, 1BRS, 1DFJ) is significantly under-estimated (Fig. 22. These three complexes involve an inhibitor taking the place of a cognate nucleic acid. Such complexes typically involve strong electrostatic interactions 42, 30, which are overlooked by our models. It could also be the case that such complexes manage to limit the entropic loss upon binding, possibly by transferring the dynamics of interfacial atoms to the protein's non interacting atoms.

On the other hand, predictions are excellent for several flexible cases, in particular 1F6M and 2I9B (Fig. 24. Complex 1F6M consists of a thioredoxin reductase in flavin-reducing conformation with its substrate. The reductase switches between bound and unbound conformations using a hinge-like motion. Complex 2I9B consists of a urokinase plasminogen activator receptor and its associated ligand. There is a global conformational change of the receptor upon binding (RMSD $2.657 \AA$ ) but no obvious hinge motion. It is the only complex with an iRMSD greater than $3 \AA$ and a prediction error below $1.4 \mathrm{kcal} / \mathrm{mol}$.

Classically for complexes with large interfaces, affinity predictions based on the BSA often result in overestimates. Beyond a certain interface size, the affinity no longer increases as much with the interface size, a behavior which could be related to a non-uniform atomic packing at the interface 30. However, the packing distribution of large interfaces matches that of the remaining ones (supplemental Fig. 6), and no correlation is observed between the quality of individual predictions and interface size (supplemental Fig. 77. Thus, packing heterogeneity may not account for mild to poor prediction performances in that context.

Validation on external datasets. Cross validation results obtained on on datasets of small size should be interpreted with care [25], and checks on external datasets are a must [38. We therefore ran predictions on an external dataset (supplemental Table 7), from which two striking facts emerge.

The correlations observed compare to those obtained with cross-validation, with a maximum drop of 0.11 excluding models 2 and 8 . For the latter two models, the drop reaches 0.33 and
0.25 respectively, a fact likely related to the small size of their training datasets. Second, the proportions $p_{1.4}^{\Delta \Delta G}, p_{2.8}^{\Delta \Delta G}$ and $p_{4.2}^{\Delta \Delta G}$ are smaller than their cross-validated counterparts, by a factor $1.4\left(p_{4.2}^{\Delta \Delta G}\right.$, model 8) to $9.5\left(p_{1.4}^{\Delta \Delta G}\right.$, model 7). Therefore, on this external dataset, despite being good predictors on a global level, as assessed by the correlation coefficient, our models do not always perform robustly on a per complex basis.

### 3.3 Accounting for Interface Morphology and Packing Boosts Performances

The performances of our models owe to the new variables introduced in this study (Table 2). The variable selected most often is IVW-IPL ( $6 / 8$ cases), stressing the role of the interface size (in terms of buried surface area), but also of atomic packing properties. The second variable selected most often is NIS ${ }^{\text {charged }}$ ( $5 / 8$ cases), highlighting the role of solvent interactions [30]. Two other variables selected for $3 / 8$ datasets, respectively represent volume variation at the interface rim (SVD_SO1), and solvation properties of the complex at the atomic scale (ATOM_SOLV). Interestingly, inspecting these four variables reveals a correlation between IVW-IPL and SVD_SOGT1 (supplemental Table 6), so that these variables might be used interchangeably. The same observation holds for NIS ${ }^{\text {charged }}$ and NIS ${ }^{\text {polar }}$.

Of particular interest in this context is our best overall model. This model uses variables IVW-IPL, SVD_SO1 and NIS ${ }^{\text {charged }}$ and is therefore equivalent to model 4 Not surprisingly, these variables form the top three of variables selected most often by the specific models (Table 2). Its performances, are similar to those of specific models on their own datasets (Table 3). Interestingly, it is a better predictor of flexible complexes than model 1. Finally, its results on external datasets (Tables 7 and 8 ) show that it is outperformed by specific models for four datasets, and outperforms them for two (not considering model 6).

## 4 Discussion and Outlook

This work develops sparse binding affinity predictions models, which shed new light on the hardness of affinity prediction, and improve prediction quality using variables coding enthalpic and entropic variations upon binding.

On the hardness of affinity predictions. Flexible datasets have been reported as the most challenging ones in previous studies. However, as shown here, the segregation of flexible versus rigid appears partially founded, with some easy to predict flexible complexes, and some hard to predict rigid cases. This observation is not completely surprising, since conformational changes alone tell little, in particular, on entropic changes upon binding. It also hints at the possibility of improving the quality of predictions for cases with small conformational changes upon binding, as molecular dynamics simulations in the intermediate time range may provide good estimates for the entropic penalties in those cases.

On the quality of predictions. A key achievement of this study is the quality of predictions, assessed in terms of $\Delta \Delta G_{d}$, or equivalently accuracy on $K_{d}$. To summarize, two values may be put forward, namely the fraction of cases for which $K_{d}$ is predicted within one and two orders of magnitude. For the best overall model, these fractions, corresponding to the whole SAB, are $48 \%$ and $79 \%$, respectively. For the model specific of high resolution complexes, these fractions are $62 \%$ and $89 \%$. These numbers clearly advance the state-of-the-art, and call for two comments.

First, our models do not take into account the pH , whose change by two units may alter $K_{d}$ by a factor ten or more. Given this specificity, they second the goal set in 30, namely that of approximating $K_{d}$ within two orders of magnitude.

Second, the high performance obtained for high resolution structures recalls the short range nature of selected forces-van der Waals interactions in particular, and stresses the dependence of energies on atomic details. From a quantitative standpoint, from Cruickshank's formula, the typical precision on atomic coordinates at a resolution of say $2.5 \AA$ lies in the range $[0.2,0.4]$ $\AA$ [17, 2]. At such a resolution, which is the worst used in the high resolution dataset (supp. Fig. 4), the inter-atomic distance between non covalently bonded atoms located nearby in 3D space [10] may already be spoiled by a factor circa $\sim 1 / 5$. The situation deteriorates with the resolution, with a potential significant impact on the atomic scale parameters listed in Table 1. Therefore, the incidence of resolution on prediction performance should not come as a surprise. In a more general perspective, this observation is reminiscent of the role of molecular shape in determining motions [36], and also on the importance of packing properties in protein structure 11.

One generic model versus several specific models. The diversity of the specific models may be seen as a weakness or a strength. For the former viewpoint, one may argue that thermodynamics call for a unified model. For the latter one, given the intrinsic complexity of the problem (recall that the binding affinity is inherently coupled to a thermodynamic equilibrium), and the paucity of the dataset, it is clearly beneficial to exploit specific features of datasets. Moreover, specific models are of practical interest since to predict the affinity of a complex performing a specific biological function, one may use a dataset of complexes related to that function. Further arguments to choose between these two interpretations will likely emerge upon populating the structure affinity benchmark.

