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The distribution of ascents of size d or more in compositions

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A composition of a positive integer n is a finite sequence of positive integers a_1, a_2, ..., a_k such that $a_1 + a_2 + \cdots + a_k = n$. Let d be a fixed nonnegative integer. We say that we have an ascent of size d or more if $a_{i+1} \geq a_i + d$. We determine the mean, variance and limiting distribution of the number of ascents of size d or more in the set of compositions of n. We also study the average size of the greatest ascent over all compositions of n.

Keywords: compositions, distributions, generating functions, ascents

1   Introduction

A composition of a positive integer n is a finite sequence of positive integers a_1, a_1, ..., a_k such that $a_1 + a_2 + \cdots + a_k = n$. It is well known that there are 2^{n-1} compositions of n. The compositions (denoted by $a_1 a_2 a_3 \ldots a_k$) for $n = 1, 2, \ldots, 5$ are:

n															
1	1														
2	11	2													
3	111	12	21	3											
4	1111	13	31	22	112	121	211	4							
5	11111	14	41	23	32	113	131	311	122	212	221	1112	1121	1211	2111	5

Let $d \geq 0$ be a fixed integer. We say, that we have an ascent of size d or more, whenever $a_{i+1} \geq a_i + d$. For example, there are 3 ascents of size 2 or more that occur in the compositions of 5:   1<u>4</u>, 1<u>13</u>, and 1<u>31</u> .

In Sections 2, 3 and 4, respectively, we determine the mean, variance and asymptotic distribution of the number of ascents of size d or more in compositions of n. Ordinary ascents (the cases $d = 0$ and $d = 1$) have previously been studied by Carlitz (1) and more recently by Chinn, Heubach and Grimaldi in (2). Finally in Section 5 we investigate the maximum value of d for which compositions of n can expect to have an ascent of size d. That is, we find the average value of the *largest ascent* that occurs in the compositions of n. For example, the compositions of 4 in the table have maximum ascents of sizes $0, 2, 0, 0, 1, 1, 0, 0$, respectively, giving an average largest ascent size per composition of $1/2$ when $n = 4$.

We note that the asymptotic expression for the maximum ascent size in Theorem 4 involves a fluctuating function of n of mean zero. A similar phenomenon has been observed when studying averages of certain other statistics for compositions, such as the largest part size (11) or the number of distinct part sizes (8), (7).

2 The average number of ascents of size d or more in compositions

For fixed $d \geq 0$ we wish to find the average number of ascents of size d or more per composition of n.

We use the "adding-the-slice" technique which was originally used by Flajolet and Prodinger in (4) and more recently, for example, by Knopfmacher and Prodinger in (9).

Let j be the value of the last component of the composition with k parts, i.e. $a_k = j$. We proceed from a composition with k parts to a composition with $k + 1$ parts. We denote by $f_k(z, u, v)$ the generating function where z marks the size n, u the value of j and v the number of ascents of size d or more in compositions with k parts.

In moving from a composition with k parts to a composition with $k + 1$ parts, where $a_k = j$, we have an ascent whenever the new last integer has any value from $j + d$ onwards. This gives the following rule for adding a new part or "slice" to the end of the composition:

$$u^j \longrightarrow zu + (zu)^2 + (zu)^3 + \cdots + (zu)^{j+d-1} + v \left\{ (zu)^{j+d} + (zu)^{j+d+1} + \cdots \right\}$$

$$= zu \frac{1 - (zu)^{j+d-1}}{1 - zu} + v(zu)^{j+d} \frac{1}{1 - zu}.$$

This implies that

$$f_{k+1}(z, u, v) = \frac{zu}{1 - zu} f_k(z, 1, v) - \frac{(zu)^d}{1 - zu} f_k(z, zu, v) + \frac{v(zu)^d}{1 - zu} f_k(z, zu, v)$$

$$= \frac{zu}{1 - zu} f_k(z, 1, v) - \frac{(1 - v)(zu)^d}{1 - zu} f_k(z, zu, v). \tag{2.1}$$

Now define $F(z, u, v) := \sum_{k \geq 1} f_k(z, u, v)$. Then summing (2.1) over $k \geq 1$ gives

$$F(z, u, v) - f_1(z, u, v) = \frac{zu}{1 - zu} F(z, 1, v) - \frac{(1 - v)(zu)^d}{1 - zu} F(z, zu, v),$$

so that

$$F(z, u, v) = \frac{zu}{1 - zu} F(z, 1, v) + \frac{zu}{1 - zu} - \frac{(1 - v)(zu)^d}{1 - zu} F(z, zu, v),$$

where we have used

$$f_1(z, u, v) = zu + (zu)^2 + (zu)^3 + \cdots = \frac{zu}{1 - zu}.$$

At this stage we iterate the recursion for $F(z, u, v)$.

$$
F(z, u, v) = \frac{zu}{1 - zu} F(z, 1, v) + \frac{zu}{1 - zu} - \frac{(1 - v)(zu)^d}{1 - zu} \times
$$

$$
\times \left\{ \frac{z^2 u}{1 - z^2 u} F(z, 1, v) + \frac{z^2 u}{1 - z^2 u} - \frac{(1 - v)(z^2 u)^d}{1 - z^2 u} F(z, z^2 u, v) \right\}
$$

$$
= \left[\frac{zu}{1 - zu} - \frac{(1 - v)z^2 u(zu)^d}{(1 - zu)(1 - z^2 u)} \right] [F(z, 1, v) + 1] + \frac{(1 - v)^2 (zu)^d (z^2 u)^d}{(1 - zu)(1 - z^2 u)} \times
$$

$$
\times \left\{ \frac{z^3 u}{1 - z^3 u} F(z, 1, v) + \frac{z^3 u}{1 - z^3 u} - \frac{(1 - v)(z^3 u)^d}{1 - z^3 u} F(z, z^3 u, v) \right\}
$$

$$
= \left[\frac{zu}{1 - zu} - \frac{(1 - v)z^2 u(zu)^d}{(1 - zu)(1 - z^2 u)} + \frac{(1 - v)^2 z^3 u(zu)^d (z^2 u)^d}{(1 - zu)(1 - z^2 u)(1 - z^3 u)} \right] [F(z, 1, v) + 1]
$$

$$
- \frac{(1 - v)^3 (zu)^d (z^2 u)^d (z^3 u)^d}{(1 - zu)(1 - z^2 u)(1 - z^3 u)} F(z, z^3 u, v).
$$

We keep iterating, noting that $F(z, z^m u, v) \to 0$ as $m \to \infty$ for $|z| < \frac{1}{2}$ and u, v in a suitable small neighbourhood of 1, and put $u = 1$ to obtain

$$
F(z, 1, v) = \sum_{i \geq 1} \frac{(-1)^{i-1} z^i (1 - v)^{i-1} z^{d\binom{i}{2}}}{(1 - z)(1 - z^2) \cdots (1 - z^i)} [F(z, 1, v) + 1]. \tag{2.2}
$$

By adding the term 1 for the empty composition we obtain the bivariate generating function for compositions according to the number of ascents of size d or more as

$$
F(z, v) := 1 + F(z, 1, v) = \frac{1}{1 - \tau(z, v)}, \tag{2.3}
$$

where

$$
\tau(z, v) := \sum_{i \geq 1} \frac{(-1)^{i-1} z^i (1 - v)^{i-1} z^{d\binom{i}{2}}}{(1 - z)(1 - z^2) \cdots (1 - z^i)}. \tag{2.4}
$$