On key parameters. Our models preferably select parameter IVW-IPL, and then NIS ${ }^{\text {charged }}$. The former, introduced in this work, combines the overall shape of the interface and involves atomic packing properties. It is reminiscent of cooperativity phenomena observed for weak interactions [5]. The latter, NIS ${ }^{\text {charged }}$, encodes the electrostatic properties of the non interacting surface, as recently investigated [49. The following top scorers represent volume variation at the interface rim (SVD_SO1), and solvation properties of the complex at the atomic scale (ATOM_SOLV). Among these four variables, two describe surface properties at different scales (atomic for ATOM_SOLV, and at residue level for NIS ${ }^{\text {charged }}$ ), and two encode interface properties, one static for the whole interface (IVW-IPL), and one dynamic for the outer layer of the interface (SVD_SO1).

Remarkably, these parameters are simple ones, derived from the Voronoi diagram of the solvent accessible models of the three structures involved (two for the partners, one for the complex). From a computational standpoint, processing a structure of say up to 10,000 atoms takes a handfull of seconds on a desktop computer [9].

Outlook. Estimating binding affinities is a central endeavor to understand protein - protein interactions. Strikingly, the models and variables presented here yield a prediction accuracy of 2.8 $\mathrm{kcal} / \mathrm{mol}$ per complex in $79 \%$ of cases for the whole SAB , and in $89 \%$ of cases for high resolution complexes. This represent a significant progress over previous methods. Since our methods inherently exploit static properties of crystal structures, improving results even further calls for developments in two directions. On the one hand, unveiling dynamical properties of the partners and the associated complex, by sampling and modeling the associated (potential, free) energy landscapes will undoubtedly yield enhanced predictions [50. Along the way, a central problem
to be addressed is that of the potential energy model best suited, since, as shown in this work, coarse grain descriptors can match or surpass the performances of detailed chemical ones. In this respect, our ability to accurately sample [50, 13] and compare [8 sampled energy landscapes should prove critical. On the other hand, a weakness shared by our method and previous ones is the absence of terms taking into account the pH and the ionic strength - a limitation actually accounting for the poor performances observed on complexes involving significant electrostatic interactions. For such cases, incorporating terms accounting for counter-ion condensation seems critical, yet, controlling the enthalpy - entropy balance within such models remain challenging 42, 46].

The affinity prediction problem is also of special interest from the machine learning perspective. Affinity prediction is indeed modeled here a particular instance of a problem known as regression [28]. In this setting, the data is assumed to be generated by a process and applied some random noise. The most important attribute of regressors is their consistency, i.e. their ability to converge toward the true model given data accounting for the whole space. However, for a regressor to achieve consistency, the data must satisfy some assumptions. For instance it should be well distributed over the space of possible data points. In our case, this means that the dataset should evenly represent all possible protein-protein complexes. This is most probably not the case for the SAB. The availability of larger datasets will also ease the model selection problem, undertaken by complete enumeration over the parameter set in this work. In principle, sparse least square models can be obtained using regularization techniques [48]. However, the inherent randomization used by cross-validation makes model selection unstable for small datasets, making such methods hard to use at this stage. For these reasons, sparse specific models using with relevant variables, as developed in this work, appear as a privileged solution to estimate binding affinities.

## 5 Artwork

Figure 1 Structural parameters used in this work. (A) Labeling the atoms, illustration on a fictitious 2D complex. The binding patch on each partner consists of one layer of atoms ( $\mathcal{I}$, colored solid balls), as identified by a Voronoi interface model [10, 35]. The non interface atoms $\left(\mathcal{I}^{c}\right)$ are split into those which retain solvent accessibility (SASA $>0$, dashed balls), and those which do not (SASA $=0$, dotted balls) $(\mathbf{B})$ Each interface atom is assigned an integer, its shelling order, equal to the smallest number of atoms traveled to reach an exposed non interface atom, i.e. an atom belonging to $\mathcal{I}^{c}$ and with SASA $>0$ (in grey) [5. (C,D) The volume of an atom is defined as the volume of the intersection between its ball in the solvent accessible model, and its Voronoi cell [9, a quantity well defined even if the atom retains solvent accessibility. The packing of this atom is the inverse of this volume. Practically, interfaces and binding patches are computed with Vorshell[35], while atomic surface areas and volumes are computed with Vorlume 9


Figure 2 The hardness of predicting a binding affinity does not correlate with the flexibility of the complex. $x$-axis: flexibility of the interface, expressed in terms of interface iRMSD; $y$-axis: median prediction error $\Delta \Delta G_{i}[V, \mathcal{D}]$ (Eq. (4)). Dashed, dash-dotted and dotted lines respectively show errors of $\pm 1.4, \pm 2.8, \pm 4.2 \mathrm{kcal} / \mathrm{mol}$, corresponding to $K_{d}$ approximated within one, two and three orders of magnitude.


Table 1 Parameters used to estimate binding affinities. The acronyms read as follows (see text for details): Sum of Volume Differences; Shelling Order; Inverse Volume Weighted; Internal Path Length; Non Interacting Buried/Exposed; Non Interacting Surface; Solvent Accessible Surface Area;


Table 2 Parameters used by the models, depending on the dataset (i.e. subset of the structure affinity benchmark). The parameters are those from Table 1. Black dots mark variables used by statistically significant models and white dots by other models. The last column counts the number of statistically significant models using a given parameter.

|  | $\begin{gathered} \hline \text { Model } 1 \\ \text { SAB-A } \end{gathered}$ | $\begin{array}{r} \text { Model } 22 \\ \text { SAB-A-HR } \end{array}$ | $\begin{gathered} \hline \hline \text { Model } 3] \\ \text { SAB-R } 1.0 \\ \hline \end{gathered}$ | $\begin{gathered} \hline \text { Models } 4 \text { and } 9 \\ \text { SAB-R } \\ \hline 1.1 \end{gathered}$ | $\begin{gathered} \hline \hline \text { Model } 5 \\ \text { SAB-R } 1.5 \\ \hline \end{gathered}$ | $\begin{gathered} \text { Model } \\ \text { SAB- } \end{gathered}$ | $\begin{aligned} & \hline \text { Model } 7 \\ & \text { SAB-F } \\ & \hline \end{aligned}$ | $\begin{gathered} \text { Model } 8 \\ \text { SAB-F } 1.5 \\ \hline \end{gathered}$ | Counts |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| iRMSD |  |  | $\bullet$ |  |  |  |  |  | 1 |
| IVW-IPL | - |  | $\bullet$ | $\bullet$ | $\bullet$ |  | - | - | 6 |
| SVD SO1 |  |  | $\bullet$ | $\bullet$ |  |  |  | $\bullet$ | 3 |
| SVD_SOGT1 |  | - |  |  |  |  |  |  | 1 |
| SVD_NI_E |  |  |  |  |  |  | $\bullet$ |  | 1 |
| SVD_NI_B |  | - |  |  |  | - |  |  | 1 (2) |
| NIS ${ }^{\text {polar }}$ |  |  |  |  | $\bullet$ |  |  |  | 1 |
| NIS ${ }^{\text {charged }}$ | - | - | - | $\bullet$ |  |  | $\bullet$ |  | 5 |
| $\Delta$ NIS $^{\text {polar }}$ |  |  |  |  |  | $\bigcirc$ |  |  | (1) |
| $\Delta$ NIS $^{\text {charged }}$ |  |  |  |  |  |  |  |  |  |
| ATOM_SOLV |  | - |  |  |  |  | $\bullet$ | $\bullet$ | 3 |
| POLAR_SASA |  |  |  |  |  |  |  |  |  |


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## 6 Supplemental

### 6.1 Datasets from the Structure Affinity Benchmark

Curation. To exploit variation of structural parameters between the unbound and bound form, we establish a one-to-one correspondence between the atoms of a partner (from bound to unbound). To cope with cases involving missing residues or atoms, we proceed in two stages. First, we perform an alignment and map residues of the bound and unbound chains. Second, we map atoms of paired residues. We then retain the cases for which at least $80 \%$ of atoms are paired. This procedure ruled out two cases, namely 1E6J (78\%) and 1ZLI ( $76 \%$ ). We also removed three cases (1IQD, 1NSN, 1UUG) for which an upper bound on $K_{d}$ instead of a proper value is provided for a total of 139 complexes.

Datasets. The various datasets defined in previous works from the SAB are presented on Fig. 3

Figure 3 The various datasets defined from the structure affinity benchmark (SAB), based on iRMSD between the unbound and bound structures.


The datasets depicted on Fig. 3 are defined as follows:

- SAB-A (139 complexes): all complexes.
- $\mathrm{SAB}-\mathrm{R}_{1.0}$ ( 68 complexes): (focus on ridigity, strict threshold) complexes characterized by iRMSD $<1 \AA$ [39] (33] and [51] used iRMSD $\leq 1 \AA, 69$ complexes).
- $\mathrm{SAB}-\mathrm{R}_{1.1}$ (78 complexes): (focus on ridigity, intermediate threshold) complexes characterized by iRMSD $<1.1 \AA$ [30.
- SAB-R 1. $_{1.5}$ (105 complexes): (focus on ridigity, relaxed threshold) complexes characterized by iRMSD $\leq 1.5 \AA$ [33].
- SAB-I (27 complexes): (intermediate complexes) complexes characterized by $1.1 \leq \mathrm{iRMSD} \leq$ $1.5 \AA$ 30.
- SAB-F ${ }_{1}$ ( 70 complexes): (focus on flexibility, relaxed threshold) complexes characterized by iRMSD $>1 \AA[51$ ( 39$]$ used $\mathrm{iRMSD} \geq 1 \AA, 71$ complexes)
- SAB-F 1.5 (34 complexes): (focus on flexibility, strict threshold) complexes with iRMSD > $1.5 \AA 30$ 33.

To which we add:

- SAB-A-HR (37 complexes): high resolution complexes from 21


### 6.2 Resolution of Crystal Structures in the Affinity Benchmark

The distribution of resolutions of crystal structures found in the SAB is presented on Fig. 4.
Figure 4 Resolution of the structures in the SAB. The histogram and green kernel density estimation curve are for the whole SAB , the red curve is for the complexes and the blue curve is for the unbound partners. For the whole SAB: Minimum $=0.93 \AA$, median: $=2.13 \AA$, average $=2.19 \AA$, $\max =3.5 \AA$. NB: the high resolution dataset SAB-A-HR retains only entries whose resolution is better than $2.5 \AA$ for both the complex and the individual partners [21].


### 6.3 Methods used in Previous Studies

This section reviews previous work on affinity prediction, in three respects: the type of prediction model used, the variables used, and the statistical methodology.

### 6.3.1 From reference 39

Datasets. Seven complexes were discarded from the original affinity benchmark: three because only the upper bound of the affinity was known (1UUG, 1IQD and 1NSN) and four because some features needed by the models were missing (1DE4, 1M10, 1NCA and 1NB5).

Types of models. Affinity values were predicted as the un-weighted average of the output of four different classifiers (random forest, multivariate adaptive regression splines, M5' regression trees and radial basis function interpolation). Each classifier was fed a total of 200 different and possibly correlated features. All classifiers were able to perform feature selection to some extent
and therefore, their actual number of parameter was smaller than 200. In effect, M5' trees use 84 variables, random forest uses 19 variables, MARS uses 10 variables and RBF weights all variable equally and therefore virtually uses all 200 variables. The union of the three first sets contains 94 variables.

Type of variables. The variables used fall into 7 categories:

- Statistical potentials (both atomistic and coarse-grained)
- Solvation and electrostatics (using force fields)
- Entropy terms (translational, rotational, vibrational)
- Contact potentials (H-bonds, $\pi-\pi$ ) interactions, Van der Waals, salt bridges)
- Interface properties (BSA, polarity, geometrical features, surface complementarity)
- Change between bound and unbound states for all of the above
- All of the above computed on an ensemble of structures generated with CONCOORD.

Type of cross validation. On the training set, the correlation between predictions and experimental values was computed using a cross-validation. Complexes from the test set were predicted using the model trained on the validated set. The models were trained on a subset of 57 complexes with further validated affinity values. The predictions were tested on the training dataset using leave-one-out cross-validation. They were further tested on 80 complexes. However, the reported results do not include the correlation between predictions and affinity on the test set alone. Instead, the correlations were reported for the test set + cross-validated train set.

### 6.3.2 From reference 30

Datasets. The set of rigid complexes (with iRMSD $<1.1$ ) minus six complexes (1UUG, 2PTC, 1BRS, 2BTF, 1Z0K and 1S1Q) was used to fit the model. In SAB-I, four more complexes were also removed: 1EMV, 1KXP, 1AKJ and 1WQ1.

Type of model. A linear model was fitted on the data using least-square regression.

Type of variables. Two variables were used for that model: iRMSD and the buried surface area. Both are interface properties.

Type of cross-validation. No cross-validation was involved. The correlation between fitted and experimental values was computed on various subset of the SAB and the whole SAB. The results are therefore optimistic on rigid complexes.

Remarks In various datasets, complexes which were badly predicted by the model were removed as outliers. This leads to artificially high correlation coefficients. This is denoted by yellow cells in Table 4

### 6.3.3 From reference [51]

This paper first aims at creating a scoring function for docking using statistical parameters. The correlation between the score of a complex and its affinity was also computed.

Dataset. The original full dataset consisted in 3045 complexes extracted from DOCKGROUND, and the training dataset consisted of half of it. Moreover, docking decoys were used as negative examples.

The test set consisted in the SAB without 1UUG, 1IQD, 1NSN, 1DE4, 1M10, 1NCA and 1NB5.

Type of model. The model was based on a scoring function optimized to discriminate between native complexes and decoys. This function used knowledge-based statistical potentials derived from the training set i.e. the probability of a given pair of atoms interacting in a given radius compared to that same probability for non-interacting atoms.