The expected value of the number of ascents of size d or more is $\frac{[z^n] \left. \frac{\partial F(z, v)}{\partial v} \right|_{v=1}}{2^{n-1}}$. For this we shall need

$$
\tau(z, 1) = \sum_{i \geq 1} \left. \frac{(-1)^{i-1} z^i (1 - v)^{i-1} z^{d\binom{i}{2}}}{(1 - z)(1 - z^2) \cdots (1 - z^i)} \right|_{v=1} = \frac{z}{1 - z}, \tag{2.5}
$$

and

$$
\left. \frac{\partial \tau(z, v)}{\partial v} \right|_{v=1} = \sum_{i \geq 2} \left. \frac{(-1)^i z^i (i - 1)(1 - v)^{i-2} z^{d\binom{i}{2}}}{(1 - z)(1 - z^2) \cdots (1 - z^i)} \right|_{v=1} = \frac{z^{2+d}}{(1 - z)(1 - z^2)}. \tag{2.6}
$$

In particular, the generating function for all compositions is $F(z,1) = \frac{1}{1-\tau(z,1)} = \frac{1-z}{1-2z}$.
Now

$$\left.\frac{\partial F(z,v)}{\partial v}\right|_{v=1} = \left.\frac{\frac{\partial \tau(z,v)}{\partial v}}{(1-\tau(z,v))^2}\right|_{v=1} = \frac{z^{2+d}}{(1+z)(1-2z)^2}$$

$$= z^d\left[\frac{1}{9(1+z)} + \frac{1}{6(1-2z)^2} - \frac{5}{18(1-2z)}\right].$$

So that

$$[z^n]\left.\frac{\partial F(z,v)}{\partial v}\right|_{v=1} = \frac{(-1)^{n-d}}{9} + \frac{(n-d+1)2^{n-d}}{6} - \frac{5\,2^{n-d}}{18}$$

$$= \frac{(-1)^{n-d}}{9} + \frac{(3n-3d-2)2^{n-d}}{18}.$$

After dividing by 2^{n-1}, the total number of compositions of n, we have

Theorem 1 *The expected number of ascents of size d or more in the compositions of n is*

$$\mathbb{E}(n) := \frac{2^{-d}}{9}(3n-3d-2) + \frac{2}{9}\frac{(-1)^{n-d}}{2^n}\,, \text{ for } n \geq d.$$

Hence for fixed d, as $n \to \infty$,

$$\mathbb{E}(n) = \frac{2^{-d}}{3}\,n + O(1)\,.$$

Previously Chinn, Heubach and Grimaldi found the number of ascents for $d = 0$ and $d = 1$ in (2). The case $d = 0$ corresponds to the number of rises plus the number of levels, whereas $d = 1$ corresponds to the number of rises.

3   Variance of the number of ascents of size d or more in compositions

To find the variance we first need to compute $\left.\frac{\partial^2 F(z,v)}{\partial v^2}\right|_{v=1}$. In addition to formulas (2.3) to (2.6) from Section 2 we require

$$\left.\frac{\partial^2\tau(z,v)}{\partial v^2}\right|_{v=1} = \left.\sum_{i\geq 3}\frac{(-1)^{i-1}z^i\,(i-1)(i-2)(1-v)^{i-3}z^{d\binom{i}{2}}}{(1-z)(1-z^2)\cdots(1-z^i)}\right|_{v=1} = \frac{2z^{3(1+d)}}{(1-z)(1-z^2)(1-z^3)}\,.$$

Then

$$\left.\frac{\partial^2 F(z,v)}{\partial v^2}\right|_{v=1} = \left.\frac{(1-\tau)\frac{\partial^2\tau(z,v)}{\partial v^2} + 2\tau'^{\,2}}{(1-\tau)^3}\right|_{v=1}.$$

Computing the nth coefficient of the second derivative amounts to expanding and combining binomial series.
Finally, after adding the expectation and subtracting the square of the expectation we find