Type of variables. The equivalent of variables for that model were the distance-dependent atom-pair potentials. These were based on observed and expected frequencies of occurrences of atom pairs. From 12 atom types, 78 different pairs occur, and this was computed for 14 different radii, leading to 1092 parameters

Type of cross-validation. No cross-validation was used since the training and test were assumed to be disjoint. It is worth noting however that 24 complexes from the SAB were also part of the 3045 original complexes. Since 1UUG was removed that lets at most 23 complexes shared between the training and test sets.

### 6.3.4 From reference [33]

Dataset. Only one complex was removed from the SAB, namely 2OZA because "its BSA was extraordinary large and detected as an outlier using the standard Grubbs' test"
type of model A linear model was fitted on a rigid subset of the SAB using least-square regression. This straining set was defined by complexes with an iRMSD $\leq 1 \AA$.

Type of variables. Three variables were used: the BSA which accounts for interface properties, NIS ${ }^{\text {charged }}$ and NIS ${ }^{\text {polar }}$ which account for the surface of the complex outside the interface.

Type of cross-validation 4 -fold cross-validation was performed to assess the performances of the model on the training set.

Remarks The cross-validation procedure was a less strict strategy than the standard cross validation, where the intersection between the data used to train and test was void. Namely, during the 4 -fold cross-validation, the coefficients of the four models trained on their respective folds were averaged to get a single model. The correlation coefficient of the prediction of that model with the actual values was reported. Therefore, through averaging, information about the whole dataset was used for training a model that was tested on the very same dataset, leading to overfitting. This is denoted by an orange cell in Table 4.

Moreover, dataset SAB-I is a superset of the training set. Namely, the model was trained on all complexes with iRMSD $<1 \AA$ and tested on complexes with iRMSD $<1.5 \AA$. This is another instance of overfitting. This is denoted by a cyan cell in Table 4.

### 6.3.5 From reference 21

Dataset. This paper used the SAB from which ten complexes were filtered out, namely: 1BJ1, 1F34, 1JIW, 1JMO, 1S1Q, 1XD3, 2J0T, 2TGP, 1NVU and 2OZA. It also defined a second dataset consisting of high-resolution entries, i.e. complexes for which both the individual partners and the bound form had a resolution lower than $2.5 \AA$.

Type of model A linear model was fitted on the data using least-square regression.

Type of variables. The variables used consisted in intra and inter-chain hydrogen bond potentials, geometric complementarity (Van der Waals interactions), volume of cavities at the surface (large enough to contain water molecules), iRMSD for interface, C- $\alpha$ and side-chains $\chi 1$ and $\chi 2$ dihedral angles, alanine-scanning defined hotspots, interface amino-acid propensities and electrostatics (Coulomb). In total the combinations of 13 variables were studied. The authors mentioned that adding more than four variables did not significantly improve the results, but the reported correlation coefficient seem to be for models of 7 or more variables after figure 5 of the paper. Which variables were actually selected was not mentioned.

Type of cross-validation. The reported correlation coefficients were computed using leave-one-out cross-validation.

### 6.4 Statistical Methodology

This section details our algorithms.

### 6.4.1 Algorithms

```
Algorithm 1 Computing a permutation p-value for a binding affinity prediction model specified by a set of variables \(V\). The p-value is based on a permutation test [41, which uses the prediction performances obtained on random datasets, each such dataset being obtained by permuting the dependent variable (i.e. the affinity) over the dataset.
Require: \(\mathcal{D}\) : dataset; \(V\) : a set of variables; \(c_{V}\) : median cross-validated correlation coefficient for \(V ; N_{\text {perm. }}\) : number of repetitions
for \(j \in\left\{1 \ldots N_{\text {perm. }}\right\}\) do
Randomly permute the dependent variable in \(\mathcal{D}\) (here the affinity) to obtain \(\mathcal{D}_{j}^{\text {perm }}\)
Perform 5 -fold cross-validation of linear models using the variables in \(V\) on \(\mathcal{D}_{j}^{\text {perm }}\)
Store the cross-validated correlation in \(c_{V}^{\text {perm }}\)
Report the approximate p-value for \(V\) to be \(\frac{B+1}{N_{\text {perm. }+1}}\) where \(B\) is the number of elements in
\(c_{V}^{\text {perm }}\) which are greater than or equal to \(c_{V}\).
```

```
Algorithm 2 Exhaustive evaluation of binding affinity prediction models. A subset of
the SAB is defined by its flexibility (supplemental Fig. 3) and the number of atoms which could
be matched from the bound structure to the unbound ones. Practically, the following values
were used: \(m=5, N_{X V}=10000\).
Require: \(\mathcal{D}\) : dataset; \(\mathcal{V}\) : pool of variables; \(m\) : maximum model size; \(N_{X V}\) : number of repeti-
    tions of the cross-validation procedure
    for all subsets \(V\) of \(\mathcal{V}\) of size 1 to \(m\) do
        for \(i \in\left\{1 \ldots N_{X V}\right\}\) do
            Perform 5 -fold cross-validation of linear models using the variables in \(V\) on \(\mathcal{D}\)
            Compute and store the correlation between predictions and experimental values
            Report \(c_{V}\), the median cross-validated correlation for \(V\)
            Report the p-value computed with Algorithm 1
```

```
Algorithm 3 Model selection: identifying specific models for a dataset \(\mathcal{D}\). The al
gorithm returns the index of the last model which cannot be distinguished from the best ones
given their cross-validated correlation coefficient distribution. It is assumed that the models are
sorted in non-decreasing order by their median cross-validated correlation coefficient. In short,
the algorithm executes a binary search, shrinking the interval by its end when there is a signifi-
cant difference between the models in the interval, and expanding it by its end when there is no
difference. All shrink/expand events are applied at the end of the interval to only keep the best
models in the final set. Storing the smallest upper bound encountered so far and stopping when
it is equal to the upper bound ensures that the algorithm finishes.
Require: \(C=\left\{\mathcal{C}_{V_{i}}, i \in\{1 \ldots 1585\}\right\}\) : the set of distributions of cross-validated correlation
    coefficient for each set of variables \(V_{i}\), sorted by non-increasing median value; cutoff: a cutoff
    for the p-value of the Kruskal - Wallis test.
    start: \(=0\)
    end \(:=|C|\)
    while TRUE do
        \(C_{\text {min }}:=\left\{\mathcal{C}_{V_{i}}, i \in\{\right.\) start,\(\ldots\), end \(\left.\}\right\}\)
        Perform Kruskal - Wallis' test on \(C_{m i n}\). Store the p-value in \(p\).
        if \(p<\) cutoff then
            \#\# Shrink toward best models
            end \(:=\left\lceil\left|C_{\min }\right| / 2\right\rceil\)
        else
            \#\# Expand toward worse models
            tmp := start
            start \(:=\) end
            end \(:=\) end \(+\lfloor(\) end - start \() / 2\rfloor\)
            if end \(>|C|\) then
                \#\# All models are equivalent
            return \((|C|)\)
            if start \(=\) end then
                \#\# the shrinking / expanding process has converged
            if \(p \geq\) cutoff then
                \#\# The final pivot is part of the similar distributions
                return(end)
            else
                    \#\# The final pivot is part of the outliers
                return(end -1 )
    if end \(=1\) then
        \#\# Only one remains (after a sequence of shrinkings only)
        return(end)
```


### 6.4.2 Models and their Complexity

Figure 5 Model complexity versus median correlation $C_{V}$ between predicted and experimental values.


SAB-A


SAB-R ${ }_{1.0}$


SAB-R ${ }_{1.5}$

$\mathrm{SAB}^{-F_{1}}$


SAB-F 1.5

### 6.4.3 Computing Correlation and Prediction Errors for Repeated Cross-validation

For a given model, our validation protocol results in $N_{X V}$ predictions for each complex. This can be seen as a $139 \times N_{X V}$ matrix $\hat{G}$ where each entry $\hat{g}_{i j}$ is the prediction for complex $i$ obtained at repetition $j$. From the experimental values $\Delta G_{d}^{e x p_{i}}$, there are therefore two ways to get a single value for the correlation and prediction error per complex.

As a first option, one can agglomerate all $N_{X V}$ predictions into a single value by taking their median. This is $\hat{g}_{i}$ defined in Eq. 3 and repeated hereunder for convenience:

$$
\begin{equation*}
\hat{g}_{i}=\operatorname{median}_{j} \hat{g}_{i j} . \tag{20}
\end{equation*}
$$

Then it is straightforward to compute the correlation between $\left\{\Delta G_{d}^{e x p_{i}}\right\}$ and $\left\{\hat{g}_{i}\right\}$ :

$$
\begin{equation*}
C[V, \mathcal{D}]=\operatorname{Corr}\left(\left\{\Delta G_{d}^{\exp _{i}}\right\},\left\{\hat{g}_{i}\right\}\right) \tag{21}
\end{equation*}
$$

and the prediction error for complex $i$ :

$$
\begin{equation*}
\Delta \Delta G_{i}[V, \mathcal{D}]=\Delta G_{d}^{e x p_{i}}-\hat{g}_{i} \tag{22}
\end{equation*}
$$

As a second option, one can take the median of the correlations (resp. prediction errors) over the repetitions. Let $\operatorname{Corr}_{j}$ be the correlation coefficient associated with repetition $j$, i.e. the correlation between $\left\{\Delta G_{d}^{e x p_{i}}\right\}$ and $\left\{\hat{g}_{i j}\right\}$ for a given $j$. This results in equations 2 and 4 repeated here for convenience:

$$
\begin{gather*}
C[V, \mathcal{D}]=\operatorname{median}_{j} \operatorname{Corr}_{j}  \tag{23}\\
\Delta \Delta G_{i}[V, \mathcal{D}]=\operatorname{median}_{j}\left(\Delta G_{d}^{e x p_{i}}-\hat{g}_{i j}\right) \tag{24}
\end{gather*}
$$

We chose the second method because it makes more sense to us to compute median over statistics than over predictions. Moreover, for a given complex $i$, the ordering of values $\hat{g}_{i j}$ and $\Delta G_{d}^{\text {exp }}-\hat{g}_{i j}$ is the same. We therefore have that $\operatorname{median}_{j}\left(\Delta G_{d}^{\text {exp }_{i}}-\hat{g}_{i j}\right)=\Delta G_{d}^{e x p_{i}}-\hat{g}_{i}$. For the correlation, the two methods give very similar values and give the highest value to the same model. (Fig. 8).

### 6.5 Results: Specific Models

Upon applying the methods described in 2.4, a single best specific model was obtained for each dataset.

- Model: 1. Obtained for dataset(s): SAB-A 2 variables, p-value $\leq 0.0001$.

$$
\begin{equation*}
\Delta G_{d}=\alpha+\beta \cdot \text { IVW-IPL }+\gamma \cdot \text { NIS }^{\text {charged }} \tag{25}
\end{equation*}
$$

Model: 2. Obtained for dataset(s): SAB-A-HR 4 variables, p-value $\leq 0.0001$.

$$
\begin{equation*}
\Delta G_{d}=\alpha+\beta \cdot \text { SVD_SOGT1 }^{2}+\gamma \cdot \text { SVD_NI_B }^{2}+\epsilon \cdot \text { NIS }^{\text {charged }}+\zeta \cdot \text { ATOM_SOLV } \tag{26}
\end{equation*}
$$

- Model: 3. Obtained for dataset(s): SAB- $\mathbf{R}_{1.0} 4$ variables, p -value $\leq 0.0001$.

$$
\begin{equation*}
\Delta G_{d}=\alpha+\beta \cdot \mathrm{iRMSD}+\gamma \cdot \mathrm{IVW}-\mathrm{IPL}+\delta \cdot \mathrm{SVD} \_\mathrm{SO} 1+\epsilon \cdot \mathrm{NIS}^{\text {charged }} \tag{27}
\end{equation*}
$$

- Model: 4. Obtained for dataset(s): SAB-R $R_{1.1} 3$ variables, p-value $\leq 0.0001$.

$$
\begin{equation*}
\Delta G_{d}=\alpha+\beta \cdot \text { IVW-IPL }+\gamma \cdot \text { SVD_SO1 }+\delta \cdot \text { NIS }^{\text {charged }} \tag{28}
\end{equation*}
$$

- Model: 5. Obtained for dataset(s): SAB-R $\mathrm{R}_{1.5} 2$ variables, p-value $\leq 0.0001$.

$$
\begin{equation*}
\Delta G_{d}=\alpha+\beta \cdot \text { IVW-IPL }+\gamma \cdot \text { NIS }^{\text {polar }} \tag{29}
\end{equation*}
$$

- Model: 6. Obtained for dataset(s): SAB-I 2 variables, p-value $\leq 0.090$.

$$
\begin{equation*}
\Delta G_{d}=\alpha+\beta \cdot \text { SVD_NI_B }^{2}+\gamma \cdot \Delta \mathrm{NIS}^{\text {polar }} \tag{30}
\end{equation*}
$$

- Model: 7. Obtained for dataset(s): SAB-F ${ }_{1} 4$ variables, p -value $\leq 0.0091$.

$$
\begin{equation*}
\Delta G_{d}=\alpha+\beta \cdot \text { IVW-IPL }+\gamma \cdot \text { SVD_NI_E }+\delta \cdot \text { NIS }^{\text {charged }}+\epsilon \cdot \text { ATOM_SOLV } \tag{31}
\end{equation*}
$$

- Model: 8. Obtained for dataset(s): SAB-F ${ }_{1.5} 3$ variables, p-value $\leq 0.0054$.

$$
\begin{equation*}
\Delta G_{d}=\alpha+\beta \cdot \text { IVW-IPL }+\gamma \cdot \text { SVD_SO1 }^{2} \delta \cdot \text { ATOM_SOLV }^{2} \tag{32}
\end{equation*}
$$