Theorem 2 *The variance of the expected number of ascents of size d or more in the compositions of n is*

$$\mathbb{V}(n) := 2^{-d} \left\{ \frac{-2}{9} + \frac{n}{3} - \frac{d}{3} \right\} + 2^{-2d} \left\{ \frac{80}{81} + \frac{10d}{9} + \frac{d^2}{3} - \frac{13n}{27} - \frac{2nd}{9} \right\}$$

$$+ 2^{-3d} \left\{ -\frac{352}{441} + \frac{8n}{21} - \frac{8d}{7} \right\} + 2^{-n} \left\{ (-1)^n \left(\frac{5}{27} + \frac{2n}{27} - \frac{4d}{27} \right) - \frac{1}{3} + 2\,\alpha(n) \right\}$$

$$+ 2^{-n-d}(-1)^{n-d} \left\{ \frac{4d}{27} - \frac{4n}{27} + \frac{8}{81} \right\} - 2^{-2n} \frac{4}{81},$$

for $n \geq 3d$, where

$$\alpha(n) = \begin{cases} \frac{26}{147} & \text{if } n = 3m, \\ \frac{-4}{147} & \text{if } n = 3m - 2, \\ \frac{-22}{147} & \text{if } n = 3m - 1, \quad \text{for } m \in \mathbb{N}. \end{cases}$$

For fixed d we have

$$\mathbb{V}(n) \sim n \left\{ \frac{2^{-d}}{3} - 2^{-2d} \left(\frac{13}{27} + \frac{2d}{9} \right) + \frac{2^{-3d}\,8}{21} \right\} \text{ as } n \to \infty.$$

4 Limiting distribution

We are interested in finding the limiting distribution of our random variable. We make use of Theorem IX.9 from Flajolet and Sedgewick (5). A short version is as follows:

Let $F(z, u)$ be a bivariate function that is bivariate analytic at $(z, u) = (0, 0)$ and has nonnegative coefficients there. Assume that $F(z, 1)$ is meromorphic in $z \leq r$ with only a simple pole at $z = \rho$ for some positive $\rho < r$. Then, under further conditions stated in (5), the random variable with probability generating function

$$p_n(u) = \frac{[z^n]F(z, u)}{[z^n]F(z, 1)}$$

converges in distribution to a Gaussian variable with a speed of convergence that is $O(n^{-1/2})$.

Let us introduce the notation

$$c_{i,j} := \frac{\partial^{i+j}}{\partial z^i \partial u^j} C(z, u) \Bigg|_{(\rho, 1)}. \tag{4.1}$$

From Theorem IX.9 we need to show that

$$c_{0,1} c_{1,0} \neq 0. \tag{4.2}$$

In addition, we must show that

$$\rho c_{1,0}^2 c_{0,2} - \rho c_{1,0} c_{1,1} c_{0,1} + \rho c_{2,0} c_{0,1}^2 + c_{0,1}^2 c_{1,0} + c_{0,1} c_{1,0}^2 \rho \neq 0. \tag{4.3}$$

For our specific problem

$$F(z, v) = \frac{1}{1 - \tau(z, v)} \equiv \frac{B(z, v)}{C(z, v)},$$

so that

$$C(z, v) = 1 - \sum_{i \geq 1} \frac{(-1)^{i-1} z^i (1-v)^{i-1} z^{d\binom{i}{2}}}{(1-z)(1-z^2) \cdots (1-z^i)}.$$

We have $\rho(1) = \rho = \frac{1}{2}$ and using (4.1),

$$c_{0,1} = -\frac{2^{1-d}}{3}, \qquad c_{1,0} = -4, \qquad c_{1,1} = -\frac{3d+11}{9} 2^{2-d}, \qquad c_{0,2} = -\frac{2^{4-3d}}{21}, \qquad c_{2,0} = -16.$$

We are now in a position to check the conditions listed in the theorem.
Equation (4.2) is satisfied since

$$c_{0,1} c_{1,0} = 8 \frac{2^{-d}}{3} \neq 0.$$

Equation (4.3) is equivalent to

$$2^{4-3d} \left(\frac{2^3}{21} + 2^d \frac{3d+11}{27} - \frac{2^{1+d}}{9} - \frac{2^{2+2d}}{9} \right) \neq 0$$

for non-negative integer values of d. Thus we deduce

Theorem 3 *The distribution of the number of ascents of size d or more in compositions of n converges to a Gaussian distribution with a speed of convergence of $O(n^{-1/2})$ with the mean μ_n and the variance σ_n^2 are as given in Theorems 1 and 2.*