- Model: 9. Obtained for dataset(s): All datasets 3 variables, p-value $\leq 0.0001$ for SAB-A, SAB-A ${ }_{h} r$, SAB-R ${ }_{1.0}$, SAB-R ${ }_{1.1}, \mathrm{SAB}_{1} \mathrm{R}_{1.5} ; \leq 0.6949$ for $\mathrm{SAB}-\mathrm{I} ; \leq 0.0158$ for $\mathrm{SAB}-\mathrm{F}_{1} ;$ $\leq, 0.0089$ for $\mathrm{SAB}^{2} \mathrm{~F}_{1.5}$.

$$
\begin{equation*}
\Delta G_{d}=\alpha+\beta \cdot \text { IVW-IPL }+\gamma \cdot \text { SVD_SO1 }+\delta \cdot \text { NIS }^{\text {charged }} \tag{33}
\end{equation*}
$$

### 6.6 Results: Correlation and comparison with previous work

See the detailed Table 4
Table 4 Binding affinities: correlations between predictions and measurements for all datasets and all specific models. (Whole table) Red values show the best results for a given category in the corresponding section, and cross-validated and classical correlation coefficients are treated separately in the second part of the table. Bold values in the second part of the table show the categories on which the variables selection was performed for a given model. (First section) Previous work: values published and our replica (rep). Green cells correspond to the correlation coefficient of the model over the train set (i.e. $\sqrt{R^{2}}$ from the linear regression). Purple cells correspond to either cross-validation results, prediction on a test set distinct from the trains set, or both. For the other cell colors, see details in supplemental section 6.3 Yellow cells: discrepancies between original values and our replicas - we did not remove any complex. Orange cells: the cross-validation procedure made some test data information leak into the training set. Cyan cells: overlapping training and test sets. (Second section) Eight models developed in this work. The value reported in a cell corresponds to the correlation between predictions and experimental values. For the $N_{X V}=100005$-fold cross-validation, the median of the predictions was used (see Statistical Methods section 2.4. Purple lines show the cross-validated correlation coefficient, while green lines show the classical correlation coefficient (square root of the coefficient of determination).

| Model | \#param | SAB-A | SAB-A-HR | SAB-R $\mathrm{R}_{1.0}$ | SAB-R $\mathrm{l}_{1.1}$ | SAB-R 1.5 | SAB-I | SAB-F ${ }_{1}$ | SAB-F 1.5 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| [39] | 94 (200) (see sectior 6.3.1) | 0.55 | - | 0.7 | - | - | - | 0.36 | - |
| 30] rep | 2 | 0.20 | - | - | 0.55 (0.62) | - | 0.07 (0.38) | - | 0.13 |
| 51 | $\sim 1000$ (see section 6.3.3 | 0.39 | - | 0.63 | - | - | - | 0.24 | - |
| 33] rep | 3 | 0.48 | - | 0.58 | - | 0.54 | - | - | 0.34 |
| 21] | 7 | 0.57 | 0.71 | - | - | - | - | - | - |
| Model 1. selected for datasets SAB-A |  |  |  |  |  |  |  |  |  |
| $N_{X V} \cdot 5$-fold CV | 2 | 0.48 | 0.72 | 0.64 | 0.62 | 0.46 | -0.39 | 0.27 | 0.39 |
| LOO CV | 2 | 0.48 | 0.72 | 0.63 | 0.61 | 0.46 | -0.57 | 0.24 | 0.35 |
| No CV | 2 | 0.52 | 0.76 | 0.67 | 0.64 | 0.51 | 0.20 | 0.37 | 0.59 |
| Model 2. selected for datasets SAB-A-HR |  |  |  |  |  |  |  |  |  |
| $N_{X V} \cdot 5$-fold CV | 4 | 0.44 | 0.77 | 0.52 | 0.51 | 0.40 | -0.08 | 0.24 | 0.31 |
| LOO CV | 4 | 0.44 | 0.77 | 0.51 | 0.51 | 0.39 | -0.16 | 0.22 | 0.27 |
| No CV | 4 | 0.50 | 0.83 | 0.62 | 0.60 | 0.49 | 0.40 | 0.40 | 0.59 |
| Model 3. selected for datasets SAB-R $\mathrm{R}_{1.0}$ |  |  |  |  |  |  |  |  |  |
| $N_{X V} \cdot 5$-fold CV | 4 | 0.47 | 0.71 | 0.67 | 0.62 | 0.46 | -0.21 | 0.24 | 0.42 |
| LOO CV | 4 | 0.46 | 0.71 | 0.66 | 0.62 | 0.45 | -0.31 | 0.23 | 0.39 |
| No CV | 4 | 0.52 | 0.76 | 0.72 | 0.67 | 0.53 | 0.40 | 0.41 | 0.65 |
| Model 4. selected for datasets SAB-R $\mathrm{R}_{1.1}$ |  |  |  |  |  |  |  |  |  |
| $N_{X V} \cdot 5$-fold CV | 3 | 0.47 | 0.71 | 0.64 | 0.63 | 0.46 | -0.21 | 0.28 | 0.44 |
| LOO CV | 3 | 0.47 | 0.71 | 0.63 | 0.62 | 0.45 | -0.30 | 0.26 | 0.41 |
| No CV | 3 | 0.52 | 0.76 | 0.68 | 0.67 | 0.51 | 0.35 | 0.41 | 0.65 |
| Model 5. selected for datasets SAB-R ${ }_{1.5}$ |  |  |  |  |  |  |  |  |  |
| $N_{X V} \cdot 5$-fold CV | 2 | 0.44 | 0.64 | 0.63 | 0.60 | 0.47 | -0.27 | 0.25 | 0.31 |
| LOO CV | 2 | 0.43 | 0.63 | 0.62 | 0.59 | 0.46 | -0.54 | 0.22 | 0.28 |
| No CV | 2 | 0.47 | 0.68 | 0.66 | 0.62 | 0.5 | 0.12 | 0.32 | 0.43 |
| Model 6: selected for datasets SAB-I |  |  |  |  |  |  |  |  |  |
| $N_{X V}$. 5-fold CV | 2 | 0.04 | 0.39 | 0.00 | 0.01 | 0.14 | 0.23 | -0.11 | -0.15 |
| LOO CV | 2 | -0.04 | 0.07 | 0.03 | 0.02 | 0.07 | 0.17 | -0.10 | -0.18 |
| No CV | 2 | 0.15 | 0.50 | 0.18 | 0.18 | 0.23 | 0.43 | 0.12 | 0.15 |
| Model 73 selected for datasets SAB-F 1 |  |  |  |  |  |  |  |  |  |
| $N_{X V}$. 5-fold CV | 4 | 0.46 | 0.69 | 0.59 | 0.57 | 0.43 | -0.19 | 0.31 | 0.33 |
| LOO CV | 4 | 0.46 | 0.69 | 0.59 | 0.56 | 0.43 | -0.28 | 0.29 | 0.28 |
| No CV | 4 | 0.52 | 0.77 | 0.67 | 0.65 | 0.51 | 0.34 | 0.44 | 0.60 |
| Model 8. selected for datasets SAB-F 1.5 |  |  |  |  |  |  |  |  |  |
| $N_{X V} \cdot 5$-fold CV | 3 | 0.35 | 0.46 | 0.59 | 0.53 | 0.36 | -0.04 | 0.20 | 0.46 |
| LOO CV | 3 | 0.34 | 0.43 | 0.58 | 0.52 | 0.35 | -0.12 | 0.18 | 0.45 |
| No CV | 3 | 0.40 | 0.55 | 0.64 | 0.59 | 0.42 | 0.29 | 0.33 | 0.59 |