Remark In Flajolet and Sedgewick (5) it is also shown that under the conditions of Theorem IX.9, the mean μ_n and variance σ_n^2 are of the form

$$\mu_n = \mathfrak{m} \left(\frac{\rho(1)}{\rho(u)} \right) n + O(1), \qquad \sigma_n^2 = \mathfrak{v} \left(\frac{\rho(1)}{\rho(u)} \right) n + O(1),$$

where

$$\mathfrak{m}(f) = \frac{f'(1)}{f(1)} \quad \text{and} \quad \mathfrak{v}(f) = \frac{f''(1)}{f(1)} + \frac{f'(1)}{f(1)} - \left(\frac{f'(1)}{f(1)} \right)^2.$$

These asymptotic expressions are easily checked to be in agreement with the exact results for the mean and variance found previously in Theorems 1 and 2.

5   Size of the maximum ascent

Given a composition $a_1 \, a_2 \ldots a_k$ of n we shall study the size of the maximum ascent, that is, the parameter X where

$$X := \max\{a_{i+1} - a_i \mid 1 \leq i < k \text{ and } a_{i+1} \geq a_i\}.$$

We assign a value of 0 if no pair of consecutive integers satisfies this condition. The mean value of X is given by the expression

$$\sum_{d=0}^{n} \mathbb{P}(X > d) = \sum_{d=0}^{n} (1 - \mathbb{P}(X \leq d)).$$

Therefore for each fixed d we need to compute the probability that a composition of n has maximum ascent $X \leq d$.

We already know the generating function for compositions with no ascents of size d or more, which we will denote by $F_d(z)$. For this we use (2.3) with $d + 1$ instead of d and $v = 0$, giving

$$F_{d+1}(z) := \frac{1}{1 - \tau(z, 0)} = \frac{1}{1 - \sum_{i \geq 1} \frac{(-1)^{i-1} z^{i+(d+1)\binom{i}{2}}}{(1-z)(1-z^2)\cdots(1-z^i)}}.$$

We want the generating function of compositions with $X \leq d$. This is equivalent to the generating function of the compositions with no ascent of size $d + 1$ or more, which is $F_{d+1}(z)$. We now need to study the dominant poles, ρ_d, of $F_{d+1}(z)$, that is, the dominant zeros of

$$1 - \sum_{i \geq 1} \frac{(-1)^{i-1} z^{i+(d+1)\binom{i}{2}}}{(1-z)(1-z^2)\cdots(1-z^i)} = 0. \tag{5.1}$$

We follow the approach used by Gourdon and Prodinger in (6), which is also analogous to the one found in (10). Now for $|z| \leq \frac{3}{5}$, say, using the first two terms of the series above, the root ρ_d above can be approximated by the smallest positive root of

$$1 - \frac{z}{1-z} + \frac{z^{d+3}}{(1-z)(1-z^2)} + O(z^{3d}) = 0,$$

since the omitted terms in (5.1) are $O\left(z^{3d}\right)$. This error will be majorized by subsequent O terms below. That is we want the root of

$$1 - 2z + \frac{z^{d+3}}{1-z^2} + O(z^{3d}) = 0.$$

The bootstrapping method gives a suitable approximation to ρ_d.

Let $\rho_d := \frac{1}{2} + \varepsilon_d$, then

$$1 - 2\left(\frac{1}{2} + \varepsilon_d\right) + \frac{4}{3} 2^{-d-3} = O\left(\left(\frac{3}{5}\right)^{3d}\right) + O\left(\frac{d}{2^{2d}}\right),$$

from which we find

$$\varepsilon_d = \frac{2^{-d}}{12} + O\left(\frac{d}{2^{2d}}\right) \quad \text{as } d \to \infty.$$

As $F_{d+1}(z)$ has a simple pole at ρ_d, and by means of Rouche's Theorem, see IX.6.2 in (5), by comparing $\left|\frac{1}{F_{d+1}(z)}\right|$ with $\left|1 - \frac{z}{1-z}\right|$ on the circle $|z| = \frac{3}{4}$, we see that $\frac{1}{F_{d+1}(z)}$ has no other zeros in $|z| \leq 3/4$.

It follows that

$$[z^n]F_{d+1}(z) = [z^n]\frac{A_d}{1 - z/\rho_d} + O\left(\left(\frac{4}{3}\right)^n\right) \quad \text{with } A_d = \frac{1}{\rho_d \left.\frac{d\tau(z,0)}{dz}\right|_{z=\rho_d}} = \frac{2 + O(2^{-d})}{\left.\frac{d\tau(z,0)}{dz}\right|_{z=\rho}},$$

where from (2.4) as $d \to \infty$

$$\frac{d\tau(z,0)}{dz}\Big|_{z=\rho} = \frac{\partial}{\partial z} \sum_{i \geq 1} \frac{(-1)^{i-1} z^{i+(d+1)\binom{i}{2}}}{(z;z)_i}\Big|_{z=\rho} = \left(\frac{z}{(1-z)^2} + O(z^d)\right)\Big|_{z=\rho} = 4 + O(2^{-d}).$$

Therefore as $d \to \infty$, $A_d = \frac{1}{2} + O(2^{-d})$. Let us now restrict our attention to those d for which $n^{-3} \leq 2^{-d} \leq \frac{\log n}{n}$. The probability that $X \leq d$ is then approximated as $n \to \infty$ by

$$A_d\, \rho_d^{-n} = A_d \left(\frac{1}{2} + \frac{2^{-d}}{12} + O\left(\frac{d}{2^{2d}}\right)\right)^{-n} = 2^{n-1} \exp\left(-\frac{2^{-d}\, n}{6}\right)\left(1 + O\left(2^{-d}\right) + O\left(n\, d\, 2^{-2d}\right)\right)$$

$$= 2^{n-1} \exp\left(-\frac{2^{-d}\, n}{6}\right)\left(1 + O\left(\frac{\log^3 n}{n}\right)\right).$$

So after dividing by 2^{n-1} we have for $n \to \infty$ and $n^{-3} \leq 2^{-d} \leq \frac{\log n}{n}$,

$$\mathbb{P}(X \leq d) = \exp\left(-\frac{n}{2^d\, 6}\right)\left(1 + O\left(\frac{\log^3 n}{n}\right)\right). \tag{5.2}$$

Turning now to smaller values of $d \geq 1$, that is, d such that $2^{-d} > \frac{\log n}{n}$, a similar computation shows that (5.2) remains valid in this range, although now the probabilities $\mathbb{P}(X \leq d)$ are small, since for such d, $\exp\left(-\frac{n}{2^d\, 6}\right) = O(\frac{1}{n})$ as $n \to \infty$. Finally we must consider larger values of $d \leq n$ that is, d for which $n^{-3} > 2^{-d}$, or equivalently, $d \geq 3 \log_2 n$. In this range we find that

$$\mathbb{P}(X \leq d) = 2^{n-1} \exp\left(-\frac{2^{-d}\, n}{6}\right)\left(1 + O\left(\frac{1}{n^2}\right)\right). \tag{5.3}$$

In view of the O estimates in (5.2) and (5.3) we may deduce that the mean value of X satisfies

$$\mathbb{E}_{max}(n) := \sum_{d=0}^{n}(1 - \mathbb{P}(X \leq d)) = \left(\sum_{d=0}^{n}\left(1 - \exp\left(-\frac{n}{2^d\, 6}\right)\right)\right)\left(1 + O\left(\frac{\log^4 n}{n}\right)\right) \text{ as } n \to \infty.$$

We now use the Mellin transform, (see (3)), to estimate the function

$$f(t) = \sum_{d \geq 0}\left(1 - \exp\left(-\frac{t}{2^d\, 6}\right)\right) \text{ as } t \to \infty.$$

The Mellin transform of $f(t)$ is

$$f^*(s) = -\sum_{d \geq 0}\left(2^d\, 6\right)^s \Gamma(s) = -\frac{\Gamma(s)\, 6^s}{1 - 2^s} \quad \text{for } -1 < \Re s < 0.$$