### 6.7 Results: Correlations between individual variables and/or measured affinity

Table 5 Pearson correlation coefficients between the individual variables and the affinity.

| -0.09 | -0.39 | -0.39 | -0.24 | -0.19 | -0.1 | -0.04 | 0.01 | irmsd <br> ivwipl |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.36 | 0.48 | 0.59 | 0.56 | 0.41 | 0.11 | 0.23 | 0.34 |  |
| 0.13 | 0.19 | 0.01 | -0.02 | 0.04 | 0.28 | 0.24 | 0.36 | s_diff_vol_so1 |
| -0.3 | -0.41 | -0.49 | -0.45 | -0.33 | -0.12 | -0.19 | -0.28 | s_diff_vol_sogt1 |
| 0.08 | 0.38 | 0.06 | 0.06 | 0.13 | 0.35 | 0.09 | -0.07 | s_diff_vol_not_int_bur |
| 0 | -0.1 | -0.18 | -0.16 | -0.04 | 0.25 | 0.15 | 0.15 | s_diff_vol_not_int_surf |
| 0.29 | 0.55 | 0.37 | 0.37 | 0.3 | 0.05 | 0.21 | 0.23 | nis_polar |
| -0.4 | -0.74 | -0.51 | -0.52 | -0.38 | 0.18 | -0.29 | -0.49 | nis_charged |
| -0.13 | -0.32 | -0.18 | -0.17 | -0.21 | -0.36 | -0.09 | 0.14 | nis_polar_diff |
| -0.03 | 0.01 | 0.02 | -0.03 | 0.02 | 0.18 | -0.08 | -0.18 | nis_charged_diff |
| 0.14 | 0.35 | 0.23 | 0.16 | 0.1 | -0.09 | 0.02 | 0.33 | solvation_eisen |
| -0.05 | -0.1 | 0.03 | -0.01 | -0.02 | -0.01 | -0.1 | -0.1 | polar_surf_area |
|  |  | $\frac{s_{5}^{5}}{5}$ |  |  |  |  |  |  |

Table 6 Pearson correlation coefficients between the individual variables.


### 6.8 Results: Validation on an External Dataset

Table 7 Validation of the models on an external test set. The external test set from [33, supplemental] was split using the same criteria as those used to define datasets from the structure affinity benchmark, yielding external datasets. The first part of the table displays the external dataset size, Pearson's correlation coefficients and p-value for each model on its external dataset along with $p_{1.4}^{\Delta \Delta G}, p_{2.8}^{\Delta \Delta G}$ and $p_{4.2}^{\Delta \Delta G}$. The second part show the values from table the diagonal of tables 3 and 4 for comparison.

|  | $\begin{aligned} & \hline \hline \text { Model } \\ & \text { SAB-A } \end{aligned}$ | $\sqrt[2]{2}$ | Model 3 SAB-R ${ }_{1.0}$ | Mode 4 SAB-P <br> SAB- $\mathrm{R}_{1}$ | $\begin{gathered} \hline \text { Model } 5 \\ \text { SAB-R } 51.5 \end{gathered}$ | Model 6 SAB-1 <br> SAB | $\begin{aligned} & \hline \hline \text { Model } 7 \\ & \text { SAB- } 71 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline \text { Model } 8 \\ & \text { SAB-F } 81.5 \\ & \hline \end{aligned}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| dataset size | 51 | 24 | 13 | 16 | 23 | 7 | 38 | 28 |
| p-value | 0.0004 | 0.0295 | 0.0022 | 0.0392 | 0.0170 | 0.6034 | 0.0043 | 0.2753 |
| correlation | 0.47 | 0.44 | 0.77 | 0.52 | 0.49 | 0.24 | 0.45 | 0.21 |
| $p_{1.4}^{\Delta \Delta G}, p_{2.8}^{\Delta \Delta G}, p_{4.2}^{\Delta \Delta G}$ | 11.76, 33.33, 47.06 | 20.83, 25.00, 41.67 | 30.77, 30.77, 46.15 | 31.25, 43.75, 43.75 | 26.09, 34.78, 43.48 | 0.00, 14.29, 14.29 | 5.26, 26.32, 52.63 | 17.86, 46.43, 60.71 |
| median corr. | 0.48 | 0.77 | 0.67 | 0.63 | 0.47 | 0.23 | 0.31 | 0.46 |
| $p_{1.4}^{\Delta \Delta G}, p_{2.8}^{\Delta \Delta \Delta^{G}}, p_{4.2}^{\Delta \Delta G}$ | 47.48, 78.42, 92.09 | 62.16, 89.19, 97.30 | 51.47, 82.35, 92.65 | 55.13, 79.49, 91.03 | 45.71, 80.00, 89.52 | 44.44, 70.37, 88.89 | 50.00, 78.57, 90.00 | 50.00, 79.41, 85.29 |

Table 8 Validation of the best overall model i.e. model 9 on an external test set. See Table 7 for the statistics presented.


### 6.9 Results: On the Quality of Individual Predictions

Table 9 lists the individual predictions, obtained from Eq. (3).

Table 9 Experimental affinities on a per complex basis: experimental measurements $\left(\Delta G_{d}\right)$ versus predictions ( $\hat{g}_{i}$, Eq. 3). Predictions were generated with Model 1 on dataset $\mathbf{S A B}-\mathbf{A}$. The median was taken over the $N X V$ repetitions. Blue values indicate under-predicted complexes (63) and red indicate the over-predicted ones (76). A start denotes complexes with error in the top decile.

| PDB ID | Measured | Predicted | PDB ID | Measured | Predicted | PDB ID | Measured | Predicted |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1A2K | 9.31 | 10.72 | 1I4D | 7.46 | 9.55 | 1XU1 | 11.18 | 10.66 |
| 1ACB | 13.05 | 12.63 | 1IB1 | 9.76 | 10.65 | 1YVB | 11.17 | 9.48 |
| 1AHW | 11.55 | 11.25 | 1IBR | 12.07 | 12.32 | 1Z0K | 6.98 | 10.22 |
| 1AK4 | 6.43 | 10.00 | 1IJK | 10.42 | 8.85 | 1ZHI | 9.08 | 9.09 |
| 1AKJ | 5.32 | 10.34 * | 1J2J | 8.13 | 8.6 | 1ZM4 | 8.03 | 9.26 |
| 1ATN | 12.07 | 10.29 | 1JIW | 15.55 | 12.55 | 2A9K | 10.25 | 9.93 |
| 1AVX | 12.50 | 12.66 | 1JMO | 9.47 | 11.74 | 2ABZ | 11.67 | 11.06 |
| 1AVZ | 6.55 | 10.24 | 1JPS | 13.64 | 11.95 | 2AJF | 10.63 | 10.27 |
| 1AY7 | 13.23 | 10.46 | 1JTG | 12.82 | 13.52 | 2AQ3 | 6.71 | 9.17 |
| 1B6C | 8.94 | 9.35 | 1JWH | 11.14 | 9.17 | 2B42 | 12.11 | 13.86 |
| 1BJ1 | 11.55 | 12.09 | 1K5D | 12.77 | 10.22 | 2B4J | 10.86 | 10.4 |
| 1BRS | 17.32 | 10.76 * | 1KAC | 10.68 | 11.74 | 2BTF | 7.69 | 10.68 |
| 1BUH | 9.70 | 9.91 | 1KKL | 10.02 | 9.39 | $2 \mathrm{C0L}$ | 9.82 | 10.73 |
| 1BVK | 10.53 | 10.65 | 1KLU | 7.28 | 9.75 | 2 FJU | 7.20 | 9.35 |
| 1BVN | 15.06 | 12.58 | 1KTZ | 8.92 | 9.95 | 2GOX | 12.08 | 10.01 |
| 1CBW | 10.75 | 11.88 | 1KXP | 12.34 | 12.79 | 2HLE | 10.09 | 11.19 |
| 1DE4 | 9.78 | 10.12 | 1 KXQ | 11.54 | 12.43 | 2HQS | 10.15 | 13.37 |
| 1DFJ | 18.05 | 11.34 * | 1LFD | 7.79 | 8.34 | 2HRK | 10.98 | 10.93 |
| 1DQJ | 11.67 | 12.52 | 1M10 | 11.24 | 10.7 | 2 I 25 | 12.28 | 10.84 |
| 1E4K | 7.87 | 10.33 | 1MAH | 14.51 | 11.49 | 2I9B | 12.93 | 11.81 |
| 1E6E | 8.28 | 10.28 | 1MLC | 9.61 | 11.27 | 2J0T | 13.34 | 10.53 |
| 1E96 | 7.42 | 8.82 | 1MQ8 | 7.53 | 9.02 | 2JEL | 11.59 | 11.53 |
| 1EAW | 14.06 | 12.13 | 1NB5 | 13.86 | 9.77 * | 2MTA | 7.42 | 9.73 |
| 1EER | 15.59 | 12.67 | 1 NCA | 11.02 | 11.29 | 2 NYZ | 12.69 | 13.19 |
| 1EFN | 10.12 | 10.48 | 1NVU | 7.43 | 10.94 | 2O3B | 15.68 | 11.65 * |
| 1EMV | 18.58 | 7.54 * | 1NVU | 7.80 | 12.18 * | 200B | 5.66 | 7.47 |
| 1EWY | 7.43 | 9.42 | 1NW9 | 11.19 | 10.72 | 2OOR | 10.65 | 10.72 |
| 1EZU | 13.77 | 11.23 | 1OC0 | 12.28 | 10.53 | 2OUL | 11.96 | 10.9 |
| 1F34 | 14.19 | 13 | 1OPH | 11.32 | 11.52 | 2OZA | 11.73 | 14.81 |
| 1F6M | 7.60 | 9.64 | 1P2C | 13.63 | 11.88 | 2PCB | 6.82 | 8.79 |
| 1FC2 | 10.43 | 9.68 | 1PPE | 15.56 | 12.92 | 2 PCC | 7.91 | 9.07 |
| 1FFW | 8.09 | 9.51 | 1PVH | 9.52 | 10.32 | 2PTC | 18.04 | 12.57 * |
| 1FLE | 12.28 | 13.21 | 1PXV | 12.97 | 12.63 | 2SIC | 13.84 | 13.47 |
| 1FQJ | 9.79 | 9.82 | 1QA9 | 7.16 | 8.65 | 2SNI | 15.96 | 12.15 |
| 1FSK | 13.12 | 11.48 | 1R0R | 14.17 | 13.05 | 2TGP | 7.54 | 12.7 * |
| 1GCQ | 6.51 | 9.59 | 1R6Q | 8.84 | 10.94 | 2UUY | 11.26 | 13.3 |
| 1GL1 | 13.23 | 12.18 | 1RLB | 8.18 | 8.83 | 2VDB | 13.40 | 8.42 * |
| 1GLA | 6.76 | 8.84 | 1RV6 | 13.86 | 9.93 | 2VIR | 12.28 | 11.02 |
| 1GPW | 11.32 | 10.8 | 1S1Q | 4.29 | 9.8 * | 2VIS | 7.36 | 11.34 * |
| 1GRN | 9.03 | 11.5 | 1T6B | 13.10 | 10.39 | 2WPT | 10.67 | 5.92 * |
| 1GXD | 11.30 | 10.25 | 1US7 | 8.09 | 8.24 | 3BP8 | 11.44 | 10.44 |
| 1H1V | 10.20 | 10.83 | 1VFB | 11.46 | 12.36 | 3BZD | 9.57 | 9.39 |
| 1H9D | 9.18 | 9.03 | 1WDW | 12.72 | 10.52 | 3 CPH | 8.84 | 9.55 |
| 1HCF | 13.08 | 9.82 | 1WEJ | 12.48 | 11.3 | 3SGB | 14.51 | 13.25 |
| 1HE8 | 7.37 | 8.78 | 1WQ1 | 6.62 | 11.56 * | 4 CPA | 11.32 | 11.23 |
| 1HIA | 10.76 | 11.39 | 1XD3 | 8.90 | 11.14 |  |  |  |
| 1I2M | 15.83 | 13.18 | 1XQS | 7.08 | 10.71 |  |  |  |

Figure 6 Distribution of atomic volumes i.e. volumes of Voronoi restrictions for interface atoms. Red curves denote complexes whose interface lies in the top decile in terms of size (i.e, more than 354 interface atoms).


Figure 7 The quality of individual predictions, assessed by $\Delta \Delta G_{i}[V]$, does not correlate with the interface size.


Figure 8 Comparison between two ways of computing the correlation for a given model over multiple repetitions. See details in section 6.4.3. Median of correlations: for each of the $N_{X V}$ repeats, compute the correlation between the predictions and experimental affinities. Take the median of these predictionsfor each complex. Correlation of median of predictions: compute a single prediction per complex as the median of all $N_{X V}$ predictions. Compute the correlation between those predictions and the experimental affinities. The values of all models tested on all datasets have been aggregated on this figure. The correlation between both methods is 0.997 with a median absolute difference of 0.005 . Moreover, both measures are maximal for the same model on all datasets, i.e. maximizing on either measure will result in the same model selected.


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[^0]:    * Inria
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[^1]:    ${ }^{1}$ To be precise, IPL $=\sum_{a \in \mathcal{I}} \mathrm{SO}(a)$. Note that replacing the SO of each atom by one results in the number of interface atoms, which is known to correlate with BSA for rigid cases [32].

[^2]:    ${ }^{2}$ Since we have 12 variables, one has $\sum_{k=1}^{5}\binom{12}{k}=1585$.