Next we apply the Mellin inversion formula to recover $f(t)$:

$$f(t) = -\frac{1}{2\pi i}\int_{-\frac{1}{2}-i\infty}^{-\frac{1}{2}+i\infty} \Gamma(s)\frac{6^s}{1 - 2^s}\, t^{-s}\, ds.$$

We then shift the line of integration to the right and collect the (negative) residues at $s = \frac{2k\pi i}{L}$ where $k \in \mathbb{Z}$ and $L = \log 2$.

$$f(t) = \sum_{k \geq 0} \operatorname{Res} \Gamma(s) \frac{6^s}{1 - 2^s} \, t^{-s} \bigg|_{s = \frac{2k\pi i}{L}} + \frac{1}{2\pi i} \int_{\frac{1}{2} - i\infty}^{\frac{1}{2} + i\infty} \Gamma(s) \frac{6^s}{1 - 2^s} \, t^{-s} \, ds \,.$$

To evaluate the residue at the double pole $s = 0$, we need the expansion of the terms in the integrand to two terms as $s \to 0$, (here γ denotes Euler's constant),

$$\Gamma(s) \sim \frac{1}{s} - \gamma \,,$$

$$6^s \sim 1 + s \log 6 \,,$$

$$t^{-s} \sim 1 - s \log t \,,$$

$$\frac{1}{1 - 2^s} \sim \frac{1}{2} - \frac{1}{sL} \,.$$

Hence the (negative) residue at $s = 0$ is

$$[s^{-1}] \left\{ \left(\frac{1}{2} - \frac{1}{sL} \right) (1 + s \log 6) \left(\frac{1}{s} - \gamma \right) (1 - s \log t) \right\}$$

$$= \log_2 t - \frac{1}{2} - \log_2 3 + \frac{\gamma}{L} \,.$$

For the residue at $s = \frac{2k\pi i}{L}$ for $k \neq 0$, let $\varepsilon = s - \frac{2k\pi i}{L}$ we have

$$[\varepsilon^{-1}] \Gamma\left(\frac{2k\pi i}{L} \right) \left(\frac{t}{6} \right)^{-\frac{2k\pi i}{L}} \frac{-1}{\varepsilon L} = -\frac{1}{L} \Gamma\left(\frac{2k\pi i}{L} \right) \left(\frac{t}{6} \right)^{-\frac{2k\pi i}{L}} \,.$$

The remainder integral $\frac{1}{2\pi i} \int_{\frac{1}{2} - i\infty}^{\frac{1}{2} + i\infty}$ is of smaller order, so we have found

Theorem 4 *The mean value of the size of the greatest ascent in the compositions of n satisfies as $n \to \infty$,*
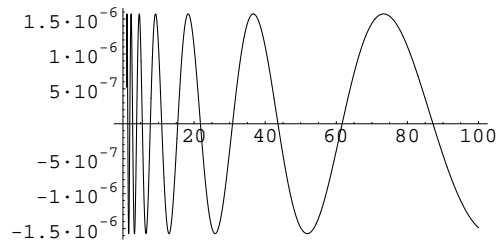
$$\mathbb{E}_{max}(n) \sim \log_2 n - \frac{1}{2} - \log_2 3 + \frac{\gamma}{L} - \delta\left(\log_2 \frac{n}{6} \right)$$

where $\delta(x)$ is a continuous periodic function of mean zero, period one and small amplitude with Fourier series

$$\delta(x) = \frac{1}{L} \sum_{k \neq 0} \Gamma\left(-\frac{2k\pi i}{L} \right) e^{2k\pi ix} \,.$$

Computations show that $\delta(x) < 1.7 \times 10^{-6}$, as a result of the fast decrease of the gamma function with imaginary argument.

The diagram below shows the plot of $\delta(\log_2 \frac{n}{6})$ for $1 \leq n \leq 100$.

Fig. 1: $\delta(\log_2 \frac{n}{6})$ for $1 \leq n \leq 100$

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